# How Well Does Paternity Confidence Match Actual Paternity? Evidence from Worldwide Nonpaternity Rates 

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The issue of paternity-whether a man really is the biological father of his supposed childrenhas long been a topic of interest to anthropologists. Evolutionary theory predicts that males will provide less parental investment for putative genetic offspring who are unlikely to be their actual offspring (e.g., Alexander 1974; Trivers 1972). Actual genetic paternity may differ from paternity confidence (a man's assessment of the likelihood that he is the father of a putative child), which must be assessed through indirect cues such as mate fidelity or child resemblance. There is great variation across cultures in beliefs about paternity (e.g., Beckerman et al. 1998; Hrdy 2000; Levine 1987), but cross-culturally, paternity confidence is positively associated with men's involvement with children or with investment or inheritance from paternal kin (e.g., Flinn 1981; Gaulin and Schlegel 1980; Greene 1978; Hartung 1985).

This paper draws on 67 studies reporting nonpaternity to examine the relationship between paternity confidence and actual paternity and to look for global variation in this relationship. I test the hypothesis that men with high paternity confidence will have higher rates of actual paternity than men with low paternity confidence by comparing nonpaternity rates from two groups of men: one biased toward high paternity confidence and the other toward low paternity confidence. The relative frequencies of men with high and low paternity confidence are generally unknown, making it difficult to estimate true nonpaternity rates for human societies.

## Nonpaternity in Cross-Cultural Perspective

Nonpaternity rates in human societies are often cited as being 10\% or greater in general populations (e.g., Alfred 2002; Cervino and Hill 2000; Stewart 1989), though little or no empirical support is generally provided for this assertion (MacIntyre and Sooman 1991). Baker and Bellis (1995) report a worldwide median nonpaternity rate of $9 \%$ from a sample of ten
studies. While little is known about global variation in nonpaternity, even less is known about cross-cultural patterns of paternity confidence. Gaulin and Schlegel (1980) used three variables measuring the degree of female sexual promiscuity among 135 societies to create a dichotomous measure of paternity confidence, and estimated that only $55 \%$ of cultures in their sample have high paternity confidence. Huber et al. (2004) used four measures of extramarital sexual activity from 57 cultures to create a 16-level measure of paternity confidence. Their results indicate that $63 \%$ of societies have paternity confidence levels at or above the median. Thus, while paternity confidence is high for many societies, a substantial minority appear have reduced levels of paternity confidence.

In recent years, the determination of nonpaternity has become a science. Modern paternity tests cannot prove paternity; instead they prove nonpaternity, by excluding men whose genotype is incompatible with that of the child in question (Pena and Chakraborty 1994; Wilson 1987). Failure to exclude a man as the father can be taken as proof of paternity if the probability of excluding non-fathers is extremely high; the probability that a man is likely to be the father is calculated using Bayesian logic, based on assumptions of the frequencies of the genotypes under consideration in the general population (for further details see Mickey, Gjertson, and Terasaki 1986; Pena and Chakraborty 1994). Contemporary paternity tests, which use DNA polymorphisms to determine nonpaternity, have probabilities of exclusion in excess of 99.99\%, so that out of 10,000 paternity tests the true non-father will be excluded as a potential father 9,999 times (Helminen et al. 1988; Jeffreys, Turner and Debenham 1991; Pena and Chakraborty 1994). Older (pre-1985) paternity tests, based on blood type or HLA antigens, had lower probabilities of exclusion, ranging from $18 \%$ (if using only the ABO blood types) to $95 \%$ or more. A lower probability of exclusion means that nonpaternity will not be established for some
non-fathers, even though they are not the actual father of the child. (For example, a man may have a blood type that is compatible with being the father of the child, even though he is not the father.) Many older studies therefore report two nonpaternity rates: the observed nonpaternity (the proportion of men excluded in the study), and the actual nonpaternity (the proportion of men who should have been excluded, not all of whom were due to limitations of the test). For example, if the probability of exclusion in a study is $50 \%$, and the study finds a $5 \%$ nonpaternity rate in the sample, then the actual nonpaternity rate is $10 \%$.

## Methods

For the present study, published data on nonpaternity rates were gathered through extensive literature searches, using online databases, bibliographies, and journal indices, resulting in a sample of 67 nonpaternity rates. While this list cannot be considered complete, it is the most extensive published list of nonpaternity rates assembled to date, far exceeding pre-existing lists (e.g., Baker and Bellis 1995; James 1993; Lucassen and Parker 2001; MacIntyre and Sooman 1991; Sasse et al. 1994).

The measures of nonpaternity used in this study were estimated in many different ways. Because older methods of establishing nonpaternity had lower probabilities of exclusion (i.e., were less likely to detect nonpaternity), the proportion of men actually excluded in older (pre1985) papers is always less than the number of non-fathers in the sample who should have been excluded. Most researchers adjust for this accordingly, presenting both the observed nonpaternity and the actual (adjusted) nonpaternity; in a few cases, actual nonpaternity was not stated, and I calculated actual nonpaternity from the stated probability of exclusion. More recent references, with greater probabilities of exclusion (greater than 95\% and typically exceeding 99.99\%), do
not distinguish between observed and actual nonpaternity, as the difference is minimal. In a few cases where nonpaternity is estimated through other methodologies (for example, Mendelian inconsistencies), there is no difference between observed and actual nonpaternity. Where possible, the actual nonpaternity rate will be used for the analysis.

On the basis of the level of presumed paternity confidence within each group, the dataset is divided into three groups: (1) men with relatively high paternity confidence, (2) men with relatively low paternity confidence, and (3) men whose paternity confidence is unknown.

1. High paternity confidence. This group includes 22 data points from genetic studies or other sources that are likely to bias the sample toward high paternity confidence (see table 1 ). None of these studies come from random samples. The nature of these studies (especially the genetic and lineage studies) will bias the samples toward men with high paternity confidence because men who do not believe they have fathered their putative children will be less likely to participate in the research. Most of these studies include mother/father/child trios, and many contain primarily or exclusively married couples. Since men in marriages are likely to have higher paternity confidence than men who father children outside of marriage (Anderson, Kaplan and Lancaster 2006a, 2006b), this will further bias the sample toward men with high paternity confidence. Some men in this sample undoubtedly do not have high paternity confidence; additionally, the studies may have included covert adoptions, misidentified stepchildren, etc., for whom paternity confidence is zero. Overall, however, these studies are likely to include men whose paternity confidence is relatively high.
[Table 1 about here]
2. Low paternity confidence. All of the 31 data points in this group come from studies of disputed paternity (for example, from paternity testing laboratories) (see table 2). The men in this
sample were sufficiently doubtful of their paternity to participate in laboratory tests to determine if they were the fathers of their putative children; thus, this sample is categorized as having low paternity confidence.
[Table 2 about here]
3. Unknown paternity confidence. This group contains 14 data points for which no conclusion can be drawn regarding the paternity confidence of the men involved (see table 3). Many are from unpublished or secondhand sources, and therefore we do not know whether the reported nonpaternity reflects observed or actual nonpaternity, or if the rates have been adjusted for laboratory error. ${ }^{1}$ One study (Baker and Bellis 1990) estimates nonpaternity through women's reports of sexual behavior; they present no data on whether the women's partners had high or low paternity confidence in any resulting pregnancies.
[Table 3 about here]

## Sidebar A: Electronic edition only

Two of the nonpaternity rates cited in the tables differ from those cited in Baker and Bellis (1995). In table 1, I cite the rate for Edwards (1957) as 3.7\% while Baker and Bellis (1995) cite $5.9 \%$. Edwards states (p. 85) that there are 17 blood group incompatibilities ("bastards"), out of 2578 matings. The exclusion rate is $18 \%$. This calculation (( $17 / 2578$ )/0.18) yields a nonpaternity rate of $3.66 \%$, which is roughly equal to the "nearly 4 percent" (p. 85) inferred by Edwards. It is unclear how Baker and Bellis arrived at their figure of 5.9\%.

In table 1 I cite nonpaternity in Salmon et al. (1980) as ranging from 6.9 to $9.4 \%$ (with a median of $8.15 \%$ used for analysis). Baker and Bellis (1995) cite the rate as being less than $14.6 \%$. Salmon et al. (1980) state that they found 25 exclusions from 171 families, but they do not state how many children were tested (which is required for the denominator of the
nonpaternity rate). Baker and Bellis divide 25 by 171 to reach an upper limit of $14.6 \%$. However, Salmon et al. state that 76 if the families have one child, and 95 have between two to ten children. Thus, the sample cannot be smaller than 266 children. If the multiple-child families have three offspring on average, then $\mathrm{N}=361$ children in the sample. I use these numbers ( $25 / 361$ and $25 / 266$ ) to create lower and upper boundaries on nonpaternity in this sample.

The paternity exclusion rate for Böök (1950) in table 1 is $100 \%$ because paternity was deduced through a Mendelian inconsistency in a simple genetic disease. Among the older (pre1985) samples, Ashton (1980) does not state the exclusion probability (table 1). However, he notes that his data have been corrected for exclusion and for laboratory error; thus, his stated nonpaternity rate is used as the actually nonpaternity. In table 3 I label the paternity exclusion rate for Baker and Bellis (1990) as not applicable rather than not stated because, alone among these estimates, Baker and Bellis do not deduce nonpaternity through DNA, blood type, HLA antigens, or other biological method. Rather, they estimate nonpaternity through information on the timing of women's sexual behavior, specifically how many women in their sample had extrapair partners around the time of conception (making assumptions about the probability of conception, etc.). Their sample was recruited from the readership of a British women's magazine; because the investigators had no data on men's assessment of paternity, and men were not involved in the sampling procedure (thereby biasing the sample towards high paternity confidence), I have put them in the unknown paternity confidence sample.

It could be argued that since the sample whose paternity confidence is unknown is unlikely to be composed of men actively disputing paternity, the cases in table 3 should be added to those of the men with high paternity confidence (table 1). This will also make my estimates of
nonpaternity more comparable to previous studies (e.g., Baker and Bellis 1995). The analyses will first examine each group separately, and then combine the high and unknown paternity confidence groups into a single group.

The data presented in tables 1-3 allow us to examine whether there is worldwide variation in nonpaternity rates by men's paternity confidence level. The data were organized geographically into three groups: United States and Canada ( $\mathrm{N}=27$ ), Europe ( $\mathrm{N}=26$ ), and elsewhere $(\mathrm{N}=14)$. The "elsewhere" category is extremely heterogeneous, as it encompasses samples from South and Central America, Africa, Israel and India; however, none of these regions have sufficient sample sizes to stand alone as separate categories. While it would be interesting to examine nonpaternity by ethnic group, the data do not allow this as most studies with multiethnic samples do not provide breakdowns by ethnic group.

Because the data are not normally distributed, comparisons between groups will be made using the nonparametric Wilcoxon rank-sum test. All analyses were done using STATA SE v. 8.2. The actual nonpaternity rates used for analysis are uncorrelated with the sample size, probability of exclusion, or year of publication associated with each study.

## [Sidebar B: electronic edition only]

The probability of exclusion is highly correlated with the year of publication (Spearman's rho $=0.7180, \mathrm{~N}=43, \mathrm{p}<0.0001$ ), but is uncorrelated with the sample size of the study (Spearman's rho $=-0.1291, \mathrm{~N}=43, \mathrm{p}=0.4095$ ). The observed nonpaternity rate (e.g., unadjusted for the probability of exclusion) is significantly correlated with the probability of exclusion (Spearman's rho $=0.5143, \mathrm{~N}=18, \mathrm{p}=0.0290)$ and with actual nonpaternity $($ Spearman's rho $=$ 0.8952, $\mathrm{N}=18, \mathrm{p}<0.0001$ ), but is uncorrelated with both year of publication (Spearman's rho $=$ $-0.2054, N=18, p=0.4136$ ) and sample size (Spearman's rho $=-0.3626, N=18, p=0.1392$ ). In
contrast, actual nonpaternity is uncorrelated with year of publication (Spearman's rho $=-0.1994$, $\mathrm{N}=67, \mathrm{p}=0.1058$ ), the probability of exclusion (Spearman's rho $=0.1506, \mathrm{~N}=43, \mathrm{p}=0.3352$ ), or sample size (Spearman's rho $=-0.0671, \mathrm{~N}=58, \mathrm{p}=0.6168$ ). Because actual nonpaternity is uncorrelated with the other variables available in the dataset, there is no need to perform multivariate analysis to control for their effects.

## Results

The median nonpaternity rate for the high paternity confidence sample is 1.7\% (range: 0.4 11.8), while median nonpaternity for the low paternity confidence sample is 29.8\% (range: 14.3 55.6). The median nonpaternity rates for these two groups are significantly different (Wilcoxon sign-rank test, $\mathrm{z}=-6.156, p<0.0001$ ). The median nonpaternity of men whose paternity confidence is unknown is 16.7\% (range: 2.0-32.0). This is significantly greater than that of the high paternity confidence sample (Wilcoxon sign-rank test, $\mathrm{z}=-4.382, p<0.0001$ ), and significantly lower than that of the low paternity confidence sample (Wilcoxon sign-rank test, z $=3.531, p=0.0004)$. When the high and unknown paternity confidence samples are combined, the median nonpaternity is $3.3 \%$ (range: $0.4-32.0$ ). This is significantly less than median nonpaternity for men with low paternity confidence (Wilcoxon sign-rank test, $\mathrm{z}=-6.099, p<$ 0.0001).

Figure 1 shows median nonpaternity by geographic location for the high paternity confidence, combined high and unknown paternity confidence, and low paternity confidence samples. Within each paternity confidence group, there is no significant geographic variation in the median values of nonpaternity (Wilcoxon sign-rank tests, results not shown, $p>0.51$ for every comparison). In other words, men with high paternity confidence have similar levels of
actual paternity in the United States and Canada, Europe, and the rest of the world; the same is true for the other two paternity confidence groups. However, for all three geographic locations nonpaternity is significantly greater in the low paternity confidence sample than in the high paternity confidence sample (Wilcoxon sign-rank tests: United States and Canada $\mathrm{z}=-3.873, p=$ 0.0001; Europe $\mathrm{z}=-3.761, p=0.0002$; elsewhere $\mathrm{z}=-2.611, p=0.0090$ ) and in the combined high/unknown paternity confidence sample (Wilcoxon sign-rank tests: United States and Canada $\mathrm{z}=-4.392, p<0.0001$; Europe $\mathrm{z}=-3.763, p=0.0002$; elsewhere $\mathrm{z}=-2.333, p=0.0196)$.
[Figure 1 about here]

## Conclusion

This survey of published estimates of nonpaternity suggests that for men with high paternity confidence, nonpaternity rates are typically $1.7 \%$ (if we exclude studies of unknown methodology) to $3.3 \%$ (if we include such studies). These figures are substantially lower than the "typical" nonpaternity rate of $10 \%$ or higher cited by many researchers, often without substantiation (e.g., Alfred 2002; Cervino and Hill 2000; Stewart 1989), or the median worldwide nonpaternity rate of 9\% reported by Baker and Bellis (1995).

Men who have low paternity confidence and have chosen to challenge their paternity through laboratory testing are much less likely than men with high paternity confidence to be the fathers of their putative children. Although these men presumably have lower paternity confidence than men who do not seek paternity tests, this group is heterogeneous; some men may be virtually certain that the putative child is not theirs, while others may simply have sufficient doubts to warrant testing. Most of these men are in fact the fathers of their putative genetic children; only $29.8 \%$ could be excluded as biological fathers of the children in question.

The results of this study raise many questions. What is the true level of nonpaternity in any particular human population? Since most if not all samples are biased toward men with either high or low paternity confidence, this question cannot yet be answered. Presumably the true level of nonpaternity is a weighted average of men from these two groups, raising the question of how many men have low versus high paternity confidence. For example, in order for the population nonpaternity rate to be $10 \%, 75 \%$ of men in the population would have to have high paternity confidence (nonpaternity $=3.3 \%$ ) and $25 \%$ have low paternity confidence (nonpaternity = 29.8\%). Anderson et al. (2006a) report that men living in Albuquerque, New Mexico do not believe that they are the father of $1.46 \%$ of pregnancies attributed to them, implying a total nonpaternity rate for that sample of $3.7 \%$. I know of no other study that has estimated the frequency of low and high paternity confidence within a particular sample, though this clearly has important implications for child wellbeing and family dynamics. Further crosscultural investigation of the relationship between paternity and paternity confidence is warranted.

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## Footnote

1. The unpublished !Kung nonpaternity rate reported in Trivers (1972) has subsequently been suggested to be due almost entirely to laboratory error, since the rate of nonmaternity in the sample was approximately the same (Howell 2000; Smith 1984).

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Table 1. Nonpaternity Rates (\%) when Paternity Confidence Is Relatively High

| Population | Actual <br> nonpaternity (\%) | Observed <br> nonpaternity (\%) <br> (\%) | Probability of <br> exclusion (\%) | Sample <br> size | Source |
| :--- | :---: | :---: | :---: | :---: | :--- |
| Sephardic Kohanim (Jewish priests) | 0.4 | not stated | not stated | 24 | Boster et al. (1999) |
| United States | 0.8 | not stated | not stated | 496 | Broman (1999) |
| Switzerland | 0.83 | not stated | 99 | 1,607 | Sasse et al. (1994) |
| Ashkenazic Kohanim (Jewish priests) | 1.2 | not stated | not stated | 44 | Boster et al. (1999) |
| Canada (Quebec) | 1.2 | not stated | not stated | 42 | Heyer et al. (1997) |
| England | 1.3 | not stated | not stated | 48 | Sykes and Irven (2000) |
| England | 1.35 | not stated | not stated | 521 | Brock and Shrimpton (1991) |
| United States (Michigan), white | 1.49 | 0.28 | 18.8 | 1,417 | Schacht and Gershowitz (1963) |
| Iceland | 1.49 | not stated | not stated | not stated | Helgason et al. (2003) |
| United Kingdom | 1.59 | not stated | not stated | 756 | Chataway et al. (1999) |
| Sweden | 1.6 | not stated | 100 | 63 | Böök (1950) |
| Canada | 1.75 | not stated | not stated | 57 | Poon et al. (1993) |
| United States (California), white | 2.1 | 0.8 | 38.1 | 6,960 | Peritz and Rust (1972) |
| United States (Hawaii) | 2.3 | not stated | not stated | 2,839 | Ashton (1980) |
| United States | 2.8 | 0.5 | 18 | 200 | Wiener et al. (1949) |
| France | 2.8 | not stated | not stated | 362 | Le Roux et al. (1992) |
| Mexico | 2.9 | 2.3 | 80.3 | 217 | Peñaloza et al. (1986) |
| United Kingdom (West London) | 3.7 | 0.7 | 18 | 2,596 | Edwards (1957) |
| France | $6.9-9.4$ | not stated | 94.4 | $266-361$ | Salmon et al. (1980) |
| Brazil/Venezuela (Yanomamo) | 6.1 | 64 | 132 | Neel and Weiss (1975) |  |
| United States (Michigan), black | 9.1 | 1.91 | 18.9 | 523 | Schacht and Gershowitz (1963) |
| Mexico (Nuevo Leon) | 10.1 | 8.1 | 64 | 396 | Cerda-Flores et al. (1999) |

Table 2. Nonpaternity Rates (\%) from Paternity Testing Laboratories

| Population | Actual nonpaternity (\%) | Observed nonpaternity (\%) | Probability of exclusion (\%) | Sample <br> size | Source |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Russia | 14.3 | not stated | 99.75 | 21 | Molyaka et al. (1997) |
| Finland ${ }^{\text {a }}$ | 15.2 | 14.3 | 94 | 35 | Helminen et al. (1992) |
| United Kingdom | 16.6 | not stated | 99.99 | 1,702 | Jeffreys et al. (1991) |
| Germany | 16.8 | not stated | 99.87 | 256 | Krawckak et al. (1993) |
| Brazil (Belo Horizonte) | 22.0 | not stated | 99.99 | 200 | Pena et al. (1993) |
| South Africa, white | 22.4 | not stated | 99.4 | 264 | Du Toit et al. (1989) |
| United States (Cleveland, OH) | 23.9 | 12 | 50 | 67 | Marsters (1957) |
| United States (Los Angeles), white | 24.9 | not stated | 97 | 1,393 | Mickey et al. (1986) |
| United States | 25.0 | not stated | 90-99 | 1,000 | Terasaki (1978) |
| United States | 25.2 | not stated | 97 | 2,500 | Houtz et al. (1982) |
| United States | 26.0 | not stated | 99.39 | 50 | Alford et al. (1994) |
| Portugal | 27.7 | not stated | 99.9 | 83 | Geada et al. (2000) |
| United States (New York City) | 28.7 | 14.3 | 50 | 300 | Sussman (1956) |
| United States (Baltimore) | 29.0 | not stated | not stated | 124 | James (1993) |
| United States (New York City) | 29.4 | not stated | 93 | 102 | Baird et al. (1986) |
| Portugal | 29.8 | not stated | 99.9 | 790 | Geada et al. (2000) |
| South Africa, Cape Malay | 30.5 | not stated | 98.0 | 59 | Du Toit et al. (1989) |
| United States (New York City), black ${ }^{\text {a }}$ | 30.6 | 15.3 | 50 | 98 | Wiener (1950) |
| United States (Cleveland, OH) | 32.0 | 16 | 50 | 200 | Marsters (1957) |
| United States (New York City), white ${ }^{\text {a }}$ | 34.4 | 17.2 | 50 | 425 | Wiener (1950) |
| Finland | 34.6 | not stated | 99 | 26 | Helminen et al. (1988) |
| United States (Illinois) | 37.0 | not stated | 99 | 753 | Strom et al. (1996) |
| France (Paris) | 38.1 | not stated | 99.99 | 543 | Rouger and van Huffel (1996) |
| Sweden | 38.7 | 33.5 | 86.5 | 5,018 | Valentin (1980) |
| South Africa, Cape Coloured | 40.1 | not stated | 99.8 | 1,156 | Du Toit et al. (1989) |
| South Africa, black | 41.1 | not stated | 99.5 | 645 | Du Toit et al. (1989) |
| United States | 42.0 | 21 | 50 | 100 | Sussman (1954) |
| Italy | 45.0 | not stated | 99.6 | 31 | Gasparini et al. (1991) |
| United States (Illinois) | 53.0 | not stated | 99.83 | 37 | Strom et al. (1996) |
| Sweden ${ }^{\text {a }}$ | 55.0 | 8.4 | 15.4 | 142 | Hirschfeld and Heiken (1963) |
| United States ${ }^{\text {a }}$ | 55.6 | 27.8 | 50 | 108 | Unger (1953) |

## a. Actual nonpaternity not calculated in the original paper;

Table 3. Nonpaternity Rates (\%) When Paternity Confidence Is Unknown
Probability of Sample

| Population | Nonpaternity (\%) | exclusion (\%) | size | Source |
| :--- | :---: | :---: | :---: | :--- |
| Southern Africa (!Kung) | 2 | not stated | not stated | Harpending (unpublished), cited in Trivers (1972) |
| United Kingdom | 4.8 | not stated | 21 | Shields (unpublished), cited in Scharfetter (1978) |
| England | $6.9-13.8$ | not applicable | 2,708 | Baker and Bellis (1990) |
| United States (rural Michigan) | "ca. 10" | not stated | not stated Chagnon (unpublished), cited in Smith (1984) |  |
| Munich and Copenhagen | "at least 10" | not stated | not stated unpublished; cited by Ritz in Grünfeld (1985) |  |
| United States | $10-30$ | not stated | not stated Reed (unpublished), cited in Allison (1996) |  |
| Italy | 13.2 | not stated | 38 | Hirsch and Vetta (1978) |
| India | $" \sim 15.3 "$ | not stated | not stated Meisner 1999, cited in Cervino and Hill (2000) |  |
| United States | 18.0 | 50 | 67 | Sussman and Schatkin (1957) |
| England (Liverpool) | $20-30$ | not stated | not stated McLaren (unpublished), cited in Cohen (1977) |  |
| United States (Michigan), black | 20.1 | 18.75 | 265 | Wiener (1966) |
| England | 30 | not stated | $200-300$ | unpublished, cited in Philipp (1973) |
| Africa | $" \sim 30 "$ | not stated | not stated Ruwende 1996, cited in Cervino and Hill (2000) |  |
| India (Vishakapatnam) | 32 | not stated | not stated Meisner 1999, cited in Cervino and Hill (2000) |  |

Figure 1. Median nonpaternity rates by paternity confidence and geographic location


