Memorandum on Differences between Carpenter/Zucker/Avorn data and Nardinelli et al Data, and Re-Analysis of Data Adjusting for These Differences, and Robustness and Sensitivity Checks on Analyses

May 17, 2008

### SUMMARY:

- There are discrepancies in safety-based withdrawals that represent two Nardinelli et al omissions, while other discrepancies involve matters of interpretation (for which the FDA's own coding has changed). In a wide variety of sensitivity analyses, the association between just-before-deadline approvals and safety-based withdrawals persists.
- Nardinelli et al's data on postmarket safety differ materially from data that (1) the agency has posted elsewhere on its website, (2) that the agency provided to other authors and that were presented at FDA-sponsored conferences, and (3) from the published medical literature.
- There are significant discrepancies in the coding of standard/priority ratings. Although we coded our data from the FDA web site, these discrepancies represent probable coding errors on our part.
- There are also significant discrepancies between our data on black-box warnings and the Nardinelli et al data. Nardinelli et al's data on black-box warnings are assembled without a clear or rigorous methodology.

# THE RELATIONSHIPS OBSERVED PERSIST EVEN WHEN USING THE NARDINELLI ET AL DATA POSTED ON APRIL 7, 2008 (see Sections 4 and 5, Sections 11-16).

- Withdrawals: If Nardinelli et al's 314 NMEs are used, and their pre-deadline approval measure is used, the relationships between pre-deadline approvals and <u>safety-based</u> withdrawals are larger and more statistically robust than reported in the original *NEJM* paper (OR = 5.89, p = 0.01 from exact logistic regression; OR = 6.09, p = 0.006 from random-effects logistic regression).
- <u>Black Box Warnings</u>: If Nardinelli et al's 314 NMEs are used, and their pre-deadline approval measure is used, the relationships between pre-deadline approvals and <u>black-box warnings</u> are slightly smaller and more statistically robust than reported in the original *NEJM* paper (OR = 3.41, *p* = 0.02 from exact logistic regression; OR = 4.53, *p* = 0.009 from random-effects logistic regression).
- <u>Dosage-Form Disc.</u> If Nardinelli et al's 314 NMEs are used, and their pre-deadline approval measure is used, the relationships between pre-deadline approvals and <u>dosage-form discontinuation</u> are similar to and more statistically robust than reported in the original *NEJM* paper (OR = 3.85, *p* = 0.007 from exact logistic regression; OR = 5.73, *p* = 0.003 from random-effects logistic regression).

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# EXECUTIVE SUMMARY May 17, 2008

After reviewing our data and correcting for factual errors, we can state with confidence that our major findings remain intact as do their implications. In particular, post-PDUFA NMEs are far more likely to be approved in the two months before the user-fee deadline than at other times, and these just-before-deadline approvals are significantly more likely to encounter postmarket safety issues once on the market. Odds ratios for the association between just-before-deadline approvals and our three postmarket safety measures remain large (> 3) and statistically significant using a variety of tests.

Our first and main measure of post-marketing safety problems was safety-based withdrawals, and we stand by our results. Seven of the eleven withdrawals in our sample were just-before-deadline approvals (alatrofloxacin, rofecoxib, cerivastatin, grepafloxacin, troglitazone, trovafloxacin, and valdecoxib). We did not originally code alosetron (Lotronex) as a safety-based withdrawal because the drug was returned to market in 2002 before the end of our study period. Levomethadyl (Orlaam) became unavailable because of reduction in supply by the manufacturer and was therefore not listed as a withdrawal in our data, nor was it listed as such in FDA publications through fall 2005; see below). Even if we add alosetron and levomethadyl to the analyses it does not change the original findings; in fact, use of the FDA's sample of NMEs and the pre-deadline measure of the FDA yields larger and more statistically significant associations between deadline approvals and safety-based withdrawals than were reported in the original article (see the embedded table). The final difference in drug withdrawals, Tegaserod (Zelnorm), was withdrawn just last March, long after our paper was already in the review process.

Re-Analysis of V	Vithdrawals and Dosage Form	Discontinuations,
Using Nardinelli et a	1's 314 NMEs and their Deadli	ne Approval Measure
Postmarket Safety Measure	Cross Tabulation	Odds Ratios from Exact Logistic Regression (ELR) and Random-Effects Logistic Regression (RLR)
Carpenter/Zucker/Avorn (CZA) Withdrawal Measure	7 of 88 deadline approvals are withdrawn; 4 of 226 non-deadline approvals.	ELR: OR = 5.89, P = 0.01 RLR: OR = 6.09, P = 0.006
CZA Withdrawal Measure, add Alosetron	7 of 88 deadline approvals are withdrawn; 5 of 226 non-deadline approvals.	ELR: OR = 4.52, P = 0.03 RLR: OR = 4.61, P = 0.01
CZA Withdrawal Measure, add Levomethadyl	7 of 88 deadline approvals are withdrawn; 5 of 226 non-deadline approvals.	ELR: OR = 4.90, P = 0.02 RLR: OR = 5.05, P = 0.01
CZA Withdrawal Measure, add Alosetron <u>and</u> Levomethadyl	7 of 88 deadline approvals are withdrawn; 6 of 226 non-deadline approvals.	ELR: OR = 3.91, P = 0.04 RLR: OR = 3.99, P = 0.02
Dosage-Form Discontinuations	16 of 88 deadline approvals are discontinued; 20 of 226 non-deadline approvals.	ELR: OR = 2.98, P = 0.008 RLR: OR = 3.13, P = 0.005
	ion year. RLR controls for submidom effects grouped by NME pri	

The data posted by Nardinelli et al on April 7, 2008 is a new dataset never before published or posted. It differs materially from data elsewhere on the FDA website, from data that the agency has published elsewhere or has provided to other authors, and from published medical literature. There are serious problems with the publicly available FDA data on withdrawals. For example, the agency's Drugs@FDA site does not list two very important withdrawn drugs, bromfenac (Duract) and mibefradil (Posicor). Another drug included in the FDA's data of April 2008, levomethadyl (Orlaam) was not mentioned in published FDA material

(and material generated by CDER) more than two years after Nardinelli et al report that it was

withdrawn. The FDA's own notice<sup>1</sup> describes it as not being a safety-based withdrawal. Still, as

<sup>&</sup>lt;sup>1</sup> See <a href="http://www.fda.gov/CDER/drug/shortages/orlaam.htm">http://www.fda.gov/CDER/drug/shortages/orlaam.htm</a>, reproduced and discussed below. See also Jim Rosack, "Med Check," *Psychiatric News* 38 (19) (October 3, 2003) <a href="http://pn.psychiatryonline.org/cgi/content/full/38/19/36">http://pn.psychiatryonline.org/cgi/content/full/38/19/36</a> [accessed April 16, 2008].

the table above suggests, recoding levomethadyl acetate as a withdrawal still does not change our original findings.

The results for our third measure – one or more dosage-form discontinuations – also hold up and are virtually identical to those in the published paper.

The Nardinelli letter mentions two other issues with our data. The first is the difference in standard and priority coding. We have recoded our drugs according to Drugs@FDA and now identify 132 priority drugs, just as Nardinelli et al report for their April 2008 data. Most of the drugs that we coded as standard that Drugs@FDA lists as priority were drugs with review times of 14 months or more, so the recoding does not affect our measure, as these are not pre-deadline approvals under any version of PDUFA.

Nardinelli et al also extrapolate from our Figure 1 and state that our data on approval times differ from theirs. This is not a matter of data but a matter of rounding. Our original coding was based on dividing the number of days of approval by 30, which produced a number of drugs that had met the review deadlines but had numerical review times that appeared to put them past the deadlines. For instance, L-Glutamine (Nutrestore; NDA # 21667) was submitted on August 8, 2003 and approved on June 10, 2004, which generates a review time of 10.23 months, which we rounded down to produce the patterns in Figure 1. This had the effect of rendering NMEs that went into the tenth month of review being coded as ninth-month approvals, which accounts for the difference. If timing calculations of Nardinelli et al is preferred, *our finding about the concentration (or "piling") of approval decisions near the deadline approvals is even stronger*.

It is for postmarket black-box warnings that the greatest discrepancies persist remain. The main difference between our data on black box warnings (BBW) and that of Nardinelli et al is the date at which coding was stopped; 17 of the 29 drugs with BBWs in the FDA data, or 58% of the BBWs in their sample, were added in the last two years. Our analysis included data through August 2005. This points to the increased use of BBWs (and reliance upon class relabeling in particular) by the FDA in recent years, which suggests that *BBWs may not be as consistent a measure of postmarket safety problems as they once were*.

In reviewing all the drugs in our data set and the Nardinelli et al dataset with boxed warnings, we have identified errors in the data provided by Nardinelli et al as well as in our own data:

- In their April 2008 posted data, Nardinelli et al have missed at least five (5) postmarket BBWs added to approval NMEs (adefovir, emtricitabine, entecavir, tenofovir, tiprinavir), and has coded as many as three other drugs as having BBWs added which contained no significant safety information. One clear example is the FDA's coding of saquinavir; an important NIH publication (for which two FDA officials consulted) did not list this drug as having a black box warning in February 2002, over a year after the FDA says that a BBW was added (see below).
- We incorrectly coded three (3) NMEs as having postmarket BBWs added. Although the boxed warning changed in all three cases, we agree that these should not have been coded as imparting major new safety information. We also omitted five (5) BBWs thatshould have been coded as having imparted important new safety information. In some cases this is because we had coded the NME as withdrawn before the BBW was added. [Another drug (celecoxib) received its BBW on July 29, 2005, likely before the drug was listed on the KUMC web database.] Dr. Carpenter assumes sole responsibility for these errors.
- For a number of other drugs, there are reasonable differences of judgment between our coding, the FDA's coding, and that of published medical literature on whether the drugs had significant postmarket black-box warnings added.
- (3) Correcting for these errors, and using either (a) Nardinelli et al's April 2008 data, or (b) combined data before August 2005 or before the post-Vioxx set of BBW relabelings, we still find large (> 3) and statistically significant odds ratios in the vast majority of replications (see <a href="http://people.hmdc.harvard.edu/~dcarpent/fdaproject/nejmresponse20080517.pdf">http://people.hmdc.harvard.edu/~dcarpent/fdaproject/nejmresponse20080517.pdf</a>), and the first two rows of results in the following table.]
- (4) Even if we accept <u>all</u> of Nardinelli et al's determinations that a given drug had a safety-based withdrawal or a postmarket BBW, and we add only those drugs that we appropriately identified as having postmarket safety problems (adefovir, alatrofloxacin, emtricitabine, entecavir, tenofovir, tipranavir, tolcapone, and trovafloxacin), there is still an appreciable (OR > 2) and statistically significant difference between just-before-deadline approvals and the combined withdrawal-or-warning measure (see

http://people.hmdc.harvard.edu/~dcarpent/fdaproject/nejmresponse20080517.pdf, the last row of results in following table, and Section 12.3 below).

Re-Analysis of Black Box Warning Measure
and Combined "Withdrawal or Warning" Measure,
using Nardinelli et al's 314 NMEs and their Deadline Approval Measure

Postmarket Safety Measure	Cross Tabulation	Odds Ratios from Exact Logistic Regression (ELR) and Random- Effects Logistic Regression (RLR)
Carpenter/Zucker/Avorn (CZA) BBW Measure, recode 6 CZA omissions, drop 3 CZA miscodes, accept saquinavir.* (21 total BBWs)	12 of 88 deadline approvals are BBWs added; 9 of 226 non-deadline approvals.	ELR: OR = 3.85, P = 0.007 RLR: OR = 5.73, P = 0.003
Carpenter/Zucker/Avorn (CZA) BBW Measure, recode 6 CZA omissions, recode 2 close cases, drop 3 CZA miscodes, accept saquinavir.* (19 total BBWs)	10 of 88 deadline approvals have BBWs added; 9 of 226 non-deadline approvals.	ELR: OR = 3.41, P = 0.03 RLR: OR = 4.52, P = 0.009
Withdrawal or Warning Measure, recode 6 CZA omissions, drop 3 CZA miscodes, accept saquinavir.*	15 of 88 deadline approvals have SBWs or BBWs; 12 of 226 non-deadline approvals.	ELR: OR = 3.93, P = 0.002 RLR: OR = 5.10, P = 0.001
Withdrawal or Warning Measure, recode 6 CZA omissions, recode 2 close cases, drop 3 CZA miscodes, accept saquinavir.*	13 of 88 deadline approvals have SBWs or BBWs; 12 of 226 non-deadline approvals.	ELR: OR = 3.58, P = 0.009 RLR: OR = 4.20, P = 0.003
Withdrawal or Warning Measure, accept all FDA positive codes of withdrawals or BBWs, extend data up through December 2007	21 of 88 deadline approvals have SBWs or BBWs; 21 of 226 non-deadline approvals.	ELR: OR = 2.29, P = 0.03 RLR: OR = 3.42, P = 0.01

Note: ELR controls for submission year. RLR controls for submission year, epidemiological covariates, and includes 135 random effects grouped by NME primary indication. \*Saquinavir represents a possible FDA error (see Section 2.3 below), but we code it as a BBW for sake of demonstration. Results are stronger if saquinavir is coded as not having a postmarket BBW (see Section 8.2 and Section 9 below).

Thus, even after addressing the issues raised by Nardinelli et al, our main findings remain valid. We did code standard and priority drugs differently from FDA, a discrepancy for which we assume responsibility but which does not materially affect our conclusions. Our determination of the presence of black-box warnings also differs from that of Nardinelli et al, and while some of these discrepancies stem from our coding errors, at least as many are the result of errors in the data of Nardinelli et al. In many other cases it is debatable which classification is more accurate. However, even after adjusting for these differences, our findings remain robust,

statistically significant, and consistent with our earlier findings. Our paper did not conclude that FDA should be freed of all deadlines. More research is needed about how best to make the American drug approval process as efficient, reliable, and credible as possible.

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- 1. Examination of Differences for Safety-Based Withdrawals, and Re-Analysis of Data
- 1.1 Coding Rules for Approval Times and Postmarket Safety Problems.

We completed our coding in the summer of 2005, and did not code any postmarketing event after August 1, 2005. The first version of this paper was submitted to the *New England Journal of Medicine* in January 2006. Except for correction of known coding errors, we did not update the data after August 1, 2005, or after the first submission of this manuscript to the journal.

Approval Times and Pre-Deadline Approval Status: We coded data from FDA approval letters and new drug applications on the FDA web site. To get a measure of the approval time for a drug, we divided the approval time in days by 30 to produce an "approval time" measure denominated in months. Our original coding was based on dividing the number of days of approval by 30 (in separate data that was sent to us by former FDA official Ed Hass, a similar measure with similar denominator was used). This produced a number of drugs that had met the review deadlines but had numerical review times that appeared to put them past the deadlines. For instance, L-Glutamine (Nutrestore; NDA # 21667) was submitted on August 8, 2003 and approved on June 10, 2004, which generates a review time of 10.23 months – so we rounded down to produce the patterns Figure 1.

In addition, we used a rounding-down rule. If a drug was approved within two weeks of the deadline for days, we coded it as a pre-deadline approval. Our intuition here is that if a drug went just a few days (or a week or two) past the deadline but was approved, there was likely to have been pressure on the FDA to approve it as close to the

deadline as possible. Significantly, we re-analyzed our results without the rounding-down rule and they were stronger. Evidence of this robustness of our results to the dropping of the two-week round-down rule is presented below. When everything else is kept the same, our observed associations are <u>stronger</u> (larger ORs, smaller p-values) when the FDA's "deadline approval" measure is used instead of ours.

Withdrawals: For safety-based withdrawals, we coded from Pharmaprojects. We coded a drug as withdrawn for safety-based reasons if it was permanently withdrawn from the U.S. market or from two or more major foreign markets (such as the E.U., Australia and Canada, for tolcapone and the E.U., Australia and other countries for alatrofloxacin and trovafloxacin). In every case of a coded withdrawal, our withdrawn drugs were removed by either the U.S. or the E.U., or both. We did not code alosetron as a withdrawal because it was not a permanent removal from the U.S. market, but we also checked to see whether a recoding of alosetron as a withdrawal produced substantively different results.

In July 2005, we received a paper from Ernst Berndt that had been recently published in *Nature Reviews Drug Discovery*. This paper had an embedded Table entitled "Safety-Based Withdrawals According to CDER" (this is reproduced as Table 3A below). We used this list to check our data to make sure that we had identified all U.S. withdrawals in our sample. Importantly, *this list did not include levomethadyl acetate (Orlaam), even though the FDA's April 2008 posted data states that this drug was withdrawn in 2003*. This list also identified alosetron as a drug that was returned to market, and that was not permanently withdrawn.

Black-Box Warnings: There is no unified database of black-box warnings added to drugs postmarket. The closest is Lasser's article in 2002, which stops coding in 1999. We used this Lasser, which was static, and KUMC. None of the black-box warnings added to drugs after August 1, 2005 were coded. (Some drugs receive multiple black-box warnings over time, and so it is possible that drugs that received an additional black-box warning after August 1, 2005 also received a postmarket boxed warning before August 1, 2005). In addition, it is possible that some of the drugs that had black-box warnings attached between January and July 2005 were not yet listed in the KUMC database.

We cannot verify this possibility as we do not have the entire KUMC database as it was posted in August of 2005.

<u>Dosage-Form Discontinuations</u>: We relied on a download from the Drugs@FDA database, which lists specific code for dosage-form discontinuations (because they require an official amendment to the drug's NDA). We do not have the exact date of this download, but it would have occurred before August 2005.

# 1.2 Specific Coding Decisions for Withdrawn Drugs

The list of safety-based withdrawals used in our article is as follows (see Section 3 for approval time data and just-before-deadline classifications for these drugs).

NDA Number	Generic Name
20760	Alatrofloxacin Mesylate
20535	Bromfenac Sodium
20740	Cerivastatin Sodium
20695	Grepafloxacin Hydrochloride
20689	Mibefradil Dihydrochloride
20984	Rapacuronium bromide
21042	Rofecoxib
20697	Tolcapone
20720	Troglitazone
20759	Trovafloxacin Mesylate
21341	Valdecoxib

The first major difference between our data and the FDA's is that we have coded alatrofloxacin mesylate and trovafloxacin mesylate as safety-based withdrawals. Numerous sources demonstrate that alatrofloxacin and trovafloxacin were withdrawn separately in the European Union, the United Kingdom, and Australia. For instance, an article in the Australian journal *Hospital Pharmacy* in 2002 states clearly that the drugs were withdrawn in Europe and Australia; the first page of the article is embedded in this document in the next page; for a URL see

http://www.factsandcomparisons.com/assets/hospitalpharm/mar2002\_newsletter.pdf (accessed April 14, 2008). The web site of the Australian National Prescribing Service also states that the drug is withdrawn there

(<a href="http://www.nps.org.au/site.php?content=/html/news.php&news=/resources/NPS\_News/news19">http://www.nps.org.au/site.php?content=/html/news.php&news=/resources/NPS\_News/news19</a>) (accessed April 13, 2008). Indeed, while we do not regard it as an authoritative source, it is interesting that Wikipedia lists the drug as withdrawn from the U.S, market (<a href="http://en.wikipedia.org/wiki/Alatrofloxacin">http://en.wikipedia.org/wiki/Alatrofloxacin</a> [accessed April 16, 2008]).

Figure 1A – Article from Hospital Pharmacy (Australia) stating withdrawal of alatrofloxacin and trovafloxacin in Australia and E.U.

Hospital Pharmacy Volume 37, Number 3, pp 326–330 2002 Facts and Comparisons

# From Your Newsletter

# Beware: Antibiotic-Induced Hepatotoxicity is Rare but Deadly

Carol Simmons, MPS\*

Hospital Pharmacy welcomes contributions to this column, in which articles oriqinally published in pharmacy department newsletters are reprinted. Material is selected because of its educational value or because it typifies the type of information that is interesting to pharmacy newsletter readers. If you would like us to consider your newspaper material for publication, mall a copy and a computer disk containing the document to: Hospital Pharmacy, Facts and Compar-Isons, 111 Westport Plaza, Sulte 300, St. Louis, MO 63146.

metabolism for virtually all ease is therefore a potential complication of nearly every medication. The majority of hepatic adverse drug reactions are idiosyncratic, rare, and unpredictable. They may be assoclated with serious morbidity and mortailty.1 The hepatotoxic potential of new MECHANISM OF drugs is often detected only after their HEPATOTOXICITY introduction to the market. They are often identified via spontaneous report- ic reactions are believed to occur via an awareness of risk factors, and reporting ing systems that have limitations.3 A Idiosyncratic immunological mechanism and recording adverse drug reactions. much published, recent example is caused by the drug itself or a metabotrovafloxacin/alatrofloxacin.

introduced worldwide in mid-to-late caused by precipitation of ceftrlax-1998. A spate of reports of hepatotoxic- one-calcium complex in the billary sysity resulted in severe restrictions on its tem; and fusidic acid-associated hepause in the US and Australia, and a sus- totoxicity, which is due to decreased bile pension of its license in Europe. It has acid excretion.4 since been withdrawn from the market In both Australia and Europe.1

Because antibiotics are the most induced liver injury is of significant clini-

he liver is the major site of callimportance. The antibiotics most frequently reported as causes of hepatodrugs. Drug-Induced liver dis- toxicity in Australia are flucioxacillin, amoxicillin/clavulanic acid, thromycin, and nitrofurantoin.1 Antibiotics in common use that have a definite causal association with acute liver. Ilver injury with prompt withdrawal of the injury are detailed in Table 1.

Trovafloxacin/alatrofloxacin was drugs listed in Table 1 are choleithiasis

#### DIAGNOSIS

The signs and symptoms of drugcommonly prescribed drugs, antibiotic- induced liver disease mimic liver disease of other etiologies. Diagnosis is

based on clinical suspicion, careful drug history, consideration of temporal relationships between drug administration and liver disease, and exclusion of other causes. It is important to remember that some reactions manifest only after discontinuation of the drug. For example, 50% of hepatotoxic reactions due to flucioxacillin and 75% of those due to amoxicilin/clavulanic acid can occur days to weeks after stopping the drug. Other hepatotoxic reactions can occur after many years of continued treatment (eq. nitrofurantoin).

#### PREVENTION AND MANAGEMENT

Prevention and early detection of offending drug are crucial. Reactions can be minimized by avoiding over use of drugs and polypharmacy, ensuring appropriate indications, keeping the Most antibiotic-induced hepatotox- duration of therapy as short as possible,

Patients should be warned to lite.12 Notable exceptions among the report unexplained nausea, malaise, abdominal pain, lethargy, or fever. Liver function tests (LFTs) should be used to follow up these symptoms. Routine screening of LFTs is not usually recommended, because the onset of liver dysfunction is usually rapid and there is uncertainty about the level of abnormal LFTs at which the drugs should be discontinued. Levels four to five times the upper limit of normal indicate a definite problem, particularly when the elevation of transaminases is involved.

> Suspected drugs should be withdrawn as early as possible. Sympto-

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The FDA has alatrofloxacin and trovafloxacin coded as drugs with a black-box warning. Note that there is no debate about whether these drugs are separate NMEs even though they often share the same trade name. They are separate molecules (separate generic names), separate NMEs, separate NMEs as listed on Drugs@FDA, and separate NMEs as listed in data. (They are also listed as separate NMEs on FDA web post of April 2008; see <a href="http://www.fda.gov/oc/pdufa/FDADrugAppSafetyData\_files/NMESafetySumm.html">http://www.fda.gov/oc/pdufa/FDADrugAppSafetyData\_files/NMESafetySumm.html</a> [accessed April 16, 2008].)

The other NME in our withdrawn sample that is not in the FDA's sample is Tolcapone, which was withdrawn in the EU, UK, Canada, and Australia. Tolcapone was not withdrawn in the U.S.

The FDA has coded three drugs as withdrawals that are not in our data. The last, tegaserod maleate (Zelnorm) was withdrawn just over a year ago on March 30, 2007, after our paper was already deep in the review process. The second discrepancy is alosetron. We did not code this as a safety-based withdrawal because the drug was withdrawn then reintroduced. Alosetron was, moreover, reintroduced in 2002, 2 years after being withdrawn, and before the end of our data collection. [An article in *Nature Genetics* remarks that it was withdrawn only in the U.S.] We believe this is a plausible and defensible interpretation, and we have checked our results repeatedly (see below).

We did not have Levomethadyl Acetate in our original dataset, in part because it not listed as a withdrawal by FDA. When the removal of levomethadyl acetate was first announced by the FDA in 2003, the drug's removal was announced not as a safety-based withdrawal but as a marketing discontinuation (see Figure 1B, next page). For a similar interpretation of levomethadyl's U.S. removal as a marketing discontinuation induced by the availability of buprenorphine, see Jim Rosack, "Med Check," *Psychiatric News* 38 (19) (October 3, 2003) <a href="http://pn.psychiatryonline.org/cgi/content/full/38/19/36">http://pn.psychiatryonline.org/cgi/content/full/38/19/36</a> [accessed April 16, 2008]. As we note below, levomethadyl acetate was not listed as a safety-based withdrawal in data provided to Ernst Berndt and co-authors in an article they published in July 2005 in Nature Reviews Drug Discovery. The results of this paper were later presented at an FDA-sponsored public meeting on the re-authorization of PDUFA.

## Figure 1B: FDA Announcement of 2003 Does Not List Levomethadyl as a Safety-Based Withdrawal, but as a Discontinuation.

(from <a href="http://www.fda.gov/CDER/drug/shortages/orlaam.htm">http://www.fda.gov/CDER/drug/shortages/orlaam.htm</a> [accessed April 12, 2008]). [See also Figure 2H below.]

Drug Shortage: Drug to be Discontinued Letter from Roxane

August 23, 2003

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#### PRODUCT DISCONTINUATION NOTICE

ORLAAM® (Levomethadyl hydrochloride acetate) Oral Solution, 10 mg/mL, CII

NDC 0054-3649-63

Dear Healthcare Professional:

Rosane Laboratories, Inc. is discontinuing the sale and distribution of ORLAAM® (Levomethadyl hydrochloride acetate) Oral Solution, 10 mg/mL after the current inventory is depleted. We estimate that this will occur early in the first quarter of 2004. Since the introduction of ORLAAM in 1995, Rosane Laboratories has received increasing reports of severe cardiac-related adverse events, including QT interval prolongation (15), Torsades de Pointes (8) and cardiac arrest (6). Other cardiac-related adverse events have also been reported, including arrhythmias, syncope, and angina. These events led to the removal of ORLAAM from the European market in March 2001, and extensive changes (including additional warnings & contraindications) were made to the US package insert in April 2001 (Dear Healthcare Professional letter dated April 11, 2001). Since these changes, the use of ORLAAM has decreased dramatically over the last two years. While there may be a very small number of patients who may benefit from ORLAAM, it is our belief that the risks of continued distribution and use in the face of less toxic treatment alternatives no longer outweigh the overall benefits.

ORLAAM is a synthetic opioid agonist solution indicated for the management of opiate dependence, reserved for the treatment of opiate-addicted patients who fail to show acceptable response to other adequate treatments for opiate addiction. Other first-line treatment options are available for the management of opiate dependence, including methadone and the recently FDA-approved buprenorphine. Methadone hydrochloride is available as an oral solution and a dispersible tablet, both which will continue to be manufactured by Roxane Laboratories and distributed by Cebert Pharmaceuticals. Buprenorphine hydrochloride is available in two sublingual formulations: one containing naloxone hydrochloride (Suboxone®, Reckitt Benckiser Pharmaceuticals) and one without naloxone (Subutew®, Reckitt Benckiser Pharmaceuticals). With these first-line agents available for the treatment of opiate addiction, it is our hope that existing patients can be converted to alternate therapies with minimal disruption to them and the centers that treat them.

Due to the forecasted unavailability shortly after the beginning of 2004, no new patients should be initiated on ORLAAM therapy. For existing ORLAAM patients, it is extremely important for healthcare providers to transfer patients to alternative treatments as soon as possible prior to the product's unavailability. To make sure this transition occurs with minimal disruption to all patients involved, we will reserve the right to limit purchase quantities based upon historical annual volumes. Careful consideration should be given to the appropriate conversion regimens. The information on the next page is from the current package insert for ORLAAM:

#### Transfer from ORLAAM® to Methadone:

Patients maintained on ORLAAM may be transferred directly to methodone. Because of the difference between the two compounds' metabolites and their pharmacological half-lives, it is recommended that methodone be started on a daily dose at 80% of the ORLAAM dose being replaced; the initial methodone dose must be given no sooner than 48 hours after the last ORLAAM dose. Subsequent increases or decreases of 5 to 10 mg in the daily methodone dose may be given to control symptoms of withdrawal or, less likely, symptoms of excessive sedation, in accordance with clinical observations.

For further information about ORLAAM, please contact our Technical Information Department at 1-800-962-8364 or our Customer Service Department at 1-800-520-1631.

Respectfully.

Michael J. Schobelock, PharmD Associate Director Medical Affairs Department Rosane Laboratories, Inc

# 2. General Notes on the Reliability of the FDA's and Nardinelli et al's Postmarket Safety Data for Safety-Based Withdrawals

It is important <u>not</u> to regard either the Drugs@FDA database or the MedWatch system data as authoritative for purposes of drug safety. The database contains some very important omissions. For instance, Posicor (mibefradil hydrochloride) and Duract [bromfenac sodium) were withdrawn for safety reasons, yet they are not listed on the Drugs@FDA webpage and are absent from the underlying database maintained by FDA. (As the following six Figures clearly show, searching for these two drugs by generic name, trade name and NDA number fails to return a record with the original NME.) These are two of the most important safety-based withdrawals of the past several decades; their absence from the FDA's own data is troubling.

Below, we attach results of a search of the Drugs@FDA database, using both the U.S. trade names for Duract and Posicor, and separate searches using the respective generic names and the respective NDA numbers. The searches were conducted on April 7, 2008.

Figure 2A – Duract Not Locatable on Drugs@FDA page using Trade Name search.



#### Search Results for 'duract'

Your search term did not return any results. Modify Your Search

- . Spelling or Formatting Problems
  - o If you are not sure of the spelling, try "Browse by Drug Name."
  - · Try putting in part of the Drug Name or Active Ingredient. You must enter at least three letters or numbers.
  - o A drug name containing a combination of letters, punctuation, and spaces has to be formatted exactly as it appears in the database. Examples: H.P. ACTHAR gel or X-TROZINE L.A. Try entering just part of the name, such as "acthar" or "trozine."
- . For Application Number searches, you must enter five digits for NDAs and ANDAs, and six digits for BLAs.

More about searching

. Dietary supplements, most biologic products, animal drugs, and certain drug products are not in Drugs@FDA. More information about the contents of Drugs@FDA

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Figure 2B – Duract Not Locatable on Drugs@FDA page using NDA Number search





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### Start Over

# Search Results for Application No. '020535'

## Your search term did not return any results.

Modify Your Search

- . Spelling or Formatting Problems
  - o If you are not sure of the spelling, try "Browse by Drug Name."
  - o Try putting in part of the Drug Name or Active Ingredient. You must enter at least three letters or numbers.
  - o A drug name containing a combination of letters, punctuation, and spaces has to be formatted exactly as it appears in the database. Examples: H.P. ACTHAR gel or X-TROZINE L.A. Try entering just part of the name, such as "acthar" or "trozine."
- For Application Number searches, you must enter five digits for NDAs and ANDAs, and six digits for BLAs.

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FDA/Center for Drug Evaluation and Research Office of Training and Communications Division of Information Services Update Frequency: Daily

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Figure 2C – Drugs@FDA Does not Have Duract (Bromfenac Sodium) NME, search using generic name [html download of April 7, 2008]. [Note that Duract, NDA # 20689 – the original NME for bromfenac sodium – does not appear but a <u>new formulation</u> of the molecule does appear. We are entirely unable to access the NME for bromfenac sodium at this time.]



## **Drug Details**

Drug Name(s) XIBROM (Brand Name Drug)

FDA Application No. (NDA) 021664

Active Ingredient(s) BROMFENAC SODIUM

Company ISTA PHARMS
Original Approval or Tentative Approval Date March 24, 2005
Chemical Type 3 New formulation

Review Classification S Standard review drug

There are no Therapeutic Equivalents

 Approval History, Letters, Reviews, and Related Documents Label Information

### Products on Application (NDA) #021664

Click on a column header to re-sort the table:

Drug Name	Active Ingredients	Strength	Dosage Form/Route	Marketing Status	RLD	TE Code
XIBROM	BROMFENAC SODIUM	0.09%	SOLUTION/DROPS; OPHTHALMIC	Prescription	Yes	None

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Figure 2D – Posicor Not Locatable on Drugs@FDA page using Trade Name search.



### Search Results for 'Posicor'

### Your search term did not return any results.

Modify Your Search

- . Spelling or Formatting Problems
  - o If you are not sure of the spelling, try "Browse by Drug Name."
  - Try putting in part of the Drug Name or Active Ingredient. You must enter at least three letters or numbers.
  - A drug name containing a combination of letters, punctuation, and spaces has to be formatted exactly as it appears in the database. Examples: H.P. ACTHAR gel or X-TROZINE L.A. Try entering just part of the name, such as "acthar" or "trozine."
- For Application Number searches, you must enter five digits for NDAs and ANDAs, and six digits for BLAs.

#### More about searching

 Dietary supplements, most biologic products, animal drugs, and certain drug products are not in Drugs@FDA. More information about the contents of Drugs@FDA

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# Figure 2E –Posicor (mibefradil hydrochloride) Not Locatable on Drugs@FDA page using NDA Number search.



## Search Results for Application No. '020689'

Your search term did not return any results.

Modify Your Search

- . Spelling or Formatting Problems
  - o If you are not sure of the spelling, try "Browse by Drug Name."
  - Try putting in part of the Drug Name or Active Ingredient. You must enter at least three letters or numbers.
  - A drug name containing a combination of letters, punctuation, and spaces has to be formatted exactly as it appears in the database. Examples: H.P. ACTHAR gel or X-TROZINE L.A. Try entering just part of the name, such as "acthar" or "trozine."
- For Application Number searches, you must enter five digits for NDAs and ANDAs, and six digits for BLAs.

More about searching

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# Figure 2F – Drugs@FDA Does Not Have Posicor (mibefradil), generic name search [download of April 7, 2008]



FDA Approved Drug Products
Start Over

### Search Results for 'mibefradil'

Your search term did not return any results.

Modify Your Search

- . Spelling or Formatting Problems
  - o If you are not sure of the spelling, try "Browse by Drug Name."
  - Try putting in part of the Drug Name or Active Ingredient. You must enter at least three letters or numbers.
  - A drug name containing a combination of letters, punctuation, and spaces has to be formatted exactly as it appears in the database. Examples: H.P. ACTHAR gel or X-TROZINE L.A. Try entering just part of the name, such as "acthar" or "trozine."
- For Application Number searches, you must enter five digits for NDAs and ANDAs, and six digits for BLAs.

#### More about searching

 Dietary supplements, most biologic products, animal drugs, and certain drug products are not in Drugs@FDA. More information about the contents of Drugs@FDA

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# Figure 2G – Drugs@FDA Does Have Vioxx (rofecoxib), generic name search [download of April 7, 2008]



#### Overview

Drug Name VIOXX

Active Ingredient(s) • ROFECOXIB

Form(s) and • SUSPENSION; ORAL: 12.5MG/5ML; 25MG/5ML Strength(s) Available • TABLET; ORAL: 0.0; 12.5MG; 25MG; 50MG

Details about drugs are organized by FDA Application Number (NDA or ANDA or BLA).

#### Click on a drug name or application number to view drug details:

Click on a column header to re-sort the table:

Drug Name and FDA Application Number	Dosage Form/Route	<u>Strength</u>	Marketing Status	Company
VIOXX (NDA # 021042)	TABLET; ORAL	Multiple Strengths	Discontinued	MERCK
VIOXX (NDA # 021052)	SUSPENSION; ORAL	Multiple Strengths	Discontinued	MERCK
VIOXX (NDA # 021647)	TABLET; ORAL	0.0	Prescription	MERCK

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# 2.1 – Significant Differences between April 2008 FDA Data and Other FDA Published and Distributed Data

The FDA's data posted on April 8, 2008 also differ materially from data that FDA published elsewhere, that were provided to other authors, and that were presented at FDA public meetings in November 2005. We note two crucial differences here.

First, the FDA data of April 8, 2008, include Orlaam (Levomethadyl Acetate Hydrochloride), and state that this drug was withdrawn on August 23, 2003. However, in data provided to Ernst Berndt, and reported in a *Nature Reviews Drug Discovery* article in July 2005.<sup>2</sup> Orlaam was not included in NME safety withdrawals "according to CDER" (see Figure on next page).

The Berndt list was based on a document published April 2004. Moreover, as late as November 2005, a presentation based on these data was given at an FDA public meeting on PDUFA re-authorization. [For the Power Point presentation at the docket, see <a href="https://www.fda.gov/ohrms/dockets/dockets/05n0410/05n-0410-ts00006-Berndt.ppt">www.fda.gov/ohrms/dockets/dockets/05n0410/05n-0410-ts00006-Berndt.ppt</a> (downloaded April 9, 2008).]

For a transcript of this meeting, see <a href="http://www.fda.gov/CBER/minutes/pdufa111405t.pdf">http://www.fda.gov/CBER/minutes/pdufa111405t.pdf</a> (downloaded April 9, 2008). Neither "Orlaam" nor "Lotronex" (nor "Levomethadyl acetate" nor "Alosetron Hydrochloride") is mentioned at this meeting. For reference, Vioxx and Bextra <a href="mailto:are-mentioned">are-mentioned</a> at this meeting.

Slide 17 of the Berndt presentation lists the source for the FDA's withdrawal data as "Source: "Center for Drug Evaluation and Research 2003 Report to the Nation: Improving Public Health through Human Drugs." U.S. Department of Health and Human Services, Food and Drug Administration, April 23, 2004." See

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<sup>&</sup>lt;sup>2</sup> Ernst R. Berndt, Adrian H. B. Gottschalk, Tomas J. Philipson and Matthew W. Strobeck, "Industry Funding of the FDA: Effects of PDUFA on Approval Times and Withdrawal Rates," *Nature Reviews Drug Discovery* 4 (July 2005) 545-554.

www.fda.gov/ohrms/dockets/dockets/05n0410/05n-0410-ts00006-**Berndt**.ppt (downloaded April 9, 2008), slide 17.

Second, as Figure 2H also makes clear, Lotronex (alosetron) was highlighted as a withdrawal that was reintroduced. (See note to Table 3 of Berndt., et al 2005; in Figure 2H, next page). Again, neither Lotronex nor its generic name was mentioned at the November 2005 FDA public meeting.

# 2.1 – Significant Differences between April 2008 FDA Data and Other FDA Published and Distributed Data (cont.)

Figure 2H. Levomethadyl Acetate (Orlaam) Missing From CDER/FDA Data Provided to Other Authors. [Table 3 from Berndt et al (2005).]

Table 3   <b>NME sa</b>	fety withdrawals a	ccording to CDE	R
Drug name	Submission date	Year approved	Year withdrawn
Fenfluramine	03 Mar 1967	1973	1997
Azaribine	17 Dec 1969	1975	1976
Ticrynafen	15 Nov 1977	1979	1980
Zomepirac	10 Nov 1978	1980	1983
Benoxaprofen	23 Jan 1980	1982	1982
Nomifensine	16 Mar 1979	1984	1986
Suprofen	10 Nov 1978	1985	1987
Terfenadine	01 Mar 1983	1985	1998
Encainide	13 Jan 1984	1986	1991
Astemizole	25 Feb 1985	1988	1999
Rosequinan	01 Oct 1990	1992	1993
Temafloxacin	30 Nov 1989	1992	1992
Cisapride	29 Aug 1991	1993	2000
Dexfenfluramine*	Unavailable	1996	1997
Bromfenac	30 Dec 1994	1997	1998
Cerivastatin	26 Jul 1996	1997	2001
Grepafloxin	08 Nov 1996	1997	1999
Mibefradil	11 Mar 1996	1997	1998
Troglitazone	01 Aug 1996	1997	2000
Rapacuronium	25 Jun 1998	1999	2001
Rofecoxib‡	23 Nov 1998	1999	2004
Alosetron <sup>6</sup>	30 Jun 1999	2000	2000

\*Not considered a new molecular entity (NME). \*Not in the Center for Drug Research and Evaluation's report. \*Returned to the market in 2002 with restrictions. Nine NMEs were withdrawn between 1980 and 1992 and include zomepirac, benoxaprofen, nomifensine, suprofen, terlenadine, eincanide, astemizole, flosequinan and temafloxacin. For the period 1993–2004, CDER reported on nine withdrawals in their report — ten when Vicox (rofecoxib) is included. However, one of the withdrawals is not considered an NME and a second drug, alsestron, that was withdrawn from market was reintroduced with distribution restrictions. Data adapted from REE 25.

Notes: (1) Levomethadyl Acetate (Orlaam) not included; (2) Alosetron (Lotronex) coded as reintroduced.

See also Figure 1B above: levomethadyl was not announced as a safety-based withdrawal by the FDA in 2003.

# 3. Data on Safety-Based Withdrawals and Re-Analysis of Data (just-before-deadline classification of Carpenter-Zucker-Avorn)

The following is the list of safety-based withdrawals as analyzed in the NEJM article.

NDA				
Number	Generic Name	Submission Date	Approval Date	JB-Deadline Approval
20760	Alatrofloxacin Mesylate	12/30/1996	12/18/1997	1
20535	Bromfenac Sodium	12/30/1994	7/15/1997	0
20740	Cerivastatin Sodium Grepafloxacin	6/26/1996	6/26/1997	1
20695	Hydrochloride	11/8/1996	11/6/1997	1
20689	Mibefradil Dihydrochloride	3/11/1996	6/20/1997	0
20984	Rapacuronium bromide	6/25/1998	8/18/1999	0
21042	Rofecoxib	11/23/1998	5/20/1999	1
20697	Tolcapone	6/3/1996	1/29/1998	0
20720	Troglitazone	8/1/1996	1/29/1997	1
20759	Trovafloxacin Mesylate	12/30/1996	12/18/1997	1
21341	Valdecoxib	1/16/2001	11/16/2001	1

Nardinelli et al claim that only five of the withdrawals in our study qualify as just-before-deadline approvals. As the previous table makes clear, this is incorrect. Alatrofloxacin, cerivastatin sodium, grepafloxacin, rofecoxib, troglitazone, trovafloxacin and valdecoxib were all approved in the final month of their PDUFA deadlines.

# 3A. Safety-Based Withdrawals -- Adding Alosetron to Carpenter/Zucker/Avorn Original Data

We acknowledge that there may be other possible codings of what counts as a safety-based withdrawal, and the FDA clearly uses alternative codes. One possibility emerges in the drug alosetron (Lotronex). Like another popular drug (Wellbutrin), alosetron was withdrawn in the United States but then returned to the market. We considered only permanent withdrawals, and so did not code alosetron as withdrawn in our original analysis.

If we use our original data of 313 drugs and code alosetron as a withdrawal, when the year of submission is controlled for in an exact logistic regression the odds ratio is well above one and is significant with an exact p-value of 0.044. It is only in a simple 2x2 analysis (not used in our study) that the p-value rises above 0.05. We attach the results of this exact logistic regression (as well as random-effects logistic regressions with more control variables) below.

Moreover [see Section 4 below] if we use the FDA's dataset of 314 NMES and the FDA's deadline approval classification, there is a larger and more statistically significant association between deadline approvals and our measure with alosetron added.

# 3A. Safety-Based Withdrawals -- Adding Alosetron to Carpenter/Zucker/Avorn Original Data (cont.)

In a simple two-way cross-tabulation using our original 313 data points, addition of alosetron as a withdrawal generates a tabulation in which the exact p-value rises just above 0.05, but this is not how we calculated our results in Figure 2, as this simple approach does not control for year of submission (as in the exact logistic regression) or any of several other important variables that we controlled for (in random-effects logistic regressions; see the on-line Appendix).

	combowit_addlotr = 20719 & ndanum2	-	• •	ar > 1992
ddlotronex	predead   0 +	1	Total	
	,   211   5	·	301 12	
Total	216	97	313	
	Fisher's exact =		0.053	
1-sided	Fisher's exact =		0.043	

However, as the next page shows, the exact p-value falls below 0.05 in an exact logistic regression that controls for submission year. In addition, mixed-effects logistic regressions that add a random effect for the primary indication of the NME, and control for epidemiological characteristics of the primary indication, yields similar results.

# **3A.** Safety-Based Withdrawals -- Adding Alosetron to Carpenter/Zucker/Avorn Original Data (cont)

**Exact Logistic Regression Results** 

# **Binary Regression**

regression (type=logit, model(sbw\_addlot = subyrctr predead), estimate(subyrctr predead), method=exact, mle=firth, profile=yes, outtype=odds):

#### **Basic Information**

Data file combowit-check-20080329.csv

**Model** sbw\_addlot(Response = 1)=%Const+subyrctr+predead

Link type Logit

Weight variable < Not Specified>
Stratum variable < Unstratified>
Analysis type Estimate :: Exact

Number of terms in model3Number of term(s) dropped0Number of observations in analysis313Number of records rejected0Number of groups24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	28.52	21	0.1259
Likelihood Ratio	338.4	3	1.394e-009

#### **Parameter Estimates**

	Point Estima	ate	Confidence In	terval and P-\	Value for Odd	s Ratio
				95	%CI	2*1-sided
Type	Odds Ratio	SE(Odds)	Type	Lower	Upper	P-Value
PMLE	0.3601	NA	Asymptotic	0.01117	11.6	0.5643
			Asymptotic(Profile)	0.01065	15.05	
PMLE	0.8559	NA	Asymptotic	0.6944	1.055	0.1449
			Asymptotic(Profile)	0.6761	1.048	
CMLE	0.8449	NA	Exact	0.6612	1.05	0.1387
PMLE	4.007	NA	Asymptotic	1.253	12.81	0.01925
			Asymptotic(Profile)	1.248	13.75	
CMLE	4.117	NA	Exact	1.038	17.85	0.04333
	PMLE PMLE CMLE PMLE	Type         Odds Ratio           PMLE         0.3601           PMLE         0.8559           CMLE         0.8449           PMLE         4.007	PMLE         0.3601         NA           PMLE         0.8559         NA           CMLE         0.8449         NA           PMLE         4.007         NA	Type         Odds Ratio         SE(Odds)         Type           PMLE         0.3601         NA         Asymptotic           Asymptotic (Profile)         Asymptotic           PMLE         0.8559         NA         Asymptotic (Profile)           CMLE         0.8449         NA         Exact           PMLE         4.007         NA         Asymptotic (Profile)           Asymptotic (Profile)         Asymptotic (Profile)	PMLE	Type

PMLE: Penalized MLE for bias correction (Firth's method).

Analysis Time = 00:00:00

# 3A. Safety-Based Withdrawals -- Adding Alosetron to Carpenter/Zucker/Avorn Original Data (cont)

# NOW ADD LOTRONEX (ALOSETRON) TO THE SAFETY-BASED WITHDRAWALS LIST, USING ORIGINAL 313 NMES

We first add control for the submission year of the NME to capture time trends, and add 134 random effects for primary indication of NME. We then add covariates for total hospitalizations associated with NME's primary indication in 1997 (= 0 when there are no HCUP hospitalizations associated with ) and average days of hospitalization, per hospitalization (= 0 when there are no HCUP hospitalizations associated with the NME's primary indication).

The estimates for the recoded predeadline measure are highlighted in yellow.

Random-effects	logistic rec	gression		Number o	of obs =	= 313
Group variable	(i): discode	9		Number o	of groups =	= 134
Random effects	u i ~ Gaussi	ian		Obs per	group: min =	= 1
	_			_	avg =	2.3
					_	16
				Wald chi	i2(4) =	= 6.65
Log likelihood	= -47.5283	33			chi2 =	
sbwlot	OR	Std. Err.	Z	P> z	[95% Conf.	. Interval]
+	OR 					
subyrctr	.8680894	.095934	-1.28	0.201	.6990315	1.078033
subyrctr   pred0410_r~e	.8680894	.095934 2.475958	-1.28 2.27	0.201 0.023	.6990315 1.208726	1.078033 13.43597
subyrctr   pred0410_r~e   hhospdisc	.8680894 4.029938	.095934 2.475958 1.59e-06	-1.28 2.27 0.78	0.201 0.023 0.436	.6990315 1.208726 .9999981	1.078033 13.43597 1.000004
subyrctr   pred0410_r~e   hhospdisc   hhosleng	.8680894 4.029938 1.000001	.095934 2.475958 1.59e-06 .0680306	-1.28 2.27 0.78	0.201 0.023 0.436	.6990315 1.208726 .9999981 .8607752	1.078033 13.43597 1.000004
subyrctr   pred0410_r~e   hhospdisc   hhosleng	.8680894 4.029938 1.000001 .9854858	.095934 2.475958 1.59e-06 .0680306	-1.28 2.27 0.78	0.201 0.023 0.436	.6990315 1.208726 .9999981 .8607752	1.078033 13.43597 1.000004 1.128265

# **3B.** Safety-Based Withdrawals -- Adding Levomethadyl Acetate (Orlaam) to Carpenter/Zucker/Avorn Original Data

Alternatively, one could add Levomethadyl Acetate (Orlaam), which was only discontinued in the U.S., not withdrawn. The effect of adding Levomethadyl Acetate is the same as that of adding Alosetron. In a simple unadjusted two-way cross-tabulation, the exact-p value rises above 0.05.

	combowit_addorl 20719 & ndanum2 ~	_	-	> 1992 &
combowit_a ddorlaam	0	1		
0		90 7	+   301   12	
Total	216	97	313	
1-sided	<pre>Fisher's exact = Fisher's exact =</pre>		0.053 0.043	

Once again, however, running an exact logistic regression on this data yields an oddsratio that is above 1.0 with an exact two-tailed probability of less than 0.05. The year of submission is a statistically significant covariate, which points to the importance of controlling for this variable.

# 3B. Safety-Based Withdrawals -- Adding Loevomethadyl Acetate to Original Carpenter/Zucker/Avorn Data, then Estimation of Exact Logistic Regression

# **Binary Regression**

regression (type=logit, model(sbw\_addori = subyrctr predead), estimate(subyrctr predead), method=exact, mle=firth, profile=yes,

#### **Basic Information**

Data file combowit-check-20080329.csv

Model sbw\_addorl(Response = 1)=%Const+subyrctr+predead

Link type Logit

Weight variable < Not Specified>
Stratum variable < Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 313
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	25.25	21	0.2365
Likelihood Ratio	340.9	3	1.748e-009

#### **Parameter Estimates**

		Point Estima	ate	Confidence In	terval and P-1	Value for Odd	s Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unner	P-Value
%Const	PMLE	1.267	NA	Asymptotic	0.03222	49.8	0.8996
				Asymptotic(Profile)	0.03277	73.19	
subyrctr	PMLE	0.7907	NA	Asymptotic	0.6291	0.9936	0.04395
				Asymptotic(Profile)	0.607	0.9826	
	CMLE	0.7762	NA	Exact	0.5896	0.9822	0.03306
predead	PMLE	4.514	NA	Asymptotic	1.391	14.65	0.01211
				Asymptotic(Profile)	1.39	15.74	
	CMLE	4.644	NA.	Exact	1.154	20.51	0.02857

PMLE: Penalized MLE for bias correction (Firth's method).

Analysis Time = 00:00:00

# 3B. Safety-Based Withdrawals -- Adding Levomethadyl Acetate to Original Carpenter/Zucker/Avorn Data (cont.), estimating Random-Effects Logistic Regression

The estimates for the recoded predeadline measure are highlighted in yellow.

Random-effects Group variable					of obs = of groups =		
Random effects	u_i ~ Gaussi	an		Obs per		1 2.3 16	
Log likelihood	= -46.42353	9			ii2(4) = chi2 =		
sbworl	OR	Std. Err.	Z	P> z	[95% Conf.	Interval]	
pred0410 r~e	4.384225	2.717897	2.38	0.017	.6332033 1.3008	14.77663	
	1.000001 .9974191				.9999981 .8798467		
/lnsig2u	-2.936907	2.711426			-8.251203	2.37739	
	.2302814		<b></b>		.0161538	3.282794	
Likelihood-rat	io test of rh	.o=0: chibar	2(01) =	0.02	Prob >= chiba	r2 = 0.447	

# 3C. Re-Examination of Safety-Based Withdrawals: A Coding Focusing only on Permanent U.S. Withdrawals

As an alternative coding, we could focus only on those drugs that were permanently removed from the U.S. market for safety-based reasons. This produces a list of the following 11 withdrawals and a cross-tabulation that is identical to Table 1 in the NEJM article.

The 11 safety-based withdrawals in the original analysis are listed here, with their priority rating, date of submission and of approval, approval time in months, and whether or not the drug was a just-before-deadline approval (7 of 11).

NDA Number	Generic Name	Date of Submission	Date of Approval	JB-Deadline Approval
20740	Cerivastatin Sodium	6/26/1996	6/26/1997	1
21341	Valdecoxib	1/16/2001	11/16/2001	1
20535	Bromfenac Sodium	12/30/1994	7/15/1997	0
20315	Levomethadyl acetate	6/21/1993	7/9/1993	0
20689	Mibefradil Dihydrochloride	3/11/1996	6/20/1997	0
20984	Rapacuronium bromide	6/25/1998	8/18/1999	0
20695	Grepafloxacin Hydrochloride	11/8/1996	11/6/1997	1
20720	Troglitazone	8/1/1996	1/29/1997	1
20759	Trovafloxacin Mesylate	12/30/1996	12/18/1997	1
20760	Alatrofloxacin Mesylate	12/30/1996	12/18/1997	1
21042	Rofecoxib	11/23/1998	5/20/1999	1

# 3C. Re-Examination of Safety-Based Withdrawals – Permanent U.S. Removals (cont.)

The two-way cross-tabulation for this variable is identical to that for the original article.

. tabulate <code>sbw\_usremoval</code> predead if(subyear > 1992 & ndanum2  $\sim$ = 20719 & ndanum2  $\sim$ = 20778), exact

sbw_usremo val	predead	1	Total
0	212   4	90	302 11
Total	216	97	313
	Fisher's exact =	=	0.0
1-sided	Fisher's exact =	=	0.0

And as the next page shows, an exact logistic regression produces an odds ratio that is <u>larger</u> than that estimated in the original article, and statistically significant.

### 3C. Re-Examination of Safety-Based Withdrawals (cont.)

## - Exact Logistic Regression with Permanent U.S. Removals

## **Binary Regression**

regression (type=logit, model(sbw\_usremo = subyrctr predead), estimate(subyrctr predead), method=exact, mle=firth, profile=yes, outtype=ordic1:

#### **Basic Information**

Data file combowit-check-20080329.csv

Model sbw\_usremo(Response = 1)=%Const+subyrctr+predead

Link type Logit

Weight variable <Not Specified>
Stratum variable <Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 313
Number of records rejected 0
Number of groups 24

### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	23.09	21	0.3392
Likelihood Ratio	348.4	3	1.492e-009

#### **Parameter Estimates**

		Point Estim	ate	Confidence In	terval and P-1	Value for Odd	ls Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unner	P-Value
%Const	PMLE	1.134	NA.	Asymptotic	0.02431	52.87	0.949
				Asymptotic(Profile)	0.02461	82.07	
subyrctr	PMLE	0.786	NA.	Asymptotic	0.6188	0.9983	0.04841
				Asymptotic(Profile)	0.5941	0.9863	
	CMLE	0.7699	NA	Exact	0.5748	0.9859	0.03665
predead	PMLE	5.601	NA	Asymptotic	1.614	19.44	0.006648
				Asymptotic(Profile)	1.638	21.57	
	CMLE	5.874	NA.	Exact	1.356	30.14	0.015

PMLE: Penalized MLE for bias correction (Firth's method).

Analysis Time = 00:00:00

# 3C. Re-Examination of Safety-Based Withdrawals (cont.)Random-Effects Logistic Regression with Permanent U.S. Removals, Additional **Control Variables**

The estimates for the recoded predeadline measure are highlighted in yellow.

Random-effects	logistic red	gression		Number o	of obs =	313	
Group variable	(i): discode	9		Number (	of groups =	134	
Random effects	u_i ~ Gaussi	ian		Obs per	group: min =	1	
					avg =	2.3	
					max =	16	
				Wald ch	i2(4) =	9.27	
og likelihood	= -42.54742	29		Prob >	chi2 =	0.0547	
sbwusrem	OR	Std. Err.	Z	P> z	[95% Conf.	Interval]	
subyrctr	.8000855	.1013168	-1.76	0.078	.6242329	1.025478	
red0410_r~e	5.574691	3.69197	2.59	0.009	1.522273	20.41498	
hhospdisc	1.000001	1.57e-06	0.87	0.385	.9999983	1.000004	
hhosleng	1.010225	.0638652	0.16	0.872	.8924961		
/lnsig2u	-2.89283	3.527802			-9.807195	4.021536	
sigma u	.2354128	.4152449			.0074198	7.469051	
	0165663	.0574746			.0000167	.9443118	

4. Re-Analysis of Association between Deadline Approvals and Safety-Based Withdrawals, Using FDA NMEs, FDA Deadline Approval Measure, and Adding pre-2007 Withdrawals as Coded by Nardinelli et al.

We now show that the associations described in our study remain large and statistically significant in analyses of the FDA's data of April 2008, with only slight (and defensible) changes to their measures. For reference, we are able to replicate the FDA cross tabulation using their 314 NMEs, their deadline approval classification, and their measure of safety-based withdrawals, as follows:

. tab sbwfda08	3 pred01fda,	exact		
sbwFDA08	pred01fda 0	a 1	Total	
0   1	220 6	83   5	303 11	
Total	226	88	314	
	ner's exact = ner's exact =		0.190 0.165	

# 4.1 Analysis of Original Carpenter-Zucker-Avorn Withdrawal Measure, using FDA NMEs and FDA Deadline Approval Measure

The difference between our analyses and their analyses clearly comes in the coding of safety-based withdrawals. We show here that using our measure, along with the 314 FDA NMEs and the FDA deadline approval measure, *yields larger and stronger results* for our hypothesis than were ported in the article.

To show this, we use Carpenter-Zucker-Avorn SBW measure, with FDA data (notice 314 NMEs) and with FDA pre-deadline approval coding ("pred01fda").

. tab sbwo	cza pred01fda, exa	ct		
sbwCZA	pred01fda   0	1	Total	
0	222	81   7	303 11	
Total	226	88	314	
1-sided	Fisher's exact = Fisher's exact =		0.013 0.013	

Note that when we use the FDA's NMEs and the FDA's deadline approval measure, but keep our withdrawal measure, the exact probability is <u>less</u> than that observed in the original NEJM article (Table 1). Similar strengthening of results is observed in exact logistic regressions, as the next page demonstrates.

# 4.1 Analysis of Original Carpenter-Zucker-Avorn Withdrawal Measure, using FDA NMEs and FDA Deadline Approval Measure (cont.)

An exact logistic regression on the same measure, controlling for year of NME submission, yields an estimated odds ratio and lower 95% confidence interval that are larger than those observed in the original NEJM article.

### **Binary Regression**

regression (type=logit, model(sbwcza = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes,

#### **Basic Information**

Data file CZAFDAcompare-dump20080410.csv

Model sbwCZA(Response = 1)=%Const+subyrctr+pred01fda
Link type Logit

Weight variable Court

Not Specified>

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 10
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	30.39	21	0.08445
Likelihood Ratio	349.1	3	1.766e-009

#### **Parameter Estimates**

		Point Estima	ate	Confidence In	terval and P-V	alue for Odd	s Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Upper	P-Value
%Const	PMLE	0.3899	NA	Asymptotic	0.0114	13.34	0.6012
				Asymptotic(Profile)	0.01096	17.94	
subyrctr	PMLE	0.8396	NA	Asymptotic	0.6793	1.038	0.106
				Asymptotic(Profile)	0.6594	1.03	
	CMLE	0.8275	NA	Exact	0.6434	1.031	0.09745
pred01fda	PMLE	5.615	NA.	Asymptotic	1.669	18.89	0.005312
				Asymptotic(Profile)	1.691	20.9	
	CMLE	5.889	NA	Exact	1.406	29.1	0.01254

PMLE: Penalized MLE for bias correction (Firth's method).

Analysis Time = 00:00:00

# 4.1 Analysis of Original Carpenter-Zucker-Avorn Withdrawal Measure, using FDA NMEs and FDA Deadline Approval Measure (cont.)

We can also estimate a random-effects logistic regression on the same data. In addition to the random effects (grouped by primary indication of the NME), the added variables are epidemiological controls for total hospitalizations of NME primary indication, and the average length of hospitalization, per hospitalization.

. xtlogit sbw	cza subyrctr	pred01fda	hhospdis	c hhosle	ng, re i(di	scoc	de) or
Random-effects	logistic red	gression		Number	of obs	=	314
Group variable	(i): discode	2		Number	of groups	=	135
Dandom officeto	Canaai	22		Oha nan	arama min	_	1
Random effects	u_1 ~ Gaussi	-all		obs per	group: min		2.3
					_		
					max	=	16
				Wald ch	i2(4)	=	9.40
Log likelihood	= -42.67760	) 9			chi2		
sbwcza	OR	Std. Err.	z	P> z	[95% Con:	 f. I	Interval]
+	.8300595	 N979313	-1 58	0 114	.6586939		1 046008
	6.089567						
hhoshdisc I	1.000001	1 586-06	0.84	0.000	9999982		1 000004
	1.001489				.8780024		
IIIO31elig	1.001405	.0072412					1.142544
/lnsig2u	-2.889474	3.575862			-9.898034		4.119086
+	.2358081	4216085			.0070904		7 842385
	.0166211				.0000153		
1110	.0100211	.0304403			.0000133		. 7472240

Again, the observed odds ratio estimate is above that reported in the NEJM article (Figure 2).

# 4.2 Adding Alosetron to Carpenter/Zucker/Avorn Withdrawals Measure, Using FDA NMEs and FDA Deadline Approval Measure

As noted previously, the FDA's data also include alosetron (Lotronex). If we add alosetron to our measure, we have 12 withdrawals, and we observe the following results for a two-way cross-tabulation.

. tab sbwl	lot pred01fda,	exact		
ماميدا م	pred01:		ma + a l	
sbwlot	U -+	1   +	Total	
0	221	81	302	
1	5 -+	7	12	
Total		88	314	
	Fisher's exact	=	0.042	
1-sided	Fisher's exact	=	0.024	

If we estimate an exact logistic regression on the same data, we retrieve the following estimates.

# 4.2 Adding Alosetron to Carpenter/Zucker/Avorn Withdrawals Measure, Using FDA NMEs and FDA Deadline Approval Measure (cont.)

An exact logistic regression estimated on this measure yields the following results.

## **Binary Regression**

regression (type=logit, model(sbwlot = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes,

#### **Basic Information**

Data file CZAFDAcompare-dump20080410.csv

Model sbwlot(Response = 1)=%Const+subyrctr+pred01fda

Link type Log

Weight variable <Not Specified>
Stratum variable <Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 10
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	31.11	21	0.07179
Likelihood Ratio	340.6	3	1.598e-009

#### **Parameter Estimates**

r drumeter Estimates							
		Point Estim	ate	Confidence In	terval and P-	Value for Odd	ls Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unner	P-Value
%Const	PMLE	0.2829	NA	Asymptotic	0.01011	7.917	0.4576
				Asymptotic(Profile)	0.009569	9.873	
subyrctr	PMLE	0.8672	NA.	Asymptotic	0.7117	1.057	0.1578
				Asymptotic(Profile)	0.6953	1.051	
	CMLE	0.8575	NA	Exact	0.6817	1.053	0.1535
pred01fda	PMLE	4,402	NA	Asymptotic	1.401	13.84	0.01119
				Asymptotic(Profile)	1.395	14.81	
	CMLE	4.522	NA.	Exact	1.164	19.18	0.02739

PMLE: Penalized MLE for bias correction (Firth's method).

### Analysis Time = 00:00:00

# 4.2 Adding Alosetron to Carpenter/Zucker/Avorn Withdrawals Measure, Using FDA NMEs and FDA Deadline Approval Measure (cont.)

A random-effects logistic regression yields similar results. In addition to the random effects (grouped by primary indication of the NME), the added variables are epidemiological controls for total hospitalizations of NME primary indication, and the average length of hospitalization, per hospitalization.

xtlogit sby	wlot subyrctr	pred01fda	hhospdis	sc hhosle	ng, re i(di	scoo	de) or
andom-effects	s logistic red	gression			of obs		
roup variable	e (i): discode	e		Number (	of groups	=	135
andom effect:	s u i ~ Gaussi	ian		Obs per	group: min	=	1
	_			-	avg	=	2.3
					max	=	16
				Wald ch	i2(4)	=	7.66
og likelihood	d = -47.05294	43		Prob > 0	chi2	=	0.1049
sbwlot	OR	Std. Err.	z	P> z	[95% Con:	 E. 1	 Interval]
	+						
subyrctr	.8603819	.0935033	-1.38	0.166	.6953214		1.064626
subyrctr pred01fda	+	.0935033 2.854321	-1.38 2.47	0.166 0.013	.6953214 1.373471		1.064626 15.51027
subyrctr pred01fda hhospdisc	+	.0935033 2.854321 1.58e-06	-1.38 2.47 0.76	0.166 0.013 0.448	.6953214 1.373471 .9999981		1.064626 15.51027 1.000004
subyrctr pred01fda hhospdisc hhosleng	.8603819   4.615507   1.000001	.0935033 2.854321 1.58e-06 .0674818	-1.38 2.47 0.76 -0.18	0.166 0.013 0.448 0.859	.6953214 1.373471 .9999981		1.064626 15.51027 1.000004 1.129494
subyrctr pred01fda hhospdisc hhosleng //lnsig2u	.8603819   4.615507   1.000001   .9879707	.0935033 2.854321 1.58e-06 .0674818	-1.38 2.47 0.76 -0.18	0.166 0.013 0.448 0.859	.6953214 1.373471 .9999981 .8641797		1.064626 15.51027 1.000004 1.129494

.

# 4.3 Adding Levomethadyl Acetate to the Carpenter/Zucker/Avorn Withdrawals Measure, Using FDA NMEs and FDA Deadline Approval Measure.

The FDA's data also include levomethadyl acetate (Orlaam), but as we noted earlier, this drug was not identified as a safety-based withdrawal by the FDA for up to two years after the market removal. If we add levomethadyl acetate to our measure, we again have 12 withdrawals, and we observe the following results for a two-way cross-tabulation.

. tab sbwo	orl pred01fda,	exact		
sbworl	pred01:	fda 1   	Total	
0	221   5	81   7	302 12	
Total		88	314	
1-sided	Fisher's exact Fisher's exact		0.042 0.024	

# 4.3 Adding Levomethadyl Acetate to the Carpenter/Zucker/Avorn Withdrawals Measure, Using FDA NMEs and FDA Deadline Approval Measure (cont.)

In an exact logistic regression on the same measure, controlling for submission year of the NME, an OR of 4.9 is observed with a p-value of 0.02.

## **Binary Regression**

regression (type=logit, model(sbworl = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtype=odds).

#### **Basic Information**

Data file CZAFDAcompare-dump20080410.csv

Model sbworl(Response = 1)=%Const+subyrctr+pred01fda

Link type L

Weight variable <Not Specified>
Stratum variable <Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 10
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	28.43	21	0.1283
Likelihood Ratio	342.9	3	1.885e-009

#### **Parameter Estimates**

i didilictei Estillates							
		Point Estim	ate	Confidence In	terval and P-	Value for Odd	ls Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Upper	P-Value
%Const	PMLE	0.8661	NA	Asymptotic	0.02645	28.36	0.9356
				Asymptotic(Profile)	0.0265	39.31	
subyrctr	PMLE	0.8089	NA.	Asymptotic	0.6535	1.001	0.05144
				Asymptotic(Profile)	0.6335	0.9918	
	CMLE	0.7962	NA	Exact	0.6177	0.9919	0.04116
pred01fda	PMLE	4.767	NA	Asymptotic	1.504	15.11	0.007965
				Asymptotic(Profile)	1.501	16.17	
	CMLE	4.899	NA.	Exact	1.252	20.96	0.02024

PMLE: Penalized MLE for bias correction (Firth's method).

Analysis Time = 00:00:00

# 4.3 Adding Levomethadyl Acetate to the Carpenter/Zucker/Avorn Withdrawals Measure, Using FDA NMEs and FDA Deadline Approval Measure (cont.)

In a random-effects logistic regression on the same measure, controlling for submission year of the NME, an OR of 5.05 is observed with a p-value of 0.009. In addition to the random effects (grouped by primary indication of the NME), the added variables are epidemiological controls for total hospitalizations of NME primary indication, and the average length of hospitalization, per hospitalization.

andom-effects l	ogistic reg	ression		Number c	of obs	=	314
roup variable (	i): discode	:		Number c	f groups	=	135
andom effects u	i ~ Gaussi	.an		Obs per	group: min	=	1
•	_			-			2.3
					_		16
				Wald chi	2(4)	=	9.37
og likelihood	= -45.89681	7			hi2		
_	10.0001			1100 / 0	,111LZ	_	0.0324
		Std. Err.					
sbworl	OR	Std. Err.	z 	P> z  0.055	[95% Con.	 f. I	 nterval]  1.005009
sbworl	OR	Std. Err.	z 	P> z  0.055	[95% Con	 f. I	 nterval]  1.005009
sbworl	OR7968601 5.059184	Std. Err. .0943533 3.158301	z -1.92 2.60	P> z  0.055	[95% Con	 f. I	 nterval]  1.005009 17.19732
sbworl   subyrctr   pred01fda	OR	Std. Err. .0943533 3.158301 1.56e-06		P> z  0.055 0.009 0.462	[95% Con .6318215 1.488333 .9999981	 f. I 	nterval] 1.005009 17.19732 1.000004
sbworl   subyrctr   pred01fda   hhospdisc	OR .7968601 5.059184 1.000001 .9997415	Std. Err. .0943533 3.158301 1.56e-06 .0634257		P> z  0.055 0.009 0.462	[95% Con .6318215 1.488333 .9999981	f. I	nterval] 1.005009 17.19732 1.000004 1.132113
sbworl   subyrctr   pred01fda   hhospdisc   hhosleng	OR .7968601 5.059184 1.000001 .9997415	Std. Err0943533 3.158301 1.56e-06 .0634257 3.648302		P> z  0.055 0.009 0.462	[95% Con .6318215 1.488333 .9999981 .8828474	f. I	nterval] 1.005009 17.19732 1.000004 1.132113 4.263154

# 4.4 Adding Alosetron <u>and</u> Levomethadyl Acetate to the Carpenter/Zucker/Avorn Withdrawals Measure, Using FDA NMEs and FDA Deadline Approval Measure

If alosetron <u>and</u> levomethadyl acetate are <u>both</u> added to the withdrawals measure, then it is equivalent to combining our measure of safety-based withdrawals with the FDA's measure, save for the exclusion of Zelnorm, which was withdrawn from the U.S. market in March 2007. This yields 13 withdrawals and the following two-way cross-tabulation.

. tab sbv	worlot pred01fda,	exact		
sbworlot	pred01fda	a 1 I	Total	
	:	+		
0	220	81	301	
1 	6 -+	7   +	13	
Total	•	88	314	
	Fisher's exact =		0.053	
1-sided	Fisher's exact =		0.041	

Here we observe a p-value from Fisher's exact test of greater than 0.05. But in controlled regressions (including exact logistic regressions with exact distributions), we observe p-values of less than 0.05 once again (see two following pages).

## 4.4 Adding Alosetron and Levomethadyl Acetate to the Carpenter/Zucker/Avorn Withdrawals Measure, Using FDA NMEs and FDA Deadline Approval Measure (cont.)

An exact logistic regression on this same measure (with 13 withdrawals), controlling for year of NME submission, yields an estimated odds ratio of 3.9 with a p-value of 0.04.

### Binary Regression

regression (type=logit, model(sbworlot = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes,

#### **Basic Information**

Data file CZAFDAcompare-dump20080410.csv Model

sbworlot(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable <Not Specified> Stratum variable <Unstratified> Analysis type Estimate :: Exact

Number of terms in model Number of term(s) dropped 314 Number of observations in analysis Number of records rejected 10 Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	28.97	21	0.1147
Likelihood Ratio	334.5	3	1.781e-009

#### **Parameter Estimates**

		Point Estima	ate	Confidence In	terval and P-\	/alue for Odd	s Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unper	P-Value
%Const	PMLE	0.5853	NA	Asymptotic	0.02194	15.61	0.7492
				Asymptotic(Profile)	0.02155	19.98	
subyrctr	PMLE	0.8379	NA	Asymptotic	0.6872	1.022	0.08035
				Asymptotic(Profile)	0.6705	1.014	
	CMLE	0.8275	NA.	Exact	0.657	1.015	0.07148
pred01fda	PMLE	3.855	NA	Asymptotic	1.281	11.6	0.01636
				Asymptotic(Profile)	1.263	12.14	
	CMLE	3.907	NA	Exact	1.055	15.04	0.04043

PMLE: Penalized MLE for bias correction (Firth's method).

Analysis Time = 00:00:00

# 4.4 Adding Alosetron <u>and</u> Levomethadyl Acetate to the Carpenter/Zucker/Avorn Withdrawals Measure, Using FDA NMEs and FDA Deadline Approval Measure (cont.)

In a random-effects logistic regression on the same measure, controlling for submission year of the NME and for epidemiological covariates, an OR of 3.98 is observed with a p-value of 0.02. In addition to the random effects (grouped by primary indication of the NME), the added variables are epidemiological controls for total hospitalizations of NME primary indication, and the average length of hospitalization, per hospitalization.

andom-effects	logistic reg	gression			of obs		
roup variable	(i): discode	•		Number o	of groups	=	135
andom effects	u i ~ Gaussi	an		Obs per	group: mir	n =	1
	_				avo	g =	2.3
					max	ζ =	16
				Wald ch	i2(4)	=	7.67
og likelihood	_ EO 1EOC5	11		Drob >	chi2	_	0 1040
og iinciincou	= -50.1596	1		PIOD > (	CIIIZ	=	0.1046
	= -50.13967						
sbworlot		Std. Err.		P> z	[95% Cor	 nf.	Interval]
sbworlot   sbworlot   subyrctr   pred01fda	OR8287204 3.985446	Std. Err. .0900371 2.357488	z -1.73 2.34	P> z   0.084 0.019	[95% Cor 	 nf.  2	Interval] 1.025387 12.70525
sbworlot   sbworlot   subyrctr   pred01fda	OR .8287204	Std. Err. .0900371 2.357488	z -1.73 2.34	P> z   0.084 0.019	[95% Cor 	 nf.  2	Interval] 1.025387 12.70525
sbworlot   subyrctr   pred01fda   hhospdisc	OR8287204 3.985446	Std. Err0900371 2.357488 1.57e-06		P> z   0.084 0.019 0.514	[95% Cor .6697742 1.250174	nf.  2 <mark>4</mark>	Interval] 1.025387 12.70525 1.000004
sbworlot   subyrctr   pred01fda   hhospdisc   hhosleng	OR .8287204 3.985446 1.000001	Std. Err0900371 2.357488 1.57e-06 .0638625		P> z   0.084 0.019 0.514	[95% Cor .6697742 1.250174 .99998 .8702681	nf. 2 1 3 1	Interval] 1.025387 12.70525 1.000004
sbworlot   subyrctr   pred01fda   hhospdisc   hhosleng   /lnsig2u	OR .8287204 3.985446 1.000001 .9878308	Std. Err0900371 2.357488 1.57e-06 .0638625 2.464286		P> z   0.084 0.019 0.514	[95% Cor .6697742 1.250174 .999998 .8702681	nf. 2 4 3 L	Interval] 1.025387 12.70525 1.000004 1.121275

.

# 4.5 Analysis of Permanent U.S. Removals, including Tegaserod Maleate, with FDA's NMEs and FDA's Deadline Approval Measure

We now return to a different but very intuitive measure that looks at permanent removals from the U.S. market that were partially safety related.

This produces the following list, which includes alatrofloxacin mesylate, levomethadyl acetate, and trovafloxacin mesylate – all of which have been permanently removed from the U.S. market – even though these were not officially a safety-based withdrawal in the U.S., they were officially safety-based withdrawals in other countries.

# 4.5.1 Cross Tabulation with Fisher's Exact Test – Taking data up through December 2007, Combine CZA and FDA Measures for Safety-Based Withdrawals.

A cross-tabulation of the permanent U.S. safety-related removals with the deadline approval measure of the FDA yields the following:

. tab sbwusrem\_addzelnorm pred01fda, exact

sbwusrem_a ddzelnorm	pred01fda   0	1	Total
0	221   5	81   7	302 12
Total	226	88	314
1 - 1 - 1	Fisher's exact =		0.042
ı-sıded	Fisher's exact =		0.024

We then write a smaller dataset into a comma-separated file, for use in exact logistic regressions with LogXact software.

. outsheet ndanum\_fda sbwusrem\_addzelnorm subyrctr pred01fda using
c:\fdatemp\sbwusrem-addzelnorm.csv, comma

.

## 4.5.2 Exact Logistic Regression – Taking Data up through December 2007, Combine CZA and FDA Measures for Safety-Based Withdrawals.

### **Binary Regression**

regression (type=logit, model(sbwusrem\_a = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, profile=yes, outtype=odds):

#### **Basic Information**

Data file sbwusrem-addzelnorm.csv

Model sbwusrem\_a(Response = 1)=%Const+subyrctr+pred01fda

Link type

Weight variable <Not Specified> Stratum variable <Unstratified> Estimate :: Exact Analysis type

Number of terms in model Number of term(s) dropped Number of observations in analysis 314 Number of records rejected 0 Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	27.71	21	0.1485
Likelihood Ratio	341.4	3	1.966e-009

#### **Parameter Estimates**

		Point Estima	ate	Confidence In	terval and P-\	Value for Odd	s Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unper	P-Value
%Const	MLE	0.4476	NA	Asymptotic	0.01251	16.02	0.6596
				Asymptotic(Profile)	0.01324	18.89	
subyrctr	MLE	0.8364	NA	Asymptotic	0.6741	1.038	0.1044
				Asymptotic(Profile)	0.6604	1.024	
	CMLE	0.8377	NA	Exact	0.6612	1.033	0.1033
pred01fda	MLE	4,728	NA	Asymptotic	1.412	15.83	0.01175
				Asymptotic(Profile)	1.426	16.9	
	CMLE	4.639	NA	Exact	1.192	19.72	0.02485

#### Analysis Time = 00:00:00

Notice that the estimated OR is slightly <u>larger</u> than that reported in Figure 2 of the NEJM article, and remains statistically significant (exact P = 0.02).

# 4.5.1 Random-Effects Logistic Regression – Taking data up through December 2007, Combine CZA and FDA Measures for Safety-Based Withdrawals.

Results for deadline approval measure highlighted in vellow.

Again, the estimated OR is slightly <u>larger</u> than that reported in Figure 2 of the NEJM article.

# Re-Analysis of Dosage-Form Discontinuations Measure – Random-Effects Logistic Regression

We have re-run our analyses of dosage-form discontinuations with the revised just-before-deadline measure according to Drugs@FDA, and with the FDA's pre-deadline approval measure, and again find estimates that are nearly identical to those of the original paper. Estimated odds ratios are still generally 3 or above (one is at 2.98), and two-tailed exact probability tests produce p-values of < 0.05 in both cross-tabulations and exact logistic regressions.

### A cross tabulation with Fisher's exact test yields:

. tab disc	cont01_null_fdadat	ta314NME	s pred01fda	, exact
ta314NMEs	pred01fda	a 1	Total	
0 1		72   16	36	
Total		88	314	
1-sided	<pre>Fisher's exact = Fisher's exact =</pre>			<mark>029</mark> 019

An exact logistic regression on the same data, controlling for year of NME submission, yields similar results and is shown on the following page.

## 5.1 Analysis of Dosage-Form Discontinuation Measure, Using FDA 314 NMEs and FDA Deadline Approval Measure (cont.)

### **Binary Regression**

regression (type=logit, model(discont01\_ = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outture=odde).

#### **Basic Information**

discont-FDA314NMEs.csv

Data file Model discont01 (Response = 1)=%Const+subyrctr+pred01fda

Logit

Link type Weight variable <Not Specified> <Unstratified> Estimate :: Exact Stratum variable Analysis type

Number of terms in model Number of term(s) dropped Number of observations in analysis Number of records rejected Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	28.95	21	0.1153
Likelihood Ratio	227.7	3	1.225e-009

#### **Parameter Estimates**

r drameter Estimates							
	Point Estimate			Confidence Interval and P-Value for Odds Ratio			
					95 %CI 2*		2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Туре	Lower	Upper	P-Value
%Const	PMLE	3.165	NA	Asymptotic	0.3603	27.79	0.2987
				Asymptotic(Profile)	0.3737	30.02	
subvrctr	PMLE	0.8134	NA	Asymptotic	0.7128	0.9282	0.00217
				Asymptotic(Profile)	0.7071	0.9234	
	CMLE	0.8092	NA	Exact	0.7019	0,9231	0.001095
pred01fda	PMLE	2.978	NA	Asymptotic	1.424	6.229	0.003744
				Asymptotic(Profile)	1.419	6.243	
	CMLE	2.977	NA	Exact	1.315	6.724	0.007818

PMLE: Penalized MLE for bias correction (Firth's method).

Analysis Time = 00:00:01

# 5.1 Analysis of Dosage-Form Discontinuation Measure, Using FDA 314 NMEs and FDA Deadline Approval Measure (cont.)

A random-effects logistic regression controlling for NME submission year, epidemiological covariates and random-effects groups by NME primary indication yields the following results. Results for the FDA-coded pre-deadline variable appear in yellow.

. xtlogit di hhosleng, re i		fdadata314N	MEs subyr	ctr pred(	)1fda hhospd	isc
Random-effects Group variable			of obs = of groups =	314 135		
Random effects u_i ~ Gaussian					group: min = avg = max =	2.3
Log likelihood	= -102.5645	59		Wald chi	2(4) = thi2 =	15.08 0.0045
discont01_~s	OR	Std. Err.	z	P> z	[95% Conf.	Interval]
<pre>pred01fda   hhospdisc  </pre>	.807728 3.131616 1.000002 .9847929	1.26217	2.83 1.56	0.005 0.120	1.421344 .9999996	6.899819 1.000004
/lnsig2u	-2.015556	3.03375			-7.961597	3.930485
	.3650291 .0389254				.0186707 .0001059	
Likelihood-rat	io test of rh	no=0: chibar	2(01) =	0.13 E	Prob >= chiba	r2 = 0.361

# 6. Standard versus Priority Classifications

The priority classification differs for 30 drugs between our data and the same NMEs as noted in Drugs@FDA. The vast majority of these are priority drugs whose reviews took a year or more, often 2-3 years. Therefore, reclassifying them as priority drugs would not affect the measure of whether they were approved just before the deadline; these reviews went well past both the standard and the priority deadlines.

Most of the original data used in our analysis came from a request made to the FDA under the Freedom of Information Act in 2000. We cannot tell whether the discrepancy in standard versus priority codes was a function of the data sent to us by the FDA's FOIA office, or whether it came from later from our own research team (this team was led by Dr. Carpenter, who assumes full responsibility for any errors in coding). For approved NMEs submitted from 2000 to 2003, we relied on approval letters from the CDER website (<a href="http://www.fda.gov/cder/foi/nda/index.htm">http://www.fda.gov/cder/foi/nda/index.htm</a>) (latest access April 15, 2008).

We now present a list of the differences between the priority coding of the FDA and the priority coding of the Carpenter-Zucker-Avorn data, followed by a full list of the drugs coded as priority in our original dataset, and a full list of the drugs coded as priority in the FDA's dataset of April 2008.

# 6.1 List of Drugs Whose Priority Classifications Differ between Carpenter/Zucker/Avorn Data and FDA Data of April 2008.

Drugs highlighted in yellow are NMEs coded as standard by CZA and coded as priority in FDA Data of April 2008.

. list  $\mbox{ ndanum2 ndanum genernam apptimex }\mbox{ priority priority\_recoded}$  if (  $\mbox{ priority\_recoded}$ )

	+					+
	ndanum2	ndanum	genernam	apptimex	priority	priori~d
6.	20315	20315	Levomethadyl acetate	.59	0	1
7.		20313	Trimetrexate glucuronate	10.49	0	1
11.	20326	20326	Nisoldipine	22.09	1	0 1
13.	20350	20350	Famciclovir	11.97	1	0 1
30.		20444	Epoprostenol	18.71	0	1 1
33.		20451	Porfimer	20.48	0	1 1
34.	20459	20459	Nalmefene	11.61	1	0 1
38.	20482	20482	Acarbose	12	1	0 1
43.	20498	20498	Bicalutamide	12.66	1	0 1
57.		20564	Lamiyudine	4.37	0	1
68.	20599	20599	Riluzole	5.46	0	1 1
95.	20687	20687	Mifepristone	55.3	0	1 1
98.	20690	20690	donepezil hydrochloride	7.92	0	1 1
106.	20715	20715	Triptorelin Pamoate	6.066667	1	0 1
194.		21060	Ziconotide	60.9	0	1 1
207.	21106	21106	Pegvisomant	27.43333	0	1 1
208.	21107	21107	Alosetron HC1	7.5	0	1
209.	21119	21119	Verteporfin	8.066667	0	1
220.	21196	21196	Sodium Oxybate	21.83333	0	1 1
222.	21200	21200	Tegaserod Maleate	29.8	0	1
226.	21223	21223	Zoledronic Acid	20.26667	0	1
227.	21226	21226	Lopinavir; Ritonavir	3.566667	0	1
229.	21232	21232	Nitisinone	25.1	0	1
232.	21257	21257	Travoprost	8.433333	0	1
233.	21264	21264	Apomorphine Hydrochloride	15.86667	0	1 i
235.	21272	21272	Treprostinil Sodium	19.4	0	1
242.	21320	21320	Abarelix	35.96667	0	1
249.	21345	21345	Fondaparinux Sodium	9.833333	0	1
264.	21431	21431	Acamprosate Calcium	31.5	0	1
268.	21446	21446	Pregabalin	14.23333	0	1
271.	21462	21462	Pemetrexed Disodium	4.266667	0	1
279.	21506	21506	Micafungin Sodium	35.06667	0	1
286.	21572	21572	Daptomycin	8.9	0	1
292.	21640	21640	Ovine Hyaluronidase	9.166667	0	1
294.	21665	21665	Hyaluronidase	16.93333	0	1
296.	21670	21670	Trypan Blue	13.96667	0	1
297.	21673	21673	Clofarabine	9.133333	0	1
301.	21743	21743	Erlotinib Hydrochloride	3.733333	0	1 1
302.	21749	21749	Pentetate Calcium Trisodium	4.4	0	1
303.	21751	21751	Pentetate Zinc Trisodium	4.266667	0	1
312.	50678	50678	Dirithromycin	25.68	1	0
321.	50794	50794	Azacitidine	4.833333	0	1
	+					

# 6.1 List of Drugs Whose Priority Classifications Differ between Carpenter/Zucker/Avorn Data and FDA Data of April 2008 (cont.)

There would appear to be 35 drugs that we coded as standard but which the FDA codes as priority NMES, and 7 drugs which we coded as priority but that the FDA codes as standard NMES.

```
. count if( priority == 0 & priority_recoded == 1)
35
. count if( priority == 1 & priority_recoded == 0)
7
```

So the total difference is 28 drugs, which if added to our 102 priority NMEs in the original data set makes 130.

In addition, the FDA data has five NMEs that are not in our data set, two of which – hyaluronidase (NDA # 21716) and Fludeoxyglucoase F 18 (NDA # 20306) – are coded by the FDA as priority NMEs. Adding this to the 130 priority NMEs in our revised data brings the total to 132, which is the FDA's count according to the Nardinelli, et al. letter.

# 6.2. LIST OF DRUGS THAT CZA CODED STANDARD THAT FDA CODES AS PRIORITY, AS OF APRIL 2008

The majority of the discrepancies are drugs that we coded as priority that FDA coded as standard. In the next section, we show that most of these drugs were approved in well over one year of review time, and many others were approved well before the two-month before deadline window. Hence very few of our pre-deadline classifications were affected by the discrepancy.

Drugs with 14 months or greater approval time are highlighted in yellow.

. list  $\mbox{ ndanum2 ndanum genernam apptimex }\mbox{ priority_recoded if( priority == 0 & priority_recoded == 1)}$ 

	+   ndanum2	ndanum	genernam	apptimex	priority	   priori~d 
6.	20315	20315	Levomethadyl acetate	.59	 0	   1
7.	20315	20326	Trimetrexate glucuronate	10.49	0	1
30.		20444	Epoprostenol	18.71	0	1
33.	20451	20451	Porfimer	20.48	0	1 i
57.	20564	20564	Lamivudine	4.37	0	1
68.	20599	20599	Riluzole	5.46	0	 1
95.	20687	20687	Mifepristone	55.3	0	1
98.	20690	20690	donepezil hydrochloride	7.92	0	1
194.	21060	21060	Ziconotide	60.9	0	1
207.	21106	21106	Pegvisomant	27.43333	0	1
208.	21107	21107	Alosetron HC1	7.5	0	 1
209.	21119	21119	Verteporfin	8.066667	0	1 i
220.	21196	21196	Sodium Oxybate	21.83333	0	1
222.	21200	21200	Tegaserod Maleate	29.8	0	1
226.	21223	21223	Zoledronic Acid	20.26667	0	1
227.	21226	21226	Lopinavir;Ritonavir	3.566667	0	   1
229.	21232	21232	Nitisinone	25.1	0	1
232.	21257	21257	Travoprost	8.433333	0	1
233.	21264	21264	Apomorphine Hydrochloride	15.86667	0	1
235.	21272	21272	Treprostinil Sodium	19.4	0	1
242.		21320	Abarelix	35.96667	0	1
249.	21345	21345	Fondaparinux Sodium	9.833333	0	1
264.		21431	Acamprosate Calcium	31.5	0	1
268.		21446	Pregabalin	14.23333	0	1
271.	21462	21462	Pemetrexed Disodium	4.266667	0	1   1
279.	21506	21506	Micafungin Sodium	35.06667	0	1
286.	21572	21572	Daptomycin	8.9	0	1
292.	21640	21640	Ovine Hyaluronidase	9.166667	0	1
294.	21665	21665	Hyaluronidase	16.93333	0	1
296.	21670	21670	Trypan Blue	13.96667	0	1
297.	21673	21673	 Clofarabine	9.133333	0	 1
301.	21743	21743	Erlotinib Hydrochloride	3.733333	0	⊥ I 1 I
301.	21743	21743	Pentetate Calcium Trisodium	4.4	0	⊥ I 1 I
302.	21749	21749	Pentetate Zinc Trisodium	4.266667	0	1 1
303.	50794	50794	Azacitidine	4.833333	0	⊥ I 1 I
JZI.	+	JU194 	AZACILIQINE			± +

# 6.2. LIST OF DRUGS THAT CZA CODED STANDARD THAT FDA CODES AS PRIORITY, AS OF APRIL 2008

Comparing the drugs that we coded as standard but which the FDA codes as priority, shows that there is a material difference in average approval time between the drugs with discrepancies and the sample of drugs coded as priority by the FDA. In particular, the average approval time for the total sample of priority NMES as coded in the FDA's April 2008 data is 11.33 months, whereas the discrepancy drugs average 16.5 months of review time. Again, Dr. Carpenter and his research team assume responsibility for any errors – which seem to have come from our coding of FDA approval letters on the FDA web site – but the lengthier review of these drugs should be noted.

. sum apptimex it	f( priority	== 0 & pr	iority_recoded	d == 1)		
Variable	Obs	Mean	Std. Dev.	Min	Max	
apptimex		16.50533	14.19526	.59	60.9	
. sum apptime_fo	da if(prior	ity_fda01 =	= 1)			
Variable	Obs	Mean	Std. Dev.	Min	Max	
apptime_fda		11.33333	9.658439	.6	60.1	

Note that fully 17 of these FDA's priority drugs that we coded as standard NMEs had approval times of 14 months or greater.

## 7. Timing of Approvals

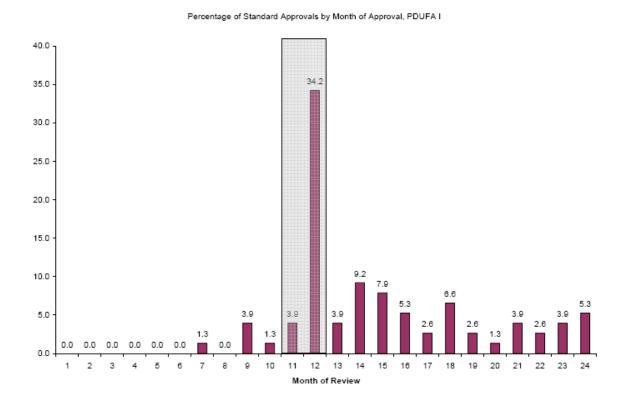
Nardinelli et al. also argue that our approval time data are faulty, pointing in particular to 25 standard NMEs under PDUFA II and after that are approved before the tenth month in our data, whereas they observe only five such NMEs in their data. This difference is simply a matter of rounding. Our original coding was based on dividing the number of days of approval by 30 (in separate data that was sent to us by former FDA official Ed Hass, a similar measure with similar denominator was used). This produced a number of drugs that had met the review deadlines but had numerical review times that appeared to put them past the deadlines. For instance, L-Glutamine (Nutrestore; NDA # 21667) was submitted on August 8, 2003 and approved on June 10, 2004, which generates a review time of 10.23 months – so we rounded down to produce the patterns Figure 1. This had the effect of rendering NMEs that went into the tenth month of review being coded as ninth-month approvals, which accounts for the difference.

We note that, to the extent that the FDA's rounding is preferred, our argument about the concentration (or "piling") of approval decisions near the deadline approvals is even stronger. If for the 10-month deadline period (standard NMEs submitted under PDUFA II and afterwards) there are more tenth-month approvals and fewer ninth-month approvals, then a greater proportion of NMES are being approved immediately before the operative deadline. What is more, since the differences are being driven by distinctions between ninth and tenth month approvals, none of these discrepancies affects our coding of just-before-deadline approvals.

## 7.1. Timing of Approvals (cont.)

# COUNT OF APPROVAL TIMES, ROUNDING TO NEAREST INTEGER INSTEAD OF ROUNDING DOWN

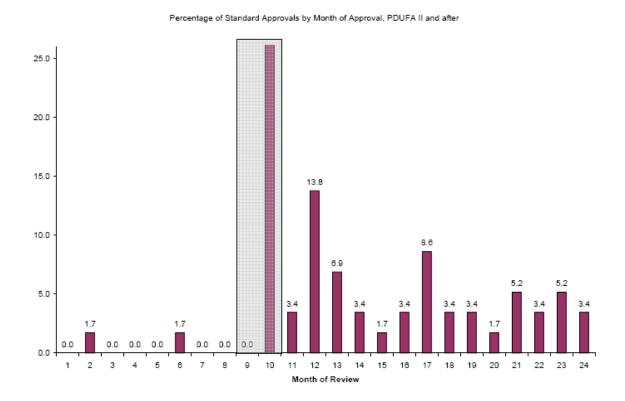
First do standard NMEs under PDUFA I



## 7.2. Timing of Approvals (cont.)

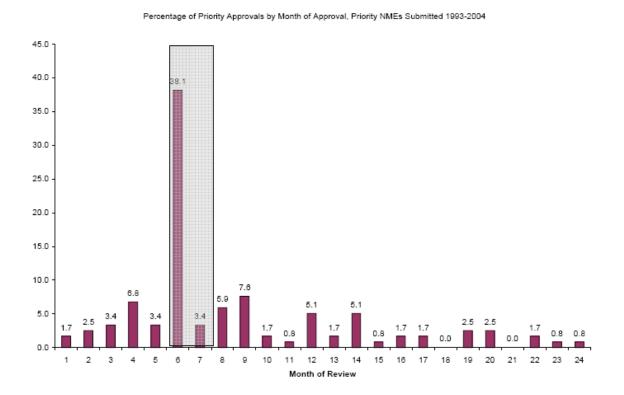
## **RE-TABULATION USING DIFFERENT ROUNDING RULE (cont.)**

Next do standard NMEs under PDUFA II and after. This produces just 2 NMEs approved before the  $10^{\rm th}$  month.



## 7.1. Timing of Approvals (cont.)

Now re-tabulate priority NMEs using rounding to nearest integer.



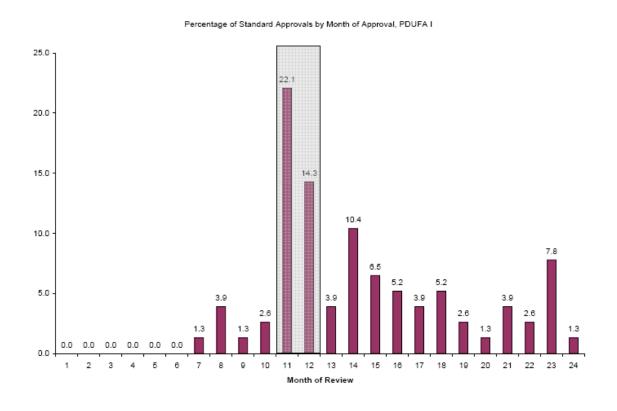
### 7.2. Timing of Approvals (Using Floor Function)

# RE-TABULATION OF APPROVAL TIMES BY MONTH, USING "FLOOR" FUNCTION.

We then tried another rounding function, which appears to produce numbers closer to the FDA for standard NMEs since PDUFA II (see below). Again, using this rounding function does not change our coding of pre-deadline measures and does not change our substantive results.

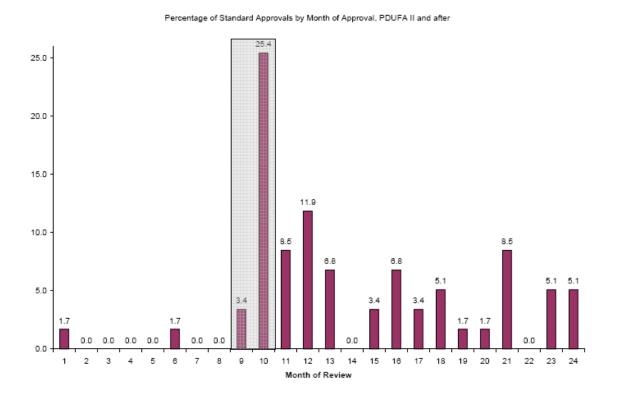
```
In STATA SE9, floor(x) returns the integer n such that n < x < n+1. floor(x) returns x(not ".") if x is missing, meaning floor(.a) = .a. gen apptime_floor = floor(apptimex)
```

### First do standard NMEs under PDUFA I



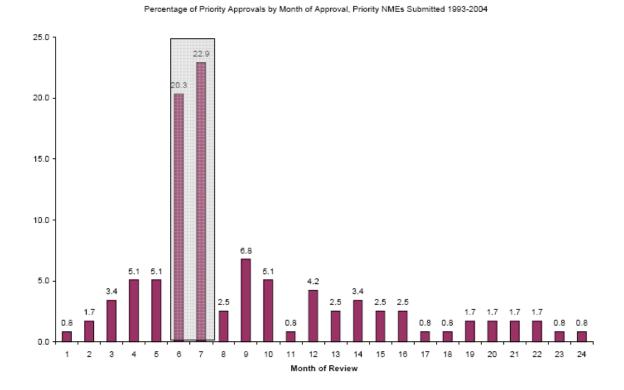
## 7.2. Timing of Approvals (Using Floor Function) (cont.)

Now do standard NMEs under PDUFA II and after, for first 24 months of review cycle. This produces 4 drugs, which is very close to the 5 drugs described in the Nardinelli et al. letter.



## 7.2. Timing of Approvals (Using Floor Function) (cont.)

Now tabulate priority NMEs for first 24 months of review cycle.



Memorandum to *NEJM* Editors on Differences between Carpenter/Zucker/Avorn data and FDA data on Black-Box Warnings, and Re-Analysis of Data (cont.)

# 8. Coding Black-Box Warnings<sup>3</sup>

We completed our coding in the summer of 2005; the first version of this paper was submitted to the *New England Journal of Medicine* in January 2006.

Black-Box Warnings: There is no unified database of black-box warnings added to drugs postmarket. The closest is Lasser's paper of 2002, which stops coding in 1999. We used this Lasser article [JAMA [2002; 287(17) May 1:2215-20], which was static, and the KUMC database, a comprehensive list of black-box warnings added to drugs that is regularly updated (http://www.formularyproductions.com/master/showpage.php?dir=blackbox&whichpage=9 (most recent download May 17, 2008). Postmarket black box warnings listed on the KUMC were coded until July 2005, when Mr. Zucker stopped working on the project; nowork was completed on or after August 1, 2005. (Some drugs receive multiple black-box warnings over time, and so it is possible that drugs that received an additional black-box warning after August 1, 2005 also received a postmarket boxed warning before August 1, 2005). In addition it is possible that some of the drugs that had black-box warnings attached between January and July 2005 were not yet listed in the KUMC database. We cannot verify this possibility as we do not have the entire KUMC database as it was posted in August of 2005.

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 $<sup>^{3}</sup>$  For a review of our coding rules in general, see Section 1.1 of the document.

9. Differences between the April 2008 Nardinelli et al data and the Carpenter, Zucker and Avorn (CZA) data.

# 9.1 Catalog of Black-Box Warning (BBW) Differences

The following is the list of NMEs with a postmarket black-box warning as coded by Carpenter, Zucker and Avorn.

. list	ndanum_fc	da genernam_fda if(bbwcza == 1)	
<del> </del>	⊦   ndanum~a	genernam_fda	<del>+</del> 
3.	21341	Valdecoxib	 
42.	21492	Oxaliplatin	
104.	20430	Danaparoid Sodium	
106.	21345	Fondaparinux Sodium	
112.	21500	Emtricitabine	
120.	21356	Tenofovir Disoproxil Fumarate	 
124.	21814	Tipranavir	[
137.	20759	Trovafloxacin Mesylate	
177.	21427	Duloxetine Hydrochloride	[
221.	20697	Tolcapone	
232.	20720	Troglitazone	 
300.	21797	Entecavir	
301.	21449	Adefovir Dipivoxil	1
311.	21618	Tinidazole	
+	+		+

#### 9.1 Catalog of Black-Box Warning (BBW) Differences (cont.)

The following is the complete list of NMEs with postmarket BBWs as coded by Nardinelli et al in their April 2008 web-posted data.

. list	ndanum_fda	<pre>genernam_fda bbwdate_fda if</pre>	( bbwfda08 ==
+	 ndanum~a 	genernam_fda	bbwdate~a
2.	20998	Celecoxib	29-Jul-05
3.	21341	Valdecoxib	24-Nov-04
4.	20938	Meloxicam	11-Aug-05
14.	20535	Bromfenac Sodium	31-Mar-98
33.	20896	Capecitabine	7-Sep-01
54.	20169	Nilutamide	29-Sep-00
104.	20430	Danaparoid Sodium	30-Jan-98
108.	20564	Lamivudine	15-Dec-97
114.	20628	Saquinavir	14-Nov-00
134. <sub> </sub>	21144	Telithromycin	12-Feb-07
137.	20759	Trovafloxacin Mesylate	9-Jun-99
138.	20760	Alatrofloxacin Mesylate	9-Jun-99
175. I	20822	Citalopram Hydrobromide	18-Feb-05
176.	20415	Mirtazapine	12-Jan-05
177.	21427	Duloxetine Hydrochloride	18-Feb-05
183.	20592	Olanzapine	16-Feb-06
184.	21436	Aripiprazole	16-Feb-06
185.	20639	Quetiapine Fumarate	20-Sep-06
186.	20825	Ziprasidone Hydrochloride	17-Aug-05
<mark>189.  </mark>	21302	Pimecrolimus	19-Jan-06
221.	 20697	Tolcapone	 16-Nov-98
228.	21073	Pioglitazone Hydrochloride	14-Aug-07
232.	20720	Troglitazone Troglitazone	15-Dec-97
236. I	21071	Rosiglitazone Maleate	14-Aug-07
244.	21411	Atomoxetine Hydrochloride	8-Nov-05
260.	20937	Gadoversetamide	4-Sep-07
266.		Perflutren Lipid Microsphere	10-Oct-07
273.	20815	Raloxifene Hydrochloride	13-Sep-07
284.	21107	Alosetron Hydrochloride	7-Jun-02
+			

Of these 29 NMEs, 13 (highlighted in yellow) had their warnings attached since the beginning of August 2005, while four more (in green), and 17 have had their black-box warnings added since

January 2005, in the wake of the Vioxx withdrawal. Thus, 58% of the postmarket black-box warnings listed by Nardinelli et al were added since the withdrawal of Vioxx, and 45 percent were added after August 1, 2005, when our coding stopped. These calculations assume that all of the FDA's NMEs represent methodologically sound coding of post-market BBWs.

The pattern of adding black-box warnings changed appreciably since the withdrawal of rofecoxib (Vioxx) in September 2004. This consideration did not drive our decision to terminate coding in at the end of July 2005; this is simply when our data collection stopped. However, the shift that occurs during the period September 2004 to January 2005 is noticeable and dramatic. Nearly six in ten of the FDA's BBWs occur in the last one-fifth of the FDA's time window. Put differently, a highly disproportionate share of the FDA's coded BBWs are added from 2005 through 2007. No such discontinuity is evident for other periods of time, especially as demonstrated by Lasser and collagues (*JAMA* 2002). The high prevalence of molecular-class relabelings in recent years may mean that post-Vioxx BBWs reflect an aspect of FDA's postmarket warnings policies that are different from the dimensions captured by pre-Vioxx BBWs.

At least one significant difference between class BBW relabelings and more specific BBW relabelings is the specificity of the evidence concerned. Class relabelings occur with indirect evidence from other molecules that rely upon similar mechanisms; they are not based upon direct evidence of studies of the molecule itself. Such patterns are observed for the BBWs added to "atypical antipsychotic" drugs in recent years (apiprazole, olanzipine, quetiapine, ziprasidone).

Hence the more recent BBW data may not be comparable to the BBW data of the sort that Lasser and others collected. Further research may be necessary to examine how post-Vioxx black-box warning decisions differ from those before the rofecoxib withdrawal. To assess what happens when we exclude those NMEs whose black-box warnings are based upon indirect clinical

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<sup>&</sup>lt;sup>4</sup> Note that since many of the 314 NMEs in the FDA's April 2008 data were not approved until well after 1993, it is <u>not</u> true that 58 percent of these NMEs received a black-box warning in the last three years of their market life. Still, there is a stark and rather odd discontinuity in the FDA's black-box warning data, as we document below.

evidence and upon therapeutic class and molecular mechanism similarities to other NMEs, see Section 15 below.

#### 9.2 Erroneous Coding in the FDA's Data of April 2008

The FDA's April 2008 data do not include at least five NMEs where a black-box warning has been added or significantly modified. These five are

- **Tenofovir** (Viread)
- Adefovir (Hepsera)
- Entecavir (Baraclude)
- Tiprinavir (Aptivus)
- **Emtricitabine** (Emtriva)

We have conducted year-by-year searches of these drugs in the *Physicians' Desk Reference* since their FDA approval. While we are not certain that we have captured <u>all</u> of the changes to the BBWs for these drugs,<sup>5</sup> we are confident that we have identified <u>at least one</u> significant labeling boxed warning change in each of them, and are confident that the FDA has thus erred in their April 2008 coding.

There are also cases where we believe that the FDA has coded a drug as receiving a black box warning, when in fact no warning was added or it is not clear that significant new safety information was conveyed by the labeling change. The clearest case of an error comes in the drug **saquinavir** (**Invirase**). While this drug did receive a new black-box warning in February 2008 (see Roche "Dear Doctor Letter" of February 2008, available at <a href="http://www.rocheusa.com/products/invirase/DM11.1">http://www.rocheusa.com/products/invirase/DM11.1</a> Invirase Dear Doctor Letter.pdf (accessed May 12, 2008)]), a February 2002 publication by the National Institutes of Health and the AHRQ states clearly that no Black Box Warning was evident for saquinavir at that time (see

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<sup>&</sup>lt;sup>5</sup> Both the Lasser 2002 article in *JAMA* and a communication by physicians Charles Bennett and Oliver Sartor [*Annals of Internal Medicine* 139 (6) (September 16, 2003) 529] identify possible discrepancies between information in the PDR and other sources. It is clear, however, that significant safety information in boxed warnings was added to the labels in at least one instance for each of these five drugs.

Figure 9A, page after next). This is more than a year after the date for which the FDA's April 2008 data has the warning added.

This omission is odd, given that two FDA officials (Heidi Jolson and Jeffrey Murray) appear to have served as participants in the creation of this report. (See Figure 9B, below, page after Figure 9A.)

In other cases, the FDA codes drugs that as having postmarket BBWs where there were not significant changes to the label or where the changes were directed to very small patient populations relative to the target population. The first of these cases is **nilutamide** (**Nilandron**) and **capecitabine** (**Xeloda**), both of them oncologic drugs that are used concomitantly with other therapies. In a communication to the *Annals of Internal Medicine*, Drs. Charles Bennett and Oliver Sartor note that small-scale changes in evidence spurred the labeling modifications for nilutamide, and that different sources code these changes quite variably. ["Pneumonitis with Anti-Androgens," *Annals of Internal Medicine* 139 (6) (September 16, 2003) 529]. Dr. Bennett is one of the nation's most respected authorities on postmarket risks of oncologic drugs.

The second of these cases is and **capecitabine** (**Xeloda**). Here the labeling change drew attention to interactions with warfarin in a small number of patients. [Laurel M. Janney and Nancee V. Waterbury, "Capecitabine-Warfarin Interaction," *The Annals of Pharmacotherapy* 39 (9) 1546-1551.]

Both of these drugs – nilutamide (Nilandron) and capecitabine (Xeloda) – are oncologic drugs that are used concomitantly with other therapies; as Bennett has noted, postmarket surveillance and coding of safety problems for oncologic therapies is highly problematic ["The Research on Adverse Drug Events and Reports (RADAR) Project," *JAMA*.2005; 293: 2131-2140].

Figure 9A: NIH Publication of February 2002 Lists Saquinavir as Not Having BBW.

Nevirapine (Viramune™)	<ul> <li>Severe, life-threatening hepatotoxicity, including fulminant and cholestatic hepatitis, hepatic necrosis, and hepatic failure. Patients should be advised to seek medical evaluation immediately should signs and symptoms of hepatitis occur.</li> <li>Severe, life-threatening, and even fatal skin reactions, including Stevens-Johnson syndrome, toxic epidermal necrolysis, and hypersensitivity reactions characterized by rash, constitutional findings, and organ dysfunction have occurred with nevirapine treatment</li> <li>Patients should be monitored intensively during the first 12 weeks of nevirapine therapy to detect potentially life-threatening hepatotoxicity or skin reactions.</li> <li>A 14-day lead-in period with nevirapine 200mg daily must be strictly followed.</li> <li>Nevirapine should not be restarted after severe hepatic, skin, or hypersensitivity reactions</li> </ul>
Ritonavir (Norvir <sup>TM</sup> )	Co-administration of ritonavir with certain medications may result in potentially serious and/or life-threatening adverse events due to effects of ritonavir on hepatic metabolism of certain drugs
Saquinavir (Fortovase <sup>TM</sup> , Invirase <sup>TM</sup> )	None
Stavudine (Zerit <sup>TM</sup> )	<ul> <li>Lactic acidosis and severe hepatomegaly with steatosis, including fatal cases, have been reported with the use of antiretroviral nucleoside analogues alone or in combination</li> <li>Fatal lactic acidosis has been reported in pregnant women who received a combination of stavudine and didanosine along with other antiretroviral combinations         <ul> <li>Stavudine and didanosine combination should only be used during pregnancy if the potential benefit clearly outweighs the potential risks</li> </ul> </li> <li>Fatal and non-fatal pancreatitis have occurred when stavudine was part of a combination regimen with didanosine with or without hydroxyurea</li> </ul>

http://www.ncbi.nlm.nih.gov/books/bv.fcgi?rid=hstat2.table.10535 (accessed May 4, 2008).

Figure 9B: Two FDA Officials Were Participants in the February 2002 NIH/AHRQ Study that Listed Saquinavir as not having a BBW.

#### Participants from the Department of Health and Human Services:

Victoria Cargill National Institutes of Health Oren Cohen National Institutes of Health Mark Dybul National Institutes of Health

T. Randolph Graydon Centers for Medicare and Medicaid Services

Heidi Jolson Food and Drug Administration

Jonathan Kaplan Centers for Disease Control and Prevention
Abe Macher Health Resources and Services Administration

Henry Masur National Institutes of Health
Lynne Mofenson National Institutes of Health
Jeffrey Murray Food and Drug Administration

Joseph O'Neill Health Resources and Services Administration

Alice Pau National Institutes of Health

Lucille C. Perez Substance Abuse and Mental Health Services Administration

http://www.ncbi.nlm.nih.gov/books/bv.fcgi?rid=hstat2.chapter.10256 (accessed May 4, 2008).

#### 9.3 Erroneous Codings in the Carpenter Data

Thorough re-evaluation of all our coding has revealed some instances of errors in coding of the Black-Box Warning (BBW) data. First, some drugs that received postmarket black-box warnings well before August 2005 were not coded as such: citalopram (Celexa), lamivudine (Epivir), bromfenac (Duract), alatrofloxacin (Trovan) and alosetron (Lotronex). In two of these cases – alatrofloxacin and bromfenac– we had correctly coded these drugs as withdrawn from the market, but failed to detect the BBW added before the withdrawal. Dr. Carpenter assumes sole responsibility for these two errors.

Another drug – celecoxib (Celebrex) – appears to have had its BBW added on July 29, 2005. This occurred while Mr. Zucker completed his work during the middle of the summer and was omitted. It is quite possible, although we cannot say for sure, that the KUMC web database had not yet updated its list by July 31, 2005. Again, Dr. Carpenter assumes sole responsibility for this error.

We coded three drugs as having postmarket BBWs added when less than significant safety information was communicated in the labeling changes: oxaliplatin (Eloxatin), fondaparinux (Arixtra) and tinidazole (Tindamax). In all three cases the label's black box warning did change but we did not distinguish adequately between formatting changes and information-based changes. Dr. Carpenter assumes sole responsibility for these errors.

<sup>&</sup>lt;sup>6</sup> Alosetron did not have a black-box warning added before its 2000 withdrawal, but did have a black-box warning when it was reintroduced in 2002, which we failed to detect. In hindsight, we should have coded alosetron as having had a black-box warning added.

## 10. Additional General Notes on the Reliability of the FDA's Postmarket Safety Data, esp as concerns Black Box Warnings

More broadly, we believe that the FDA's data of April 2008 are methodologically flawed and empirically inconsistent with numerous sources, including the FDA's own publications.<sup>7</sup>

(a) It is important <u>not</u> to regard either the Drugs@FDA database or the MedWatch system data as authoritative for purposes of drug safety. The database contains some very important omissions. For instance, **Posicor** (mibefradil hydrochloride) and Duract [bromfenac sodium) were withdrawn for safety reasons, yet they are not listed on the Drugs@FDA webpage and are absent from the underlying database maintained by FDA. (As Figures 2A through 2F above clearly show, searching for these two drugs by generic name, trade name and NDA number fails to return a record with the original NME.) These are two of the most important safety-based withdrawals of the past several decades; their absence from the FDA's own data is troubling.

Regarding Black Box Warnings in particular, not every addition to an existing black-box warning conveys important new safety information.

The FDA letter to the Editor cites "FDA data," but we know of no single, accessible database where such data are kept. Indeed, the data posted to the FDA website on April 7, 2008 (<a href="http://www.fda.gov/oc/pdufa/FDADrugAppSafetyData\_files/NMESafetySumm.html">http://www.fda.gov/oc/pdufa/FDADrugAppSafetyData\_files/NMESafetySumm.html</a>), are the first such publication of FDA data by the agency that combine approval time information with safety-related postmarket regulatory events.

Moreover, the information posted by Narindelli et al is not an authoritative database. As footnote 5 of the FDA letter states, "The FDA data is based on major safety labeling changes posted on

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<sup>&</sup>lt;sup>7</sup> We refer to the FDA's "data of April 2008" because this data differs so appreciably from other data on the FDA website, other data that the FDA has published, and other data that the FDA has provided to other authors and which has been presented at FDA-sponsored conferences. While Nardinelli et al refer to "FDA data," it is important to point out that there is no unified, consistent "FDA data" on approvals and postmarket safety problems.

Medwatch and Drugs@FDA." There is no other methodology noted or acknowledged to compile these data. How for instance is a "major" safety labeling change defined? Does the "and" in the sentence mean that the safety labeling change must be posted on both Medwatch and Drugs@FDA in order to qualify?

For these and related reasons, we used independent sources – principally the article of Lasser (*JAMA* [2002; 287(17) May 1:2215-20]) – to code post-market safety events. We also used an independent source, a list maintained by the Kansas University Medical Center (<a href="http://www.formularyproductions.com/master/showpage.php?dir=blackbox&whichpage=9">http://www.formularyproductions.com/master/showpage.php?dir=blackbox&whichpage=9</a> [most recent access May 17, 2008]). These sources existed years before the FDA's data were assembled, they are publicly accessible in single data form, and in both cases (especially the Lasser article) a clear methodology is used.

Nardinelli et al do not acknowledge previous research methods in defining their new classification of black-box warnings. In 2006 FDA Nardinelli and colleagues circulated a paper under the National Bureau of Economic Research (NBER) Working Paper Series that analyzed NMEs and black-box warnings. [We do not know of its current publication status, but cannot find evidence of its publication. We used an April 7, 2008 download from the nber.org website, but cannot find this paper on the FDA website.] Among the many glaring inaccuracies in the paper, **the Lasser article (JAMA 2002) was not even cited**. Since we suspect that much of the data that FDA is using was also used in this paper, this omission underscores the fundamental methodological weaknesses of the FDA's black-box warning data.

#### Figure 10A – Title Page of Begosh, et al Working paper

NBER WORKING PAPER SERIES

### BLACK BOX WARNINGS AND DRUG SAFETY: EXAMINING THE DETERMINANTS AND TIMING OF FDA WARNING LABELS

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NATIONAL BUREAU OF ECONOMIC RESEARCH 1050 Massachusetts Avenue Cambridge, MA 02138 December 2006

The views expressed in this paper are those of the authors and do not necessarily reflect those of the institutions with which the authors are affiliated. Vernon is also a professor at the University of Connecticut and a Faculty Research Fellow with the National Bureau of Economic Research (NBER).

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Figure 10B – Page 21 ("References") of Begosh et al Working Paper, with Eight Total References. Lasser and Other Medical Literature Not Cited.

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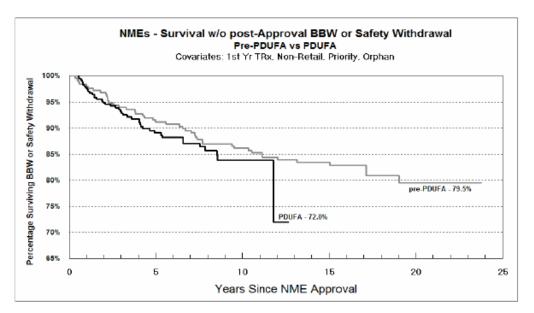
Public Citizen Health Research Group (2002), "Comments on the Prescription Drug User Fee Act (PDUFA)" Docket No. 01N-0450, January 25, 2002. Available at: http://www.citizen.org/publications/.

Vernon, J. Golec, J. Lutter, R. Nardinelli, C. (2006), "FDA New Drug Approval Times, Prescription Drug User Fees, and R&D Spending," *AEI-Brookings Joint Center for Regulatory Studies*.

There are other worrisome aspects to this paper. There appears to be a large, discontinuous change in FDA black-box warnings, comprising 12 percent of the data or more, added at exactly the same time (see Figure 2 of the Begosh et al paper, reproduced below). The simultaneous occurrence of such a large fraction of BBWs or withdrawals should raise concerns about the quality of the data.

Figure 2 of Begosh et al, working paper (p. 14, download of April 7, 2008).





This simultaneous recoding does not appear in our black-box warning data (and no such discontinuity appears in the data presented by Lasser and colleagues in 2002). We nonetheless reproduce it here to note that it is highly improbable that a rigorously-coded boxed warning measure would yield such a discontinuity. Such a discontinuity would require that 12 percent of the NMEs in the post-PDUFA sample (288 drugs, or either 34 or 35 drugs) would have to have received black-box warnings *at the exact same year* + *month interval* after having been approved. If for instance the discontinuity appears at 11 years and 10 months, for instance, it would have to be the case that at least 34 drugs received a black-box warning exactly 11 years and 10 months after having been coded.

(e) FDA Postmarket Safety Data May Substantially Overcount Black-Box Warnings, Compared to Published Sources. [An Example, from Comparison of the FDA's data with Lasser from 1981 to 1992]

In their 2006 paper, for the "pre-PDUFA period" 1981 to 1992, FDA counts 21% of 228 drugs, or 47 NMEs, as having had a black-box warning added by the end of their sample period (pp. 11-12).

For approved NMEs submitted in the same period (1981 to 1992), Lasser and colleagues count just 28 (see Table 1, Lasser 2002). This is a difference of 68 percent. Some of the difference could be the more recent data of the FDA, but this seems unlikely to count for all of the difference.

These issues raise important doubts about the quality of the black-box warnings presented by Nardinelli et al as "FDA data."

<sup>&</sup>lt;sup>8</sup> "Among the pre-PDUFA NMEs, 79 percent survived to the end of the data sample period without new BBWs" (Begosh, et al., 11).

<sup>&</sup>lt;sup>9</sup> If one focuses on the approval year as opposed to the submission year, it is possible to retrieve an estimate of 30 drugs so coded from Lasser, Table 1, Yet drugs must be coded as to submission date if an accurate coding of pre-PDUFA versus post-PDUFA drugs is to be achieved.

11. Reanalysis with Corrected CZA Measure of Black Box Warnings (BBWs), and Analyses Combining Corrected CZA Measure with Nardinelli et al Measure.

#### 11.1.1 Combined Measure with Corrections for Carpenter-Zucker-Avorn (CZA) errors.

We first generated combined measure as of August 2005, but corrected for three NMEs that represent CZA error (fondaparinux sodium, tinidazole, oxaliplatin).

```
. gen CZAFDAbbw_august2005_corrected = CZAFDAcombobbw_august2005
. replace CZAFDAbbw_august2005_corrected = 0 if(ndanum_fda == 21345)
(1 real change made) {THIS IS FONDAPARINUX SODIUM (ARIXTRA)}
. replace CZAFDAbbw_august2005_corrected = 0 if(ndanum_fda == 21618)
(1 real change made) {THIS IS TINIDAZOLE (TINDAMAX)}
. replace CZAFDAbbw_august2005_corrected = 0 if(ndanum_fda == 21492)
(1 real change made) {THIS IS OXALIPLATIN (ELOXATIN)}
```

A cross-tabulation of this measure upon the FDA's deadline approval measure, using the FDA's 314 NMEs, yields the following:

To generate a shorter name for this variable, to be used in LogXact8 (which has max 9 characters in variable names), we used the following code:

```
. gen combcorr1 = CZAFDAbbw_august2005_corrected
```

This variable is used as the dependent variable in the exact logistic regressions on the following page.

### 11.1.2 Exact Logistic Regression on Combined CZA-FDA Measure, dropping Erroneous CZA Positive Codings, stop coding August 2005

#### **Binary Regression**

regression (type=logit, model(combcorr1 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes,

#### **Basic Information**

Data file combocorrectbbws20080513.csv

Model combcorr1(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable <Not Specified>
Stratum variable <Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	26.23	21	0.1977
Likelihood Ratio	289.5	3	1.244e-009

#### **Parameter Estimates**

	Point Estima	ate	Confidence Interval and P-Value for Odds Ratio				
				95	2*1-sided		
Type	Odds Ratio	SE(Odds)	Type	Lower	Unner	P-Value	
PMLE	0.05751	NA	Asymptotic	0.004505	0.7342	0.02797	
			Asymptotic(Profile)	0.004188	0.7388		
PMLE	0.9855	NA	Asymptotic	0.8545	1.136	0.8403	
			Asymptotic(Profile)	0.85	1.136		
CMLE	0.9832	NA	Exact	0.8447	1.139	0.8561	
PMLE	3,802	NA	Asymptotic	1.551	9.319	0.003502	
			Asymptotic(Profile)	1.555	9.594		
CMLE	3.851	NA.	Exact	1.4	11	0.007451	
	PMLE PMLE CMLE PMLE	Type         Odds Ratio           PMLE         0.05751           PMLE         0.9855           CMLE         0.9832           PMLE         3.802	PMLE         0.05751         NA           PMLE         0.9855         NA           CMLE         0.9832         NA           PMLE         3.802         NA	Type         Odds Ratio         SE(Odds)         Type           PMLE         0.05751         NA         Asymptotic           Asymptotic(Profile)         Asymptotic           PMLE         0.9855         NA         Asymptotic           CMLE         0.9832         NA         Exact           PMLE         3.802         NA         Asymptotic           Asymptotic(Profile)         Asymptotic(Profile)	Type	Type   Odds Ratio   SE(Odds)   Type   Lower   Unper	

PMLE: Penalized MLE for bias correction (Firth's method).

Analysis Time = 00:00:00

This OR of 3.8 is substantial. It is 0.6 units below the OR we originally reported (see Figure 2), but the lower CI on the OR is also 0.2 units above the lower CI reported in the paper (1.2) and the associated exact p-value is lower (= 0.007).

## 11.1.3 Random Effects Logistic Regression on Combined CZA-FDA Measure, dropping Erroneous CZA Positive Codings, stop coding August 2005

#### Results for FDA Deadline Approval Measure highlighted in yellow.

. xtlogit CZA i(discode) or	FDAbbw_august	2005_correct	ted subyı	ctr pred01	fda hhospd	isc hhosleng,	re
Random-effects	logistic reg	ression		Number of	obs =	314	
Group variable	(i): discode	!		Number of	groups =	135	
Random effects	u i ~ Gaussi	an		Obs per q	roup: min =	1	
	_					2.3	
					_	16	
				Wald chi2	(4) =	9.24	
Log likelihood	= -68.88941	7		Prob > ch	(4) = = = = = = = = = = = = = = = = = = =	0.0555	
CZAFDAbbw_~d	OR	Std. Err.	 Z	P>   z	 [95% Conf.	Interval]	
subvrctr	.9430763	.0865306	 -0.64	0.523	 .7878546	1.128879	
pred01fda	5.729092	3.407662	2.93	0.003	1.78563	18.38147	
hhospdisc	1.000001	2.09e-06	0.37	0.709	.9999967	1.000005	
	1.051545						
 /lnsig2u	.848002	.4883908			1092263	1.80523	
 sigma u	1.528063	.3731459			.9468514	2.466044	
	.4151184				.2141526		
 Likelihood-rat	io test of rh	.o=0: chibar2	2(01) =	5.55 Pr	ob >= chiba	r2 = 0.009	

.

This OR is larger than that reported in the paper (5.7 vs 4.4), and the lower limit of the CI is higher.

#### 11.2.1 Combined Measure with Corrections, Drop Entecavir and Tiprinavir from Measure.

We then generated a more conservative combined measure, dropping 2 more NMEs that are clearly postmarket BBWs but for which *Physicians' Desk Reference* has labeling revision after our data collection stopped in August 2005. Put differently, we recoded Entecavir and Tiprinavir as <u>not</u> having had BBWs added postmarket, even though BBWs have been added.

```
. gen CZAFDAbbw_august2005_corrected2 = CZAFDAbbw_august2005_corrected
. replace CZAFDAbbw_august2005_corrected2 = 0 if(ndanum_fda == 21797)
(1 real change made) {THIS IS ENTECAVIR}
. replace CZAFDAbbw_august2005_corrected2 = 0 if(ndanum_fda == 21814)
(1 real change made) {THIS IS TIPRINAVIR}
```

. tab CZAI	FDAbbw_august2005_	correcte	ed2 pred01fda,	exact	
CZAFDAbbw_ august2005 corrected					
_ 2	0	1	Total		
0	217	78   10	295 19		
Total	226	88	314		
	Fisher's exact =		0.031		
1-sided	Fisher's exact =		0.017	1	

To generate a shorter name for this variable, to be used in LogXact8 (which has max 9 characters in variable names), we used the following code:

```
. gen bbw0805 corr2 = CZAFDAbbw_august2005_corrected2
```

This variable is used as the dependent variable in the exact logistic regressions on the following page.

## 11. 2. 2. Exact Logistic Regression on Combined Measure with Corrections, omitting Entecavir and Tiprinavir from Measure, Using FDA's NMEs and FDA's Deadline Approval Measure

#### **Binary Regression**

regression (type=logit, model(bbw0805\_co = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes,

#### **Basic Information**

Data file BBWchecks-20080528.csv

Model bbw0805\_co(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable <Not Specified>
Stratum variable <Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	26.33	21	0.1943
Likelihood Ratio	298.4	3	1.581e-009

#### Parameter Estimates

		Point Estima	ate	Confidence Interval and P-Value for Odds Ratio				
					95 %CI		2*1-sided	
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower Upper P-		P-Value	
%Const	PMLE	0.179	NA	Asymptotic	0.01242	2.58	0.2063	
				Asymptotic(Profile)	0.01187	2.754		
subyrctr	PMLE	0.922	NA.	Asymptotic	0.7901	1.076	0.3021	
				Asymptotic(Profile)	0.7825	1.073		
	CMLE	0.9175	NA	Exact	0.7753	1.075	0.3052	
pred01fda	PMLE	3,383	NA	Asymptotic	1.333	8.588	0.01034	
				Asymptotic(Profile)	1.322	8.791		
	CMLE	3.406	NA.	Exact	1.169	10.11	0.02308	

PMLE: Penalized MLE for bias correction (Firth's method).

#### Analysis Time = 00:00:00

The OR is still above 3 and is still statistically significant using an exact distribution. However, it is 1 unit lower than reported in the paper, and the lower bound of the CI is 0.05 units lower.

## 11. 2. 3. Random-Effects Logistic Regression on Combined Measure with Corrections, Drop Entecavir and Tiprinavir from Measure, Using FDA's NMEs and FDA's Deadline Approval Measure

Results for FDA Deadline Approval Measure highlighted in yellow.

ndom-offoata	logistic ro	rroccion		Number	of obs	_	21/
andom-effects roup variable							314 135
roup variable	(1). 0150000	-		Number c	or groups		133
andom effects	u i ~ Gauss:	ian		Obs per	group: mi	n =	1
	_				av	g =	2.3
					ma	x =	16
				Wald chi	2(4)	=	8.17
	66 10001	= 7					
og likelihood	= -66.10893	O /		Prob > c	chi2	=	0.0857
og likelinood	= -66.10893	) <i> </i>		Prob > c	chi2	=	0.0857
og 11ke11n00d bw0805_co~2							
 bw0805_co~2   +-	OR	Std. Err.		P>   z	[95% Co	 nf.	Interval]
bw0805_co~2   + subyrctr	OR .8884057	Std. Err. .0830225	-1.27	P> z  0.205	[95% Co	 nf. 	Interval] 1.066982
bw0805_co~2  + subyrctr   pred01fda   hhospdisc	OR8884057 4.525702 1.000001	Std. Err0830225 2.597366 1.86e-06	-1.27 2.63 0.41	P> z  0.205 0.009 0.680	[95% Co .73971 1.46949 .999997	 nf.  7 8	Interval] 1.066982 13.93808 1.000004
bw0805_co~2  + subyrctr   pred01fda   hhospdisc	OR8884057 4.525702	Std. Err0830225 2.597366 1.86e-06	-1.27 2.63 0.41	P> z  0.205 0.009 0.680	[95% Co .73971 1.46949 .999997	 nf.  7 8	Interval] 1.066982 13.93808 1.000004
bw0805_co~2   subyrctr   pred01fda   hhospdisc   hhosleng	OR8884057 4.525702 1.000001	Std. Err0830225 2.597366 1.86e-06 .0559989	-1.27 2.63 0.41	P> z  0.205 0.009 0.680 0.369	.73971 1.46949 .999997 .94489	 nf.  7 8 1 8	Interval] 1.066982 13.93808 1.000004
bw0805_co~2   subyrctr   pred01fda   hhospdisc   hhosleng	OR .8884057 4.525702 1.000001 1.049108	Std. Err0830225 2.597366 1.86e-06 .0559989 .5647479	-1.27 2.63 0.41	P> z  0.205 0.009 0.680 0.369	.95% Co .73971 1.46949 .999997 .94489	 nf. 7 8 1 8 	1.066982 13.93808 1.000004 1.16481

This OR of 4.53 is very close to (though slightly larger than) that reported in Figure 2 of the paper and the lower bound of the CI is also slightly larger than that reported in Figure 2 of the paper.

#### 12. Results for Combined "Withdrawal or Black-Box Warning" Measure

### **12.1.1.** Generate Combined SBW and BBW Measures, Using Corrected BBW Measures as Above.

```
. gen sbworbbw_fda0805andCZA_correct1 = 0
. replace sbworbbw_fda0805andCZA_correct1 = 1 if( sbwfdas05 == 1)
(10 real changes made)
. replace sbworbbw_fda0805andCZA_correct1 = 1 if( sbwcza == 1)
(3 real changes made)
. replace sbworbbw_fda0805andCZA_correct1 = 1 if(
CZAFDAbbw_august2005_corrected == 1)
(14 real changes made)
```

This produces the following list of drugs that had either a safety-based withdrawal or a postmarket black-box warning.

. list	t ndanum_fda	<pre>genernam_fda pred01fda if(sbworbbw_</pre>	_fda0805and0	CZA_correct1 =	== 1)
	ndanum~a	genernam_fda	pred01~a	<del>-</del>   	
6.   7.   13.   42.   47.	21449   20760   21107   20535   20896	Adefovir Dipivoxil Alatrofloxacin Mesylate Alosetron Hydrochloride Bromfenac Sodium Capecitabine	1   1   0   0   1		
52.   53.   59.   66.   82.	20822	Celecoxib Cerivastatin Sodium Citalopram Hydrobromide Danaparoid Sodium Duloxetine Hydrochloride	1   1   0   0		
87.   90.   129.   152.   160.	20695 20564	Emtricitabine Entecavir Grepafloxacin Hydrochloride Lamivudine Levomethadyl Acetate Hydrochloride	1   1   1   1   0		
172.   177.   188.   226.   238.	20169 20984	Mibefradil Dihydrochloride Mirtazapine Nilutamide Rapacuronium Bromide Rofecoxib	0   0   0   0   1		
245.   269.   278.   281.   292.	21814   20697	Saquinavir Tenofovir Disoproxil Fumarate Tipranavir Tolcapone Troglitazone	0   1   1   0   1		
294.   299.	   20759   21341 +	Trovafloxacin Mesylate Valdecoxib	1   1   1	 	

### **12.1.2.** Cross-Tabulation with Fisher's Exact Test – Combined SBW and BBW Measures, Using Corrected BBW Measures, Using FDA's NMEs and FDA's Deadline Approval Measure

```
sbworbbw fda0805andCZA correct1 pred01fda, exact
. tab
sbworbbw f |
da0805andC |
ZA correct |
                  pred01fda
        0 |
                  214
                              73 |
                                         287
        1 |
                              15 I
                                          27
                   12
    Total |
                  226
                              88 |
                                         314
          Fisher's exact =
                                           0.003
  1-sided Fisher's exact =
                                           0.002
```

To generate a shorter name for this variable, to be used in LogXact8 (which has max 9 characters in variable names), we used the following code:

```
. gen combcom1 = sbworbbw fda0805andCZA correct1
```

This variable is used as the dependent variable in the exact logistic regressions on the following page.

### 12.1.3. Exact Logistic Regression – Combined SBW and BBW Measures, Using Corrected BBW Measures, Using FDA's NMEs and FDA's Deadline Approval Measure

#### Binary Regression

regression (type=logit, model(combcom1 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, cuttive==odds);

#### **Basic Information**

Data file combocorrectSBWorBBW-20080513.csv

Model combcom1(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable <Not Specified>
Stratum variable <Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	26.22	21	0.1982
Likelihood Ratio	262	3	1.093e-009

#### **Parameter Estimates**

		Point Estimate Confidence Interval and P-Value for Odds		s Ratio			
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Upper	P-Value
%Const	PMLE	0.1681	NA.	Asymptotic	0.01688	1.675	0.1284
				Asymptotic(Profile)	0.0164	1.707	
subyrctr	PMLE	0.941	NA.	Asymptotic	0.8251	1.073	0.365
				Asymptotic(Profile)	0.8208	1.071	
	CMLE	0.9384	NA	Exact	0.8161	1.073	0.3692
pred01fda	PMLE	3.898	NA	Asymptotic	1.737	8.747	0.0009725
				Asymptotic(Profile)	1.743	8.908	
	CMLE	3.932	NA.	Exact	1.602	9.883	0.002076
M.E. D		0.50-	NA				0.002076

PMLE: Penalized MLE for bias correction (Firth's method).

Analysis Time = 00:00:00

This OR of 3.9 is slightly lower than that reported in Figure 2 of the paper; the lower bound of the CI is slightly higher than reported in Figure 2 of the paper (that is, the result is more statistically significant than the result originally reported).

### 12.1.4. Random-Effects Logistic Regression – Combined SBW and BBW Measures, Using Corrected BBW Measures, Using FDA's NMEs and FDA's Deadline Approval Measure

Results for FDA Deadline Approval Measure in yellow.

-	•					
_i ~ Gaussi	Lan		Obs per	avg =	2.3	
= -83.82037	<i>1</i> 5		Wald chi Prob > c	2(4) = hi2 =	11.33	
OR	Std. Err.	 Z	P> z	[95% Conf.	Interval]	
5.101227	2.574711	3.23	0.001	1.896937	13.71818	
1.000001	1.63e-06	0.85	0.394	.9999982	1.000005	
.4489665	.4762334			4844338	1.382367	
	(i): discode  a_i ~ Gaussi  = -83.82037  OR  .9112509  5.101227  1.000001  1.01765  .4489665	OR Std. Err.	OR Std. Err. z  .9112509 .0737148 -1.15  5.101227 2.574711 3.23 1.000001 1.63e-06 0.85 1.01765 .0567741 0.31  .4489665 .4762334	OB Std. Err. z P> z   9112509 .0737148 -1.15 0.251  5.101227 2.574711 3.23 0.001  1.000001 1.63e-06 0.85 0.394  1.01765 .0567741 0.31 0.754  .4489665 .4762334  1.251676 .2980449	Number of groups	Number of groups = 135  A_i ~ Gaussian  Obs per group: min = 1

This OR of 5.1 is slightly higher than that reported in Figure 2 of the paper; the lower bound of the CI is slightly higher than reported in Figure 2 of the paper (that is, the result is more statistically significant than the result originally reported).

## **12.2.1** Generate Combined SBW and BBW Measures, Using Corrected BBW Measures, omitting Entecavir and Tiprinavir from Measure, Using FDA's NMEs and FDA's Deadline Approval Measure

```
. gen sbworbbw_fda0805andCZA_correct2 = 0
. replace sbworbbw_fda0805andCZA_correct2 = 1 if(sbwfdas05 == 1)
(10 real changes made)
. replace sbworbbw_fda0805andCZA_correct2 = 1 if(sbwcza == 1)
(3 real changes made)
. replace sbworbbw_fda0805andCZA_correct2 = 1
if(CZAFDAbbw_august2005_corrected2 == 1)
(12 real changes made)
```

. list ndanum\_fda genernam\_fda pred01fda if(sbworbbw\_fda0805andCZA\_correct2
== 1)

-	+			+
	ndanum~a	genernam_fda	pred01~a	pred01~A
6.	   21449	Adefovir Dipivoxil	1	1
7.	20760	Alatrofloxacin Mesylate	1	1
13.	21107	Alosetron Hydrochloride	0	0
42.	20535	Bromfenac Sodium	0	0
47.	20896	Capecitabine	1	1
52.	   20998	Celecoxib	 1	1
53.	20740	Cerivastatin Sodium	1	1
59.	20822	Citalopram Hydrobromide	0	0
66.	20430	Danaparoid Sodium	0	0
82.	21427	Duloxetine Hydrochloride	0	0
87.	   21500	Emtricitabine	1	1
129.	20695	Grepafloxacin Hydrochloride	1	1
152.	20564	Lamivudine	1	1
160.	20315	Levomethadyl Acetate Hydrochloride	0	0
172.	20689	Mibefradil Dihydrochloride	0	0
177.	   20415	Mirtazapine	0	0
188.	20169	Nilutamide	0	0
226.	20984	Rapacuronium Bromide	0	0
238.	21042	Rofecoxib	1	1
245.	20628	Saquinavir	0	0
269.	   21356	Tenofovir Disoproxil Fumarate	 1	 1
281.	20697	Tolcapone	0	0
292.	20720	Troglitazone	1	1
294.	20759	Trovafloxacin Mesylate	1	1
299.	21341	Valdecoxib	1	1
-	+			+

**12.2.2.** Cross-Tabulation with Fisher's Exact Test – Combined SBW and BBW Measures, Using Corrected BBW Measures, Drop Entecavir and Tiprinavir from Measure, Using FDA's NMEs and FDA's Deadline Approval Measure

```
. tab sbworbbw fda0805andCZA correct2 pred01fda, exact
sbworbbw f |
da0805andC |
ZA correct |
               pred01fda
       2 |
                           1 |
               214
       0 |
                          75 |
                                   289
                                   25
       1 |
               12
                          13 |
               226
    Total |
                          88 |
                                   314
         Fisher's exact = 0.009
  1-sided Fisher's exact =
                                     0.007
```

To generate a shorter name for this variable, to be used in LogXact8 (which has max 9 characters in variable names), we used the following code:

```
. gen WOW0805_corr2 = sbworbbw_fda0805andCZA_correct2
```

This variable is used as the dependent variable in the exact logistic regressions on the following page.

**12.2.3.** Exact Logistic Regression – Combined SBW and BBW Measures, Using Corrected BBW Measures, Drop Entecavir and Tiprinavir from Measure, Using FDA's NMEs and FDA's Deadline Approval Measure

#### **Binary Regression**

regression (type=logit, model(wow0805\_co = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtyne=ordic):

#### **Basic Information**

Data file BBWchecks-20080528.csv

Model WOW0805\_co(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable < Not Specified>
Stratum variable <Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	26.43	21	0.1906
Likelihood Ratio	270.8	3	1.396e-009

#### Parameter Estimates

		Point Estima	ate	Confidence Interval and P-Value for Odds Ratio			
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unper	P-Value
%Const	PMLE	0.4475	NA	Asymptotic	0.04026	4.974	0.5129
				Asymptotic(Profile)	0.04	5.289	
subyrctr	PMLE	0.8876	NA	Asymptotic	0.7704	1.023	0.0988
				Asymptotic(Profile)	0.7641	1.019	
	CMLE	0.8834	NA	Exact	0.7581	1.02	0.0929
pred01fda	PMLE	3,559	NA.	Asymptotic	1.543	8.207	0.002899
				Asymptotic(Profile)	1.54	8.33	
	CMLE	3.579	NA.	Exact	1.401	9.266	0.006492

PMLE: Penalized MLE for bias correction (Firth's method).

#### Analysis Time = 00:00:00

This OR of 3.6 is 0.8 units lower than that reported in Figure 2 of the paper; the lower bound of the CI is nearly identical with the one reported in Figure 2 of the paper.

**12.2.4.** Random Effects Logistic Regression – Combined SBW and BBW Measures, Using Corrected BBW Measures, Drop Entecavir and Tiprinavir from Measure, Using FDA's NMEs and FDA's Deadline Approval Measure

. $xtlogit WOW0805\_corr2 subyrctr pred01fda hhospdisc hhosleng, re i(discode) or$									
Random-effects logistic regression Number of obs = 314									
135	of groups =	Number		Э	(i): discode	roup variable			
1	Random effects u i ~ Gaussian Obs per group: min =								
2.3		_			_				
16									
10.84	i2(4) =	Wald ch							
	chi2 =			39	= -80.77063	og likelihood			
Interval]	[95% Conf.	P> z	Z	Std. Err.	OR	OW0805_co~2			
1.017398	.7375022	0.080	-1.75	.0710962	.8662177	subyrctr			
1.017398 10.88476	.7375022 1.624271	0.080	 -1.75 2.96	.0710962 2.040519	.8662177 4.204735	subyrctr   pred01fda			
1.017398 10.88476	.7375022	0.080	 -1.75 2.96	.0710962 2.040519	.8662177 4.204735	+ subyrctr   pred01fda			
1.017398 10.88476 1.000004	.7375022 1.624271	0.080 0.003 0.387	-1.75 2.96 0.87	.0710962 2.040519 1.46e-06	.8662177 4.204735 1.000001	subyrctr   pred01fda   hhospdisc			
1.017398 10.88476 1.000004 1.121039	.7375022 1.624271 .9999984	0.080 0.003 0.387 0.701	-1.75 2.96 0.87 0.38	.0710962 2.040519 1.46e-06 .049659	.8662177 4.204735 1.000001	subyrctr   pred01fda   hhospdisc   hhosleng			
1.017398 10.88476 1.000004 1.121039	.7375022 1.624271 .9999984 .9260825	0.080 0.003 0.387 0.701	-1.75 2.96 0.87 0.38	.0710962 2.040519 1.46e-06 .049659 	.8662177 4.204735 1.000001 1.018908	subyrctr   pred01fda   hhospdisc   hhosleng			

.

This OR of 4.2 is slightly lower than that reported in Figure 2 of the paper; the lower bound of the CI is slightly above that reported in Figure 2 of the paper.

### 12.3 Combined Measure Accepting <u>All</u> of FDA's Positive Coding Decisions and Coding Through December 2007

We then took data through December 2007, combined CZA and FDA Measures for both SBWs and BBWs, and dropped erroneous BBW codings from Carpenter-Zucker-Avorn (fondaparinux sodium, oxaliplatin, tinidazole).

```
. gen CZAFDAcomboSBWorBBW dec2007 = 0
. replace CZAFDAcomboSBWorBBW dec2007 = 1 if(sbwcza == 1)
(11 real changes made) {This adds the 11 NMEs with SBWs from the original
paper.}
. replace CZAFDAcomboSBWorBBW dec2007 = 1 if(sbwfda08 == 1)
(3 real changes made) {This adds the 3 NMEs coded as SBWs by FDA that are not
in the CZA list - Zelnorm, Orlaam, and alosetron.
. replace CZAFDAcomboSBWorBBW dec2007 = 1 if(bbwfda08 == 1)
(22 real changes made) {This adds the 22 NMEs that FDA has as BBWS that are
not already coded as SBWs.}
. replace CZAFDAcomboSBWorBBW dec2007 = 1 if(bbwcza == 1)
(8 real changes made) {This adds the NMEs that CZA originally coded as BBWs
that were not already coded as SBWS and which the FDA had not coded as BBWs.}
. replace CZAFDAcomboSBWorBBW dec2007 = 0 if(ndanum fda == 21345)
(1 real change made) { THIS DROPS FONDAPARINUX SODIUM }
. replace CZAFDAcomboSBWorBBW dec2007 = 0 if(ndanum fda == 21618)
(1 real change made) {THIS DROPS TINIDAZOLE}
. replace CZAFDAcomboSBWorBBW dec2007 = 0 if(ndanum fda == 21492)
(1 real change made) {THIS DROPS OXALIPLATIN}
```

. list  $ndanum_fda genernam_fda pred01fda if(CZAFDAcomboSBWorBBW_dec2007 == 1)$ 

+----+

	ndanum~a	genernam_fda	pred01~a
6. 7. 13. 23. 27.	21449   20760   21107   21436   21411	Adefovir Dipivoxil Alatrofloxacin Mesylate Alosetron Hydrochloride Aripiprazole Atomoxetine Hydrochloride	1   1   0   0   0
42. 47. 52. 53.		Bromfenac Sodium Capecitabine Celecoxib Cerivastatin Sodium Citalopram Hydrobromide	0   1   1   1   0

66.	20430	Danaparoid Sodium	0
82.	21427	Duloxetine Hydrochloride	0
87.	21500	Emtricitabine	1
90.	21797	Entecavir	1
119.	20937	Gadoversetamide	0
129.	20695	Grepafloxacin Hydrochloride	1
152.	20564	Lamivudine	1
160.	20315	Levomethadyl Acetate Hydrochloride	0
167.	20938	Meloxicam	0
172.	20689	Mibefradil Dihydrochloride	0
177.	20415	Mirtazapine	0
188.	20169	Nilutamide	0
193.	20592	Olanzapine	0
213.	21064	Perflutren Lipid Microsphere	0
214.	21302	Pimecrolimus	0
215.	21073	Pioglitazone Hydrochloride	1
222.	20639	Quetiapine Fumarate	0
224.	20815	Raloxifene Hydrochloride	1
226.	20984	Rapacuronium Bromide	0
238.	21042	Rofecoxib	1
241.	21071	Rosiglitazone Maleate	1
245.	20628	Saquinavir	0
265.	21200	Tegaserod Maleate	0
266.	21144	Telithromycin	0
269.	21356	Tenofovir Disoproxil Fumarate	1
278.	21814	Tipranavir	1
281.	20697	Tolcapone	0
292.	20720	Troglitazone	1
294.	20759	Trovafloxacin Mesylate	1
299.	21341	Valdecoxib	1
311.	20825	Ziprasidone Hydrochloride	0   

.

Adopting this measure accepts virtually all the assumptions made in the FDA's April 2008 argument. All of the FDA's April 2008 coding decisions are used, the coding is not stopped in August 2005 but runs through December 2007 (long after analysis had stopped for the original paper), and only eight additional drugs (adefovir pipivoxil, alatrofloxacin, emtricitabine, entecavir, tenofovir disoproxil, tipranavir, tolcapone, and trovafloxacin) are coded as having had postmarket regulatory events.

Even with all of these assumptions made, there is a statistically significant difference between just-before-deadline approvals and all other drugs in the postmarket regulatory problems they experience.

**12.3.1** Cross Tabulation with Fisher's Exact Test – Taking data up through December 2007, Combine CZA and FDA Measures for both SBWs and BBWs, and dropping CZA's erroneous BBW codings (fondaparinux sodium, oxaliplatin, tinidazole).

. tab CZAFDAcomboSBWorBBW_dec2007 pred01fda, exact										
CZAFDAcomb oSBWorBBW_ dec2007	pred01fda	1	Total							
0	203	70	273 41							
Total	·	88	314							
	Fisher's exact = Fisher's exact =		0.0							

To generate a shorter name for this variable, to be used in LogXact8 (which has max 9 characters in variable names), we used the following code:

```
. gen SBWBBWd07 = CZAFDAcomboSBWorBBW dec2007
```

This variable is used as the dependent variable in the exact logistic regressions on the following page.

**12.3.2 Exact Logistic Regression** – **Taking data up through December 2007**, Combine CZA and FDA Measures for both SBWs and BBWs, and dropping CZA's erroneous BBW codings (fondaparinux, oxaliplatin, tinidazole).

#### **Binary Regression**

regression (type=logit, model(sbwbbwd07 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtype=odds):

#### **Basic Information**

Data file combocorrectSBWorBBW-dec2007-20080513.csv

Model SBWBBWd07(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable <Not Specified>
Stratum variable <Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	26.22	21	0.1982
Likelihood Ratio	197.4	3	9.234e-010

#### **Parameter Estimates**

tranicia Estinaces										
		Point Estima	ate	Confidence Interval and P-Value for Odds Ratio						
					95	%CI	2*1-sided			
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unper	P-Value			
%Const	PMLE	0.1371	NA	Asymptotic	0.02025	0.928	0.0417			
				Asymptotic(Profile)	0.01978	0.92				
subyrctr	PMLE	0.9908	NA	Asymptotic	0.8897	1.103	0.8657			
				Asymptotic(Profile)	0.8881	1.102				
	CMLE	0.9897	NA	Exact	0.8854	1.104	0.8777			
pred01fda	PMLE	2,295	NA	Asymptotic	1.162	4.533	0.01673			
				Asymptotic(Profile)	1.155	4.52				
	CMLE	2.287	NA.	Exact	1.079	4.798	0.02995			

PMLE: Penalized MLE for bias correction (Firth's method).

#### Analysis Time = 00:00:03

Even using this measure, which again as constructed accepts all of the coding decisions in the FDA's April 2008 argument, there is a statistically significant difference between deadline approvals and other NMEs. The OR is smaller than reported in the paper (2.3 versus 4.4), but it is statistically significant with an exact p-value of 0.03. However, it is not clear that all of the assumptions made by Nardinelli et al should be accepted, for reasons stated above.

## **12.3.3 Random-Effects Logistic Regression – Taking data up through December 2007**, Combine CZA and FDA Measures for both SBWs and BBWs, and dropping CZA's erroneous BBW codings (fondaparinux sodium, oxaliplatin, tinidazole).

#### Results for FDA Deadline Approval Measure highlighted in yellow.

. xtlogit i(discode)		AcomboSBWor	BBW_dec2007	subyrctr	pred01fda	hhospdisc	hhosleng,	re
Random-effe Group varia		-				obs = groups =		
Random effe	ects u	_i ~ Gaussi	an		Obs per g	_	1 2.3 16	
Log likelih	ıood	= -111.6155	9		Wald chi2 Prob > chi	(4) = i2 =	8.00 0.0916	
CZ~W_dec200	7	OR	Std. Err.	z	P> z	[95% Conf.	Interval]	
subyrct	r	.9654842	.069295	-0.49	0.625	.8387886	1.111317	
pred01fd	la	3.425254	1.649158	2.56	0.011	1.333106	8.800779	
hhospdis	c	1.000002	1.77e-06	1.08	0.280	.9999984	1.000005	
hhoslen	ıg	1.030768	.0606853	0.51	0.607	.9184329	1.156843	
/lnsig2	u	1.143569	.4641163			.2339177	2.05322	
sigma	u	1.771425	.4110737			1.124073	2.791587	
		.4881826				.2774933		
 Likelihood-	ratio	test of rh	 o=0: chibar2	2(01) =	10.92 Pro	ob >= chiba	r2 = 0.000	

Again using a combined measure which accepts all of the positive coding decisions in the FDA's April 2008 argument, there is a statistically significant difference between deadline approvals and other NMEs. The OR is one unit less than reported in the paper (3.4 versus 4.4), but is large (>3) and is statistically significant with a p-value of 0.01.

# 13. Re-Analysis of Association between Deadline Approvals and Black-Box Warnings, Using FDA NMEs, FDA Deadline Approval Measure, and Nardinelli et al Black-Box Warning Measures

In this section we examine what happens to the analysis when we rely entirely on the FDA's BBW codings but stop coding earlier. We use the 314 NMEs posted in the FDA's April 2008 data, combined with the FDA's "deadline approval" measure, and re-examine the pattern of postmarket black-box warnings. For the sake of hypothesis and demonstration, we do not correct here for clear FDA errors identified in Section 9 (above). For reference purposes, here is the replicated cross-tabulation result from the FDA's communication.

. tab bbwf	da08 pr	ed01fda,	exact		
bbwFDA08	 	pred01fda	a 1	Total	
0 1	•	207 19	78 10	+	
Total		226	88	314	
1-sided	Fisher's exact = 1-sided Fisher's exact =			-	.395 .270

We now discuss the variable results of statistical analysis when coding of these warnings is stopped in August 2005 (when we stopped coding) or after Vioxx (when it appears that the FDA began to change its policy for issuing black-box warnings).

#### 13.2 Using the FDA's BBW Measure, Stopping Coding in August 2005

The following is the list of NMEs with a post-market BBW as coded by the FDA, if coding is stopped in August 2005. Notice that we have not recoded saquinavir (see Section 9.2 above).

. list	ndanum_fc	da genernam_fda bbwdate_fo	da if( bbwfdas0	5 == 1)
+	ndanum~a	genernam_fda	+ bbwdate~a	
2.	20998	Celecoxib	29-Jul-05	
3.	21341	Valdecoxib	24-Nov-04	
14.	20535	Bromfenac Sodium	31-Mar-98	
33.	20896	Capecitabine	7-Sep-01	
54.	20169	Nilutamide	29-Sep-00	
104.	20430	Danaparoid Sodium	30-Jan-98	
108.	20564	Lamivudine	15-Dec-97	
114.	20628	Saquinavir	14-Nov-00	
137.	20759	Trovafloxacin Mesylate	9-Jun-99	
138.	20760	Alatrofloxacin Mesylate	9-Jun-99	
175.	20822	Citalopram Hydrobromide	18-Feb-05	
176.	20415	Mirtazapine	12-Jan-05	
177.	21427	Duloxetine Hydrochloride	18-Feb-05	
221.	20697	Tolcapone	16-Nov-98	
232.	20720	Troglitazone	15-Dec-97	
284.	21107	Alosetron Hydrochloride	7-Jun-02	

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#### 13.2 Using the FDA's BBW Measure, Stopping Coding in August 2005 (cont.)

If we then <u>assume</u> that the FDA's coding is correct (and we have already argued that it is not reliable; see Sections 9 and 10), and we analyze the warnings data as stopped in August 2005, we get the following two-way tabulation.

. tab bbwf	fdas05 pred01fda,		
bbwFDAs05	0	1	Total
0 1	217	81   7	298 16
Total	226	88	314
	Fisher's exact =		0.16
1-sided	Fisher's exact =		0.12

This two-way tabulation yields an exact p-value of 0.16, though the exact p-value is less than half that of the original FDA calculation (0.395). An exact logistic regression yields an odds ratio of 2.6 with a p-value of 0.14. However, a random-effects logistic regression analysis yields an odds-ratio of 3.4 with a p-value of less than 0.05.

### 13.2 Exact Logistic Regression Analysis of Nardinelli BBW Measure, stopping Coding in August 2005

#### Binary Regression

regression (type=logit, model(bbwfdas05 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtwe=odds);

#### **Basic Information**

Data file bbwdump20080412.csv

Model bbwfdas05(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable <Not Specified>
Stratum variable <Unstratified>
Analysis type Estimate :: Exact

Analysis type Estin
Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	21.16	21	0.4492
Likelihood Ratio	314.6	3	1.619e-009

#### **Parameter Estimates**

Point Estimate			Confidence Interval and P-Value for Odds Ratio			
			95 %CI		2*1-sided	
Type	Odds Ratio	SE(Odds)	Type	Lower	Unner	P-Value
PMLE	0.7845	NA.	Asymptotic	0.03945	15.6	0.8736
			Asymptotic(Profile)	0.03923	18.79	
PMLE	0.843	NA.	Asymptotic	0.7037	1.01	0.06369
			Asymptotic(Profile)	0.6906	1.003	
CMLE	0.8347	NA.	Exact	0.6798	1.004	0.05556
PMLE	2,577	NA.	Asymptotic	0.9367	7.092	0.06675
			Asymptotic(Profile)	0.9049	7.15	
CMLE	2.552	NA.	Exact	0.7593	8.268	0.1396
	PMLE PMLE CMLE PMLE	Type         Odds Ratio           PMLE         0.7845           PMLE         0.843           CMLE         0.8347           PMLE         2.577	Type         Odds Ratio         SE(Odds)           PMLE         0.7845         NA           PMLE         0.843         NA           CMLE         0.8347         NA           PMLE         2.577         NA	Type         Odds Ratio         SE(Odds)         Type           PMLE         0.7845         NA         Asymptotic           Asymptotic (Profile)         Asymptotic           PMLE         0.843         NA         Asymptotic           CMLE         0.8347         NA         Exact           PMLE         2.577         NA         Asymptotic           Asymptotic(Profile)         Asymptotic(Profile)	Type	Type   Odds Ratio   SE(Odds)   Type   Lower   Unper

PMLE: Penalized MLE for bias correction (Firth's method).

#### Analysis Time = 00:00:00

#### 13.2 Using the Nardinelli et al BBW Measure, Stopping Coding in August 2005 (cont.)

A random-effects logistic regression that controls for submission year and epidemiological covariates yields an odds ratio of 3.4 which is statistically significant at the 0.05 level. The estimates for the FDA's deadline approval measure are highlighted in yellow.

	Number o	of obs	= 314	
	Number o	of groups	= 135	
	Obs per			
		avg	= 2.3	
		max	= 16	
	Wald chi	12(4)	= 7.39	
z 	P> z	[95% Conf	. Interval]	
-2.03	0.042	.6414047	.9923188	
		.6414047 1.013789		
1.98	0.047		11.36727	
1.98 0.45	0.047 0.656	1.013789	11.36727 1.000005	
1.98 0.45	0.047 0.656	1.013789 .9999971	11.36727 1.000005 1.161393	
1.98 0.45	0.047 0.656	1.013789 .9999971 .9344825 	11.36727 1.000005 1.161393	
		Number of Obs per  Wald chirther of the Prob > o	Number of groups  Obs per group: min avg max  Wald chi2(4)  Prob > chi2	Number of groups = 135

Hence in a controlled statistical analysis, there is a large and statistically differentiable relationship *even if we rely entirely on the problematic data posted by Nardinelli et al.* 

## 13.3 Using the Nardinelli et al BBW Measure, Stopping Coding after Vioxx withdrawal (cont.)

If we now use the FDA's BBW measure but stop coding in January 2005 (equivalently, after the Vioxx withdrawal), we get the following list.

. list	ndanum_fo	da genernam_fda bbwdate_fo	da if( bbwfd
-	ndanum~a	genernam_fda	bbwdate~a
3.	   21341	Valdecoxib	24-Nov-04
14.	20535	Bromfenac Sodium	31-Mar-98
33.	20896	Capecitabine	7-Sep-01
54.	20169	Nilutamide	29-Sep-00
104.	20430	Danaparoid Sodium	30-Jan-98
108.	   20564	Lamivudine	15-Dec-97
114.	20628	Saquinavir	14-Nov-00
137.	20759	Trovafloxacin Mesylate	9-Jun-99
138.	20760	Alatrofloxacin Mesylate	9-Jun-99
221.	20697	Tolcapone	16-Nov-98
232.	   20720	Troglitazone	15-Dec-97
284.	21107	Alosetron Hydrochloride	7-Jun-02
+	·		

## 13.3 Using the Nardinelli et al BBW Measure, Stopping Coding After Vioxx Withdrawal (cont.)

. tab bbwi	fdaw05 pred01fda,	exact		
bbwFDAw05	pred01fda	a 1	Total	
0 1	220   6	82   6	302 12	
Total	226	88	314	
	Fisher's exact =		0.103	
1-sided	Fisher's exact =		0.085	

This crude two-way cross-tabulation yields an exact p-value of greater than 0.10. An exact logistic regression with one control for submission year yields a *p*-value of 0.08. The more adequately controlled analysis is presented below.

#### 13.3 Using the Nardinelli et al BBW Measure, Stopping Coding After Vioxx Withdrawal (cont.)

#### **Binary Regression**

regression (type=logit, model(bbwfdaw05 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtype=odds).

#### **Basic Information**

bbwdump20080412.csv Data file

Model Link type bbwfdaw05(Response = 1)=%Const+subyrctr+pred01fda

Logit

Weight variable Stratum variable <Not Specified> <Unstratified> Analysis type Estimate :: Exact

Number of terms in model Number of term(s) dropped Number of observations in analysis 314 Number of records rejected 24 **Number of groups** 

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	22.35	21	0.3797
Likelihood Ratio	341.4	3	1.98e-009

#### **Parameter Estimates**

		Point Estima	ate	Confidence In	terval and P-\	/alue for Odd	s Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unper	P-Value
%Const	PMLE	1.526	NA	Asymptotic	0.04373	53.24	0.8156
				Asymptotic(Profile)	0.04459	77.33	
subyrctr	PMLE	0.789	NA	Asymptotic	0.6331	0.9834	0.03496
				Asymptotic(Profile)	0.6118	0.973	
	CMLE	0.7753	NA	Exact	0.5954	0.9725	0.0253
pred01fda	PMLE	3.513	NA	Asymptotic	1.121	11.01	0.03112
				Asymptotic(Profile)	1.089	11.44	
	CMLE	3.528	NA	Exact	0.8824	14.21	0.0775

PMLE: Penalized MLE for bias correction (Firth's method).

## 13.3 Using the Nardinelli et al BBW Measure, Stopping Coding after Vioxx withdrawal (cont.)

Yet again, analysis of the data in a random-effects logistic regression with controls for submission year and epidemiological covariates generates a p-value of less than 0.05. The estimates for the FDA's deadline approval measure are highlighted in yellow.

Random-effects		•			of obs			
Group variable	(i): discode	2		Number o	of groups	=	135	
Random effects	u_i ~ Gaussi	an		Obs per	group: m	in =	1	
					a	vg =	2.3	
					ma	ax =	16	
				Wald chi	12(4)	=	8.38	
Log likelihood	= -46.28369	91		Prob > c	chi2	=	0.0786	
bbwfdaw05	OR	Std. Err.	z	P> z	[95% C	onf.	Interval]	
subyrctr	.7729781	.0953021	-2.09	0.037	.60704	 48	.9842686	
subyrctr   pred01fda	.7729781 3.784404	.0953021 2.353177	-2.09 2.14	0.037	.60704	 48 <mark>97</mark>	.9842686 12.80214	
subyrctr   pred01fda	.7729781	.0953021 2.353177	-2.09 2.14	0.037	.60704	 48 <mark>97</mark>	.9842686 12.80214	
subyrctr   pred01fda   hhospdisc	.7729781 3.784404	.0953021 2.353177 1.65e-06	-2.09 2.14 0.44	0.037 0.032 0.659	.60704 1.11869	 48 <mark>97</mark> 75	.9842686 12.80214 1.000004	
subyrctr   pred01fda   hhospdisc   hhosleng	.7729781 3.784404 1.000001 1.046485	.0953021 2.353177 1.65e-06 .0468186	-2.09 2.14 0.44	0.037 0.032 0.659	.60704 1.1186 .99999 .9580	 48 <mark>97</mark> 75 63	.9842686 12.80214 1.000004	
subyrctr   pred01fda   hhospdisc   hhosleng   /lnsig2u	.7729781 3.784404 1.000001 1.046485	.0953021 2.353177 1.65e-06 .0468186	-2.09 2.14 0.44	0.037 0.032 0.659	.60704 1.1186 .99999 .9580	 48 <mark>97</mark> 75 63  84	.9842686 12.80214 1.000004 1.142391	

Hence in a controlled statistical analysis, there is a large (OR > 3) and statistically differentiable relationship even if we rely entirely on the FDA's posted data. This relationship is associated with a p = 0.08 estimate in an exact logistic regression and with a p = 0.03 estimate in a random-effects logistic regression. These p-values are far below those of the FDA analysis, and these analyses abandon our data entirely and assume that the FDA data are correct.

#### 13.4 Use of a Combined BBW Measure, with coding stopped August 1, 2005.

If we create a combined measure, coded 1 if the NME was coded *either* by Carpenter, Zucker and Avorn or as coded by Nardinelli et al as having a black-box warning added, but again stop coding at August 2005, we get the following.

```
. gen CZAFDAcombobbw_august2005 = 0
. replace CZAFDAcombobbw_august2005 = 1 if( bbwcza == 1)
(14 real changes made)
. replace CZAFDAcombobbw_august2005 = 1 if( bbwfdas05 == 1)
(10 real changes made)
```

This recoding produces the following list of BBWs.

		duces the following list of BBWs.	
. list	t ndanum_fo	da genernam_fda if( CZAFDAcombo	bbw_august2005 == 1)
-	+		•
	ndanum~a	genernam_fda 	
2.	20998	Celecoxib	
3.	21341	Valdecoxib	
14.	20535	Bromfenac Sodium	
33.	20896	Capecitabine	
42.	21492	Oxaliplatin	
54.	20169	Nilutamide	
104.	20430	Danaparoid Sodium	
106.	21345	Fondaparinux Sodium	
108.	20564	Lamivudine	
112.	21500	Emtricitabine	
114.	20628	Saquinavir	
120.	21356	Tenofovir Disoproxil Fumarate	
124.	21814	Tipranavir	
137.	20759	Trovafloxacin Mesylate	
138.	20760	Alatrofloxacin Mesylate	
175.	20822	Citalopram Hydrobromide	
176.	20415	Mirtazapine	
177.	21427	Duloxetine Hydrochloride	
221.	20697	Tolcapone	
232.	20720	Troglitazone	
284.	21107	Alosetron Hydrochloride	
300.	21797	Entecavir	
301.	21449	Adefovir Dipivoxil	
311.	21618	Tinidazole	
-	+		-

.

#### 13.4 Using a Combined Measure, with coding stopped August 1, 2005 (cont.)

If we then analyze this combined measure statistically, using the FDA's pre-deadline approval measure, we get the following results. A two-way cross tabulation with Fisher's exact test produces the following.

. tab CZAI	FDAcombobbw_augus	st2005 pred	l01fda, exa
CZAFDAcomb obbw_augus t2005	pred01fo	da 1   	Total
0	215   11	75   13	290 24
Total	·	88	314
	Fisher's exact		0.0
1-sided	Fisher's exact =	=	0.0

#### 13.4 Using a Combined Measure, with coding stopped August 1, 2005 (cont.)

An exact logistic regression yields the following.

#### **Binary Regression**

regression (type=logit, model(bbwcombo08 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=vec\_outtyne=odds).

#### **Basic Information**

Data file Model Link type bbwdump20080412.csv bbwcombo08(Response = 1)=%Const+subyrctr+pred01fda

Logit

<Not Specified> <Unstratified> Weight variable Stratum variable Analysis type Estimate :: Exact

Number of terms in model Number of term(s) dropped Number of observations in analysis 314 Number of records rejected Number of groups 24

**Summary Statistics** 

Statistics	Value	DF	P-Value
Deviance	25.25	21	0.2366
Likelihood Ratio	274	3	1.425e-009

#### **Parameter Estimates**

		Point Estima	ate	Confidence In	terval and P-1	Value for Odd	ls Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unper	P-Value
%Const	PMLE	0.02426	NA	Asymptotic	0.002139	0.2752	0.002688
				Asymptotic(Profile)	0.001956	0.2669	
subyrctr	PMLE	1.047	NA	Asymptotic	0.9169	1.195	0.4976
				Asymptotic(Profile)	0.9147	1.197	
	CMLE	1.046	NA	Exact	0.911	1.201	0.5304
pred01fda	PMLE	3.162	NA	Asymptotic	1.362	7.34	0.007376
				Asymptotic(Profile)	1.359	7.467	
	CMLE	3.182	NA	Exact	1.234	8.346	0.01499

PMLE: Penalized MLE for bias correction (Firth's method).

#### 13.4 Using a Combined Measure, with coding stopped August 1, 2005 (cont.)

A random-effects logistic regression with controls for epidemiological covariates produces the following. The estimates for the FDA's deadline approval measure are highlighted in yellow.

Random-effects	logistic red	gression		Number o	of obs	=	314
Group variable	(i): discode	2		Number o	of group	s =	135
Random effects	u_i ~ Gaussi	lan		Obs per	group:	min =	1
						avg =	2.3
						max =	16
				Wald chi	2(4)	=	9.01
Log likelihood	= -76.25782	2.4					0.0609
 CZAFDAc~2005	OR	Std. Err.	z	P> z	[95%	Conf.	Interval]
subyrctr	1.033437	.0869026	0.39	0.696	.8764	064	1.218603
subyrctr   pred01fda	1.033437 4.488479	.0869026 2.493757	0.39	0.696 0.007	.8764 1.510		1.218603 13.33579
subyrctr   pred01fda	1.033437	.0869026 2.493757	0.39	0.696 0.007	.8764 1.510		1.218603 13.33579
subyrctr   pred01fda   hhospdisc	1.033437 4.488479	.0869026 2.493757 2.03e-06	0.39 2.70 0.38	0.696 0.007 0.704	.8764 1.510 .9999	.064 .705 .968	1.218603 13.33579 1.000005
subyrctr   pred01fda   hhospdisc   hhosleng	1.033437 4.488479 1.000001	.0869026 2.493757 2.03e-06 .061535	0.39 2.70 0.38	0.696 0.007 0.704 0.307	.8764 1.510 .9999 .947	.064 .705 .968 .059	1.218603 13.33579 1.000005
subyrctr   pred01fda   hhospdisc   hhosleng   /lnsig2u	1.033437 4.488479 1.000001 1.061063	.0869026 2.493757 2.03e-06 .061535	0.39 2.70 0.38	0.696 0.007 0.704 0.307	.8764 1.510 .9999 .947	064 1705 1968 1059 1114	1.218603 13.33579 1.000005 1.188791

Hence combining the FDA's codes of before August 2005 with the Carpenter-Zucker-Avorn codes of before August 2005 yields estimated odds ratios of greater than three (3), and in all three tests the *p*-value is below 0.05 for a two-tailed test, and in both regressions the estimated odds ratio is above three (3).

### 14. Re-Analysis of Association between Deadline Approvals and Combined Withdrawal-or-Warning Measure, Using FDA NMEs, FDA Deadline Approval Measure, and FDA Warning Measures

In this section we proceed as we did in Section 7, but focus on the combined "withdrawal or warnings" measure that we used in our paper. Again, for the sake of hypothesis and demonstration, we do not correct here for clear FDA errors identified in Section 9 (above) or in Section 2 (above). If we then proceed according to Carpenter, Zucker and Avorn 2008, and we create a combined withdrawal or black-box warning measure, and if we rely entirely on the FDA's data and stop coding in August 2005, we get the following list of postmarket safety events for NMEs.

list +	ndanum_fo	da genernam_fda sbwdate_fda bbwdate_fo	da if( sbwork	bbw_fda0805 
į	ndanum~a	genernam_fda	sbwdate~a	bbwdate~a
1.	21042	Rofecoxib	30-Sep-04	
2.	20998	Celecoxib		29-Jul-05
3.	21341	Valdecoxib	6-Apr-05	24-Nov-04
4.	20535	Bromfenac Sodium	22-Jun-98	31-Mar-98
3.	20896	Capecitabine		7-Sep-01
4.	20169	Nilutamide		29-Sep-00
4.	20740	Cerivastatin Sodium	8-Aug-01	
5.	20689	Mibefradil Dihydrochloride	8-Jun-98	
4.	20430	Danaparoid Sodium		30-Jan-98
8.	20564	Lamivudine		15-Dec-97
4.	20628	Saguinavir		14-Nov-00
7.	20759	Trovafloxacin Mesylate		9-Jun-99
8.	20760	Alatrofloxacin Mesylate		9-Jun-99
9.	20695	Grepafloxacin Hydrochloride	27-Oct-99	
5.	20822	Citalopram Hydrobromide		18-Feb-05
6.	20415	Mirtazapine		12-Jan-05
7.	21427	Duloxetine Hydrochloride		18-Feb-05
2.	20315	Levomethadyl Acetate Hydrochloride	23-Aug-03	
6.	20984	Rapacuronium Bromide	27-Mar-01	
1.	20697	Tolcapone		16-Nov-98
2.	20720	Troglitazone	22-Mar-00	15-Dec-97
4.	21107	Alosetron Hydrochloride	28-Nov-00	7-Jun-02

#### 14.1 Combined Withdrawal and Black-Box Warning Measure, Using FDA Data Only, **August 2005 Stop Coding (cont.)**

A two-way cross-tabulation yields a p-value of 0.08, but an exact logistic regression on the same data (controlling for submission year) yields a p-value of 0.04.

. tab sbwoı	cbbw_fda0805	pred01fda,	exact	
da0805	•	1	Total	
0	214   12	78 10	22	
Total	226	88	+ 314	
4 ' 1 1	Fisher's exa			.082
ı-sıded	Fisher's exa	CT =	U .	.054

An exact logistic regression yields the following results.

#### Binary Regression

 $regression\ (type=logit,\ model(sbworbbw\_f=subyrctr\ pred01fda),\ estimate(subyrctr\ pred01fda),\ method=exact,\ profile=yes,\ outbyne=odds).$ 

#### **Basic Information**

Data file bbwdump20080412-2.csv

Model sbworbbw\_f(Response = 1)=%Const+subyrctr+pred01fda Link type Logit

Weight variable <Not Specified> Stratum variable <Unstratified> Analysis type Estimate :: Exact

Number of terms in model Number of term(s) dropped Number of observations in analysis 314 Number of records rejected 0 24 **Number of groups** 

#### **Summary Statistics**

	Statistics	Value	DF	P-Value
De	viance	22.5	21	0.3714
Lik	elihood Ratio	285.7	3	1.657e-009

#### **Parameter Estimates**

		Point Estima	ate	Confidence In	terval and P-\	Value for Odd	s Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unner	P-Value
%Const	MLE	1.676	NA.	Asymptotic	0.1088	25.82	0.7114
				Asymptotic(Profile)	0.1154	28.71	
subyrctr	MLE	0.8155	NA.	Asymptotic	0.6901	0.9638	0.0167
				Asymptotic(Profile)	0.6819	0.9552	
	CMLE	0.8169	NA	Exact	0.6829	0.961	0.01288
pred01fda	MLE	2,922	NA	Asymptotic	1.174	7.269	0.02113
				Asymptotic(Profile)	1.156	7.309	
	CMLE	2.891	NA.	Exact	1.039	7.941	0.04139

## 14.1 Combined Withdrawal and Black-Box Warning Measure, Using FDA Data Only, August 2005 Stop Coding (cont.)

A random-effects logistic regression of the same data on submission year and epidemiological covariates yields a p-value of 0.02. The estimates for the FDA's deadline approval measure are highlighted in yellow.

istic reg	ression		Number o	f obs	= 314	
: discode			Number o	f groups	= 135	
~ Gaussi	an		Obs per	group: min	= 1	
				_		
			Wald chi	2 (4)	= 10.31	
-72.94882	2					
OR	Std. Err.	z	P> z	[95% Con	f. Interval]	
7822211	.0782345	-2.46	0.014	.6429777	.9516191	
					.9516191 10.68308	
.660941	2.000369	2.37	0.018	1.254553	.9516191 10.68308 1.000005	
.660941 .000001	2.000369 1.63e-06	2.37 0.90	0.018 0.368	1.254553 .9999983	10.68308	
.660941 .000001 .009056	2.000369 1.63e-06	2.37 0.90	0.018 0.368	1.254553 .9999983 .9033836 	10.68308 1.000005 1.127089  1.450255	
.660941 .000001 .009056 	2.000369 1.63e-06 .0569524	2.37 0.90	0.018 0.368	1.254553 .9999983 .9033836 	10.68308 1.000005 1.127089	
	: discode	istic regression : discode	: discode  ~ Gaussian  -72.948822	: discode Number o  ~ Gaussian Obs per o  Wald chi:  -72.948822 Prob > c	: discode Number of groups  ~ Gaussian Obs per group: min avg max  Wald chi2(4)  Prob > chi2	: discode Number of groups = 135 ~ Gaussian Obs per group: $min = 1$ avg = 2.3 max = 16 Wald chi2(4) = 10.31

In summary: If we rely entirely upon the FDA's data and we stop coding when Carpenter, Zucker and Avorn stop coding, we still find large (ORs > 2.8) and statistically significant relationships between pre-deadline approvals and post-market safety problems.

## 14.2 Combined Withdrawal and Black-Box Warning Measure, Using FDA Data Only, Stop Coding After Vioxx Withdrawal

If we then proceed as according to Carpenter, Zucker and Avorn 2008, and we create a combined withdrawal or black-box warning measure, and <u>if we rely entirely on the FDA's data</u> and stop coding in after the Vioxx withdrawal, we get the following list of postmarket safety events for NMEs.

bbwdate~a	sbwdate~a	genernam_fda	ndanum~a
	30-Sep-04	Rofecoxib	21042
29-Jul-05		Celecoxib	20998
24-Nov-04	6-Apr-05	Valdecoxib	21341
31-Mar-98	22-Jun-98	Bromfenac Sodium	20535
7-Sep-01		Capecitabine	20896
29-Sep-00		Nilutamide	20169
	8-Aug-01	Cerivastatin Sodium	20740
	8-Jun-98	Mibefradil Dihydrochloride	20689
30-Jan-98		Danaparoid Sodium	20430
15-Dec-97		Lamivudine	20564
14-Nov-00		Saquinavir	20628
9-Jun-99		Trovafloxacin Mesylate	20759
9-Jun-99		Alatrofloxacin Mesylate	20760
	27-Oct-99	Grepafloxacin Hydrochloride	20695
18-Feb-05		Citalopram Hydrobromide	20822
12-Jan-05		Mirtazapine	20415
18-Feb-05		Duloxetine Hydrochloride	21427
	23-Aug-03	Levomethadyl Acetate Hydrochloride	20315
	27-Mar-01	Rapacuronium Bromide	20984
16-Nov-98		Tolcapone	20697
15-Dec-97	22-Mar-00	Troglitazone	20720
7-Jun-02	28-Nov-00	Alosetron Hydrochloride	21107

### 14.2 Combined Withdrawal and Black-Box Warning Measure, Using FDA Data Only, post-Vioxx Stop Coding (cont.)

A two-way cross-tabulation yields a *p*-value of 0.054, but an exact logistic regression on the same data (controlling for submission year) yields a *p*-value of 0.04.

. tab sbwor	rbbw_fda0105 pred0	)1fda, ex	xact	
$da01\overline{0}5$	pred01fda   0	1	Total	
0		79   9	296 18	
Total	226	88	314	
	Fisher's exact =		0.05	<mark>54</mark>
1-sided	Fisher's exact =		0.03	35

An exact logistic regression yields the following results, with an estimated odds ratio of 3.68 and an exact p-value of 0.02.

#### **Binary Regression**

regression (type=logit, model(comb0105 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtwoe=odds):

#### **Basic Information**

Data file SBWorBBWdump4-20080412.csv

Model comb0105(Response = 1)=%Const+subyrctr+pred01fda
Link type Loqit

Weight variable <Not Specified>

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 0
Number of groups 24

**Summary Statistics** 

Statistics	Value	DF	P-Value
Deviance	23.52	21	0.317
Likelihood Ratio	309.6	3	1.553e-009

#### Parameter Estimates

. arameter asemiates							
		Point Estim	ate	Confidence In	terval and P-	Value for Odd	s Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	Unper	P-Value
%Const	PMLE	2.544	NA.	Asymptotic	0.1257	51.5	0.5429
				Asymptotic(Profile)	0.1318	64.69	
subyrctr	PMLE	0.7829	NA	Asymptotic	0.6496	0.9434	0.01012
				Asymptotic(Profile)	0.6358	0.9347	
	CMLE	0.7734	NA	Exact	0.6246	0.9338	0.005584
pred01fda	PMLE	3,663	NA	Asymptotic	1.387	9.676	0.00879
				Asymptotic(Profile)	1.373	9.881	
	CMLE	3.68	NA.	Exact	1.197	11.43	0.02143

PMLE: Penalized MLE for bias correction (Firth's method).

### 14.2 Combined Withdrawal and Black-Box Warning Measure, Using FDA Data Only, post-Vioxx Stop Coding (cont.)

A random-effects logistic regression of the same data on submission year and epidemiological covariates yields a *p*-value of 0.01. The estimates for the FDA's deadline approval measure are highlighted in yellow.

Group variable	s logistic rec				of obs =	
-					2 1	
Random effects	3 u_i ~ Gauss:	ian		Obs per	group: min =	1
						2.3
					max =	16
				Wald ch	12(4) =	11.71
Log likelihood	d = -62.42623	16		Prob > 0	chi2 =	0.0196
sbworbb~0105			Z	P> z	[95% Conf.	Interval]
subyrctr	.7727803	.0783442	-2.54	0.011	.6335217	.9426503
pred01fda	3.803343	1.962027	2.59	0.010	1.383748	10.45379
hhospdisc	1.000001	1.33e-06	0.87	0.384	.9999985	1.000004
hhosleng	1.012535	.0473796	0.27	0.790	.923804	1.109789
		1.109992			-5.288712	9376236
/lnsig2u	-3.113168					
	-3.113168 +   .2108551	.1170238			.0710511	.6257453

Again, if we rely entirely upon the FDA's data and we merely stop coding boxed warnings and withdrawals after the withdrawal of Vioxx, we observe large (ORs > 3.6) and statistically significant relationships between pre-deadline approvals and post-market safety problems.

## 15. Notes on Reanalysis of BBWs without NMEs Changed through Indirect Class Relabeling

At least one significant difference between class BBW relabelings and more specific BBW relabelings is the specificity of the evidence concerned. Class relabelings occur with indirect evidence from other molecules that rely upon similar mechanisms; they are not based upon direct evidence of studies of the molecule itself. Such patterns have been observed for the BBWs added to "atypical antipsychotic" drugs in recent years (aripiprazole, olanzapine).

#### First we generate a new variable that excludes indirect class relabelings, as follows:

```
. gen BBWcombo excludeclassrelabel = CZAFDAcombobbw dec2007
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 20998)
(1 real change made) {THIS IS CELECOXIB}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 20937)
(1 real change made) {THIS IS GADOVERSETAMIDE}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 20639)
(1 real change made) {THIS IS QUETIAPINE FUMARATE}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 20592)
(1 real change made) {THIS IS OLANZAPINE}
. replace BBWcombo_excludeclassrelabel = 0 if( ndanum_fda == 21436)
(1 real change made) {THIS IS ARIPIPRAZOLE}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 20825)
(1 real change made) {THIS IS ZIPRASIDONE}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 20938)
(1 real change made) {THIS IS MELOXICAM}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 20822)
(1 real change made) {THIS IS CITALOPRAM}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 21427)
(1 real change made) {THIS IS DULOXETINE}
. replace BBWcombo_excludeclassrelabel = 0 if( ndanum fda == 20415)
(1 real change made) {THIS IS MIRTAZAPINE}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 21449)
(1 real change made) {THIS IS ADEFOVIR DIPIVOXIL}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 21500)
(1 real change made) {THIS IS EMTRICITABINE}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 21618)
(1 real change made) {THIS IS TINIDAZOLE, CZA Positive Coding Error}
. replace BBWcombo excludeclassrelabel = 0 if( ndanum fda == 21492)
(1 real change made) {THIS IS ELOXATIN, CZA Positive Coding Error}
. replace BBWcombo_excludeclassrelabel = 0 if( ndanum_fda == 21345)
(1 real change made) {This is Fondaparinux Sodium, CZA Positive Coding Error}
```

## **15.1.2.** Revised List of BBWs after Exclusion of Indirect Therapeutic Class Relabeling. [Notice Saquinavir Included; see Section 15.2 and Section 16 for analyses where it is not coded as BBW.]

. list ndanum\_fda genernam\_fda pred01fda if( BBWcombo\_excludeclassrelabel == 1)

_			+
	ndanum~a	genernam_fda	pred01~a
7.	20760	Alatrofloxacin Mesylate	1
13.	21107	Alosetron Hydrochloride	0
27.	21411	Atomoxetine Hydrochloride	0
42.	20535	Bromfenac Sodium	0
47.	20896	Capecitabine	1
66.	20430	Danaparoid Sodium	0
90.	21797	Entecavir	1
152.	20564	Lamivudine	1
188.	20169	Nilutamide	0
213.	21064	Perflutren Lipid Microsphere	0
214.	21302	Pimecrolimus	0
215.	21073	Pioglitazone Hydrochloride	1
224.	20815	Raloxifene Hydrochloride	1
241.	21071	Rosiglitazone Maleate	1
245.	20628	Saquinavir	0
266.	21144	Telithromycin	0
269.	21356	Tenofovir Disoproxil Fumarate	1
278.	21814	Tipranavir	1
281.	20697	Tolcapone	0
292.	20720	Troglitazone	1
294. 299.	20759	Trovafloxacin Mesylate Valdecoxib	 1   1

.

#### 15.1.3. List of NMEs Recoded on the Basis of Indirect Class Relabeling

To see the time distribution of the drugs that were recoded, we reprint the exclusions along with generic name and the date of BBW. Notice that 10 (**34 percent**) of the 29 NMEs that Nardinelli et al coded in April 2008 as having black-box warnings added were based partially or wholly on indirect class relabeling, and notice that two others (which the FDA missed) also qualify.

```
. list ndanum_fda genernam_fda pred01fda bbwdate_fda if(
BBWcombo_excludeclassrelabel == 0 & CZAFDAcombobbw_dec2007 == 1)
```

	+-·	 ndanum~a 	genernam_fda	pred01~a	+ bbwdate~a   
6.	1	21449	Adefovir Dipivoxil	1	- 1
23.		21436	Aripiprazole	0	16-Feb-06
52.		20998	Celecoxib	1	29-Jul-05
59.		20822	Citalopram Hydrobromide	0	18-Feb-05
82.		21427	Duloxetine Hydrochloride	0	18-Feb-05
87.	1	21500	Emtricitabine	1	1
119.		20937	Gadoversetamide	0	4-Sep-07
167.		20938	Meloxicam	0	11-Aug-05
177.		20415	Mirtazapine	0	12-Jan-05
193.		20592	Olanzapine	0	16-Feb-06
222.		20639	Quetiapine Fumarate	0	20-Sep-06
311.		20825	Ziprasidone Hydrochloride	0	17-Aug-05

.

## 15.2.1. Cross-Tabulation of BBW Measure without Indirect Therapeutic Class Relabelings, corrected CZA Measure, with Fisher's Exact Test

. tab BBWo	. tab BBWcombo_excludeclassrelabel pred01fda,				
BBWcombo_e xcludeclas srelabel	pred01fo	1	Total		
0	216   10	76   12	292 22		
Total	226	88	314		
	Fisher's exact =	=	0.	007	
1-sided	Fisher's exact =	=	0.	006	

To generate a shorter name for this variable, to be used in LogXact8 (which has max 9 characters in variable names), we used the following code:

```
. gen BBWnoclass = BBWcombo excludeclassrelabel
```

This variable is used as the dependent variable in the exact logistic regressions on the following page.

### 15.2.2. Exact Logistic Regression Analysis of BBW Measure without Indirect Therapeutic Class Relabelings, corrected CZA Measure, with Fisher's Exact Test

#### **Binary Regression**

 $regression \ (type=logit,\ model(bbwnoclass=subyrctr\ pred01fda),\ estimate(subyrctr\ pred01fda),\ method=exact,\ mle=firth,\ profile=yes,\ outtyne=odde).$ 

#### **Basic Information**

Data file comboBBWnoclass.csv

Model BBWnoclass(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	26.1	21	0.2027
Likelihood Ratio	283.2	3	1.291e-009

#### **Parameter Estimates**

Turdineter Estimates										
		Point Estima	ate	Confidence In	terval and P-\	Value for Odd	s Ratio			
					95	95 %CI 2*1-sid				
Model Term	Type	Odds Ratio	SF(Odds)	Type	Lower	Upper	P-Value			
%Const	PMLE	0.05982	NA	Asymptotic	0.00495	0.723	0.02675			
				Asymptotic(Profile)	0.004619	0.7256				
subyrctr	PMLE	0.9891	NA	Asymptotic	0.8603	1.137	0.8778			
				Asymptotic(Profile)	0.8561	1.137				
	CMLE	0.987	NA	Exact	0.8511	1.14	0.8933			
pred01fda	PMLE	3.409	NA	Asymptotic	1.422	8.175	0.005985			
				Asymptotic(Profile)	1.419	8.354				
	CMLE	3.437	NA	Exact	1.279	9.45	0.01272			

PMLE: Penalized MLE for bias correction (Firth's method).

## 15.2.3. Random-Effects Logistic Regression Analysis of BBW Measure without Indirect Therapeutic Class Relabelings, corrected CZA Measure, with Fisher's Exact Test

Results for FDA deadline approval measure highlighted in yellow.

. xtlogit BBWc i(discode) or	combo_excluded	lassrelabel	subyrctr	pred01fda	hhospdisc	hhosleng,	re	
Random-effects logistic regression  Number of obs = 314  Group variable (i): discode  Number of groups = 135								
Random effects u_i ~ Gaussian  Obs per group: min = 1  avg = 2.3  max = 16								
Log likelihood	= -75.60214	8		Wald chi2( Prob > chi	(4) =	8.03 0.0904		
BBWcombo_e~1		Std. Err.			[95% Conf.	Interval]		
subyrctr   pred01fda	.9872738 3.538954 .9999999	.0734065	-0.17 2.73	0.863	.8533917 1.428535	1.14216 8.767157		
	.9999999 1.035291							
/lnsig2u	-2.848745	4.281551		-	11.24043	5.542939		
	.2406594 .0173001				.0036239 3.99e-06			
 Likelihood-rat	io test of rh	o=0: chibar	2(01) =	0.01 Pro	bb >= chiba	r2 = 0.458		

Hence if we use a corrected measure and exclude only class relabelings that are based on indirect evidence, we observe ORs above 3.4 in exact logistic and random-effects logistic regressions, with p-values less than 0.02.

### 15.3.1. Now use the same measure excluding therapeutic class relabelings, but drop Saquinavir from count, as FDA Coding Error

```
. gen BBWcombonoclassbbw_dropsaquin = BBWcombo_excludeclassrelabel
. replace BBWcombonoclassbbw_dropsaquin = 0 if(ndanum_fda == 20628)
(1 real change made)
```

. tab BBWco	mbonoclassbbw_c	lropsaqu	in pred01fd	da, exac
BBWcombono   classbbw_d   ropsaquin	pred01fda 0	1	Total	
0   1	217 9	76   12	293 21	
Total	226	88	314	
	<pre>sher's exact = sher's exact =</pre>			.004 .004

To generate a shorter name for this variable, to be used in LogXact8 (which has max 9 characters in variable names), we used the following code:

```
gen BBWnocnos = BBWcombonoclassbbw_dropsaquin
```

This variable is used as the dependent variable in the exact logistic regressions on the following page.

### 15.3.2. Exact Logistic Regression of BBW Measure without Indirect Therapeutic Class Relabelings, corrected CZA Measure, drop Saquinavir

#### **Binary Regression**

regression (type=logit, model(bbwnocnos = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtbroe=odde):

#### **Basic Information**

Data file comboBBWnoclass.csv

Model BBWnocnos(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable Stratum variable Stratum variable Stratum variable Stimate :: Exact

Number of terms in model3Number of term(s) dropped0Number of observations in analysis314Number of records rejected0Number of groups24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	28.78	21	0.1193
Likelihood Ratio	289.5	3	1.226e-009

#### **Parameter Estimates**

Tarameter Estimates										
		Point Estima	ate	Confidence In	terval and P-\	Value for Odd	s Ratio			
					95	%CI	2*1-sided			
Model Term	Type	Odds Ratio	SF(Odds)	Type	Lower	Upper	P-Value			
%Const	PMLE	0.04356	NA	Asymptotic	0.003401	0.558	0.01602			
				Asymptotic(Profile)	0.003133	0.5559				
subyrctr	PMLE	1.001	NA	Asymptotic	0.869	1.154	0.985			
				Asymptotic(Profile)	0.865	1.154				
	CMLE	0.9995	NA	Exact	0.8599	1.157	1			
pred01fda	PMLE	3.727	NA	Asymptotic	1.522	9.13	0.004			
				Asymptotic(Profile)	1.524	9.397				
	CMLE	3.775	NA	Exact	1.372	10.77	0.008435			

PMLE: Penalized MLE for bias correction (Firth's method).

## 15.3.3. Random-Effects Logistic Regression of BBW Measure without Indirect Therapeutic Class Relabelings, corrected CZA Measure, drop Saquinavir from BBW list

Results for FDA deadline approval measure highlighted in vellow.

. xtlogit BI	BWcombonoclassk	bw_dropsaqu	in subyro	ctr pred01	lfda hhosp	disc	hhosleng,	re i(	discode)	or
Random-effect	ts logistic red	gression		Number o	of obs	=	314			
Group variable (i): discode Number of groups = 135										
Random effects u i ~ Gaussian Obs per group: min = 1										
	_			-			2.3			
					_		16			
				Wald chi	12(4)	=	8.71			
Log likeliho	od = $-72.66301$	12		Prob > 0	chi2	=	0.0687			
	OR			P> z	[95% Con	 f. In	terval]			
	-+   .9996823		-0.00	0.997	.8626872	1	.158432			
pred01fda	3.878498	1.838523	2.86	0.004	1.531683	9	.821057			
	.9999999									
hhosleng	1.028765									
/lnsig2u	-2.936272	2.717393								
sigma u	.2303545	.3129819			.0160647	3				
rho	.0158733				.0000784					
Likelihood-ra	atio test of rh	no=0: chibar	2(01) =	0.01 H	?rob >= chil	 bar2	= 0.466			

.

Hence if we use a corrected measure and exclude only class relabelings that are based on indirect evidence, and recode saquinavir as not having had a BBW, we observe ORs above 3.7 in exact logistic and random-effects logistic regressions, with p-values less than 0.01.

## Section 16: Notes on Reanalysis of BBWs, Drop Saquinavir (FDA Coding Error)

First begin with the combined black-box warning measure that stops coding in August 2005, then recode only saquinavir.

```
. replace CZAFDAbbw0805_corrdropsaquin = 0 if(ndanum_fda == 20628)
(1 real change made) {THIS IS SAQUINAVIR}

. save "F:\fdadata\FDAData-314NMEs-20080509.dta", replace
file F:\fdadata\FDAData-314NMEs-20080509.dta saved

. save "c:\fdatemp\FDAData-314NMEs-20080509.dta", replace
file c:\fdatemp\FDAData-314NMEs-20080509.dta saved
```

### 16.1.1. Cross-Tabulation of Revised BBW Measure with Stopped Coding in August 2005, Using FDA's NMES and FDA's Deadline Approval Measure, with Fisher's Exact Test

### 16.1.2 Exact Logistic Regression Analysis of Revised BBW Measure with Stopped Coding in August 2005, Using FDA's NMES and FDA's Deadline Approval Measure

#### **Binary Regression**

regression (type=logit, model(bbw0805 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtype=odde):

#### **Basic Information**

Data file analysesdroppingsaquinavir20080509.csv

Model bbw0805(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable < Not Specified>
Stratum variable < Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 314
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	25.87	21	0.2116
Likelihood Ratio	295.9	3	1.249e-009

#### **Parameter Estimates**

Tarameter Estimates										
		Point Estim	ate	Confidence In	terval and P-\	Value for Odd	s Ratio			
					95	%CI	2*1-sided			
Model Term	Type	Odds Ratio	SF(Odds)	Type	Lower	Upper	P-Value			
%Const	PMLE	0.04114	NA	Asymptotic	0.003026	0.5593	0.01656			
				Asymptotic(Profile)	0.002774	0.5583				
subyrctr	PMLE	0.9981	NA	Asymptotic	0.8634	1.154	0.9791			
				Asymptotic(Profile)	0.859	1.154				
	CMLE	0.9961	NA	Exact	0.8536	1.157	0.9958			
pred01fda	PMLE	4.201	NA	Asymptotic	1.671	10.56	0.002282			
				Asymptotic(Profile)	1.683	10.97				
	CMLE	4.28	NA	Exact	1.513	12.79	0.004698			

PMLE: Penalized MLE for bias correction (Firth's method).

## 16.1.3 Random-Effects Logistic Regression Analysis of Revised BBW Measure with Stopped Coding in August 2005, Using FDA's NMES and FDA's Deadline Approval Measure

Now we use the same corrected BBW variable with coding stopped in August 2005, and analyze the association between this variable and just-before-deadline approvals, using the Nardinelli et al "deadline approval" measure. Results for Deadline Approval Variable highlighted in yellow.

xtlogit CZAE (discode) or	FDAbbw0805_cd	orrdropsaqui	n subyrct	er pred01fd	a hhospdis	sc hhosleng,	re
andom-effects	logistic red	gression		Number of	obs =	314	
roup variable	(i): discode	÷		Number of	groups =	= 135	
andom effects	u i ~ Gaussi	lan		Obs per q	roup: min =	= 1	
	_				-	2.3	
					_	= 16	
				Wald chi2	(4) =	10.35	
						0.0349	
og likelihood	= -66.29453	39		Prob > Cn	.12 =	0.0349	
og likelihood ZAFDAbbw0~n							
ZAFDAbbw0~n   	OR 9638772	Std. Err.	z 	P> z  0.692	[95% Conf.	Interval]	
ZAFDAbbw0~n   +- subyrctr   pred01fda	OR9638772 6.973425	Std. Err0894364 4.342355	z  -0.40 3.12	P> z  0.692 0.002	[95% Conf.  .8036007 2.057797	Interval] 1.156121 23.63141	
ZAFDAbbw0~n   +- subyrctr   pred01fda	OR 9638772	Std. Err0894364 4.342355	z  -0.40 3.12	P> z  0.692 0.002	[95% Conf.  .8036007 2.057797	Interval] 1.156121 23.63141	
ZAFDAbbw0~n   subyrctr   pred01fda   hhospdisc	OR9638772 6.973425	Std. Err0894364 4.342355 2.11e-06		P> z   0.692 0.002 0.698	[95% Conf.  .8036007 2.057797 .9999967	Interval] 1.156121 23.63141 1.000005	
ZAFDAbbw0~n   subyrctr   pred01fda   hhospdisc   hhosleng	OR9638772 6.973425 1.000001	Std. Err0894364 4.342355 2.11e-06 .0661873		P> z  0.692 0.002 0.698 0.433	[95% Conf.  .8036007 2.057797 .9999967	Interval] 1.156121 23.63141 1.000005 1.188707	
ZAFDAbbw0~n   subyrctr   pred01fda   hhospdisc   hhosleng   /lnsig2u	OR .9638772 6.973425 1.000001 1.050633	Std. Err0894364 4.342355 2.11e-06 .0661873 .4872433		P> z  0.692 0.002 0.698 0.433	[95% Conf. .8036007 2.057797 .9999967 .9285977	Interval] 1.156121 23.63141 1.000005 1.188707	

### 16.2 Analysis of Combined CZA and FDA Measure Extending until December 2007. Drop Saquinavir.

Now we use the same corrected BBW variable with coding stopped in August 2005, and analyze the association between this variable and just-before-deadline approvals, using the Nardinelli et al "deadline approval" measure. We also recode saquinavir as not having had a BBW (see Section 9).

```
. gen CZAFDAcombobbw1207_dropsaquin = CZAFDAcombobbw_dec2007
. replace CZAFDAcombobbw1207_dropsaquin = 0 if(ndanum_fda == 20628)
(1 real change made)
```

#### A cross-tabulation with Fisher's exact test yields the following.

#### 16.2.2 Exact Logistic Regression Analysis of Combined CZA and FDA Measure Extending until December 2007. Drop Saquinavir.

Now we use the same corrected BBW variable with coding stopped in August 2005, and analyze the association between this variable and just-before-deadline approvals, using the Nardinelli et al "deadline approval" measure. We also recode saquinavir as not having had a BBW (see Section 9).

#### **Binary Regression**

regression (type=logit, model(bbw1207 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtype=odds):

#### **Basic Information**

analysesdroppingsaquinavir20080509.csv Data file

Model bbw1207(Response = 1)=%Const+subyrctr+pred01fda

Logit

Link type Weight variable <Not Specified> Stratum variable <Unstratified> Analysis type Estimate :: Exact

Number of terms in model Number of term(s) dropped Number of observations in analysis 314 Number of records rejected Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	20.12	21	0.5134
Likelihood Ratio	218	3	1.084e-009

#### **Parameter Estimates**

		Point Estima	ate	Confidence In	terval and P-\	alue for Odd	s Ratio
					95	%CI	2*1-sided
Model Term	Type	Odds Ratio	SF(Odds)	Type	Lower	Upper	P-Value
%Const	PMLE	0.03096	NA	Asymptotic	0.00399	0.2402	0.000886
				Asymptotic(Profile)	0.003772	0.2321	
subyrctr	PMLE	1.069	NA	Asymptotic	0.956	1.195	0.2418
				Asymptotic(Profile)	0.9554	1.196	
	CMLE	1.069	NA	Exact	0.9531	1.199	0.2578
pred01fda	PMLE	2.112	NA	Asymptotic	1.033	4.32	0.04054
				Asymptotic(Profile) 1.023 4.305			
	CMLE	2.102	NA	Exact	0.9474	4.597	0.06903

PMLE: Penalized MLE for bias correction (Firth's method).

## 16.2.3 Random-Effects Logistic Regression Analysis of Combined CZA and FDA Measure Extending until December 2007. Drop Saquinavir.

Now we use the same corrected BBW variable with coding stopped in August 2005, and analyze the association between this variable and just-before-deadline approvals, using the Nardinelli et al "deadline approval" measure. We also recode saquinavir as not having had a BBW (see Section 9). Results for Deadline Approval Variable highlighted in yellow.

. xtlogit CZAFI i(discode) or	DAcombobbw12	07_dropsaqui	n subyro	tr pred01f	da hhospdi	sc hhosleng,	re
Random-effects	-				obs =		
Group variable	(I): alscode			Number of	groups =	135	
Random effects	u i ~ Gaussi	an		Obs per g	roup: min =	1	
	_				avg =	2.3	
					max =	16	
				Wald chi2	(4) =	9.53	
Log likelihood	= -99.94002	2		Prob > ch	i2 =	0.0492	
CZAFDAcomb~n	OR	Std. Err.	z	P>   z	[95% Conf.	Interval]	
		.0797693					
pred01fda							
hhospdisc	1.000001	1.95e-06	0.68	0.498	.9999975	1.000005	
hhosleng	1.072317	.0626088	1.20	0.232	.9563666	1.202325	
/lnsig2u	1.200583	.4364999			.3450585	2.056107	
sigma u	1.82265	.3977932			1.188307	2.795618	
	.5024338				.300317	.7037583	
Likelihood-ratio	o test of rh	o=0: chibar2	(01) =	12.44 Pr	ob >= chiba	r2 = 0.000	

.

### 16.3 Analysis of Combined Withdrawal or Black-Box Warnings Measure, Stop Coding in August 2005, Drop Saquinavir

We now take the combined "Withdrawals or Warnings" Measure and exclude saquinavir from the BBW count.

```
. gen sbworbbwFDACZA0805\_dropsaquin = CZAFDAbbw\_august2005\_corrected
```

## 16.3.1. Cross-Tabulation of Combined Withdrawal or Black-Box Warnings Measure, Stop Coding in August 2005, Drop Saquinavir; Fisher's Exact Test.

. tab sbwo	rbbwFDACZA0805_	dropsaquin	pred01fda,	exact	
ropsaquin	pred01	1	Total		
0	218   8	76   12	294 20		
Total	-+   226	88	314		
	Fisher's exact			003	
1-sided	Fisher's exact	; =	0.	002	

<sup>.</sup> replace sbworbbwFDACZA0805\_dropsaquin = 0 if(ndanum\_fda == 20628)
(1 real change made)

### 16.3.2. Exact Logistic Regression Analysis of Combined Withdrawal or Black-Box Warnings Measure, Stop Coding in August 2005, Drop Saquinavir.

#### **Binary Regression**

regression (type=logit, model(sbwbbw081 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtype=odds):

#### **Basic Information**

**Data file** analysesdroppingsaquinavir20080509.csv

Model sbwbbw081(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable < Not Specified>
Stratum variable < Unstratified>
Analysis type Estimate :: Exact

Number of terms in model3Number of term(s) dropped0Number of observations in analysis314Number of records rejected0Number of groups24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	25.87	21	0.2116
Likelihood Ratio	295.9	3	1.249e-009

#### **Parameter Estimates**

		Point Estima	ate	Confidence In	terval and P-\	/alue for Odd	s Ratio
				95 %CI			2*1-sided
Model Term	Type	Odds Ratio	SF(Odds)	Type	Lower	Upper	P-Value
%Const	PMLE	0.04114	NA	Asymptotic	0.003026	0.5593	0.01656
				Asymptotic(Profile)	0.002774	0.5583	
subyrctr	PMLE	0.9981	NA	Asymptotic	0.8634	1.154	0.9791
				Asymptotic(Profile)	0.859	1.154	
	CMLE	0.9961	NA	Exact	0.8536	1.157	0.9958
pred01fda	PMLE	4.201	NA	Asymptotic	1.671	10.56	0.002282
				Asymptotic(Profile)	1.683	10.97	
	CMLE	4.28	NA	Exact	1.513	12.79	0.004698

PMLE: Penalized MLE for bias correction (Firth's method).

#### Analysis Time = 00:00:00

Taking the combined "Withdrawals or Warnings" Measure and excluding saquinavir from the BBW count yields an OR of 4.3 (just less than reported in the original article) and a lower bound on the exact CI of 1.51 (nearly identical to that in the original article).

## 16.3.3. Random-Effects Logistic Regression Analysis of Combined Withdrawal or Black-Box Warnings Measure, Stop Coding in August 2005, Drop Saquinavir.

Results for Deadline Approval Variable highlighted in yellow.

(discode) or	orbbwFDACZA08(	)5_dropsaqui	n subyrct	tr pred011	fda hhospd	isc hhosleng	, re
	s logistic req					= 314 = 135	
.oup variable	e (I): alscode	=		Number (	or groups	- 155	
andom effects	s u i ~ Gaussi	ian		Obs per	group: min	= 1	
	_			-		= 2.3	
					max	16	
				Wald ch	i 2 (4)	= 10.35	
211 211		2.0		- 1	1 10	0 0010	
og likelihood	d = -66.29453	39		Prob > 0	chi2	= 0.0349	
	d = -66.29453			Prob > 0	chi2 	= 0.0349	
oworbbwFD~n	OR	Std. Err.		Prob > 0	chi2  [95% Con	= 0.0349 f. Interval]	
oworbbwFD~n	OR +	Std. Err. .0894364	z 	Prob > 0 	ehi2 [95% Con  .8036007	= 0.0349  f. Interval]  1.156121	
oworbbwFD~n   subyrctr   pred01fda   hhospdisc	OR	Std. Err0894364 4.342355 2.11e-06	z 	Prob > 0  P> z   0.692  0.002  0.698	[95% Con .8036007 2.057797 .9999967	f. Interval] 1.156121 23.63141 1.000005	
oworbbwFD~n   subyrctr   pred01fda   hhospdisc	OR +	Std. Err0894364 4.342355 2.11e-06	z 	Prob > 0  P> z   0.692  0.002  0.698	[95% Con .8036007 2.057797 .9999967	f. Interval] 1.156121 23.63141 1.000005	
oworbbwFD~n   subyrctr   pred01fda   hhospdisc   hhosleng	OR	Std. Err0894364 4.342355 2.11e-06 .0661873	z 	Prob > 0  P> z   0.692  0.002  0.698	[95% Con .8036007 2.057797 .9999967 .9285977	= 0.0349  f. Interval]  1.156121 23.63141 1.000005 1.188707  1.813351	
oworbbwFD~n   subyrctr   pred01fda   hhospdisc   hhosleng   /lnsig2u	OR .9638772 6.973425 1.000001 1.050633	Std. Err0894364 4.342355 2.11e-06 .0661873 .4872433	z 	Prob > 0  P> z   0.692  0.002  0.698	[95% Con .8036007 2.057797 .9999967 .9285977	f. Interval]  1.156121 23.63141 1.000005 1.188707	

.

Taking the combined "Withdrawals or Warnings" Measure and excluding saquinavir from the BBW count yields an OR of 6.7 (significantly above that reported in the original article) and a lower bound on the 95% exact CI of 2.05 (well above that in the original article).

## 16.4 Now Re-Analyze 12/2007 Cutoff Data, combine Warnings and Withdrawals, and Drop Saquinavir from the BBW Count as an FDA Coding Error

Now we return to the combined "withdrawals and warnings" measure examined in Section 12.3, and we drop saquinavir from the BBW count.

```
. gen CZAFDASBWorBBW1207_dropsaquin = CZAFDAcomboSBWorBBW_dec2007
. replace CZAFDASBWorBBW1207_dropsaquin = 0 if(ndanum_fda == 20628)
(1 real change made)
```

#### 16.4.1. Cross-Tabulation with Fisher's Exact Test.

. tab CZA	FDASBWorBBW1207_c	dropsaquin	pred01fda,	, exact	
CZAFDASBWo rBBW1207_d ropsaquin	pred01fo	1	Total		
0 1	204	70   18	274 40		
Total	•	88	314		
	Fisher's exact :	=	0.0	014	
1-sided	Fisher's exact	=	0.0	011	

## 16.4.2. Exact Logistic Regression Analysis of Combined and Corrected Warnings and Withdrawals Measure, Dropping Saquinavir from the BBW Count as an FDA Coding Error

#### **Binary Regression**

regression (type=logit, model(sbwbbw12 = subyrctr pred01fda), estimate(subyrctr pred01fda), method=exact, mle=firth, profile=yes, outtype=odds):

#### **Basic Information**

**Data file** analysesdroppingsaquinavir20080509.csv

Model sbwbbw12(Response = 1)=%Const+subyrctr+pred01fda

Link type Logit

Weight variable < Not Specified>
Stratum variable < Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	26.8	21	0.1777
Likelihood Ratio	201.8	3	1.433e-009

#### Parameter Estimates

arameter Leannace									
		Point Estima	ate	Confidence Interval and P-Value for Odds Ratio					
					95	%CI	2*1-sided		
Model Term	Type	Odds Ratio	SF(Odds)	Type	Lower	Upper	P-Value		
%Const	PMLE	0.1154	NA	Asymptotic	0.01668	0.7984	0.02866		
				Asymptotic(Profile)	0.01624	0.7894			
subyrctr	PMLE	0.9979	NA	Asymptotic	0.8953	1.112	0.9693		
				Asymptotic(Profile)	0.8938	1.111			
	CMLE	0.9969	NA	Exact	0.8911	1.113	0.9821		
pred01fda	PMLE	2.387	NA	Asymptotic	1.203	4.739	0.01288		
				Asymptotic(Profile)	1.196	4.731			
	CMLE	2.38	NA	Exact	1.117	5.031	0.02355		

PMLE: Penalized MLE for bias correction (Firth's method).

#### Analysis Time = 00:00:03

Hence using the combined "withdrawals and warnings" measure that includes all Nardinelli et al determinations of a BBW or SBW (this was examined in Section 12.3), and dropping saquinavir from the BBW count, yields an OR of 2.4 with a exact p-value of 0.02 for a two-tailed test.

## 16.4.3. Random-Effects Logistic Regression Analysis of Combined and Corrected Warnings and Withdrawals Measure, Dropping Saquinavir from the BBW Count as an FDA Coding Error

Results for Deadline Approval Variable highlighted in vellow.

i(discode) or	ASBWorBBW12	207_dropsaqu	in subyrd	ctr pred01f	Eda hhospdi	sc hhosleng,	re
Random-effects l	ogistic reg	gression		Number of	f obs =	314	
Group variable (	i): discode	9		Number of	groups =	135	
Random effects u	_i ~ Gaussi	ian		Obs per g	group: min =	1	
						2.3	
					max =	16	
						0 00	
					= (4)		
Log likelihood	= -109.4686	59			2(4) = ni2 =		
CZAFDASBWo~n	OR			Prob > ch	ni2 =	0.0630	
CZAFDASBWo~n	OR	Std. Err.	Z	Prob > ch	= = = = = = = = = = = = = = = = = = =	0.0630  Interval]	
CZAFDASBWo~n   subyrctr   pred01fda	OR9789434 3.87044	Std. Err0708445	z -0.29 2.74	Prob > ch  P> z   0.769 0.006	ni2 = [95% Conf	0.0630 Interval]  1.128126 10.19017	
CZAFDASBWo~n   subyrctr   pred01fda	OR9789434 3.87044	Std. Err0708445	z -0.29 2.74	Prob > ch  P> z   0.769 0.006	ni2 = [95% Conf	0.0630 Interval]  1.128126 10.19017	
CZAFDASBWo~n	OR	Std. Err0708445 1.911668 1.79e-06	z -0.29 2.74 1.09	Prob > ch 	[95% Conf. .8494886 1.470074 .9999984	0.0630  Interval]  1.128126 10.19017 1.000005	
CZAFDASBWo~n	OR .9789434 3.87044 1.000002 1.029629	Std. Err0708445 1.911668 1.79e-06 .0616835	z -0.29 2.74 1.09	Prob > ch 	[95% Conf. .8494886 1.470074 .9999984	0.0630 Interval]  1.128126 10.19017 1.000005 1.15791	
CZAFDASBWo~n   subyrctr   pred01fda   hhospdisc   hhosleng	OR9789434 3.87044 1.000002 1.029629 1.161936	Std. Err0708445 1.911668 1.79e-06 .0616835 .4548797	z -0.29 2.74 1.09	Prob > ch P> z  0.769 0.006 0.274 0.626	[95% Conf. .8494886 1.470074 .9999984 .9155596	0.0630 Interval]  1.128126 10.19017 1.000005 1.15791  2.053484	

Hence using the combined "withdrawals and warnings" measure that includes all Nardinelli et al determinations of a BBW or SBW (this was examined in Section 12.3), and dropping saquinavir from the BBW count, yields an OR of 3.9 with a p-value of 0.006 for a two-tailed test.

# 17. Re-Analysis of CZA Black-Box Warnings Measure, Using Recoded CZA Pre-Deadline Approval Measure, then FDA Pre-Deadline Approval Measure.

Finally, we used our original dataset of 313 drugs and a revised just-before-deadline measure based upon a recoding of the standard-vs.-priority discrepancies. If we re-analyze the original Carpenter/Zucker/Avorn (CZA) postmarket black-box warning measure and use a recoded just-before-deadline approval measure (based upon the reclassified standard and priority drugs), we obtain the following results for a cross-tabulation and for Fisher's exact test.

. tab combo	obbw pred0410_rec	ode, exac	t	
riginal		1	Total	
0 1	214   5	85   9	299 14	
Total	219	94	313	
	Fisher's exact =		0.007	
1-sided	Fisher's exact =		0.007	

## 17.1. Re-Analysis of Black-Box Warnings Measure, Using Recoded Pre-Deadline Approval Measure and FDA Deadline Approval Measure (cont.)

An exact logistic regression on the same data yields the following results, with a p-value of just above 0.05 for a two-tailed test.

#### Binary Regression

regression (type=logit, model(combobbw\_o = subyrctr pred0410\_r), estimate(subyrctr pred0410\_r), method=exact, mle=firth,

#### **Basic Information**

Data file COMBOBBW-dump20080413.csv

Model combobbw\_o(Response = 1)=%Const+subyrctr+pred0410\_r

Link type Logi

Weight variable <Not Specified>
Stratum variable <Unstratified>
Analysis type Estimate :: Exact

Number of terms in model 3
Number of term(s) dropped 0
Number of observations in analysis 313
Number of records rejected 0
Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	24.49	21	0.2702
Likelihood Ratio	334.2	3	1.665e-009

#### **Parameter Estimates**

rarameter Estimates								
		Point Estim	ate	Confidence Interval and P-Value for Odds Ratio				
					95	%CI	2*1-sided	
Model Term	Type	Odds Ratio SE(Odds)		Type	Lower	Unper	P-Value	
%Const	PMLE	0.0003102	NA.	Asymptotic	8.626e-006	0.01115	1e-005	
				Asymptotic(Profile)	5.783e-006	0.009312		
subyrctr	PMLE	1.273	NA	Asymptotic	1.064	1.524	0.00838	
				Asymptotic(Profile)	1.067	1.545		
	CMLE	1,282	NA	Exact	1.064	1.567	0.007871	
pred0410 r	PMLE	3,416	NA	Asymptotic	1.14	10.24	0.02826	
				Asymptotic(Profile)	1.153	11.02		
	CMLE	3.538	NA.	Exact	0.9935	14.2	0.05137	

PMLE: Penalized MLE for bias correction (Firth's method).

#### Analysis Time = 00:00:00

As the next page shows, a random-effects logistic regression on the same data yields a slightly larger odds ratio estimate, and a p-value of 0.04.

## 17.1. Re-Analysis of Black-Box Warnings Measure, Using Recoded Pre-Deadline Approval Measure and FDA Deadline Approval Measure (cont.)

Random-effects	logistic red	gression		Number o	of obs	=	313
Group variable	(i): discode	Number o	of groups	=	134		
Random effects	u i ~ Gaussi	ian		Obs per	group: mir	n =	1
	_			-			2.3
						_	16
				Wald chi	i2(4)	=	13.20
Log likelihood	= -48.15915	55			chi2		
combobbw_o~1	OR	Std. Err.	Z	P>   z	[95% Cor	nf.	Interval]
+	OR 1.311151						
+ subyrctr   pred0410 r~e	1.311151 3.773193	.1387747	2.56 2.09	0.010 0.037	1.065516 1.084578	 6 3	1.613413 13.12676
+ subyrctr   pred0410 r~e	1.311151	.1387747	2.56 2.09	0.010 0.037	1.065516 1.084578	 6 3	1.613413 13.12676
subyrctr   pred0410_r~e   hhospdisc	1.311151 3.773193	.1387747 2.400124 1.99e-06	2.56 2.09 0.80	0.010 0.037 0.426	1.065516 1.084578 .9999977	 6 <mark>3</mark> 7	1.613413 13.12676 1.000005
subyrctr   pred0410_r~e   hhospdisc   hhosleng	1.311151 3.773193 1.000002	.1387747 2.400124 1.99e-06 .063383	2.56 2.09 0.80	0.010 0.037 0.426 0.290	1.065516 1.084578 .9999977 .947736	 6 3 7 7 6 	1.613413 13.12676 1.000005 1.196757
subyrctr   pred0410_r~e   hhospdisc   hhosleng	1.311151 3.773193 1.000002 1.064993	.1387747 2.400124 1.99e-06 .063383	2.56 2.09 0.80	0.010 0.037 0.426 0.290	1.065516 1.084578 .9999977 .947736	6 3 7 6 	1.613413 13.12676 1.000005 1.196757

.

### 17.2. Analysis of Carpenter/Zucker/Avorn BBW Measure, Using FDA Deadline Approval Measure.

If we re-analyze the original Carpenter/Zucker/Avorn postmarket black-box warning measure and use a recoded (based upon the reclassified standard and priority drugs), we obtain the following results for a cross-tabulation and for Fisher's exact test.

. tab combo	obbw pred01fda_C	ZAdata, e	xact			
combobbw_o riginal		1	Total			
0 1	220	79   9	299 14			
Total	-+   225	88	313			
	Fisher's exact =	's exact = 0.004				
1-sided	Fisher's exact =		0.004			

The next page shows the results of an exact logistic regression (controlling for submission year of the NME) on the same data.

#### 17.2. Analysis of Carpenter/Zucker/Avorn BBW Measure, Using FDA Deadline Approval Measure (cont.)

An exact logistic regression on the same Carpenter/Zucker/Avorn measure, but using the FDA's deadline approval measure, yields an estimated odds ratio of 3.8 and a p-value of 0.04 for a twotailed test.

#### **Binary Regression**

regression (type=logit, model(combobbw\_o = subyrctr pred01fda\_), estimate(subyrctr pred01fda\_), method=exact, mle=firth,

#### Basic Information

Data file COMBOBBW-dump20080413-2.csv

Model Link type combobbw\_o(Response = 1)=%Const+subyrctr+pred01fda\_

Logit

Weight variable Stratum variable <Not Specified> <Unstratified> Analysis type Estimate :: Exact

Number of terms in model Number of term(s) dropped Number of observations in analysis 313 Number of records rejected 0 Number of groups 24

#### **Summary Statistics**

Statistics	Value	DF	P-Value
Deviance	24.47	21	0.2707
Likelihood Ratio	334.7	3	1.896e-009

#### **Parameter Estimates**

	Point Estimate Confidence Interval and P-Value for					/alue for Odd	Odds Ratio	
					95	2*1-sided		
Model Term	Type	Odds Ratio	SE(Odds)	Type	Lower	P-Value		
%Const	PMLE	0.0003543	NA	Asymptotic	1e-005	0.01255	1.283e-005	
				Asymptotic(Profile)	6.686e-006	0.01046		
subyrctr	PMLE	1.263	NA.	Asymptotic	1.055	1.512	0.01084	
				Asymptotic(Profile)	1.059	1.534		
	CMLE	1.272	NA.	Exact	1.056	1.555	0.01017	
pred01fda	PMLE	3.675	NA	Asymptotic	1.22	11.07	0.02072	
				Asymptotic(Profile)	1.233	11.91		
	CMLE	3.803	NA	Exact	1.06	15.35	0.03914	

PMLE: Penalized MLE for bias correction (Firth's method).

## 17.2. Analysis of Carpenter/Zucker/Avorn BBW Measure, Using FDA Deadline Approval Measure (cont.)

A random-effects logistic regression on the same Carpenter/Zucker/Avorn measure, but using the FDA's deadline approval measure, yields an estimated odds ratio of 3.8 and a p-value of 0.04 for a two-tailed test.

The estimates for the FDA's "deadline approval" measure are highlighted in yellow. The random-effects logistic regression shows an estimated odds ratio of just over 4 with a p-value of 0.03 for a two-tailed test.

Random-effects	Number o	f obs	=	313				
Group variable (i): discode				Number of groups			134	
Random effects	Obs per group: min			1				
					avo	g =	2.3	
					max	=	16	
				Wald chi	2 (4)	=	13.62	
Log likelihood		Prob > cl	ni2	=	0.0086			
combobbw_o~1	OR	Std. Err.	z	P> z	[95% Cor	nf.	Interval]	
subyrctr	1.301012	.1376919	2.49	0.013	1.057292	2	1.600913	
pred01fda_~a	4.069912	2.607341	2.19	0.028	1.159505	5	14.28556	
hhosleng	1.065602	.06325	1.07	0.284	.9485739	)	1.197069	
	1.000002							
/lnsig2u	2140948	.745051			-1.674368	3	1.246178	
sigma u	.8984831	.3347079			.4329279	)	1.864679	
	.197033						.5138292	
Likelihood-rati	lo test of rh	.o=0: chibar2	2(01) =	0.72 P	rob >= chi	ba:	c2 = 0.197	

All of the results in this Section 17, of course, use the black-box warning measure of Carpenter/Zucker/Avorn for purposes of demonstration. As we have shown in Sections 11 through 16, our results hold using a revised and corrected measure, using the FDA's 314 NMEs, using the FDA's deadline approval measure, and using the FDA's own withdrawal and BBW measures.