# 2012 NATIONAL SURVEY ON DRUG USE AND HEALTH 

QUESTIONNAIRE DWELLING UNIT-LEVEL AND PERSON PAIRLEVEL SAMPLING WEIGHT CALIBRATION

Prepared for the 2012 Methodological Resource Book<br>RTI Project No. 0212800.001.107.005<br>Contract No. HHSS283201000003C<br>Deliverable No. 41<br>Project Director: Thomas G. Virag

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February 2014

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## Preface

This report documents the method of weight calibration used for producing the final set of questionnaire dwelling unit (QDU) and pair weights for the National Survey on Drug Use and Health (NSDUH) data from 2012. The weighting team faced several challenges in this task and was able to address them by resorting to innovative modifications of certain basic statistical ideas, which are listed below.

- Under Brewer's method, high weights may occur due to small pair selection probabilities. In any calibration exercise, some treatment of extreme value (ev) in weights is needed, but there is a danger of introducing too much bias by overtreatment. In the generalized exponential model (GEM), which is described in detail in Chen et al. (2014), extreme value control is built in, but one needs to define suitable ev domains so that not too many evs are defined. If too many design variables are used to define ev domains, then each domain will be very sparse and will not be of much use in defining thresholds for ev. As in past surveys, a hierarchy of domains was defined using pair age (each pair member being in one of the three categories: 12 to 25,26 to 49 , and $50+$ ) and number of persons aged 12 to 25 in the household, State, and clusters of States (see Section 5.2 for details).
- Control of extreme values in weights helps reduce instability of estimates to some extent, but there is a need for methods that do not introduce much bias. Following the famous suggestion of Hajek (1971) in his comments on Basu's fabled example of circus elephants, we performed ratio adjustment (a form of poststratification) to estimated totals obtained from the household data on the number of persons belonging to the pair domain of interest. This was implemented in a multivariate manner to get one set of final weights.
- In the absence of a suitable source of poststratification controls for the person pairlevel weights and the household-level weights, the inherent two-phase nature of the survey design was capitalized upon to estimate these controls from the first phase of the large screener sample. The first-phase sample weight was poststratified to personlevel U.S. Census Bureau counts to get more efficient estimated counts for pair and household data.
- The problem of multiplicities complicated the issue of providing one set of final weights. When dealing with person-level parameters involving drug-related behaviors among members of the same household, it is possible for an individual to manifest himself or herself in the pair sample through different pairs. To avoid overcounting, the pair weights have to be divided by multiplicity factors, which tend to be domain specific. For this reason, multiplicity factors for a key set of pair analysis domains also are produced along with a set of final calibrated pair weights.
- Missing items in the respondent questionnaire led to imputation for deriving pair relationships, multiplicity factors, and household counts for Hajek adjustments.

The calibration task described in this document has been in place, with minor modifications, since the 1999 version of NSDUH, which was then called the National Household Survey on Drug Abuse (NHSDA). ${ }^{1}$ Results from this calibration applied to an earlier survey year were presented at the 2001 Joint Statistical Meetings. The procedures described in the proceedings papers from these presentations can serve as useful supplemental reference material on estimation in the presence of multiplicities and extreme weights (Chromy \& Singh, 2001) and on GEM calibration of pair weights (Penne, Chen, \& Singh, 2001). The experience of using GEM with person weights is described in an earlier proceedings paper (Chen, Penne, \& Singh, 2000). This work was completed for the Substance Abuse and Mental Health Services Administration (SAMHSA), Center for Behavioral Health Statistics and Quality (CBHSQ), by RTI International (a trade name of Research Triangle Institute), Research Triangle Park, North Carolina, under Contract No. HHSS283201000003C. The authors would like to take this opportunity to thank a number of individuals for useful discussions and suggestions: Joe Gfroerer and Art Hughes of SAMHSA and Jim Chromy of RTI.

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## List of Terms and Abbreviations

| DU | Dwelling unit. |
| :--- | :--- |
| Ev | Extreme value. See Sections 5.1 and 5.2 for more detail. |
| GEM | Generalized exponential model. See Chapter 3 for more detail. |
| Household-level | The number of pairs associated with a given domain in a given household. <br> person count <br> These counts are used as control totals in the poststratification step. See <br> Chapter 11 in the editing and imputation report (Carpenter et al., 2014) for <br> details on how these counts are created, and Chapter 4 for details on their <br> use in poststratification. |
| IQR | Interquartile range. |
| Multiplicity factor | The number of pairs associated with a given respondent in a given domain. <br> See Carpenter et al. (2014) for more detail. |
| Nr | Nonresponse. |
| Outwinsor | The proportion of weights trimmed after extreme value adjustment via <br> winsorization. |
| Pair domain relationship | A pair relationship where the target population is defined by one of the pair <br> members, conditional on the attributes of the other pair member. <br> The relationship between selected pair members. |
| Parent-child | A pair relationship where either both pair members identify the other as <br> part of a parent-child relationship, or both pair members otherwise are <br> determined to form a parent-child pair (either through other evidence or <br> through imputation). |
| poststratification. |  |


| res.pr.ps | Respondent pair poststratification adjustment step. See Section 6.3 .4 for <br> more detail. |
| :--- | :--- |
| res.qdu.ps | Respondent questionnaire dwelling unit poststratification adjustment step. <br> See Section 6.2 .4 for more detail. |
| SDU | Screener dwelling unit: a household where screener information is <br> available. |
| sel.pr.ps | Selected person pair poststratification adjustment step. See Section 6.3 .2 <br> for more detail. |
| sel.qdu.ps | Selected questionnaire dwelling unit poststratification adjustment step. See <br> Section 6.2 .2 for more detail. |
| Sibling-sibling | A pair relationship where the pair members are siblings (either reported to <br> be so, or otherwise determined to be so). |
| Spouse-spouse | A pair relationship where the pair members are either married or living <br> together as though married (either reported to be so, or otherwise <br> determined to be so). |
| SS | State sampling. |
| UWE | Unequal weighting effect. It refers to the contribution in the design effect <br> due to unequal selection probability and is defined as $1+[(n-1) / n] * C V^{2}$, <br> where $C V=$ coefficient of variation of weights and $n$ is the sample size. |
| Winsorization | A method of extreme value adjustment that replaces extreme values with <br> the critical values used for defining low and high extreme values. |

## 1. Introduction

Traditionally, most household surveys have been designed either to measure characteristics of the entire household or to focus on a randomly selected respondent from among those determined to be eligible for the survey. Selecting more than one person from the same household is generally avoided since persons from the same household often exhibit the same or similar characteristics and behavioral patterns. The intra-class correlation found among members of the same household leads to a clustering effect on the variance of estimates resulting in less precise estimates compared with estimates of the same sample size from a simple random sample. Selecting only one person per household avoids this clustering effect on the variance. The "one person per household" sampling approach, however, precludes the opportunity to gather information about the relationships among household members. In the National Survey on Drug Use and Health (NSDUH), ${ }^{2}$ we allow for a richer analytic capability of a survey designed to ensure a positive pairwise probability of selection among all eligible household members in each sample household. Achieving positive probabilities for all pairs within sampled households permits unbiased estimation of the within-dwelling-unit component of variance. Besides providing efficient data collection, this sampling method also facilitates the study of the relationships of social behaviors among members of the same household. This report documents the methodology and development of calibrated weights for the second objective, the study of behavioral relationships among persons residing in the same household. The report also describes the development of questionnaire dwelling unit (QDU) weights, which are of independent interest for studying household-level characteristics and also are needed for producing household count estimates of the number of persons belonging to pair relationship domains for use as poststratification controls for pair weights.

NSDUH allows for estimating characteristics at the person level, pair level, and household or QDU level. This report describes the weight calibration methods used for the pairand QDU-level respondents. As described in the person-level report, NSDUH is an annual survey of about 67,500 persons selected from the civilian, noninstitutionalized population aged 12 or older from all 50 States and the District of Columbia. Based on a composite size measure, States were geographically partitioned into roughly equal-sized regions according to population. The 42 smaller States and the District of Columbia were partitioned into 12 State sampling (SS) regions, whereas the eight largest States were divided into 48 SS regions. Therefore, the partitioning of the United States resulted in the formation of a total of 900 SS regions. Under a stratified design with States serving as the primary strata and SS regions serving as the secondary strata, census tracts, segments within census tracts, and dwelling units (DUs) within segments were each selected using probability proportional to size sampling. NSDUH is sometimes referred to as a two-phase sample where the first phase consisted of a large number of screener dwelling units (SDUs, about 200,000) selected to ensure that various age groups (five in all: 12 to 17,18 to 25,26 to 34,35 to 49 , and $50+$ ) of eligible individuals were represented adequately in the second phase. Information collected from SDUs also provided estimates of population controls (as in two-phase sampling) for calibration at levels (such as pair and QDU) for which suitable U.S. Census Bureau-based controls were not available. The second phase consisted of

[^1]the selection of zero, one, or two persons from each selected SDU using a modification of Brewer's method such that prescribed sampling rates for the five age groups in each State were achieved with high selection rates for youths (12 to 17) and young adults (18 to 25). Table 1.1 shows the eligible number of selected and responding SDUs, QDUs, pairs, and persons for each of the 5 years (2008-2012). The distribution of pair data for different pairs of age groups may vary considerably (see Chapter 2 for details). It is seen that for certain age group domains, the realized sample size may not be sufficient to yield reliable estimates. Also, there may be problems of extreme weights due to small pair selection probabilities under Brewer's method that may cause instability of estimates. These and some other estimation issues related to pair data are discussed below, along with some adopted solutions.

Table 1.1 2008-2012 NSDUH Sample Sizes

| Sample Unit |  | $\mathbf{2 0 0 8}$ | $\mathbf{2 0 0 9}$ | $\mathbf{2 0 1 0}$ | $\mathbf{2 0 1 1}$ | $\mathbf{2 0 1 2}$ |
| :--- | :--- | :--- | ---: | ---: | ---: | ---: |
| SDU | Selected | 160,114 | 161,377 | 166,532 | 179,293 | 178,586 |
|  | Completed | 142,159 | 142,933 | 147,010 | 156,048 | 153,873 |
| QDU | Selected | 58,942 | 58,288 | 58,702 | 61,441 | 60,621 |
|  | Completed | 48,180 | 48,088 | 48,113 | 50,133 | 48,850 |
| Pair | Selected | 26,769 | 26,497 | 26,295 | 27,095 | 27,035 |
|  | Completed | 19,748 | 19,919 | 19,691 | 19,976 | 19,459 |
| Person | Selected | 85,711 | 84,785 | 84,997 | 88,536 | 87,656 |
|  | Completed | 67,928 | 68,007 | 67,804 | 70,109 | 68,309 |

Note: The 2008-2010 sample sizes reflect the removal of falsified cases found in Pennsylvania and Maryland. The 2011 and 2012 samples were not affected. For additional information, see Section B.3.5 in Appendix B of the 2011 and 2012 national findings reports (Center for Behavioral Health Statistics and Quality, 2012; 2013).

First, we note that for studying drug-related behavioral relationships among members of the same household, pair data is required because the outcome variable generally is defined with respect to the specific other member selected from the household. However, the parameter of interest is generally at the person level and is not at the pair level. For example, in the parentchild pairs, one may be interested in the proportion of children that have used drugs in the past year who have parents that report talking to their child about drugs. Here the target population consists only of children, and not all possible pairs. Note that the pair-level (two persons per QDU) sample forms a subsample of the larger person-level (one or two persons per QDU) sample, with the QDUs themselves selected from the larger sample of SDUs. NSDUH has features of a two-phase design, which turns out to be useful for estimating calibration controls for poststratification of household-level weights and person pair-level weights. No other outside source is available for obtaining these controls. For this purpose, the screener-level household weights are poststratified to person-level census counts to obtain more efficient estimated controls for pair and household data.

In estimation for pair domains, two major problems arise: one is that of multiplicities because, for a given domain defined by the pair relationship, when the parameter of interest is at the person level, several pairs in the household could be associated with the same person, For example, analysts are interested in an outcome at the person level, the proportion of children who use drugs and whose parents report talking to them about drugs, where the focus is on the child in a parent-child pair. Several parent-child pairs in the household could be associated with the same child. If the household has two parents, the selected child has two inclusion possibilities
(one with each parent) in the set of all such parent-child pairs (Carpenter et al., 2014). The other problem is that of extreme weights that may arise due to small selection probabilities for certain pair age groups, which may lead to unstable estimates. Each of these issues is discussed in turn.

If several pairs in the household are associated with the same person, it is necessary to use the average measure of behavior relationships for each member, which gives rise to multiplicities. Thus, the pair weights need to be divided by the person-level multiplicity factors for each domain of interest, and, therefore, multiplicity factors need to be produced along with the final set of calibrated weights. Because it is not straightforward to create these multiplicities, analyses would have to be necessarily limited to pair relationships where the multiplicities were produced a priori. It was anticipated that analyses of interest would be limited to 14 pair domains, listed in Table 1.2. Since no multiplicity was necessary for the spouse-spouse/partnerpartner pair relationships (by definition, each pair member could have only one partner or one spouse), multiplicity factors were produced for only 12 of these domains. Note that a single pair relationship might have two domains associated with it, since the parameter of interest might be associated with only one member of the pair (the "focus" member), and the multiplicity would differ depending upon which pair member was the focus member.

Table 1.2 Pair Domains

| Pair Relationship | Focus |
| :--- | :--- |
| Parent-child: parent, child aged 12-14 | Parent |
| Parent-child: parent, child aged 12-14 | Child |
| Parent-child: parent, child aged 12-17 | Parent |
| Parent-child: parent, child aged 12-17 | Child |
| Parent-child: parent, child aged 12-20 | Parent |
| Parent-child: parent, child aged 12-20 | Child |
| Parent-child: parent, child aged 15-17 | Parent |
| Parent-child: parent, child aged 15-17 | Child |
| Sibling-sibling: older sibling 15-17, younger sibling 12-14 | Older sibling |
| Sibling-sibling: older sibling 15-17, younger sibling 12-14 | Younger sibling |
| Sibling-sibling: older sibling 18-25, younger sibling 12-17 | Older sibling |
| Sibling-sibling: older sibling 18-25, younger sibling 12-17 | Younger sibling |
| Spouse-spouse or partner-partner, with or without children | No multiplicity necessary |
| Spouse-spouse or partner-partner, with children aged 0-17 | No multiplicity necessary |

Some of the multiplicities, including counts of all possible pairs in a household for a given domain, were used for poststratification. Details are provided in Chapter 4. Prior editions of this report contained a chapter on editing and imputation of pair relationships, multiplicity factors, and household-level person counts for poststratification, but this information has been removed in the interest of consolidating the imputation documentation and can now be found in the editing and imputation report (Carpenter et al., 2014). Special consideration is required for analysis of pair-level data, and details can be found in How To Prepare and Analyze Pair Data in the National Survey on Drug Use and Health (Frechtel, Warren, \& Porter, in press). See Appendix M for an excerpt from this manual on inferential population and multiplicities.

A resolution to the extreme weight problem is to use a Hajek-type modification (Hajek, 1971). This modification essentially entails calibration (like poststratification) to controls for the number of persons in households belonging to each domain of interest. These controls can be obtained from the larger sample of singles and pairs (i.e., one or two persons selected from DUs). Note, however, that the multiplicity factor, being domain specific, renders the calibration adjustment factor domain specific. This raises the question of finding one set of calibration weights for use with all domains or outcome variables. To get around this problem, we performed a multivariate calibration with respect to a key set of pair domains. This type of poststratification then was followed by a repeat poststratification to further control the extreme weights by imposing separate bound restrictions on the initially identified extreme weights.

The generalized exponential model (GEM) method (Folsom \& Singh, 2000) was used for calibration of both QDU- and pair-level design weights through several steps of adjustment as shown in Exhibit 1.1. In GEM, treatment of extreme value (ev) weights is built in via the definition of lower and upper bounds for the extreme weights. For pair data, there was a problem defining suitable domains for defining extreme weights, as explained in the following paragraphs.

Exhibit 1.1 QDU and Pair Sampling Weight Calibration Steps


In dealing with extreme weights, it is assumed that they arise due to design (due to an imperfect frame, assignment of very small selection probabilities to some units, or a small weight adjustment factor after calibration) so that they make the sample representative of the population and, hence, do not introduce bias. The only problem is that they may lead to highly unstable
estimates similar to the problem of Basu's circus elephants ${ }^{3}$ (Hajek, 1971). So, we need to perform some treatment (such as winsorization ${ }^{4}$ ) within suitably defined extreme weight domains such that these domains contain units possibly from different strata but with similar sample selection probabilities to avoid the occurrence of extreme weights due to a mix of different designs. The domains must be large enough (e.g., at least size 30) to be able to define extreme values according to the domain-specific weight distribution. Any extreme value treatment to increase precision of estimates would introduce some bias. However, this bias can be reduced considerably if the ev treatment is performed under calibration controls. This is what the built-in ev control in GEM tries to accomplish.

It follows that the definition of extreme weight domains should depend on factors that affect the selection probabilities of units in the sample, such as State- and age-specific sampling rates, segment selection probabilities, pair age-specific selection probabilities, and household composition. If one tries to define extreme weight domains by taking account of all these factors via cross-classification, it will lead to too many domains with insufficient observations. That is why it is difficult to define suitable extreme weight domains for pair data. In the case of personlevel weights it was less difficult, since State by age group suitably captured the extreme weight domain requirements. The definition of extreme weight domains used in the 2012 survey was the same as the one used in the 1999-2011 surveys. The domains were defined as the crossclassification of State, pair age, ${ }^{5}$ and number of persons aged 12 to 25 in a household. In particular, the pair age was defined by the age groups of each pair member according to the age categories of 12 to 25 , 26 to 49 , and 50 or older (resulting in six pair age categories), and the number of persons aged 12 to 25 were categorically defined as zero, one, and two or more. For more details, see Chapter 5.

[^2]
## 2. Questionnaire Dwelling Unit and Pair Selection Probabilities

Similar to the 1999-2001 National Household Surveys on Drug Abuse (NHSDAs) and the 2002-2011 National Surveys on Drug Use and Health (NSDUHs), ${ }^{6}$ the 2012 NSDUH had a two-phase design and used a computer-assisted interviewing (CAI) method. There were four stages of selection: census tracts, segments within census tracts, dwelling units (DUs) within segments, and persons within dwelling units. Any two survey eligible persons had some nonzero chance of being selected and, when both were selected, they formed a within household pair. This design feature is of interest to NSDUH researchers because, for example, it allows analysts to examine how the drug use propensity of an individual (in a family) relates to the drug use propensity of other members residing in the same dwelling unit (Morton, Martin, Shook-Sa, Chromy, \& Hirsch, 2013).

For the 1999-2001 surveys, the method used for selecting pairs was as follows. For a given DU, if the sum of the age-specific selection probabilities was larger than 2, then the individual person-selection probabilities were ratio adjusted downward to make the sum equal to 2. If the sum was less than 2, the difference between 2 and the sum of the probabilities was evenly distributed over three dummy persons so that the sum of the person probabilities was made to equal 2. Brewer's method was then applied to select a person pair. If the selected pair consisted of two real persons, then both persons were selected. If the selected pair consisted of one real person and one dummy person, then the real person was selected. If the selected pair consisted of two dummy persons, no one was selected from that DU.

Starting with the 2002 NSDUH and continuing through 2012, the pair-sampling algorithm was modified to increase the number of pairs selected in the sample. Dwelling units with the sum of person selection probabilities greater than or equal to 2 were treated the same as in previous survey years. However, DUs where the sum of person-level selection probabilities was less than 2 received a slightly different treatment that increased the chance for selecting a pair of real persons. Section 2.1 describes the selection process for both types of DUs.

Table 2.1 provides a summary of these NSDUH sampling units: eligible and completed screening dwelling units (SDUs), selected and completed questionnaire dwelling units (QDUs), selected and completed person interviews, and selected and completed person pairs, as well as their response rates. Using Brewer's method, zero, one, or two individuals were selected per household. Those SDUs where at least one person was selected were counted as the selected QDUs. A QDU where two persons were selected and both had completed interviews was considered to be a completed person pair. The table provides a breakdown by age group at the person level and age group by selection group (none, single, or pair) at the person pair level.

[^3]
### 2.1 Pair Selection Probability

### 2.1.1 Case I: DUs with $S \geq 2$

For a given DU, if the sum of the age-specific person selection probabilities $(S)$ was larger than 2 , then the selection probability was ratio adjusted by a multiplicative adjustment factor so that all probabilities were scaled down to sum to exactly 2 . Now, Brewer's method sets the pairwise selection probabilities at

$$
\begin{equation*}
P_{h(i j)}=\left[\frac{P_{h(i)} P_{h(j)}}{K}\right]\left[\frac{1}{1-P_{h(i)}}+\frac{1}{1-P_{h(j)}}\right] \tag{2.1}
\end{equation*}
$$

by setting $K$ at

$$
\begin{equation*}
K=2+\sum \frac{P_{h(i)}}{1-P_{h(i)}} \tag{2.2}
\end{equation*}
$$

where $i=i^{\text {th }}$ person in household $h$ (whose selection probability depends on his or her age category: 1, 2, 3, 4, or 5) and
$j=j^{t h}$ person in household $h$ (whose selection probability depends on his or her age category: $1,2,3,4$, or 5),
where age category 1 corresponds to persons aged 12 to 17,2 to persons aged 18 to 25,3 to persons aged 26 to 34,4 to persons aged 35 to 49 , and 5 to persons aged 50 or older.

The sum of the pairwise selection probabilities taken over all unique pairs will be guaranteed to be exactly 1 .

$$
\begin{equation*}
\sum_{i} \sum_{j>i} P_{h(i j)}=1 \tag{2.3}
\end{equation*}
$$

It also guarantees that the sum of the pairwise selection probabilities for an individual is equal to the individual's selection probability

$$
\begin{equation*}
\sum_{j \neq i} P_{h(j)}=P_{h(i)} \tag{2.4}
\end{equation*}
$$

for all values of $i$.
Note the above scheme always selects a pair of two eligible persons.

### 2.1.2 Case II: DUs with $S<2$

If the sum $S$ of person-level selection probabilities was less than 2 , the method used in survey years 1991 to 2001 consisted of dividing $2-S$ equally among the three dummy persons
added to the household, and then used Brewer's method to select a pair, as in Case I. However, if the household had two or more persons, we preferred a pair of real persons to have a greater chance of being selected. To achieve this goal, the individual selection probabilities, $P_{h(i)}$, were scaled upward by the factor $F_{s}$ such that their sum came close to but did not exceed 2 and such that each person selection probability did not exceed the maximum allowed probability of 0.99 . Thus, denoting the revised person selection probabilities by $P^{\prime}{ }_{h(i)}$, the factor $F_{s}$ is given by

$$
\begin{equation*}
F_{s}=\operatorname{Min}\left\{\frac{T(\lambda)}{S}, \frac{0.99}{\max \left\{P_{h(i)}\right\}}\right\} \tag{2.5}
\end{equation*}
$$

where $T(\lambda)=S+\lambda(2-S)$ and $\lambda$ is set to 0.5 . Note that if $\lambda$ is chosen as 0 , then $F_{s}=1$ and the selection scheme would follow that of Case I. The individual person probabilities are scaled upward by the factor $F_{s}$ so they either sum to 2 or sum as close to 2 as possible. Denote $S^{\prime \prime}$ as the sum of the selection probability after scale adjustment by $F_{s}$. If $S^{\prime}$ is exactly 2 , then dummy persons are not needed. If $S^{\prime}$ is less than 2 , then three dummy persons are added to the DU .

Now, for Brewer's method, we set the pairwise selection probabilities similar to (2.1), as

$$
\begin{equation*}
P_{h(i)}^{\prime}=\left[\frac{P_{h(i)}^{\prime} P_{h(j)}^{\prime}}{K^{\prime}}\right]\left[\frac{1}{1-P_{h(i)}^{\prime}}+\frac{1}{1-P_{h(j)}^{\prime}}\right] \tag{2.6}
\end{equation*}
$$

by setting $K^{\prime}$ at

$$
\begin{equation*}
K^{\prime}=2+\sum_{i} \frac{P_{h(i)}^{\prime}}{1-P_{h(i)}^{\prime}}, \tag{2.7}
\end{equation*}
$$

where $P^{\prime}{ }_{h(i)}$ and $P_{h(j)}^{\prime}$ are the selection probabilities adjusted by the scaling factor $F_{s}$,
where $i=i^{\text {th }}$ person in the household (whose selection probability depends on his or her age category: $0,1,2,3,4$, or 5 ),
$j=j^{\text {th }}$ person in the household (whose selection probability depends on his or her age category: $0,1,2,3,4$, or 5), and
where age category 0 corresponds to dummy persons, and categories 1 to 5 are defined as in Case I.

Note that we now have $\sum_{j \neq i} P_{h(i j)}^{\prime}=P_{h(i)}^{\prime}$. To maintain the original person selection probabilities despite the scale adjustment by $F_{s}$, we modified Brewer's method as follows. First, draw a random number, $R$, from a uniform $(0,1)$ distribution. If $R \leq 1 / F_{s}$, then select a pair using Brewer's method based on formula (2.6). However, if $R>1 / F_{s}$, then no persons are selected from
the household. In this way, the probability for selecting a pair $(i, j)$ in household $h$ becomes $P_{h(i j)}^{*}$ $=P_{h(i j)}^{\prime} / F_{s}$, which, in turn, gives the original person selection probabilities, $P_{h(i)}$. Unlike Case I, where a pair of eligible persons was always selected, this adjusted selection scheme allows for zero, one, or two persons to be selected from a DU.

### 2.2 Questionnaire Dwelling Unit Selection Probability

A dwelling unit was considered a selected QDU if it had completed the screening interview and had at least one person selected for the questionnaire interview. QDUs with at least one respondent were considered respondent QDUs.

The QDU selection probability was defined as

$$
\begin{equation*}
P_{h}^{*}=\left(1-P_{h(00)}^{*}\right), \tag{2.8}
\end{equation*}
$$

where $P^{*}{ }_{h(00)}$ is the probability of not selecting any person. For the DUs with an unadjusted sum of age-specific selection probabilities larger than or equal to 2 (Case I), $P^{*}{ }_{h(00)}$ is 0 . It follows from Section 2.1, under Case II, $P^{*}{ }_{h(00)}$ can be calculated as

$$
\begin{equation*}
P_{h(00)}^{*}=\left(1-\frac{1}{F_{s}}\right)+\frac{3}{F_{s}}\left[\frac{P_{h(0)}^{\prime} P_{h(0)}^{\prime}}{K^{\prime}}\right]\left[\frac{1}{1-P_{h(0)}^{\prime}}+\frac{1}{1-P_{h(0)}^{\prime}}\right], \tag{2.9}
\end{equation*}
$$

where $P^{\prime}{ }_{h(0)}$ is the selection probability of a dummy person when person selection probabilities are adjusted by $F_{s}$.

Table 2.1 Building Blocks of the QDU and Person Pair Samples: Dwelling Units and Persons in the 2008-2012 NSDUHs

| Domain | 2008 |  |  | 2009 |  |  | 2010 |  |  | 2011 |  |  | 2012 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Sel. ${ }^{1}$ | Resp. ${ }^{2}$ | \% Rate ${ }^{\text {3 }}$ | Sel. ${ }^{1}$ | Resp. ${ }^{2}$ | \% Rate ${ }^{\text {3 }}$ | Sel. ${ }^{1}$ | Resp. ${ }^{2}$ | \% Rate ${ }^{3}$ | Sel. ${ }^{1}$ | Resp. ${ }^{2}$ | \% Rate ${ }^{3}$ | Sel. ${ }^{1}$ | Resp. ${ }^{2}$ | \% Rate ${ }^{\text {3 }}$ |
| DUs |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Total DUs Screened | 160,114 | 142,159 | 88.79 | 161,377 | 142,933 | 88.57 | 166,532 | 147,010 | 88.28 | 179,293 | 156,048 | 87.04 | 178,586 | 153,873 | 86.16 |
| QDUs |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Total QDUs | 58,942 | 48,180 | 81.74 | 58,288 | 48,088 | 82.50 | 58,702 | 48,113 | 81.96 | 61,441 | 50,133 | 81.60 | 60,621 | 48,850 | 80.58 |
| Persons |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Total Persons | 85,711 | 67,928 | 79.25 | 84,785 | 68,007 | 80.21 | 84,997 | 67,804 | 79.77 | 88,536 | 70,109 | 79.19 | 87,656 | 68,309 | 77.93 |
| 12-17 | 26,228 | 22,250 | 84.83 | 26,157 | 22,398 | 85.63 | 25,908 | 21,960 | 84.76 | 27,911 | 23,510 | 84.23 | 27,147 | 22,473 | 82.78 |
| 18-25 | 28,793 | 22,875 | 79.45 | 28,158 | 22,686 | 80.57 | 28,164 | 22,793 | 80.93 | 28,589 | 22,876 | 80.02 | 28,639 | 22,529 | 78.67 |
| 26-34 | 8,337 | 6,560 | 78.69 | 8,242 | 6,591 | 79.97 | 8,545 | 6,780 | 79.34 | 8,323 | 6,543 | 78.61 | 8,304 | 6,484 | 78.08 |
| 35-49 | 12,995 | 9,674 | 74.44 | 12,855 | 9,616 | 74.80 | 12,979 | 9,668 | 74.49 | 12,220 | 9,149 | 74.87 | 12,364 | 9,076 | 73.41 |
| 50+ | 9,358 | 6,569 | 70.20 | 9,373 | 6,716 | 71.65 | 9,401 | 6,603 | 70.24 | 11,493 | 8,031 | 69.88 | 11,202 | 7,747 | 69.16 |
| Pairs |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Total Pairs ${ }^{4}$ | 26,769 | 19,748 | 73.77 | 26,497 | 19,919 | 75.17 | 26,295 | 19,691 | 74.88 | 27,095 | 19,976 | 73.73 | 27,035 | 19,459 | 71.98 |
| $0,0{ }^{5}$ | 83,217 | N/A | N/A | 84,645 | N/A | N/A | 88,308 | N/A | N/A | 94,607 | N/A | N/A | 93,252 | N/A | N/A |
| 0, 12-17 | 8,634 | 8,094 | 93.75 | 8,432 | 7,936 | 94.12 | 8,595 | 7,906 | 91.98 | 9,402 | 8,651 | 92.01 | 9,017 | 8,277 | 91.79 |
| 0, 18-25 | 9,932 | 9,213 | 92.76 | 9,870 | 9,081 | 92.01 | 10,093 | 9,270 | 91.85 | 10,306 | 9,497 | 92.15 | 10,325 | 9,461 | 91.63 |
| 0, 26-34 | 3,771 | 3,303 | 87.59 | 3,798 | 3,318 | 87.36 | 3,914 | 3,418 | 87.33 | 3,930 | 3,386 | 86.16 | 3,856 | 3,327 | 86.28 |
| 0, 35-49 | 4,601 | 3,835 | 83.35 | 4,565 | 3,810 | 83.46 | 4,659 | 3,843 | 82.49 | 4,431 | 3,704 | 83.59 | 4,368 | 3,645 | 83.45 |
| 0, 50+ | 5,235 | 4,013 | 76.66 | 5,126 | 4,042 | 78.85 | 5,146 | 3,998 | 77.69 | 6,277 | 4,919 | 78.37 | 6,020 | 4,681 | 77.76 |
| 12-17, 12-17 | 4,319 | 3,654 | 84.60 | 4,288 | 3,648 | 85.07 | 4,160 | 3,525 | 84.74 | 4,649 | 3,885 | 83.57 | 4,507 | 3,668 | 81.38 |
| 12-17, 18-25 | 3,449 | 2,668 | 77.36 | 3,595 | 2,852 | 79.33 | 3,444 | 2,718 | 78.92 | 3,756 | 2,921 | 77.77 | 3,627 | 2,759 | 76.07 |
| 12-17, 26-34 | 897 | 690 | 76.92 | 872 | 724 | 83.03 | 922 | 752 | 81.56 | 834 | 685 | 82.13 | 825 | 658 | 79.76 |
| 12-17, 35-49 | 3,944 | 3,012 | 76.37 | 3,979 | 3,061 | 76.93 | 3,948 | 3,044 | 77.10 | 3,855 | 2,918 | 75.69 | 3,813 | 2,812 | 73.75 |
| 12-17, 50+ | 666 | 482 | 72.37 | 703 | 532 | 75.68 | 679 | 493 | 72.61 | 766 | 565 | 73.76 | 851 | 631 | 74.15 |
| 18-25, 18-25 | 5,872 | 4,261 | 72.56 | 5,588 | 4,168 | 74.59 | 5,502 | 4,165 | 75.70 | 5,476 | 4,015 | 73.32 | 5,476 | 3,901 | 71.24 |
| 18-25, 26-34 | 1,103 | 821 | 74.43 | 1,102 | 820 | 74.41 | 1,140 | 851 | 74.65 | 1,049 | 768 | 73.21 | 1,079 | 794 | 73.59 |
| 18-25, 35-49 | 1,706 | 1,112 | 65.18 | 1,576 | 1,059 | 67.20 | 1,639 | 1,098 | 66.99 | 1,469 | 994 | 67.67 | 1,582 | 1,053 | 66.56 |
| 18-25, 50+ | 859 | 546 | 63.56 | 839 | 547 | 65.20 | 844 | 537 | 63.63 | 1,057 | 666 | 63.01 | 1,074 | 660 | 61.45 |
| 26-34, 26-34 | 890 | 614 | 68.99 | 886 | 635 | 71.67 | 903 | 621 | 68.77 | 858 | 599 | 69.81 | 880 | 604 | 68.64 |
| 26-34, 35-49 | 545 | 380 | 69.72 | 447 | 312 | 69.80 | 512 | 354 | 69.14 | 492 | 314 | 63.82 | 469 | 320 | 68.23 |
| 26-34, 50+ | 241 | 145 | 60.17 | 251 | 149 | 59.36 | 251 | 162 | 64.54 | 302 | 192 | 63.58 | 315 | 177 | 56.19 |
| 35-49, 35-49 | 873 | 535 | 61.28 | 917 | 569 | 62.05 | 886 | 542 | 61.17 | 748 | 474 | 63.37 | 833 | 487 | 58.46 |
| 35-49, 50+ | 453 | 269 | 59.38 | 454 | 239 | 52.64 | 449 | 245 | 54.57 | 477 | 271 | 56.81 | 466 | 272 | 58.37 |
| 50+, 50+ | 952 | 559 | 58.72 | 1,000 | 604 | 60.40 | 1,016 | 584 | 57.48 | 1,307 | 709 | 54.25 | 1,238 | 663 | 53.55 |

$\mathrm{DU}=$ dwelling unit; $\mathrm{N} / \mathrm{A}=$ not applicable; $\mathrm{QDU}=$ questionnaire dwelling unit.
Note: The 2008-2010 sample sizes reflect the removal of falsified cases found in Pennsylvania and Maryland. The 2011 and 2012 samples were not affected. For additional information, see Section B. 3.5 in Appendix B of the 2011 and 2012 national findings reports (Center for Behavioral Health Statistics and Quality, 2012; 2013).
${ }^{1}$ Selected pairs are based on the screener age.
${ }^{2}$ Respondent pairs are based on the questionnaire age and are comprised only of respondent persons.
${ }^{3}$ These rates are unweighted and based only on the total selected and total responding counts of pairs.
${ }^{4}$ Total pairs excludes dummy person pairs.
${ }^{5}$ Among the completed screening dwelling units, no person was selected in this dwelling unit.

## 3. Brief Description of the Generalized Exponential Model

In survey practice, design-based weights are typically adjusted in three steps: (1) for extreme values (ev) via winsorization, (2) for nonresponse ( nr ) via weighting classes, and (3) for poststratification (ps) via raking ratio adjustments. If weights are not treated for extreme values, the resulting estimates, although unbiased, will tend to have low precision. The bias introduced by winsorization is alleviated to some extent through ps . The nr adjustment is a correction for bias introduced in estimates based only on responding units, and ps is an adjustment for coverage (typically undercoverage) bias and variance reduction due to correlation between the study and control (usually demographic) variables.

There are limitations in the existing methods of weight adjustment for ev , nr , and ps . It would be desirable to adjust for bias introduced in the ev step (when extreme weights are treated via winsorization) in that the sample distribution for various demographic characteristics is preserved. For the nr step, there are general raking type methods, such as the scaled constrained exponential model developed by Folsom and Witt (1994), where the lower and upper bounds can be suitably chosen by use of a separate scaling factor. The factor is set as the inverse of the overall response propensity. It would be desirable to have a model for the nr adjustment factor so that the desired lower and upper bounds on the factor are part of the model. Note that the lower bound on the nr adjustment factor should be one, as it is interpreted as the inverse of the probability of response for a particular unit. For the ps step, on the other hand, the general calibration methods of Deville and Särndal (1992), such as the logit method, allow for built-in lower (L) and upper ( U ) bounds (for ps, typically $\mathrm{L}<1<\mathrm{U}$ ). However, it would be desirable to have nonuniform bounds ( $L_{k}, U_{k}$ ) depending on the unit $k$ such that the final adjusted weight, $w_{k}$, could be controlled within certain limits. An important application of this feature would be weight adjustments in the presence of ev to allow some control on the final adjustment of the initially identified extreme values.

A modification of the earlier method of the scaled constrained exponential model of Folsom and Witt (1994), termed as the method of the generalized exponential model (GEM) and proposed by Folsom and Singh (2000), provides a unified approach to the three weight adjustments for ev , nr , and ps , and it has the desired features mentioned above. The functional form of the GEM adjustment factor is provided in Appendix A. It generalizes the logit model of Deville and Särndal (1992), typically used for ps, such that the bounds (L, U) may depend on $k$. Thus, it provides a built-in control on ev during both ps and nr adjustments. In addition, the bounds are internal to the model and can be set to chosen values (e.g., $L_{k}=1$ in the nr step). If there is a low frequency of ev in the final ps, then a separate ev step may not be necessary.

In fitting GEM to a particular problem, the choice of a large number of predictor variables along with tight bounds will have an impact on the resulting unequal weighting effect (UWE) and the proportion of extreme values. In practice, this leads to somewhat subjective considerations of trade-off between the target set of bounds for a given set of factor effects and the target UWE and the target proportion of extreme values. It also may be beneficial to look at
the proportion of "outwinsors" (a term coined to signify the extent of residual weights after winsorization), which is probably more realistic in determining the robustness of estimates in the presence of extreme values.

A large increase in the number of predictor variables in GEM typically would result in a higher UWE, thus indicating a possible loss in precision. This was checked by comparing SUDAAN-based standard errors of a key set of estimates computed from two sets of calibration models, one baseline using only the main effects and the other using the final model. The results are presented in Chapter 7.

To implement GEM, several steps need to be followed: (1) define and create all the covariates; (2) define the extreme weights; (3) fit the GEM model. The details of practical aspects of GEM implementation can be found in Chapters 4 and 5 of this report and Chapter 4 of Chen et al. (2014).

# 4. Predictor Variables for the Questionnaire Dwelling Unit and Pair Weight Calibration via the Generalized Exponential Model 


#### Abstract

We note that unlike the person-level weight calibration, the control totals for the questionnaire dwelling unit (QDU)-level and person pair-level poststratification are not available from the U.S. Census Bureau. A way around this problem is to take advantage of the two-phase nature of the design, in which the screener data provides a large sample containing demographic information that can be used to derive control totals for the QDU-level and person pair-level sampling weight calibrations, as well as for the selected person poststratification adjustment. The stability of control totals from the screener dwelling unit (SDU)-level data can be improved by poststratification of the SDU sample using person-level counts from the census. This was indeed done and is documented in the person-level weight calibration report (Chen et al., 2014).


### 4.1 Questionnaire Dwelling Unit Weight Calibration

After the nonresponse and poststratification adjustments at the SDU level, which are common to the person-level weight calibration, the QDU sample weights were adjusted in three steps: poststratification of selected QDUs, nonresponse adjustment of respondent QDUs, and poststratification of respondent QDUs. The set of initially proposed predictor variables for these adjustments using generalized exponential model (GEM) were set to be common and to correspond to those used for the SDU nonresponse and poststratification adjustments. The variables are of two types: Those used for SDU nonresponse adjustment are $0 / 1$ indicators, while those used for SDU poststratification adjustment are counting variables. The variables of the first type ( $0 / 1$ indicators) are population density, group quarters, race/ethnicity of householder, percentage of persons in segment who are black or African American, percentage of persons in segment who are Hispanic or Latino, percentage of owner-occupied dwelling units (DUs) in segment, segment-combined median rent and housing value, and household type. Variables of the second type (counting variables) represent the number of eligible persons within each DU who fall into the various demographic categories of race, age group, Hispanicity, and gender. Note that the State and quarter variables are represented as both binary and counting variables. Thus, not only are DU counts within a specific State or quarter in the QDU sample controlled to the corresponding totals obtained from the SDU sample, but also counts of persons living in the DUs in the QDU sample are controlled to totals from the SDU sample. These person-level totals match the census estimates because of the SDU-level poststratification to census counts. It may be noted that in the poststratification of selected QDUs and the nonresponse adjustment of the respondent QDUs steps, demographic information from screener data was used in defining covariates, whereas in the poststratification of the selected QDUs step, questionnaire demographic information was used.

Exhibit 4.1 lists all predictor variables proposed for QDU-level calibration and identifies them as counting, binary, or both. Various main effects and higher level factor effects based on
the predictor variables were included in the GEM modeling. As stated previously, all adjustment steps at the QDU level used a common set of proposed predictor variables.

### 4.2 Pair Weight Calibration

Like QDU, the initial set of weight components in pair weight calibration are the same as the set obtained from the SDU-level weight calibration. The SDU-calibrated weight is multiplied by the pair-level design weight, which in turn was adjusted in four steps: poststratification of selected pairs, nonresponse adjustment of respondent pairs, poststratification of respondent pairs, and the extreme weight adjustment of respondent pairs. All the adjustment steps for pair weights utilized the same set of initially proposed predictor variables, which included a subset of those used for the person-level nonresponse adjustment. This included segment characteristic variables, such as population density, percentage of persons in segment who are black or African American, percentage of persons in segment who are Hispanic or Latino, percentage of owneroccupied DUs in segment, and segment-combined median rent and housing value. Also included were pair-specific covariates, such as the demographic characteristics of pair age, pair race/ethnicity, and pair gender, as well as dwelling unit characteristics, such as race/ethnicity of householder, household type, household size, and group quarters indicators. State and quarter indicators were included as well. However, for two-factor effects, instead of individual State, State/region was used due to insufficient sample size. This resulted in a 12-level variable where the eight large sample States were kept separate, and the remainder of States were grouped according to the four census regions. All variables were defined as $0 / 1$ indicators. These proposed predictor variables and their levels are shown in Exhibit 4.2.

In the poststratification of selected pairs and the nonresponse adjustment of respondent pairs, screener data were used in the definition of the pair-specific variables such as pair age, pair race/ethnicity, and pair gender, whereas in the poststratification and extreme weight adjustment of respondent pairs, these variables were obtained from the questionnaire. For the latter case, in addition to the variables described above, indicator covariates corresponding to selected pair domains were included to perform Hajek-type ratio adjustments via weight calibration, as mentioned in Chapter 1. The selected pair domains were limited to 10 of the 14 pair domains listed in Chapter 1. (Parent-child pairs where the child was in the $15-$ to 17 -year-old age range and sibling-sibling-younger sibling focus pairs were not included in the poststratification.) The inclusion of these pair domain covariates led to the use of two sets of control totals in the modeling. Details of the construction of these control totals can be found in Appendix B.

Exhibit 4.1 Definitions of Levels for QDU-Level Calibration Modeling Variables

## Age ${ }^{\text {b }}$

1: 12-17, 2: 18-25, 3: 26-34, 4: 35-49, 5: 50+ ${ }^{1}$
Gender ${ }^{\text {b }}$
1: Male, 2: Female ${ }^{1}$
Group Quarter Indicator ${ }^{\text {a }}$
1: College Dorm, 2: Other Group Quarter, 3: Non-Group Quarter ${ }^{1}$
Hispanicity ${ }^{\text {b }}$
1: Hispanic or Latino, 2: Non-Hispanic or Latino ${ }^{1}$
Household Size ${ }^{\text {b }}$
Continuous Variable Count of Individuals Rostered with DU
Household Type (Ages of Persons Rostered within DU) ${ }^{\text {a }}$
1: 12-17, 18-25, 26+; 2: 12-17, 18-25; 3: 12-17, 26+; 4: 18-25, 26+; 5: 12-17, 6: 18-25; 7: $26+^{1}$
Percentage of Owner-Occupied Dwelling Units in Segment (\% Owner-Occupied) ${ }^{\text {a }}$ 1: 50-100\%, ${ }^{1}$ 2: $10-<50 \%, 3: 0-<10 \%$
Percentage of Segments That Are Black or African American ${ }^{\text {a }}$ 1: $50-100 \%, 2: 10-<50 \%, 3: 0-<10 \%{ }^{1}$
Percentage of Segments That Are Hispanic or Latino ${ }^{\text {a }}$
$1: 50-100 \%, 2: 10-<50 \%, 3: 0-<10 \%^{1}$
Population Density ${ }^{\text {a }}$
1: MSA $1,000,000$ or More, 2: MSA Less than $1,000,000,3$ : Non-MSA Urban, 4: Non-MSA Rural ${ }^{1}$
Quarter ${ }^{\text {a,b }}$
1: Quarter 1, 2: Quarter 2, 3: Quarter 3, 4: Quarter $4{ }^{1}$
Race (3 Levels) ${ }^{\text {b }}$
1: White, ${ }^{1}$ 2: Black or African American, 3: Other
Race ( 5 Levels) ${ }^{\text {b }}$
1: White, ${ }^{1}$ 2: Black or African American, 3: American Indian or Alaska Native, 4: Asian, 5: Two or More Races
Race/Ethnicity of Householder ${ }^{\text {a }}$
1: Hispanic or Latino White, ${ }^{1}$ 2: Hispanic or Latino Black or African American, 3: Hispanic or Latino Other, 4: Non-Hispanic or Latino White, 5: Non-Hispanic or Latino Black or African American, 6: Non-Hispanic or Latino Other
Relation to Householder ${ }^{\text {a }}$
1: Householder or Spouse, 2: Child, 3: Other Relative, 4: Nonrelative ${ }^{1}$
Segment-Combined Median Rent and Housing Value (Rent/Housing) ${ }^{\text {a,2 }}$
1: First Quintile, 2: Second Quintile, 3: Third Quintile, 4: Fourth Quintile, 5: Fifth Quintile ${ }^{1}$
States ${ }^{\text {a, }, \text {, }, 3}$
Model Group 1: 1: Connecticut, 2: Maine, 3: Massachusetts, ${ }^{1}$ 4: New Hampshire, 5: New Jersey, 6: New York, 7: Pennsylvania, 8: Rhode Island, 9: Vermont
Model Group 2: 1: Illinois, 2: Indiana, 3: Iowa, 4: Kansas, 5: Michigan, 6: Minnesota, 7: Missouri, 8: Nebraska, 9: North Dakota, 10: Ohio, 11: South Dakota, 12: Wisconsin ${ }^{1}$
Model Group 3: 1: Alabama, 2: Arkansas, 3: Delaware, 4: District of Columbia, 5: Florida, 6: Georgia, 7: Kentucky, 8: Louisiana, 9: Maryland, 10: Mississippi, 11: North Carolina, ${ }^{1}$ 12: Oklahoma, 13: South Carolina, 14: Tennessee, 15: Texas, 16: Virginia, 17: West Virginia
Model Group 4: 1: Alaska, 2: Arizona, ${ }^{1}$ 3: California, 4: Colorado, 5: Idaho, 6: Hawaii, 7: Montana, 8: Nevada, 9: New Mexico, 10: Oregon, 11: Utah, 12: Washington, 13: Wyoming
State/Region ${ }^{\text {a, }}{ }^{3}$
Model Group 1: 1: New York, 2: Pennsylvania, 3: Other ${ }^{1}$
Model Group 2: 1: Illinois, 2: Michigan, 3: Ohio, 4: Other ${ }^{1}$
Model Group 3: 1: Florida, 2: Texas, 3: Other ${ }^{1}$
Model Group 4: 1: California, 2: Other ${ }^{1}$
DU = dwelling unit; MSA = metropolitan statistical area; QDU = questionnaire dwelling unit.
${ }^{1}$ The reference level for this variable. This is the level against which effects of other factor levels are measured.
${ }^{2}$ Segment-Combined Median Rent and Housing Value is a composite measure based on rent, housing value, and percentage owner-occupied.
${ }^{3}$ The States or district assigned to a particular model is based on census regions.
${ }^{a}$ Binary variable.
${ }^{\mathrm{b}}$ Counting variable.

## Exhibit 4.2 Definitions of Levels for Pair-Level Calibration Modeling Variables

## Group Quarter Indicator

1: College Dorm, 2: Other Group Quarter, 3: Non-Group Quarter ${ }^{1}$

## Household Size

1: DU with 2 Persons, ${ }^{1}$ 2: DU with 3 Persons, 3 : DU with $\geq 4$ Persons
Pair Age ( 15 Levels)
1: 12-17 and 12-17, ${ }^{1} 2: 12-17$ and $18-25,3: 12-17$ and $26-34,4: 12-17$ and $35-49,5: 12-17$ and $50+, 6: 18-25$ and 18-25, 7: 18-25 and 26-34, 8: 18-25 and 35-49, 9: 18-25 and 50+, 10: 26-34 and 26-34, 11: 26-34 and 35-49,
12: 26-34 and 50+, 13: 35-49 and 35-49, 14: 35-49 and 50+, 15: 50+ and 50+
Pair Age (6 Levels)
1: 12-17 and 12-17, ${ }^{1} 2: 12-17$ and 18-25, 3: 12-17 and $26+, 4: 18-25$ and 18-25, 5: 18-25 and 26+, 6: 26+ and 26+
Pair Age (3 Levels)
1: 12-17 and 12-17, ${ }^{1}$ 2: 12-17 and 18+, 3: 18+ and 18+
Pair Gender
1: Male and Female, ${ }^{1}$ 2: Female and Female, 3: Male and Male
Pair Race/Ethnicity (10 Levels)
1: White and White, ${ }^{1}$ 2: White and Black or African American, 3: White and Hispanic or Latino, 4: White and Other, 5: Black or African American and Black or African American, 6: Black or African American and Hispanic or Latino, 7: Black or African American and Other, 8: Hispanic or Latino and Hispanic or Latino, 9: Hispanic or Latino and Other, 10: Other and Other
Pair Race/Ethnicity (5 Levels)
1: Two or More Races Pair, 2: Hispanic or Latino Pair, 3: Black or African-American Pair, 4: White Pair, ${ }^{1}$
5: Other Pair
Pair Race/Ethnicity (4 Levels)
1: Two or More Races Pair or Other and Other, 2: Hispanic or Latino Pair, 3: Black or African-American Pair,
4: White Pair ${ }^{1}$
Percentage of Owner-Occupied Dwelling Units in Segment (\% Owner-Occupied)
1: 50-100\%, ${ }^{1} 2: 10-<50 \%, 3: 0-<10 \%$
Percentage of Segments That Are Black or African American
1: $50-100 \%, 2: 10-<50 \%, 3: 0-<10 \%{ }^{1}$
Percentage of Segments That Are Hispanic or Latino
1: $50-100 \%, 2: 10-<50 \%, 3: 0-<10 \%{ }^{1}$
Segment-Combined Median Rent and Housing Value (Rent/Housing) ${ }^{2}$
1: First Quintile, 2: Second Quintile, 3: Third Quintile, 4: Fourth Quintile, 5: Fifth Quintile ${ }^{1}$
Population Density
1: MSA $1,000,000$ or More, 2: MSA Less than $1,000,000,3$ : Non-MSA Urban, 4: Non-MSA Rural ${ }^{1}$
Quarter
1: Quarter 1, 2: Quarter 2, 3: Quarter 3, 4: Quarter $4^{1}$
Race/Ethnicity of Householder
1: Hispanic or Latino White, ${ }^{1}$ 2: Hispanic or Latino Black or African American, 3: Hispanic or Latino Other,
4: Non-Hispanic or Latino White, 5: Non-Hispanic or Latino Black or African American, 6: Non-Hispanic or Latino Other

## Exhibit 4.2 Definitions of Levels for Pair-Level Calibration Modeling Variables (continued)

```
State/Region
    Model Group 1:1: Connecticut, Maine, Massachusetts, New Hampshire, New Jersey, Rhode Island, Vermont;
                                    2: Alabama, Arkansas, Delaware, District of Columbia, Georgia, Kentucky, Louisiana,
                                    Maryland, Mississippi, North Carolina, Oklahoma, South Carolina, Tennessee, Virginia, West
                                    Virginia;}\mp@subsup{}{}{1}3: New York; 4: Pennsylvania; 5: Florida; 6: Texas
    Model Group 2: 1: Indiana, Iowa, Kansas, Minnesota, Missouri, Nebraska, North Dakota, South Dakota,
                Wisconsin;}\mp@subsup{}{}{1}2: Alaska, Arizona, Colorado, Idaho, Hawaii, Montana, Nevada, New Mexico Oregon, Utah, Washington, Wyoming; 3: Michigan; 4: Illinois; 5: Ohio; 6: California
States \({ }^{3}\)
Model Group 1: 1: Alabama, 2: Arkansas, 3: Connecticut, 4: Delaware, 5: District of Columbia, 6: Florida, 7: Georgia, 8: Kentucky, 9: Louisiana, 10: Maine, 11: Maryland, \({ }^{1}\) 12: Massachusetts, 13: Mississippi, 14: New Hampshire, 15: New Jersey, 16: New York, 17: North Carolina, 18: Oklahoma, 19: Pennsylvania, 20: Rhode Island, 21: South Carolina, 22: Tennessee, 23: Texas, 24: Vermont, 25: Virginia, 26: West Virginia
Model Group 2: 1: Alaska, 2: Arizona, \({ }^{1}\) 3: California, 4: Colorado, 5: Idaho, 6: Illinois, 7: Indiana, 8: Iowa, 9: Hawaii, 10: Kansas, 11: Michigan, 12: Minnesota, 13: Missouri, 14: Montana, 15: Nebraska, 16: Nevada, 17: New Mexico, 18: North Dakota, 19: Ohio, 20: Oregon, 21: South Dakota, 22: Utah, 23: Washington, 24: Wisconsin, 25: Wyoming
Pair Relationship Associated with Multiplicity
1: Parent-Child (12-14)*
2: Parent-Child (12-17)*
3: Parent-Child (12-10)*
4: Parent*-Child (12-14)
5: Parent*-Child (12-17)
6: Parent*-Child (12-20)
7: Sibling (12-14)-Sibling (15-17)
8: Sibling (12-17)-Sibling (18-25)
9: Spouse-Spouse/Partner-Partner
10: Spouse-Spouse/Partner-Partner with Children (younger than 18)
```

DU = dwelling unit; MSA = metropolitan statistical area; QDU = questionnaire dwelling unit.
${ }^{1}$ The reference level for this variable. This is the level against which effects of other factor levels are measured.
${ }^{2}$ Segment-Combined Median Rent and Housing Value is a composite measure based on rent, housing value, and percentage owner-occupied.
${ }^{3}$ The States or district assigned to a particular model is based on combined census regions.

* The pair member focused on.


## 5. Definition of Extreme Weights

An important feature of the generalized exponential model (GEM) is the built-in provision of extreme value (ev) treatment. Sampling weights are often classified as extreme (high or low) if they fall outside the interval, median $\pm 3 \times$ interquartile range (IQR). The interval is set for prespecified domains defined usually by design variables corresponding to deep stratification. ${ }^{7}$ Similar to previous NSDUHs, for the GEM modeling used in the 2012 NSDUH, a more conservative (narrower) interval was defined, median $\pm 2.5 \times \mathrm{IQR}$. The narrower interval better prevents the adjusted weights from crossing the standard interval boundaries by treating weights near but not outside the commonly used boundaries (i.e., those that have the most potential to become extreme) as extreme as well.

Denote the interval boundaries (or critical values) for low and high extreme values by $b_{k(1)}$ and $b_{k(w)}$, respectively. For implementing extreme value control via GEM, the variable $m_{k}$ was defined as the minimum of $b_{k(u)} / w_{k}$ and one for high extreme weights, and the maximum of $b_{k(l)} / w_{k}$ and one for low extreme weights, where $w_{k}$ represents the sampling weight before adjustment, and $b_{k(u)}$ and $b_{k(l)}$ denote the critical values for the extreme weights. Note that under this definition, for high extreme weights, the more extreme the weight is, the smaller $m_{k}$ will be, and, conversely, for low extreme weights, the more extreme the weight is, the bigger $m_{k}$ will be. Nonextreme weights had a value of one for $m_{k}$. The upper and lower bounds for the adjustment factors were defined, respectively, as the product of $m_{k}$ and the upper and lower boundary parameters of GEM. GEM allows inputs of up to three different upper and lower boundary parameters ( $L_{1}$ and $U_{1}, L_{2}$ and $U_{2}, L_{3}$ and $U_{3}$ ) for high, non-, and low extreme weights. By applying a small upper boundary parameter for high extreme weights and a large lower boundary parameter for low extreme weights, the extreme weights can be controlled in the modeling process.

### 5.1 Questionnaire Dwelling Unit Extreme Weight Definition

For the questionnaire dwelling unit (QDU)-level weight adjustment, extreme weights were defined using a nested hierarchy of six domains:

1. State;
2. State sampling region;
3. State by household type;

Levels of household type indicate whether the household has members who are youths, young adults, or adults, where youth signifies 12 - to 17 -year-olds, young adult 18 - to 25 -year-olds, and adult 26 years or older.

[^4]a. Youth, Young Adult, Adult;
b. Youth, Young Adult;
c. Youth, Adult;
d. Young Adult, Adult;
e. Youth Only;
f. Young Adult Only; and
g. Adult Only.
4. State sampling region by household type;
5. State by household type by household size ( $1,2,3,4+$ ); and
6. State sampling region by household type by household size.

The hierarchy is used to satisfy the minimum of 30 observations for defining the boundaries for extreme values. If this sample size requirement is not met at the lower level, then the next level up in the hierarchy is used.

### 5.2 Person Pair Extreme Weight Definition

The pair selection probability is a function of the selection probability of each person in the pair given by formula (2.1) or (2.6), depending on the sum of the person selection probabilities within the household as discussed in Section 2.1. This probability can be very small if the selection probabilities of individual members are small. For example, consider a particular selected dwelling unit (DU) from the 2012 survey. This DU gave rise to a selected pair of respondents, both aged 50 or older. The selection probability in this DU was 0.10207 for a respondent aged 50 or older. Using the formula (2.6) in Chapter 2, the pair selection probability was computed to be 0.000227082 . Therefore, the inverse of the selection probability, the pairlevel design weight, was 4403.70 . Thus, a small pair selection probability can create a high initial weight, which is the product of the screener dwelling unit (SDU) weight and the person pair design-based weight.

As mentioned in the introduction, it turns out to be difficult to select suitable domains for defining extreme weights for pair-level data. However, as was done for the 1999-2012 surveys, the extreme weight definition was based on the following hierarchy of domains:

1. Pair age group ${ }^{8}$ (with three age categories, 12 to 25,26 to 49 , and $50+$ ) by number ( 0 , $1,2+$ ) of persons aged 12 to 25 in the household;
2. State cluster (with five levels [explained below]) by pair age group by number ( 0,1 , $2+$ ) of persons aged 12 to 25 in the household;
3. State cluster (with three levels [explained below]) by pair age group by number ( 0,1 , $2+$ ) of persons aged 12 to 25 in the household; and
4. State by pair age group by number of persons aged 12 to $25(0,1,2+)$ in the household.
[^5]The hierarchy was used to satisfy the minimum of 30 observations for defining the boundaries for extreme values. If this sample size requirement was not met at the lower level, then the next level up in the hierarchy was used.

We now briefly introduce the considerations behind the above definition for extreme weight domains. The sample design prespecified the person-level selection probability within State by five age groups ( 12 to 17,18 to 25,26 to 34,35 to $49,50+$ ). Age groups 12 to 17 and 18 to 25 have a relatively similar selection probability, and the same is true for age groups 26 to 34 and 35 to 49 . The $50+$ group, however, has a quite different selection probability from the other groups. Furthermore, since the 12 to 17 and 18 to 25 age groups have large selection probabilities, they have a very high chance of being selected if the household has persons in these age groups. Therefore, the number of persons aged 12 to 25 in the household has a significant impact on the type of pair selected and the pair selection probability. Taking into consideration these design-related features, a suitable domain to define the pair-level extreme weight seems to be given by State by pair age group by number of persons aged 12 to 25 in the household.

The hierarchy of domains mentioned above was used to satisfy the minimum of 30 observations. However, it was found that for many ev domains the minimum sample size requirement was not met. To alleviate this problem, States were grouped into a small number of clusters, such as three or five. The assignment of States to clusters was determined by the clustering algorithm in PROC CLUSTER in SAS, where the clustering variable was defined as the average person-level weight (ANALWT) for each of the five age groups within each State. The choice of the average person-level weight for each group for each State was motivated from the objective of finding a single variable that would reflect the design-based difference in pair selection probabilities across States. Even with clustering of States, the ev domain sample size was insufficient in some cases, so the most general level of the hierarchy, the national level, was required. Furthermore, at the national level, we had to collapse some pair age categories in forming domains of reasonable sample size to define extreme weights. More specifically, for the national level, we collapsed all levels of number of persons aged 12 to 25 for the pair age groups of $50+, 50+$ and 26 to $49,50+$. In addition, levels 1 and $2+$ of number of persons aged 12 to 25 were combined for the pair age group of 26 to 49,26 to 49 .

# 6. Weight Calibration at Questionnaire Dwelling Unit and Pair Levels 

The 2012 National Survey on Drug Use and Health (NSDUH) was based on probability sampling so that valid inferences can be made from survey findings about the target population. Probability sampling refers to sampling in which every unit on the frame is given a known, nonzero probability for inclusion in the survey. This is required for unbiased estimation of the population total. The assumption of nonzero inclusion probability for every pair of units in the frame also is required for unbiased variance estimation. The basic sampling plan involved four stages of selection across two phases of design: within Phase I, (1) the selection of census tracts within each State sampling (SS) region, (2) the selection of subareas or segments (comprised of U.S. Census Bureau blocks) within SS regions; (3) the selection of dwelling units (DUs) within these subareas; and, finally, within Phase II, (4) the selection of eligible individuals within DUs. Specific details of the sample design and selection procedures for the sample can be found in the 2012 NSDUH sample design report (Morton, Martin, Shook-Sa, Chromy, \& Hirsch, 2013).

As part of the postsurvey data-processing activities, analysis weights that reflected the selection probabilities from various stages of the sample design were calculated for respondents. These sample weights were adjusted at the DU (screening sample), questionnaire dwelling unit (QDU), person, and paired respondent levels (the latter three all based around the drug questionnaire sample) to account for bias due to extreme values (ev), nonresponse (nr), and coverage.

The final sample weights for Phase I screener dwelling units (SDU) and Phase II QDU, person, and pair levels for the 2012 samples consisted of products of several factors, each representing either a probability of selection at some particular stage or some form of ev , nr , or ps calibration adjustment. In the following sections, we describe the QDU and pair weight components in greater detail. In summary, the first 10 factors were defined for all SDUs and reflected the fully adjusted SDU sample weight. The remaining components branched to reflect QDU and pair selection probabilities, as well as additional adjustments for ev, nr , and ps. Note that the final QDU and pair weights for the 2012 survey sample are the product of all weight components for each type of sample, illustrated in Exhibits 6.1 and 6.2.

For QDU data, generalized exponential modeling (GEM) calibration modeling was applied by partitioning the data into four groups of States: Northeast, South, Midwest, and West, based on census regions in the interest of computational feasibility. Previous experience showed that with current computing power, the large number of variables and records prevented any further reduction of modeling groups.

For pair data, GEM modeling was initially applied by partitioning the pair data into four groups based on census regions. However, there were not enough observations in each group to fit a comprehensive model to reduce bias. Alternatively, a single model was attempted for the whole pair data, but it was rejected as not practical due to computational limitations. A compromise approach was adopted by combining census regions into two groups: Northeast with South and Midwest with West. This grouping proved both manageable and desirable as it
assisted in bias reduction, ease of modeling, and workload reduction. Exhibit 6.3 provides more details of the data partition for GEM modeling. The resulting sample sizes of selected and respondent units for the pair and QDU data partitions are shown for the 2008-2012 surveys in Table 6.1.

It may be noted that for the pair data in the 1999, 2000, and 2001 surveys, the built-in ev control feature of GEM was not used until the final respondent pair ev adjustment step. The reason for this is that the definition for ev domain was not finalized before the pair data calibration process was begun. However, for the 2002-2012 survey pair data, the built-in ev control feature was used for each adjustment step.

Exhibit 6.1 Summary of 2012 NSDUH QDU Sample Weight Components
Phase I Screener Dwelling Unit Level

| Design Weight Components |  |
| :--- | :--- |
| $\# 1$ | Inverse Probability of Selecting Census Tract |
| $\# 2$ | Inverse Probability of Selecting Segment |
| $\# 3$ | Quarter Segment Weight Adjustment |
| $\# 4$ | Subsegmentation Inflation Adjustment |
| $\# 5$ | Inverse Probability of Selecting SDU |
| $\# 6$ | Subsampling of Added SDU Adjustment |
| $\# 7$ | SDU Release Adjustment |
|  |  |
| Weight Adjustment* |  |
| $\# 8$ | SDU Nonresponse Adjustment (res.sdu.nr) |
| $\# 9$ | SDU Poststratification Adjustment (res.sdu.ps) |
| $\# 10$ | SDU Extreme Value Adjustment (res.sdu.ev) |
|  |  |

Phase II Questionnaire Dwelling Unit Level

| Design Weight Component |  |  |
| :--- | :---: | :---: |
| $\# 11$ Inverse of Selection Probability of at Least One Person in the Dwelling Unit |  |  |
| Weight Adjustment* |  |  |

\#12 Selected QDU Poststratification to SDU-Based Control Totals (sel.qdu.ps)
\#13 Respondent QDU Nonresponse Adjustment (res.qdu.nr)
\#14 Respondent QDU Poststratification to SDU-Based Control Totals (res.qdu.ps)
\#15 Respondent QDU Extreme Value Adjustment (res.qdu.ev)
QDU = questionnaire dwelling unit; $\operatorname{SDU}=$ screener dwelling unit.

* These adjustments use the generalized exponential model (GEM), which also involves pre- and postprocessing in addition to running the GEM macro. See Exhibit 4.1 (Chen et al., 2014). For computational feasibility, all weight adjustments were done using the four model groups based on census regions defined in Exhibit 6.3.

Exhibit 6.2 Summary of 2012 NSDUH Person Pair Sample Weight Components
Phase I Screener Dwelling Unit Level

|  | Design Weight Components |
| :--- | :--- |
| $\# 1$ | Inverse Probability of Selecting Census Tract |
| $\# 2$ | Inverse Probability of Selecting Segment |
| $\# 3$ | Quarter Segment Weight Adjustment |
| $\# 4$ | Subsegmentation Inflation Adjustment |
| $\# 5$ | Inverse Probability of Selecting SDU |
| $\# 6$ | Subsampling of Added SDU Adjustment |
| $\# 7$ | SDU Release Adjustment |
|  |  |
|  |  |
| $\# 8$ | SDU Nonresponse Adjustment (res.sdu.nr) |
| $\# 9$ | SDU Poststratification Adjustment (res.sdu.ps) |
| $\# 10$ | SDU Extreme Value Adjustment (res.sdu.ev) |

Phase II Person Pair Level

| Design Weight Component |  |
| :--- | :--- |
| $\# 11$ | Inverse of Selection Probability of a Person Pair in SDU |
|  |  |
| Weight Adjustment* |  |
| $\# 12$ | Selected Pair Poststratification to SDU-Based Control Totals (sel.pr.ps) |
| $\# 13$ | Respondent Pair Nonresponse Adjustment (res.pr.nr) |
| $\# 14$ | Respondent Pair Poststratification Adjustment to SDU-Based Control Totals <br> (res.per.ps) |

## \#15 Respondent Pair Extreme Value Adjustment (res.per.ev)

$\mathrm{QDU}=$ questionnaire dwelling unit; $\mathrm{SDU}=$ screener dwelling unit.

* These adjustments use the generalized exponential model (GEM), which also involves pre- and postprocessing in addition to running the GEM macro. See Exhibit 4.1 (Chen et al., 2014). For computational feasibility, all weight adjustments were done using the four model groups based on census regions defined in Exhibit 6.3.

Exhibit 6.3 U.S. Census Bureau Regions/Model Groups

| Model Group | Census Region |
| :---: | :---: |
| QDU |  |
| 1 | Northeast (9 States) |
|  | Connecticut, Maine, Massachusetts, New Hampshire, New Jersey, New York, Pennsylvania, Rhode Island, Vermont |
| 2 | Midwest (12 States) |
|  | Illinois, Indiana, Iowa, Kansas, Michigan, Minnesota, Missouri, Nebraska, North Dakota, Ohio, South Dakota, Wisconsin |
| 3 | South (16 States and the District of Columbia) |
|  | Alabama, Arkansas, Delaware, District of Columbia, Florida, Georgia, Kentucky, Louisiana, Maryland, Mississippi, North Carolina, Oklahoma, South Carolina, Tennessee, Texas, Virginia, West Virginia |
| 4 | West (13 States) |
|  | Alaska, Arizona, California, Colorado, Idaho, Hawaii, Montana, Nevada, New Mexico, Oregon, Utah, Washington, Wyoming |
| Pair |  |
| 1 | Northeast + South (25 States and the District of Columbia) |
|  | Alabama, Arkansas, Connecticut, Delaware, District of Columbia, Florida, Georgia, Kentucky, Louisiana, Maine, Massachusetts, Maryland, Mississippi, New Hampshire, New Jersey, New York, North Carolina, Oklahoma, Pennsylvania, Rhode Island, South Carolina, Tennessee, Texas, Vermont, Virginia, West Virginia |
| 2 | Midwest + West (25 States) |
|  | Alaska, Arizona, California, Colorado, Hawaii, Idaho, Illinois, Indiana, Iowa, Kansas, Michigan, Minnesota, Missouri, Montana, Nebraska, Nevada, New Mexico, North Dakota, Ohio, Oregon, South Dakota, Utah, Washington, Wisconsin, Wyoming |

Table 6.1 Sample Size, by Model Group at QDU and Pair Levels

| Model Group | 2008 |  | 2009 |  | 2010 |  | 2011 |  | 2012 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Selected QDUs | Completed QDUs | Selected QDUs | Completed QDUs | Selected QDUs | Completed QDUs | Selected QDUs | Completed QDUs | Selected QDUs | Completed QDUs |
| QDU |  |  |  |  |  |  |  |  |  |  |
| Northeast | 11,490 | 9,216 | 11,605 | 9,340 | 11,627 | 9,339 | 11,997 | 9,456 | 12,616 | 9,917 |
| South | 17,703 | 14,747 | 17,756 | 14,909 | 17,880 | 14,857 | 19,690 | 16,487 | 18,345 | 15,019 |
| Midwest | 16,763 | 13,650 | 16,382 | 13,491 | 16,670 | 13,686 | 17,045 | 13,752 | 16,984 | 13,687 |
| West | 12,986 | 10,567 | 12,545 | 10,348 | 12,525 | 10,231 | 12,709 | 10,438 | 12,676 | 10,227 |
| Total | 58,942 | 48,180 | 58,288 | 48,088 | 58,702 | 48,113 | 61,441 | 50,133 | 60,621 | 48,850 |
|  | 2008 |  | $2009$ |  | 2010 |  | 2011 |  | 2012 |  |
| Model Group | Selected Pairs | Completed Pairs | Selected Pairs | $\begin{aligned} & \text { Completed } \\ & \text { Pairs } \end{aligned}$ | Selected Pairs | Completed Pairs | Selected Pairs | $\begin{gathered} \text { Completed } \\ \text { Pairs } \end{gathered}$ | Selected Pairs | Completed Pairs |
| Pair |  |  |  |  |  |  |  |  |  |  |
| Northeast + South | 13,060 | 9,700 | 13,058 | 9,806 | 12,872 | 9,590 | 13,686 | 10,127 | 13,619 | 9,723 |
| Midwest + West | 13,709 | 10,048 | 13,439 | 10,113 | 13,423 | 10,101 | 13,409 | 9,849 | 13,416 | 9,736 |
| Total | 26,769 | 19,748 | 26,497 | 19,919 | 26,295 | 19,691 | 27,095 | 19,976 | 27,035 | 19,459 |

$\mathrm{QDU}=$ questionnaire dwelling unit.
Note: The 2008-2010 sample sizes reflect the removal of falsified cases found in Pennsylvania and Maryland. The 2011 and 2012 samples were not affected. For additional information, see Section B.3.5 in Appendix B of the 2011 and 2012 national findings reports (Center for Behavioral Health Statistics and Quality, 2012; 2013).

### 6.1 Phase I SDU-Level Weight Components

A total of 10 weight components for the SDU level correspond to selection probabilities and nonresponse, poststratification, and extreme value adjustment factors. Note that this differs from previous NHSDAs and NSDUHs in that a new design-based component was incorporated at the beginning of the process so that corresponding weight component numbers are incremented by one when compared to previous survey years with an otherwise similar weighting scheme. The first seven components in the Phase I sample weights reflect the probability of selecting the DUs. These components were derived from (1) the probability of selecting the census tract within each State sampling (SS) region, (2) the probability of selecting the geographic segment within each SS region, (3) a quarter segment weight adjustment, (4) a subsegmentation inflation factor, (5) the probability of selecting a DU from within each counted and listed sampled segment, (6) the probability of inclusion of added DUs, and (7) DU percent release adjustment. The three remaining weight components, \#8 through \#10, are GEM calibration adjustments accounting for (8) DU nonresponse at the screening level, (9) DU poststratification to census controls, and (10) DU-level extreme value adjustment, although in 2009 extreme value adjustment at this stage was deemed unnecessary, and thus Weight Component \#10 was set to one for all respondent DUs. The person-level, QDU-level, and person pair-level weights use the product of the above 10 weight components as the common initial weight before further adjustments. For more detailed information on Weight Components \#1, \#2, and \#4 through \#7, refer to the 2012 NSDUH sample design report (Morton et al., 2013), and for more detail on Weight Components \#3 and \#8 through \#10, see the 2012 person-level sampling weight calibration report (Chen et al., 2014).

Note that from 2008 to 2012, there was an occasional second subsegmentation step when the initial partitioning of segments was insufficient due to out-of-date census counts or the segment was still too large to list after the original subsegmentation. This second partitioning was not accounted for in the weighting over these survey years. A comparison was done to evaluate the effect of this omission, and it was determined that the missing second subsegmenting factor in the analysis weight had minimal impact on estimates. Therefore, weights for these years were not reproduced. Additional detail can be found in the 2012 NSDUH sample design report (Morton et al., 2013).

### 6.2 QDU Weight Components

### 6.2.1 QDU Weight Component \#11: Inverse of Selection Probability of at Least One Person in the Dwelling Unit

The selection of a questionnaire dwelling unit from all completed SDUs is based on the outcome of a variant of Brewer's method, which may select zero, one, or two persons. Any pair of survey eligible residents within the dwelling unit had some known, nonzero chance of being selected for the survey. The value for Weight Component \#11 is equal to the inverse of the probability that at least one person in the dwelling unit is selected (see Section 2.2 for details).

### 6.2.2 QDU Weight Component \#12: Selected QDU Poststratification to SDU-Based Control Totals

This poststratification factor adjusts the weights for selected QDUs to the SDU-based control totals. The SDU-based control totals are obtained by using the calibrated SDU weights. This adjustment step provides more stable controls for the subsequent nonresponse adjustment (Weight Component \#13). Exhibit 4.1 lists the initially proposed variables for GEM modeling. The predictor variables are either $0 / 1$ indicators or counting variables representing the number of persons who fall into a given demographic domain. The counting variables are derived from the screener demographic information. It may be noted that during screening, the only required demographic information was the age of each person rostered. Thus, other demographic information necessary for weight calibration, such as race/ethnicity and gender may be missing for certain rostered eligible persons, and so imputation was done to replace this missing data. For more details on the imputation of screener demographic information, see Chen et al. (2014).

The details on the predictor variables retained in the model and model summary statistics can be found in Appendix C.

### 6.2.3 QDU Weight Component \#13: Respondent QDU Nonresponse Adjustment

This nonresponse adjustment step accounts for the failure to obtain respondent person(s) from each and every selected QDU. The same set of initially proposed predictor variables were used as for the previous adjustment (\#12).

See Appendix C for more details on the predictor variables retained in the model and model summary statistics.

### 6.2.4 QDU Weight Component \#14: Respondent QDU Poststratification to SDU-Based Control Totals

This final poststratification for all respondent QDUs utilized the same set of initially proposed predictor variables as previous adjustments. The corresponding control totals were obtained from the SDU-level sample, as was done for Weight Component \#12.

See Appendix C for more details on the predictor variables retained in the model and model summary statistics.

### 6.2.5 QDU Weight Component \#15: Respondent QDU Extreme Value Adjustment

The extreme weight proportions for the final poststratified weights were acceptably low, eliminating the need for the extreme value adjustment. Weight Component \#15 was set to one for each responding QDU.

### 6.3 Pair-Level Weight Components

Exhibit 4.2 lists the initially proposed predictor variables for the following adjustment steps via GEM.

### 6.3.1 Pair Weight Component \#11: Inverse of Selection Probability of a Person Pair in the Dwelling Unit

Selection of pairs of individuals from all eligible persons residing within the dwelling unit is based on the outcome of a variant of Brewer's method, which may select zero, one, or two persons. Any pair of survey eligible residents within the dwelling unit has some known, nonzero chance of being selected for the survey. When two persons are selected, a pair is formed. The pair selection probability is determined by either formula (2.1) or formula (2.6) in Chapter 2. This weight component is the inverse of the selection probability discussed above.

### 6.3.2 Pair Weight Component \#12: Selected Pair Poststratification to SDU-Based Control Totals

Similar to QDU Weight Component \#12, this step was motivated by the consideration that the larger sample of all possible pairs provides more stable control totals for the respondent pair nonresponse adjustment. The weights of selected pairs were poststratified to the control totals that derived from calibrated SDU weights of all possible pairs. The pair-level demographic variables for all selected pairs, such as pair age group, pair race/ethnicity, etc., were derived from screener demographic information.

The details on the predictor variables retained in the model and model summary statistics can be found in Appendix H.

### 6.3.3 Pair Weight Component \#13: Respondent Pair Nonresponse Adjustment

If both persons in the selected pair completed interviews successfully, the pair then was considered a respondent pair. This adjustment step accounts for failure to obtain respondent pairs
from all selected pairs. In this step, respondent pair weights were adjusted to the control totals based on the full sample of selected pairs. Due to the low response rate of person pairs, this step had a relatively large adjustment on the weights. The same set of proposed predictor variables was used as for Weight Component \#12. Similar to Weight Component \#12, the pair-level demographic variables for all selected pairs, such as pair age group, pair race/ethnicity, etc., were derived from screener demographic information.

See Appendix H for more details on the predictor variables retained in the model and model summary statistics.

### 6.3.4 Pair Weight Component \#14: Respondent Pair Poststratification to SDU-Based Control Totals

This final poststratification utilized the same set of initially proposed predictor variables as previous adjustment steps. In addition, 10 pair relationship domain-level indicator variables were added to the set of covariates. The control totals for GEM calibration were derived from the SDU sample of all possible pairs of eligible persons, as was done for Weight Component \#12. The calibration control totals for these 10 domains used household-level person counts and the final QDU weights. As mentioned in the introduction, use of these household-level count totals for pair relationship domains in GEM calibration provided Hajek-type weight adjustment in the interest of obtaining more stable estimates. In setting up calibration covariates, multiplicity factors were needed. These factors, as discussed in the introduction, are used in constructing estimates for person-level parameters based on pair-related drug behavior. The factors depend on the pair domains of interest. For a selected set of pair domains, multiplicity factors are provided along with the pair-level analysis weights. See Chapter 11 in the editing and imputation report (Carpenter et al., 2014) for more detail on the creation of and imputation of missing values in the pair relationship, multiplicity, and household-level person counts. See Chapter 4 for more detail on the use of multiplicities and household-level person counts in poststratification.

Unlike Weight Components \#12 and \#13, demographic covariates were based on data from the questionnaire instead of information pulled from the dwelling unit screener.

For more details on the predictor variables retained in the GEM model and model summary statistics, see Appendix H.

### 6.3.5 Pair Weight Component \#15: Respondent Pair Extreme Weight Adjustment

We checked the extreme weight proportions for the weights up to Weight Component \#14, using the extreme weight domains (see Section 5.2). Even though the previous adjustment steps utilized the built-in extreme weight control feature of GEM, the extreme weight proportions were still high enough to cause concern that they might produce unreliable estimates. Therefore, the extreme weight adjustment via GEM was implemented, using the same final set of predictor variables kept in the model for Weight Component \#15. This step was successful in reducing the extreme weight proportion in all model groups. For details, see Appendix J.

## 7. Evaluation of Calibration Weights

During the weight calibration process, several criteria for quality control were implemented to assess model adequacy. In this chapter, we describe the individual procedures and a summary of their results. All tables referred to in this chapter can be found in Appendices D through G and I through L.

### 7.1 Response Rates

Table D. 1 in Appendix D displays the final selected and responding questionnaire dwelling unit (QDU) sample sizes from the 2012 National Survey on Drug Use and Health (NSDUH) for various national domains. This table also shows the weighted response rates. Most domains reflect the overall 75.87 percent response rate, with most rates relatively close to 80 percent, although the highest response rate is 93.20 percent, from the Group level of the Group Quarters category. The lowest response rate came from Race/Ethnicity of Householder NonHispanic or Latino Other, with 72.15 percent.

Table I. 1 in Appendix I displays the final selected and responding pair-level sample sizes from the 2012 survey for various national domains. Due to the nature of the pair data, the response rates were lower in all domains examined than at the QDU level, with an overall response rate of 63.76 percent. The response rates range from a low of 37.40 percent in the pair race/ethnicity Black or African American and Other category to a high of 81.51 percent from the Pair Age Group 12-17,12-17. This extreme range of response rates is probably due to a combination of small sample sizes and response burden as a result of selection of pairs within households among various domains. Like at the QDU level, the top response rates are among the younger respondents (as measured by household type for the QDU data and pair age for the pair data). This pattern may be related to the relatively high response rates in the group level of the variable group quarters since it includes college dormitories.

### 7.2 Proportions of Extreme Values and Outwinsors

During the stages of modeling adjustments (i.e., nonresponse [nr] and poststratification [ps]), one major issue of concern when deciding the adequacy of a particular model was the extent of the resulting proportions of extreme value (ev) and outwinsor weights (see Sections 5.1 and 5.2 for these definitions). For each weight adjustment step, these proportions are computed before and after the step for various domains. Prior to adjustment, the product of all weight components is used to compute proportions of evs and outwinsors, while after the adjustment the product includes the new adjustment factor. If the proportion of evs and outwinsors are deemed high, a separate ev treatment step after ps could be performed. This was done for the pair-level weights. Details of this step are explained in Section 6.3.5. A separate ev treatment step was deemed unnecessary for the QDU-level weights.

Tables E. 1 and E. 2 and Tables J. 1 through J. 3 present percentages of evs at the QDU level and the pair level, respectively, for various domains. Unweighted percentages are the percentage of actual counts of units defined as evs relative to the total sample size. Weighted percentages reflect the percentage of total ev weights relative to the total sample weight, while
outwinsor percentages represent the total amount of residual weight when the weights are trimmed to the critical values (used for ev definition) relative to the total sample weight. For evaluation purposes, the outwinsor percentage is considered the most important of the three percentages, as this gave a measure of the impact of winsorization (or trimming) of ev weights (if we performed this treatment). See Sections 5.1 and 5.2 for the domains that were used to define extreme values.

### 7.3 Slippage Rates

The slippage rate for a given domain is defined as the relative percentage difference between the sampling weights and the external control totals, both before and after ps. The control totals for QDU and person pair ps are derived from the screener dwelling unit (SDU) weights, which were poststratified to U.S. Census Bureau population estimates (Chen et al., 2014). Table F. 1 displays QDU national domain-specific weight sums for both before and after ps , as well as the desired totals to be met through ps. Table K. 1 shows the same for the pair sample. These tables also show the relative percentage difference, or the amount of adjustment necessary (positive or negative) to meet the desired totals. The first relative difference is used explicitly during the ps modeling procedure to identify potential problems for convergence. Large differences in domains with relatively small sample sizes are indicative of potential large adjustment factors, which may cause problems in convergence while satisfying bound constraints. The reason is that adjustments required for one domain may have an adverse effect on another domain when a unit belongs to both.

As an example, consider that Table F.1, for the 2012 QDU domain household size of two, indicates a sample size of 16,974 with a total design-based weight of $54,585,381$ and a census total of $54,556,667$ with an initial slippage rate of -0.07 percent, which would imply a common weight adjustment approximately equal to 1.000681 , if this were the only calibration control. Similarly, looking at pair data in Table K.1, the pair domain category of pair age 18-25, 18-25 has a sample size of 3,901 , a design-based weight of $12,270,700$, and a census total of $12,545,793$, showing an initial slippage of -2.69 percent. The resultant required adjustment would be approximately equal to 1.027695 , if this were the only control. However, in the generalized exponential model (GEM), all controls are simultaneously satisfied under a complex algorithm that allows for different adjustment factors for different units.

### 7.4 Weight Adjustment Summary Statistics

Tables G.1, G.2, and L. 1 through L. 3 display summary statistics on the product of weight components before and after all stages of adjustment for the QDU and person pair, respectively. The summary statistics include sample size ( $n$ ), minimum (min), maximum (max), median (med), $25^{\text {th }}$ percentile (Q1), $75^{\text {th }}$ percentile (Q3), and the unequal weighting effect (UWE). Note that in Tables L. 2 and L. 3 the sample size for pair age group, pair race/ethnicity, and pair gender are slightly different. This is because those variables were defined using screening demographic information in the nonresponse adjustment of respondent pairs, while in the poststratification of respondent pairs, they were defined from questionnaire demographic information. Because UWE is directly affected by weight adjustment factors and extreme weights, these values-along with the percentage of extreme weights as noted in Section 7.2-were used as guidelines for determining model adequacy.

### 7.5 Sensitivity Analysis of Drug Use Estimates

It is known that, in general, there is a trade-off between bias reduction and variance reduction. For instance, with GEM (for nr or ps ), enlarging a simple model (such as the one with only main effects) has the potential of further reducing the bias. At the same time, this enlargement also may be associated with a corresponding increase in the variance of the estimate due to additional variability caused by estimating the model parameters. To check for possible overfitting of the GEM model, we conducted a sensitivity analysis for respondent QDU poststratification for the QDU weights, respondent pair poststratification, and extreme weight adjustment for the person pair weights. A simple baseline model was fitted with the same bounds and maximum number of iterations as was used for the chosen (more complex) final model. We then looked for substantial changes in point estimates and standard errors (SEs). For the QDU weights, some household-level characteristics were selected such as family income, number of youths in the household, whether the household had health insurance coverage, and number of elders living in the household. The estimates and SEs are displayed in Table 7.1. For the person pair weights, selected licit and illicit drug use prevalence rates of 12- to 17-year-olds were calculated from parent-child pairs, and estimates and SEs of the estimates based on pair weights are shown in Tables 7.2a to 7.7b.

As seen in Table 7.1, the estimates and their SEs for the two models (baseline and the final) are generally similar to each other for the QDU weights. However, among the person pair estimates and SEs, there are some differences, but they do not seem significant in general.

Since the sensitivity analyses for both QDU- and pair-level calibrated weights seem to indicate that adding more covariates does not introduce an undesirable degree of instability in the estimates or their SEs, the final, more complex GEM models were deemed reasonable.

Table 7.1 Estimates of Totals and SEs for Domains of Interest Based on QDU Sample: 2012

| Domain | 2012 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $n$ | Baseline (B) ${ }^{1}$ | Final (F) ${ }^{2}$ | (B-F)/F\% (Estimate) | (B-F)/F\% (SE) |
| Households with Family Income |  |  |  |  |  |
| \$0-< \$10,000 | 5,078 | 10,460,715 (329,056) | 10,444,004 (329,693) | 0.16 | -0.19 |
| \$10,000 - < \$20,000 | 6,555 | 15,622,561 (398,493) | 15,638,194 (400,681) | -0.10 | -0.55 |
| \$20,000 - < \$30,000 | 5,748 | 14,195,740 (359,770) | 14,214,304 (360,747) | -0.13 | -0.27 |
| \$30,000 - < \$40,000 | 5,331 | 13,379,640 (327,509) | 13,364,505 (327,477) | 0.11 | 0.01 |
| \$40,000 - < \$50,000 | 5,080 | 12,604,296 (332,618) | 12,606,932 (331,778) | -0.02 | 0.25 |
| \$50,000 - < \$75,000 | 7,703 | 19,169,185 (415,719) | 19,159,643 (415,892) | 0.05 | -0.04 |
| \$75,000 - < \$100,000 | 5,265 | 13,710,548 (383,622) | 13,716,657 (385,188) | -0.04 | -0.41 |
| \$100,000+ | 8,090 | 20,005,297 (507,343) | 20,003,744 (507,248) | 0.01 | 0.02 |
| Households with Number of Youths (<18) |  |  |  |  |  |
| 0 | 20,419 | 77,402,125 (1,069,348) | 77,400,559 (1,071,586) | 0.00 | -0.21 |
| 1 | 11,659 | 17,372,185 (274,586) | 17,379,411 (276,588) | -0.04 | -0.72 |
| 2 | 9,719 | 15,069,648 (275,565) | 15,063,588 (274,382) | 0.04 | 0.43 |
| 3 | 4,603 | 6,422,304 (162,360) | 6,404,101 (161,754) | 0.28 | 0.37 |
| 4+ | 2,450 | 2,881,721 (96,317) | 2,900,323 (99,755) | -0.64 | -3.45 |
| Households with Insurance Coverage |  |  |  |  |  |
| Yes | 39,384 | 99,535,440 (1,147,242) | 99,551,648 (1,148,703) | -0.02 | -0.13 |
| No | 9,466 | 19,612,543 (381,660) | 19,596,335 (381,456) | 0.08 | 0.05 |
| Households with Number of Older Adults (65+) |  |  |  |  |  |
| 0 | 43,406 | 89,339,074 (983,199) | 89,341,651 (986,485) | -0.00 | -0.33 |
| 1 | 3,721 | 19,513,672 (525,590) | 19,511,535 (525,200) | 0.01 | 0.07 |
| 2 | 1,676 | 10,100,864 (375,736) | 10,094,422 (376,016) | 0.06 | -0.07 |
| 3+ | 47 | 194,373 (42,011) | 200,375 (44,298) | -3.00 | -5.16 |

[^6]Table 7.2a Percentages of Youths (12 to 17) Reporting Lifetime, Past Year, and Past Month Use of Alcohol and Tobacco among Mother-Child (12 to 17) Pairs, by Mother Use: 2012


Note: Standard errors of prevalence estimates are provided in parentheses.
${ }^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last two steps of calibration, res.pr.ps and res.pr.ev, and a full model for preceding steps.
${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

Table 7.2b Percentages of Youths (12 to 17) Reporting Lifetime, Past Year, and Past Month Use of Alcohol and Tobacco among Father-Child (12 to 17) Pairs, by Father Use: 2012

| Drug | Father User | 2012 |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | $n$ | Baseline ${ }^{1}$ | Final ${ }^{2}$ |
| Alcohol |  |  |  |  |
| Lifetime | Yes | 1,385 | 31.37 (2.19) | 31.58 (2.28) |
|  | No | 99 | 15.24 (4.48) | 15.06 (4.48) |
|  | Overall | 1,484 | 30.19 (2.08) | 30.30 (2.16) |
| Past Year | Yes | 1,178 | 26.45 (2.22) | 26.59 (2.33) |
|  | No | 306 | 11.82 (2.29) | 11.44 (2.29) |
|  | Overall | 1,484 | 23.51 (1.85) | 23.44 (1.92) |
| Past Month | Yes | 971 | 15.55 (2.21) | 15.62 (2.37) |
|  | No | 513 | 6.45 (1.28) | 6.24 (1.25) |
|  | Overall | 1,484 | 12.38 (1.53) | 12.32 (1.62) |
| Cigarettes |  |  |  |  |
| Lifetime | Yes | 1,059 | 20.53 (2.20) | 20.29 (2.27) |
|  | No | 425 | 10.73 (2.76) | 10.87 (3.02) |
|  | Overall | 1,484 | 17.11 (1.75) | 16.96 (1.83) |
| Past Year | Yes | 428 | 16.48 (3.02) | 16.40 (3.14) |
|  | No | 1,056 | 10.97 (1.88) | 10.95 (1.93) |
|  | Overall | 1,484 | 12.28 (1.60) | 12.24 (1.64) |
| Past Month | Yes | 377 | 13.01 (3.27) | 13.55 (3.47) |
|  | No | 1,107 | 3.51 (0.89) | 3.43 (0.89) |
|  | Overall | 1,484 | 5.41 (1.00) | 5.43 (1.03) |

Note: Standard errors of prevalence estimates are provided in parentheses.
${ }^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last two steps of calibration, res.pr.ps and res.pr.ev, and a full model for preceding steps.
${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

Table 7.3a Percentages of Youths (12 to 17) Reporting Lifetime, Past Year, and Past Month Use of Any Illicit Drug or Marijuana among Mother-Child (12 to 17) Pairs, by Mother Use: 2012


Note: Standard errors of prevalence estimates are provided in parentheses.
${ }^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last two steps of calibration, res.pr.ps and res.pr.ev, and a full model for preceding steps.
${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

Table 7.3b Percentages of Youths (12 to 17) Reporting Lifetime, Past Year, and Past Month Use of Any Illicit Drug or Marijuana among Father-Child (12 to 17) Pairs, by Father Use: 2012


Note: Standard errors of prevalence estimates are provided in parentheses.
${ }^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last two steps of calibration, res.pr.ps and res.pr.ev, and a full model for preceding steps.
${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

Table 7.4 Percentages of Youths (12 to 17) Living with a Parent Reporting Lifetime, Past Year, and Past Month Use of Alcohol and Tobacco among Parent-Child (12 to 17) Pairs, Asked Whether Their Parents Had Spoken to Them about the Dangers of Tobacco, Alcohol, or Drug Use within the Past 12 Months: 2012


[^7]Table 7.5 Percentages of Youths (12 to 17) Living with a Parent Reporting Lifetime, Past Year, and Past Month Use of Any Illicit Drug and Marijuana among Parent-Child (12 to 17) Pairs, Asked Whether Their Parents Had Spoken to Them about the Dangers of Tobacco, Alcohol, or Drug Use within the Past 12 Months: 2012

| Drug | Parent Talked about <br> Dangers with Child | 2012 |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | $n$ | Baseline ${ }^{1}$ | Final ${ }^{2}$ |
| Any Illicit |  |  |  |  |
| Lifetime | Yes | 2,227 | 20.16 (1.35) | 20.25 (1.41) |
|  | No | 1,468 | 24.77 (2.20) | 24.87 (2.21) |
|  | Overall | 3,695 | 22.06 (1.23) | 22.15 (1.25) |
| Past Year | Yes | 2,227 | 15.29 (1.23) | 15.28 (1.28) |
|  | No | 1,468 | 17.02 (1.93) | 17.07 (1.92) |
|  | Overall | 3,695 | 16.00 (1.09) | 16.01 (1.11) |
| Past Month | Yes | 2,227 | 8.46 (1.00) | 8.51 (1.06) |
|  | No | 1,468 | 8.64 (1.37) | 8.65 (1.39) |
|  | Overall | 3,695 | 8.53 (0.83) | 8.56 (0.87) |
| Marijuana |  |  |  |  |
| Lifetime | Yes | 2,227 | 14.16 (1.17) | 14.23 (1.22) |
|  | No | 1,468 | 15.90 (1.81) | 15.83 (1.79) |
|  | Overall | 3,695 | 14.88 (1.01) | 14.89 (1.03) |
| Past Year | Yes | 2,227 | 11.56 (1.11) | 11.56 (1.17) |
|  | No | 1,468 | 12.65 (1.73) | 12.60 (1.71) |
|  | Overall | 3,695 | 12.01 (0.96) | 11.99 (0.98) |
| Past Month | Yes | 2,227 | 6.42 (0.90) | 6.49 (0.97) |
|  | No | 1,468 | 6.38 (1.09) | 6.38 (1.10) |
|  | Overall | 3,695 | 6.40 (0.70) | 6.45 (0.73) |

[^8]Percentages of Youths (12 to 17) Reporting Lifetime, Past Year, and Past Month Use of Alcohol and Tobacco among Mother-Child (12 to 17) Pairs, for Mother in the Pair, Asked Whether She Had Spoken to Her Children about the
Dangers of Tobacco, Alcohol, or Drug Use within the Past 12 Months: 2012

|  |  | Mother Talked about |  | 2012 |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Drug | Dangers with Child | $n$ | Baseline ${ }^{1}$ | Final ${ }^{2}$ |
|  | Alcohol |  |  |  |  |
|  | Lifetime | 0 times | 171 | 23.88 (5.10) | 24.15 (5.25) |
|  |  | 1-2 times | 310 | 24.65 (3.85) | 24.16 (3.91) |
|  |  | A few times | 546 | 25.42 (3.53) | 24.67 (3.50) |
|  |  | Many times | 1,147 | 33.49 (2.25) | 33.44 (2.30) |
|  |  | Overall | 2,174 | 29.36 (1.65) | 29.14 (1.68) |
|  | Past Year | 0 times | 171 | 15.35 (3.74) | 15.15 (3.78) |
|  |  | 1-2 times | 310 | 18.83 (3.49) | 18.70 (3.60) |
|  |  | A few times | 546 | 22.43 (3.49) | 21.91 (3.45) |
|  |  | Many times | 1,147 | 26.16 (2.00) | 26.07 (2.06) |
|  |  | Overall | 2,174 | 23.17 (1.53) | 22.99 (1.56) |
|  | Past Month | 0 times | 171 | 7.53 (2.63) | 7.77 (2.81) |
| $\pm$ |  | 1-2 times | 310 | 8.17 (2.15) | 7.91 (2.07) |
|  |  | A few times | 546 | 10.60 (2.76) | 10.15 (2.67) |
|  |  | Many times | 1,147 | 13.75 (1.54) | 13.95 (1.65) |
|  |  | Overall | 2,174 | 11.58 (1.10) | 11.58 (1.13) |
|  | Cigarettes |  |  |  |  |
|  | Lifetime | 0 times | 171 | 11.53 (4.89) | 10.97 (4.87) |
|  |  | 1-2 times | 310 | 11.01 (2.48) | 10.62 (2.38) |
|  |  | A few times | 546 | 11.05 (2.55) | 11.14 (2.59) |
|  |  | Many times | 1,147 | 16.32 (1.51) | 16.20 (1.52) |
|  |  | Overall | 2,174 | 13.83 (1.16) | 13.70 (1.17) |
|  | Past Year | 0 times | 171 | 5.50 (2.14) | 4.86 (2.00) |
|  |  | 1-2 times | 310 | 7.70 (2.25) | 7.27 (2.11) |
|  |  | A few times | 546 | 9.09 (2.50) | 9.19 (2.53) |
|  |  | Many times | 1,147 | 11.52 (1.23) | 11.44 (1.24) |
|  |  | Overall | 2,174 | 9.81 (0.97) | 9.68 (0.97) |
|  | Past Month | 0 times | 171 | 4.12 (1.99) | 3.62 (1.87) |
|  |  | 1-2 times | 310 | 5.36 (1.96) | 5.10 (1.85) |
|  |  | A few times | 546 | 5.14 (2.06) | 5.00 (2.01) |
|  |  | Many times | 1,147 | 6.12 (0.94) | 6.06 (0.94) |
|  |  | Overall | 2,174 | 5.59 (0.78) | 5.45 (0.76) |

[^9]Table 7.6b Percentages of Youths (12 to 17) Reporting Lifetime, Past Year, and Past Month Use of Alcohol and Tobacco among Father-Child (12 to 17) Pairs, for Father in the Pair, Asked Whether He Had Spoken to His Child about the Dangers of Tobacco, Alcohol, or Drug Use within the Past 12 Months: 2012

| Drug | Father Talked about Dangers with Child | 2012 |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | $n$ | Baseline ${ }^{1}$ | Final ${ }^{2}$ |
| Alcohol |  |  |  |  |
| Lifetime | 0 times | 174 | 20.01 (5.17) | 21.59 (5.66) |
|  | 1-2 times | 247 | 23.73 (4.69) | 23.00 (4.75) |
|  | A few times | 459 | 34.16 (4.00) | 34.70 (4.24) |
|  | Many times | 497 | 30.66 (3.24) | 30.67 (3.30) |
|  | Overall | 1,377 | 29.29 (2.11) | 29.53 (2.19) |
| Past Year | 0 times | 174 | 13.01 (4.46) | 13.61 (4.84) |
|  | 1-2 times | 247 | 21.30 (4.61) | 20.68 (4.67) |
|  | A few times | 459 | 24.11 (3.49) | 24.44 (3.79) |
|  | Many times | 497 | 24.52 (2.84) | 24.34 (2.85) |
|  | Overall | 1,377 | 22.43 (1.83) | 22.45 (1.91) |
| Past Month | 0 times | 174 | 3.04 (1.48) | 2.91 (1.45) |
|  | 1-2 times | 247 | 10.85 (2.56) | 10.10 (2.45) |
|  | A few times | 459 | 12.73 (3.03) | 13.32 (3.39) |
|  | Many times | 497 | 12.55 (2.37) | 12.33 (2.43) |
|  | Overall | 1,377 | 11.18 (1.42) | 11.16 (1.52) |
| Cigarettes |  |  |  |  |
| Lifetime | 0 times | 174 | 13.46 (4.95) | 14.89 (5.44) |
|  | 1-2 times | 247 | 14.25 (4.31) | 13.89 (4.37) |
|  | A few times | 459 | 16.12 (3.12) | 16.25 (3.42) |
|  | Many times | 497 | 18.51 (2.61) | 17.85 (2.61) |
|  | Overall | 1,377 | 16.31 (1.70) | 16.23 (1.79) |
| Past Year | 0 times | 174 | 10.70 (4.82) | 11.79 (5.33) |
|  | 1-2 times | 247 | 10.94 (4.16) | 11.04 (4.28) |
|  | A few times | 459 | 9.63 (2.41) | 9.40 (2.35) |
|  | Many times | 497 | 14.42 (2.49) | 13.92 (2.50) |
|  | Overall | 1,377 | 11.71 (1.55) | 11.62 (1.57) |
| Past Month | 0 times | 174 | 6.61 (4.10) | 7.29 (4.54) |
|  | 1-2 times | 247 | 3.58 (1.15) | 3.54 (1.16) |
|  | A few times | 459 | 4.13 (1.28) | 4.13 (1.31) |
|  | Many times | 497 | 6.48 (1.89) | 6.41 (1.89) |
|  | Overall | 1,377 | 5.16 (0.97) | 5.21 (1.01) |

[^10]Table 7.7a Percentages of Youths (12 to 17) Reporting Lifetime, Past Year, and Past Month Use of Any Illicit Drug and Marijuana among Mother-Child (12 to 17) Pairs, for Mother in the Pair, Asked Whether She Had Spoken to Her Child about the Dangers of Tobacco, Alcohol, or Drug Use within the Past 12 Months: 2012

| Drug | Mother Talked about Dangers with Child | 2012 |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | $n$ | Baseline ${ }^{1}$ | Final ${ }^{2}$ |
| Any Illicit |  |  |  |  |
| LifetimePast YearPast Month | 0 times | 171 | 16.81 (5.15) | 16.94 (5.22) |
|  | 1-2 times | 310 | 19.66 (3.65) | 19.03 (3.54) |
|  | A few times | 546 | 19.04 (2.99) | 19.41 (3.12) |
|  | Many times | 1,147 | 22.10 (1.82) | 22.29 (1.86) |
|  | Overall | 2,174 | 20.52 (1.39) | 20.64 (1.42) |
|  | 0 times | 171 | 5.67 (1.77) | 5.74 (1.76) |
|  | 1-2 times | 310 | 11.75 (2.64) | 11.38 (2.57) |
|  | A few times | 546 | 13.90 (2.70) | 13.97 (2.76) |
|  | Many times | 1,147 | 17.25 (1.71) | 17.45 (1.75) |
|  | Overall | 2,174 | 14.55 (1.19) | 14.65 (1.20) |
|  | 0 times | 171 | 2.73 (1.28) | 2.52 (1.13) |
|  | 1-2 times | 310 | 6.27 (1.88) | 6.20 (1.86) |
|  | A few times | 546 | 7.58 (2.39) | 7.72 (2.44) |
|  | Many times | 1,147 | 8.34 (1.28) | 8.44 (1.32) |
|  | Overall | 2,174 | 7.32 (0.93) | 7.39 (0.95) |
| Marijuana |  |  |  |  |
| Lifetime | 0 times | 171 | 11.81 (4.81) | 11.89 (4.83) |
|  | 1-2 times | 310 | 11.66 (2.72) | 11.39 (2.66) |
|  | A few times | 546 | 13.64 (2.62) | 13.78 (2.68) |
|  | Many times | 1,147 | 15.11 (1.46) | 15.15 (1.49) |
|  | Overall | 2,174 | 13.93 (1.15) | 13.96 (1.16) |
| Past Year | 0 times | 171 | 4.60 (1.62) | 4.51 (1.53) |
|  | 1-2 times | 310 | 9.29 (2.55) | 9.03 (2.47) |
|  | A few times | 546 | 10.87 (2.49) | 10.89 (2.52) |
|  | Many times | 1,147 | 12.66 (1.38) | 12.69 (1.40) |
|  | Overall | 2,174 | 10.98 (1.03) | 10.97 (1.03) |
| Past Month | 0 times | 171 | 1.42 (0.69) | 1.41 (0.66) |
|  | 1-2 times | 310 | 5.14 (1.82) | 5.13 (1.80) |
|  | A few times | 546 | 5.13 (2.12) | 5.25 (2.15) |
|  | Many times | 1,147 | 6.22 (0.94) | 6.23 (0.97) |
|  | Overall | 2,174 | 5.36 (0.76) | 5.40 (0.77) |

[^11]Table 7.7b Percentages of Youths (12 to 17) Reporting Lifetime, Past Year, and Past Month Use of Any Illicit Drug and Marijuana among Father-Child (12 to 17) Pairs, for Father in the Pair, Asked Whether He Had Spoken to His Child about the Dangers of Tobacco, Alcohol, or Drug Use within the Past 12 Months: 2012

| Drug | Father Talked about Dangers with Child | 2012 |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | $n$ | Baseline ${ }^{1}$ | Final ${ }^{2}$ |
| Any Illicit |  |  |  |  |
| Lifetime | 0 times | 174 | 17.12 (4.72) | 18.72 (5.16) |
|  | 1-2 times | 247 | 21.40 (4.75) | 21.62 (4.90) |
|  | A few times | 459 | 19.79 (3.31) | 20.46 (3.62) |
|  | Many times | 497 | 27.18 (3.10) | 26.96 (3.24) |
|  | Overall | 1,377 | 22.43 (1.91) | 22.83 (2.03) |
| Past Year | 0 times | 174 | 13.47 (4.54) | 15.02 (4.99) |
|  | 1-2 times | 247 | 15.00 (4.36) | 14.70 (4.45) |
|  | A few times | 459 | 15.20 (3.09) | 15.84 (3.43) |
|  | Many times | 497 | 19.45 (2.67) | 18.69 (2.65) |
|  | Overall | 1,377 | 16.48 (1.72) | 16.56 (1.82) |
| Past Month | 0 times | 174 | 7.52 (3.71) | 7.71 (3.71) |
|  | 1-2 times | 247 | 5.29 (1.57) | 4.99 (1.48) |
|  | A few times | 459 | 9.20 (2.61) | 9.74 (3.02) |
|  | Many times | 497 | 12.46 (2.30) | 11.90 (2.24) |
|  | Overall | 1,377 | 9.43 (1.33) | 9.39 (1.42) |
| Marijuana |  |  |  |  |
| Lifetime | 0 times | 174 | 6.97 (3.05) | 8.19 (3.79) |
|  | 1-2 times | 247 | 14.49 (4.33) | 14.92 (4.50) |
|  | A few times | 459 | 14.38 (3.09) | 14.96 (3.43) |
|  | Many times | 497 | 16.33 (2.54) | 15.99 (2.58) |
|  | Overall | 1,377 | 14.24 (1.63) | 14.54 (1.75) |
| Past Year | 0 times | 174 | 6.55 (3.04) | 7.76 (3.78) |
|  | 1-2 times | 247 | 12.46 (4.28) | 12.28 (4.35) |
|  | A few times | 459 | 12.65 (3.05) | 13.35 (3.40) |
|  | Many times | 497 | 13.15 (2.39) | 12.65 (2.39) |
|  | Overall | 1,377 | 12.08 (1.58) | 12.25 (1.69) |
| Past Month | 0 times | 174 | 2.57 (1.47) | 2.62 (1.55) |
|  | 1-2 times | 247 | 5.00 (1.54) | 4.70 (1.46) |
|  | A few times | 459 | 8.32 (2.59) | 8.90 (3.01) |
|  | Many times | 497 | 8.53 (2.03) | 8.27 (2.01) |
|  | Overall | 1,377 | 7.10 (1.19) | 7.16 (1.30) |

[^12]
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# Appendix A: Technical Details about the Generalized Exponential Model 

## Appendix A: Technical Details about the Generalized Exponential Model

## A. 1 Distance Function

Let $\Delta(w, d)$ denote the distance between the initial weights $d=\left\{d_{k}: k \in s\right\}$ and the adjusted weights $w$, with $k$ being the $k^{\text {th }}$ unit in the sample and $s$ being the sample selected. The distance function minimized under the generalized exponential model (GEM), subject to calibration constraints, is given by

$$
\begin{equation*}
\Delta(w, d)=\sum_{k \in s} \frac{d_{k}}{A_{k}}\left\{\left(a_{k}-\ell_{k}\right) \log \frac{a_{k}-\ell_{k}}{c_{k}-\ell_{k}}+\left(u_{k}-a_{k}\right) \log \frac{u_{k}-a_{k}}{u_{k}-c_{k}}\right\}, \tag{A.1.1}
\end{equation*}
$$

where $a_{k}=w_{k} / d_{k}, A_{k}=\left(u_{k}-\ell_{k}\right) /\left[\left(u_{k}-c_{k}\right)\left(c_{k}-\ell_{k}\right)\right]$ and $\ell_{k}, c_{k}$, and $u_{k}$ are prescribed real numbers. Let $T_{x}$ denote the $p$-vector of control totals corresponding to predictor variables ( $\left.x_{1}, \ldots, x_{p}\right)$. Then, the calibration constraints for the above minimization problem are

$$
\begin{equation*}
\sum_{k \in s} x_{k} d_{k} a_{k}=T_{x} . \tag{A.1.2}
\end{equation*}
$$

The solution for the above minimization problem, if it exists, is given by a GEM with model parameters $\lambda$; that is,

$$
\begin{equation*}
a_{k}(\lambda)=\frac{\ell_{k}\left(u_{k}-c_{k}\right)+u_{k}\left(c_{k}-\ell_{k}\right) \exp \left\{A_{k} x_{k}^{\prime} \lambda\right\}}{\left(u_{k}-c_{k}\right)+\left(c_{k}-\ell_{k}\right) \exp \left\{A_{k} x_{k}^{\prime} \lambda\right\}} \tag{A.1.3}
\end{equation*}
$$

Note that the number of parameters in the GEM should be $\leq n$, where $n$ is the size of the sample $s$. This is also the dimension of vectors $d$ and $w$. It follows from Equation A.1.3 that

$$
\begin{equation*}
\ell_{k}<a_{k}<u_{k}, k=1, \ldots, n . \tag{A.1.4}
\end{equation*}
$$

The usual raking ratio method (Singh \& Mohl, 1996) of weight adjustment is a special case of the GEM, noting that for $\ell_{k}=0, u_{k}=\infty, c_{k}=1$, and $k=1, \ldots, n$, we have

$$
\begin{equation*}
\Delta(w, d)=\sum_{k \in s} d_{k} a_{k} \log a_{k}-\sum_{k \in s} d_{k}\left(a_{k}-1\right) \tag{A.1.5}
\end{equation*}
$$

and $a_{k}(\lambda)=\exp \left(x_{k}^{\prime} \lambda\right)$.
The logit method of Deville and Särndal (1992) is also a special case of the GEM, by setting $\ell_{k}=\ell, u_{k}=u$, and $c_{k}=1$ for all $k$. The new method was introduced by Folsom and Singh (2000).

## A. 2 GEM Adjustments for Extreme Value Treatment, Nonresponse, and Poststratification

By choosing the user-specified parameters $\ell_{k}, c_{k}$, and $u_{k}$ appropriately, the unified GEM formula (A.1.3) can be justified for all three types of adjustment: extreme value treatment, nonresponse, and poststratification. For extreme value treatment via winsorization, denote the winsorized weights by $\left\{b_{k}\right\}$, where $b_{k}=d_{k}$ if $d_{k}$ is not an extreme weight, and $b_{k}=\operatorname{med}\left\{d_{k}\right\} \pm 3 * \mathrm{IQR}$ if $d_{k}$ is an extreme weight, where IQR denotes the interquartile range, and the median and quartiles for the weights are defined with respect to a suitable design-based stratum.

For the nonresponse adjustment, the sample is first divided into two parts: the nonextreme weight subsample and the extreme weight subsample. For nonextreme weights, the following are set: $\ell_{2}=1, c_{2}=\rho^{-1}, u_{2}=u>\rho^{-1}$, where $\rho$ is the overall response propensity. For extreme weights with high weights, $\ell_{k}=\ell_{1} m_{k}, c_{k}=\rho^{-1} m_{k}$, and $u_{k}=u_{1} m_{k}$, where $m_{k}=b_{k} / d_{k}$ and $1 \leq \ell_{1}<\rho^{-1}=c_{1}<u_{1}$ are prescribed numbers. Similarly, for extreme weights with low weights, $\ell_{k}=\ell_{3} m_{k}, c_{k}=\rho^{-1} m_{k}, u_{k}=u_{3} m_{k}$, and $1 \leq \ell_{3}<\rho^{-1}=c_{3}<u_{3}$.

For the poststratification adjustment, the following weights are set: for nonextreme weights, $\ell_{k}=\ell_{2}, c_{k}=c_{2}=1$, and $u_{k}=u_{2}$; for high extreme weights, $\ell_{k}=\ell_{1} m_{k}, c_{k}=m_{k}$, and $u_{k}=u_{1} m_{k}$; and similarly, for low extreme weights, $\ell_{k}=\ell_{3} m_{k}, c_{k}=m_{k}$, and $u_{k}=u_{3} m_{k}$. The extreme value adjustment is identical to poststratifcation, except for tighter bounds on extreme weights resulting from the final poststratification.

Notice that the GEM allows the flexibility of specifying different bounds for different subsamples. In addition, the lower bound (in the case of nonresponse adjustments) can be made to equal one by choosing the center $c_{k}>1$.

## A. 3 Newton-Raphson Steps

Let $X$ denote the $n \times p$ matrix of predictor values, and for the $v^{t h}$ iteration,

$$
\Gamma_{\phi v}=\operatorname{diag}\left(d_{k} \phi_{k}^{(v)}\right), \phi_{k}^{(o)}=1
$$

where $\phi_{k}^{(v)}=\left[\left(u_{k}-a_{k}^{(v)}\right)\left(a_{k}^{(v)}-\ell_{k}\right)\right] /\left[\left(u_{k}-c_{k}\right)\left(c_{k}-\ell_{k}\right)\right]$.

Then, for the Newton-Raphson iteration $v$, the value of the $p$-vector $\lambda$ is adjusted as

$$
\lambda^{(v)}=\lambda^{(v-1)}+\left(X^{\prime} \Gamma_{\phi, v-1} X\right)^{-1}\left(T_{x}-\hat{T}_{x}^{(v-1)}\right),
$$

where $\lambda^{(0)}=1$.

The convergence criterion is based on the Euclidean distance $\left\|T_{x}-\hat{T}_{x}^{(v)}\right\|$, which is defined as $\sqrt{\left(T_{x}-\hat{T}_{x}^{(v)}\right)^{\prime}\left(T_{x}-\hat{T}_{x}^{(v)}\right)}$. At each iteration, it is checked to determine whether it is decreasing or not. If not, a half step ${ }^{1}$ is used in the iteration increment.

## A. 4 Scaled Constrained Exponential Model

In National Household Surveys on Drug Abuse (NHSDAs) ${ }^{2}$ prior to 1999, constrained exponential models (CEMs) were used for poststratification, and scaled CEMs were used for nonresponse adjustments. The CEM refers to the logit model of Deville and Särndal (1992), in which lower and upper bounds do not vary with $k$; that is, $\ell_{k}=\ell, u_{k}=u$, and $c_{k}=c=1$, such that $\ell<1<u$. Thus, the CEM is a special case of the GEM. For the nonresponse adjustment, Folsom and Witt (1994) modified the CEM estimating equations by a scaling factor ( $\rho^{-1}$, the inverse of the overall response propensity), such that $1<\rho^{-1} a_{k}<\rho^{-1} u$. This implies that choosing $\ell$ in the CEM as $\rho$ ensures that the scaled adjustment factor for nonresponse is at least one.

[^13]
# Appendix B: Derivation of Poststratification Control Totals 

## Appendix B: Derivation of Poststratification Control Totals

Unlike the person-level poststratification adjustment, the control totals for questionnaire dwelling unit (QDU)-level and person pair-level weight calibration could not be derived from the U.S. Bureau of the Census directly. Estimates of the number of households and person pairs were not available at the domains that we wanted to control, and person pair population estimates were not available even at a national level. However, by taking advantage of the two-phase design of the National Survey on Drug Use and Health (NSDUH), the screener dwelling unit (SDU) sample weights could be poststratified to census population estimates. The calibrated SDU weights then could be used as stable control totals for the QDU- and person pair-level sample weights. In addition to the SDU weights, the person pair-level weights were calibrated to a second set of controls derived from the questionnaire, called household-level person counts. These controls were applied to pairs that were members of the 10 selected pair domains given below.

1. Parent-child pairs, child aged 12 to 14 , target population is parents whose children aged 12 to 14 live with them;
2. Parent-child pairs, child aged 12 to 14 , target population is children aged 12 to 14 living with their parents;
3. Parent-child pairs, child aged 12 to 17 , target population is parents whose children aged 12 to 17 live with them;
4. Parent-child pairs, child aged 12 to 17 , target population is children aged 12 to 17 living with their parents;
5. Parent-child pairs, child aged 12 to 20 , target population is parents whose children aged 12 to 20 live with them;
6. Parent-child pairs, child aged 12 to 20 , target population is children aged 12 to 20 living with their parents;
7. Sibling-sibling pairs, older sibling aged 15 to 17 , younger sibling aged 12 to 14 , target population is siblings aged 15 to 17 whose siblings are aged 12 to 14 ;
8. Sibling-sibling pairs, older sibling aged 18 to 25 , younger sibling aged 12 to 17 , target population is siblings aged 18 to 25 whose siblings are aged 12 to 17 ;
9. Spouse-spouse and partner-partner pairs; and
10. Spouse-spouse and partner-partner pairs with children younger than the age of 18 living in the household.

## B. 1 Derivation of QDU-Level Poststratification Controls

The derivation of QDU-level poststratification controls was not directly possible. Instead, it had to be based on work done for the person-level calibration. At the person level, weights were calibrated to the control totals that we wished to reach. These weights then were altered in order to conform to use with QDU-level data.

## B.1.1 Person Level

## B.1.1.1 Receiving and Deriving Person-Level Poststratification Control Totals

Civilian, noninstitutionalized population estimates for ages 12 or older were provided by the Population Estimates Branch of the U.S. Bureau of the Census. We received two files, one at the national level and the other at the State level, each containing estimates of the population broken down by levels of month (12), Hispanicity (2), race (6), gender (2), and age (11).

The breakdown received from the census did not match the levels of the domains that we wanted to control. To account for this, we collapsed levels. From this altered data, we created datasets with model group-specific control totals. Observations in these datasets corresponded to a breakdown by quarter (4), Hispanicity (2), race (5), gender (2), age (11), and number of States ${ }^{1}$ in the model group (number of States varied according to which census region was represented in the model group).

## B.1.1.2 Adjusting SDU Data to the Control Totals

In the person-level weighting, the SDU weights were poststratified to meet control totals based on the population estimates received from the census. For NSDUH weighting, GEM was utilized to calibrate sample weights to multiple control totals. In doing so, each SDU received an adjustment factor, which, when multiplied by the initial weight, produced a final weight. The sum of all final weights corresponded to the civilian, noninstitutionalized population estimate for ages 12 or older, and the sum of all final weights in a domain corresponded to the control total for that domain. Note that there were a number of controls being calibrated to for each SDU, depending upon the domains to which the SDU belonged. The adjusted SDU weight reflected the civilian, noninstitutionalized population estimates for ages 12 or older and could be utilized as a basis for constructing controls at the QDU and person pair levels.

## B.1.2 QDU Level

## B.1.2.1 Deriving QDU-Level Poststratification Control Totals from Adjusted SDU Weights

Since there were no controls for QDU-level poststratification available directly, we used the adjusted SDU weights. For these weights to be applicable at the QDU level, the SDU-level data had to be restructured by sorting and summing over the domains to be used in the QDUlevel calibration. This provided a dataset where the summed weight, which still added up to the proper population, was available for every domain to be utilized in the QDU calibration and thus could be used as a control total.

[^14]
## B.1.2 2 Adjusting QDU-Level Data to the Control Totals

As was done for the SDU data, the QDU-level data was adjusted via calibration in GEM of sample weights to multiple control totals. Each QDU received an adjustment factor, similar to that described for the SDU weight in B.1.1.2. The controls utilized in this calibration were based on the SDU weight as described in B.1.2.1 above. The adjusted weight was representative of the civilian, noninstitutionalized population estimates for ages 12 or older for all domains controlled within the modeling.

## B. 2 Derivation of Person Pair-Level Poststratification Controls

## B.2.1 Deriving Person Pair-Level Poststratification Control Totals from Adjusted SDU Weights and Household-Level Person Counts

Analogous to the QDU weights, some of the person pair controls were based on the SDU weights. However, two sets of control totals were utilized in the modeling, with one set based on the SDU weights and the other set based on the questionnaire roster.

For most pair data domains - those other than the 10 pair domains based on relationship - the control totals for the poststratification adjustments were obtained from SDU data and were based on the number of possible pairs within SDUs. In order to obtain these pair counts belonging to various sociodemographic domains, the screener roster information was used to calculate all possible pairs within SDUs. For example, consider an SDU with two persons aged 12 to 17 and three persons aged 26 to 34 . From this household composition, one can construct one pair of persons aged 12 to 17 , three pairs of persons aged 26 to 34 , and six pairs of persons aged 12 to 17 and 26 to 34. It follows that the total number of possible pairs in this SDU is 10 , from which the number of pairs belonging to the domain of interest can be obtained.

On the other hand, for the 10 selected pair domains based on relationship, the control totals for the poststratification adjustments were obtained from the questionnaire roster. This involved calibrating the pair weights to the number of persons in households belonging to each domain of interest. These controls were obtained from the larger sample of singles and pairs (i.e., one or two persons selected from dwelling units) and were calculated at the QDU (household) level. The pair weights were adjusted by the appropriate multiplicity. See Chapter 11 in Carpenter et al. (2014) for details on the multiplicity counts and household-level control totals, which are referred to as household-level person counts.

## B.2.2 Adjusting Person-Pair Level Data to the Control Totals

Like the SDU- and QDU-level data, the person pair-level data was adjusted via GEM. The use of two different types of controls required a minor modification to the GEM macro so that both sets of controls might be addressed simultaneously. Similar to the SDU- and QDUlevel poststratification steps, each pair received an adjustment factor, which, when multiplied by the initial weight, produced a final weight. The sum of all final weights corresponded to the civilian, noninstitutionalized population estimate for ages 12 or older, and the sum of all final weights in a domain corresponded to the control total for that domain.

# Appendix C: GEM Modeling Summary for the Questionnaire Dwelling Unit Weights 

# Appendix C: GEM Modeling Summary for the Questionnaire Dwelling Unit Weights 

This appendix summarizes each questionnaire dwelling unit (QDU) model group throughout all stages of weight calibration modeling. Unlike much of the other information presented in this report, this appendix provides a model-specific overview of weight calibration, as opposed to a State- or domain-specific one.

For 2012, modeling involved taking four model groups through three adjustment steps: (1) selected dwelling unit poststratification, (2) respondent dwelling unit nonresponse adjustment, and (3) respondent dwelling unit poststratification. After the final poststratification, the adjusted sampling weights were reasonably distributed and did not require the additional treatment of the extreme value step.

Model-specific summary statistics are shown in Tables C.1a through C.4b. Included in these tables, for each stage of modeling, are the number of factor effects included; the high, low, and nonextreme weight bounds set to provide the upper and lower limits for the generalized exponential model (GEM) macro; weighted, unweighted, and winsorized weight proportions; the unequal weighting effect (UWE); and weight distributions. The UWE provides an approximate partial measure of variance and provides a summary of how much impact a particular stage of modeling has on the distribution of the new product of weights. For more details on bounds, see Section 4.1. At each stage in the modeling, these summary statistics were calculated and utilized to help evaluate the quality of the current weight component under the model chosen.

Occurrences of small sample sizes and exact linear combinations in the realized data led to situations whereby inclusion of all originally proposed levels of covariates in the model was not possible. The text and exhibits in Sections C. 1 through C. 4 summarize the decisions made with regard to final covariates included in each model. For a list of the proposed initial covariates considered at each stage of modeling, see Exhibit C.2, and for the list of realized final model covariates, see Exhibits C.1.1 through C.4.3. The following sections establish a series of guidelines to assist in their interpretation.

## C. 1 Final Model Explanatory Variables

For brevity, numeric abbreviations for factor levels are established in Exhibit 4.1 (included here as Exhibit C. 1 for easy reference) in Chapter 4. There, a complete list is provided of all variables and associated levels used at any stage of modeling. Note that not all factors or levels were present in all stages of modeling, and the initial set of variables was the same across model groups but may change over stages of modeling. The initial candidates are found in any of the proposed variables columns for a particular stage of weight adjustment. Exhibits C.1.1 through C.4.3 provide lists of the proposed and realized covariates.

To help understand what effects were controlled for at each stage of the modeling, it was useful to create cross-classification tables as shown in Section C.3. Sections C. 2 and C. 3 explain how to use various exhibits for selected model variables to construct these tables.

Exhibit C. 1 Definitions of Levels for QDU-Level Calibration Modeling Variables
Age $^{\text {b }}$

$$
1: 12-17,2: 18-25,3: 26-34,4: 35-49,5: 50+{ }^{1}
$$

Gender ${ }^{\text {b }}$
1: Male, 2: Female ${ }^{1}$
Group Quarter Indicator ${ }^{\text {a }}$
1: College Dorm, 2: Other Group Quarter, 3: Non-Group Quarter ${ }^{1}$
Hispanicity ${ }^{\text {b }}$
1: Hispanic or Latino, 2: Non-Hispanic or Latino ${ }^{1}$
Household Size ${ }^{\text {b }}$
Continuous Variable Count of Individuals Rostered with DU
Household Type (Ages of Persons Rostered within DU) ${ }^{\text {a }}$
1: 12-17, 18-25, 26+; 2: 12-17, 18-25; 3: 12-17, 26+; 4: 18-25, 26+; 5: 12-17; 6: 18-25; 7: $26+{ }^{1}$
Percentage of Owner-Occupied Dwelling Units in Segment (\% Owner-Occupied) ${ }^{a}$
1: $50-100 \%,{ }^{1}{ }^{1}$ : $10-<50 \%, 3: 0-<10 \%$
Percentage of Segments That Are Black or African American ${ }^{\text {a }}$ 1: $50-100 \%, 2: 10-<50 \%, 3: 0-<10 \%^{1}$
Percentage of Segments That Are Hispanic or Latino ${ }^{\text {a }}$ 1: 50-100\%, 2: $10-<50 \%, 3: 0-<10 \%{ }^{1}$
Population Density ${ }^{\text {a }}$
1: MSA $1,000,000$ or More, 2 : MSA Less than $1,000,000,3$ : Non-MSA Urban, 4: Non-MSA Rural ${ }^{1}$
Quarter ${ }^{\text {a,b }}$
1: Quarter 1, 2: Quarter 2, 3: Quarter 3, 4: Quarter $4{ }^{1}$
Race (3 Levels) ${ }^{\text {b }}$ 1: White ${ }^{1}$, 2: Black or African American, 3: Other
Race ( 5 Levels) ${ }^{\text {b }}$ 1: White, ${ }^{1}$ 2: Black or African American, 3: American Indian or Alaska Native, 4: Asian, 5: Two or More Races Race/Ethnicity of Householder ${ }^{\text {a }}$

1: Hispanic or Latino White, ${ }^{1}$ 2: Hispanic or Latino Black or African American, 3: Hispanic or Latino Other, 4: Non-Hispanic or Latino White, 5: Non-Hispanic or Latino Black or African American, 6: Non-Hispanic or Latino Other

## Relation to Householder ${ }^{\text {a }}$

1: Householder or Spouse, 2: Child, 3: Other Relative, 4: Nonrelative ${ }^{1}$
Segment-Combined Median Rent and Housing Value (Rent/Housing) ${ }^{\text {a,2 }}$
1: First Quintile, 2: Second Quintile, 3: Third Quintile, 4: Fourth Quintile, 5: Fifth Quintile ${ }^{1}$

Exhibit C. 1 Definitions of Levels for QDU-Level Calibration Modeling Variables (continued)

| Stat ${ }^{\text {a,b,3 }}$ |  |
| :---: | :---: |
|  | Model Group 1: 1: Connecticut, 2: Maine, 3: Massachusetts, ${ }^{1}$ 4: New Hampshire, 5: New Jersey, 6: New York, 7: Pennsylvania, 8: Rhode Island, 9: Vermont |
|  | Model Group 2: 1: Illinois, 2: Indiana, 3: Iowa, 4: Kansas, 5: Michigan, 6: Minnesota, 7: Missouri, 8: Nebraska, 9: North Dakota, 10: Ohio, 11: South Dakota, 12: Wisconsin ${ }^{1}$ |
|  | Model Group 3: 1: Alabama, 2: Arkansas, 3: Delaware, 4: District of Columbia, 5: Florida, 6: Georgia, 7: Kentucky, 8: Louisiana, 9: Maryland, 10: Mississippi, 11: North Carolina, ${ }^{1}$ 12: Oklahoma, 13: South Carolina, 14: Tennessee, 15: Texas, 16: Virginia, 17: West Virginia |
|  | Model Group 4: 1: Alaska, 2: Arizona, ${ }^{1}$ 3: California, 4: Colorado, 5: Idaho, 6: Hawaii, 7: Montana, 8: Nevada, 9: New Mexico, 10: Oregon, 11: Utah, 12: Washington, 13: Wyoming |
| State/Region ${ }^{\text {a,3 }}$ |  |
|  | Model Group 1: 1: New York, 2: Pennsylvania, 3: Other ${ }^{1}$ |
|  | Model Group 2: 1: Illinois, 2: Michigan, 3: Ohio, 4: Other ${ }^{1}$ |
|  | Model Group 3: 1: Florida, 2: Texas, 3: Other ${ }^{1}$ |
|  | Model Group 4: 1: California, 2: Other ${ }^{1}$ |
| DU = dwelling unit; MSA = metropolitan statistical area; QDU = questionnaire dwelling unit. |  |
| ${ }^{1}$ The reference level for this variable. This is the level against which effects of other factor levels are measured. |  |
| ${ }^{2}$ Segment-Combined Median Rent and Housing Value is a composite measure based on rent, housing value, and percentage owner-occupied. |  |
| ${ }^{3}$ The States or district assigned to a particular model is based on census regions. |  |
| ${ }^{\text {a }}$ Binary variable. |  |
|  | Counting variable. A count of all persons in the household. |

${ }^{\mathrm{b}}$ Counting variable. A count of all persons in the household.

## C. 2 Glossary of Terms Used in the Description of the Variables in the Final Model

This glossary provides a list of general terms. Certain other terms are sometimes used within a particular section.

All levels present. All effects and all levels of the factor under consideration are in the model.

Coll. (levels). Collapse these factor effects together. Factor effects that have been collapsed with others manifest themselves jointly in the model.

Conv. If model is not convergent, dropping or collapsing of variables is performed.
Do the same for (effects). Repeat the previous step for all effect levels listed.
Drop all levels. All factor effects are completely removed from the model for all levels and any combinations involving this factor.

Drop level(s). Collapse these factor effects into the reference set. The factor effects comprising the dropped levels are manifested jointly with either some or all of the factor effects in the reference set.

Drop level(s); sing. During the modeling process the factor effects listed are removed from the model due to singularity.

Drop level(s); zero cnts. During the modeling process the factor effects listed are removed from the model due to zero sample.

Drop or Collapse using*. The asterisk is used as a wild card character to indicate all levels of the factor for that effect.

Factor effect. The factor effect represents the effects of levels considered for one factor, two factors, and higher order factors.

Hier. One or more of the factor effects in a higher order interaction is collapsed or dropped in an interaction at a lower order and the hierarchical effect carries up, either eliminating or combining factors of higher order interactions with that effect.

Reference/reference set. Factor effects composed of reference levels are not explicitly listed in the set of model variables. However, these effects manifest themselves either separately or in combination with other factors depending on the presence of other factors in the model.

## C. 3 How To Interpret Collapsing and Dropping of Factor Effects

To help visualize what effects are directly controlled for in our model, one can construct the table that reflects the collapsing scheme employed. The following is a complex example from the 2004 person-level modeling (Chen et al., 2006).

1. Locate the Factor Effect-Model 9 Person Nonresponse Adjustment.

## Three-Factor Effects

State $\times$ Age $\times$ Race (3 Levels)

## Comments

Coll. $(2,1,2) \&(2,1,3)$; hier. Repeat for all levels of age in State (2); hier. Coll. ( $1,4,2$ ) \& ( $1,4,3$ ); conv. Drop (3,4,2); sing. Drop $\left(3,{ }^{*}, *\right)$; conv. Coll. $(5,1,2) \&(5,1,3)$; conv. Repeat for all levels of age in State (5).
2. Determine the initial range of possible levels for the variables by referring to the variable definitions. See Exhibits C. 1 and H. 1 for QDU- and pair-level variable definitions. In addition, the columns "Levels," "Proposed," and "Final" will provide counts of all factor effects, all explicitly proposed factors, and all explicitly controlled factors, but these are not necessary for construction of the cross-classification table. The following example is based upon person-level variables, but the process is the same.

State (for the model group in question, in this case, Model Group 9)
Model Group 9: 1: Alaska, 2: Hawaii, 3: Oregon, 4: Washington, ${ }^{1}$ 5: California

## Age

1: 12 to $17,2: 18$ to $25,3: 26$ to $34,4: 35$ to $49,5: 50+{ }^{1}$

## Race (3 Levels)

1: White, ${ }^{1}$ 2: Black or African American, 3: Other
3. Construct the cross-classification table.

For example, the initial proposed set of covariates in Race (4 Levels) is defined this way:

| Race (4 Levels) | White | Black or African <br> American | Asian | American Indian <br> or Alaska Native |
| :---: | :---: | :---: | :---: | :---: |

Shading indicates the reference-level set.

[^15]This is the cross-classification table for the initial proposed set of covariates in State $\times$ Race (4 Levels):

| State $\times$ Race (4 Levels) | White | Black or African <br> American | Asian | American Indian <br> or Alaska Native |
| :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |
|  |  |  |  |  |
| AK |  |  |  |  |
| HI |  |  |  |  |
| OR |  |  |  |  |
| WA |  |  |  |  |
| CA |  |  |  |  |

Shading indicates the reference-level set.
The cross-classification table of interest for the initial proposed set of covariates in State $\times$ Age $\times$ Race (3 Levels) is as follows:

| State $\times$ Age $\times$ Race (3 Levels) | White | Black or African American | Other |
| :---: | :---: | :---: | :---: |
| AK $\times 12-17$ |  |  |  |
| 18-25 |  |  |  |
| 26-34 |  |  |  |
| 35-49 |  |  |  |
| 50+ |  |  |  |
| $\mathrm{HI} \times 12-17$ |  |  |  |
| 18-25 |  |  |  |
| 26-34 |  |  |  |
| 35-49 |  |  |  |
| 50+ |  |  |  |
| OR $\times 12-17$ |  |  |  |
| 18-25 |  |  |  |
| 26-34 |  |  |  |
| 35-49 |  |  |  |
| 50+ |  |  |  |
| WA $\times 12-17$ |  |  |  |
| 18-25 |  |  |  |
| 26-34 |  |  |  |
| 35-49 |  |  |  |
| 50+ |  |  |  |
| CA $\times 12-17$ |  |  |  |
| 18-25 |  |  |  |
| 26-34 |  |  |  |
| 35-49 |  |  |  |
| 50+ |  |  |  |

Shading indicates the reference-level set.
The number of respondents in the class State $\times$ Age $\times$ Race (3 Levels) at this stage of modeling would appear within each cell of the table. Construction of the other cross-
classification tables follows the same logic and is only necessary to the point of providing understanding of the final table.
4. Use the information under the "Comments" column definition to determine the combination of factors controlled.

| One-Factor Effects | Comments <br> All levels present. |
| :--- | :--- |
| State | All levels present. |
| Race (4 Levels) | All levels present. |
| Age | Comments |
| Two-Factor Effects | All levels present. |
| State $\times$ Age | Coll. (1,3) \& (1,4). Do the same for all other States except (2). Coll. <br> (2,2), (2,3), \& (2,4). |
| State $\times$ Race (4 Levels) |  |
| Age $\times$ Race (3 Levels) | All levels present. |

The reason for the hier. instruction in the three-factor effect directions is the State $\times$ Race (4 Levels) interaction. It indicates a need to maintain the collapsing scheme when setting up any three-factor crosses involving State $\times$ Race. Following these directions, the resulting two-factor table we would then have to work with is as follows:

| State $\times$ Race (4 Levels) | White | Black or African <br> American | Asian | American Indian or <br> Alaska Native |
| :---: | :---: | :---: | :---: | :---: |
| AK |  |  |  |  |
| HI |  |  |  |  |
| OR |  |  |  |  |
| WA |  |  |  |  |
| CA |  |  |  |  |

Shading indicates the reference-level set.
Returning to our instructions, we see that several other factor crosses have been affected by modeling:

## Three-Factor Effects

State $\times$ Age $\times$ Race (3 Levels)

## Comments

Coll. $(2,1,2) \&(2,1,3)$; hier. Repeat for all levels of age in State (2); hier. Coll. $(1,4,2) \&(1,4,3)$; conv. Drop $(3,4,2)$; sing. Drop (3,*,*); conv. Coll. $(5,1,2) \&(5,1,3)$; conv. Repeat for all levels of age in State (5).

Construct the complete table, and then begin combining blocks as directed. The unshaded cells represent the factors directly controlled for by the model. The shaded cells represent the composite reference set, whose values may be obtained by utilizing the marginal sums, although when changes to the initially proposed set occur, it can make certain reference cell counts indistinguishable.

After following the directions, the resulting post-modeling cross-classification table should appear as follows:

| State $\times$ Age $\times$ Race (3 Levels) | White | Black or African American | Other |
| :---: | :---: | :---: | :---: |
| AK $\times 12-17$ |  |  |  |
| 18-25 |  |  |  |
| 26-34 |  |  |  |
| 35-49 |  |  |  |
| 50+ |  |  |  |
| $\mathrm{HI} \times 12-17$ |  |  |  |
| 18-25 |  |  |  |
| 26-34 |  |  |  |
| 35-49 |  |  |  |
| 50+ |  |  |  |
| OR $\times 12-17$ |  |  |  |
| 18-25 |  |  |  |
| 26-34 |  |  |  |
| 35-49 |  |  |  |
| 50+ |  |  |  |
| WA $\times 12-17$ |  |  |  |
| 18-25 |  |  |  |
| 26-34 |  |  |  |
| 35-49 |  |  |  |
| 50+ |  |  |  |
| CA $\times 12-17$ |  |  |  |
| 18-25 |  |  |  |
| 26-34 |  |  |  |
| 35-49 |  |  |  |
| 50+ |  |  |  |

Shading indicates the reference-level set.

Exhibit C. 2 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights

| Variables | Binary | Counting | Level | Proposed |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 76 | 76 |  |
| Intercept | Yes |  | 1 | 1 |
| Population Density | Yes |  | 4 | 3 |
| Group Quarter | Yes |  | 3 | 2 |
| Race/Ethnicity of Householder | Yes |  | 6 | 5 |
| Rent/Housing | Yes |  | 5 | 4 |
| Segment \% Black or African American | Yes |  | 3 | 2 |
| Segment \% Hispanic or Latino | Yes |  | 3 | 2 |
| Segment \% Owner-Occupied | Yes |  | 3 | 2 |
| Household Type | Yes |  | 7 | 6 |
| State | Yes | Yes | Model-specific |  |
| Quarter | Yes | Yes | 4 | 3 |
| Age Group |  | Yes | 5 | 4 |
| Race |  | Yes | 5 | 4 |
| Hispanicity |  | Yes | 2 | 1 |
| Gender |  | Yes | 2 | 1 |
| Household Size |  | Yes | 1 | 1 |
| Two-Factor Effects |  |  |  |  |
| Age $\times$ Race (3 Levels) |  | Yes | $5 \times 3$ | 8 |
| Age $\times$ Hispanicity |  | Yes | $5 \times 2$ | 4 |
| Age $\times$ Gender |  | Yes | $5 \times 2$ | 4 |
| Race (3 Levels) $\times$ Hispanicity |  | Yes | $3 \times 2$ | 2 |
| Race (3 Levels) $\times$ Gender |  | Yes | $3 \times 2$ | 2 |
| Hispanicity $\times$ Gender |  | Yes | $2 \times 2$ | 1 |
| State $\times$ Age |  | Yes | Model-specific |  |
| State $\times$ Race (5 Levels) |  | Yes | Model-specific |  |
| State $\times$ Gender |  | Yes | Model-specific |  |
| State $\times$ Hispanicity |  | Yes | Model-specific |  |
| \% Black or African American $\times$ \% Owner-Occupied | Yes |  | $3 \times 3$ | 4 |
| \% Black or African American $\times$ Rent/Housing |  | Yes | $3 \times 5$ | 8 |
| \% Hispanicity $\times$ \% Owner-Occupied |  | Yes | $3 \times 3$ | 4 |
| \% Hispanicity $\times$ Rent/Housing |  | Yes | $3 \times 5$ | 8 |
| \% Owner $\times$ Rent/Housing | Yes |  | $3 \times 5$ | 8 |
| Three-Factor Effects |  |  |  |  |
| Race (3 Levels) $\times$ Age $\times$ Gender |  | Yes | 8 | 8 |
| State/Region $\times$ Age $\times$ Gender |  | Yes |  |  |
| State/Region $\times$ Age $\times$ Hispanicity |  | Yes |  |  |
| State/Region $\times$ Age $\times$ Race (3 Levels) |  | Yes |  |  |
| State/Region $\times$ Hispanicity $\times$ Gender |  | Yes |  |  |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity |  | Yes |  |  |
| State $/$ Region $\times$ Race (3 Levels) $\times$ Gender |  | Yes |  |  |

# Appendix C.1: Model Group 1: Northeast 

(Connecticut, Maine, Massachusetts, New Hampshire, New Jersey, New York, Pennsylvania, Rhode Island, Vermont)

Table C.1a 2012 QDU Weight GEM Modeling Summary (Model Group 1: Northeast)


GEM = generalized exponential model; QDU = questionnaire dwelling unit.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.
${ }^{2}$ Unequal weighting effect (UWE) defined as $1+[(n-1) / n]^{*} C V^{2}$, where $C V=$ coefficient of variation of weights.
${ }^{3}$ Number of proposed covariates on top line and number finalized after modeling.
${ }^{4}$ There are six sets of bounds for each modeling step. Nominal bounds are used in defining maximum/minimum values for the GEM adjustment factors. The realized bound is the actual adjustment produced by the modeling. The set of three bounds listed for each step correspond to the high extreme values, the nonextreme values, and the low extreme values.

Table C.1b 2012 Distribution of Weight Adjustment Factors and Weight Products (Model Group 1: Northeast)


QDU = questionnaire dwelling unit; $\operatorname{SDU}=$ screener dwelling unit.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.

## Model Group 1 Overview

## Selected Questionnaire Dwelling Unit-Level Poststratification

Out of 243 proposed effects, 242 were kept in the model, with the exception of State by Race, which combined American Indian or Alaska Native and Asian for New Hampshire.

## Respondent Questionnaire Dwelling Unit-Level Nonresponse

Out of 243 proposed effects, 236 were kept in the model. Two-factor effects were modified for State by Race, combining American Indian or Alaska Native and Asian for each of the following States: New Jersey, Connecticut, New York, and Rhode Island. Three-factor effects for State/Region by Race by Hispanicity combined Black or African American with Other.

## Respondent Questionnaire Dwelling Unit-Level Poststratification

Out of 243 proposed effects, 239 were kept in the model. Two-factor effects were modified for State by Race, combining American Indian or Alaska Native and Asian for each of the following States: New Jersey, Connecticut, Vermont, and New Hampshire. All main and three-factor effects were retained at proposed levels.

Exhibit C.1.1 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (sel.qdu.ps) Model Group 1: Northeast

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 60 | 60 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 |  | 4 | All levels present. |
| Population Density | 4 | 3 |  | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 |  |  | All levels present. |
| State (Count) | 9 | 8 | 8 | All levels present. |
| State (Binary) | 9 | 8 | 8 | All levels present. |
| Quarter (Count) | 4 | 3 | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 | 3 | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 |  | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 133 | 132 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ | 4 | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ | 1 | 1 | All levels present. |
| State $\times$ Age | $9 \times 5$ | 32 | 32 | All levels present. |
| State $\times$ Race | $9 \times 5$ | 32 | 31 | Coll. $(4,3) \&(4,4) ;$ conv. |
| State $\times$ Gender | $9 \times 2$ | 8 | 8 | All levels present. |
| State $\times$ Hispanicity | $9 \times 2$ | 8 | 8 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Hispanicity $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Hispanicity $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 50 | 50 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $3 \times 5 \times 2$ | 8 |  | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $3 \times 5 \times 3$ | 16 | 16 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $3 \times 2 \times 2$ | 2 | 2 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $3 \times 3 \times 2$ | 4 | 4 | All levels present. |
| State $/$ Region $\times$ Race (3 Levels) $\times$ Gender | $3 \times 3 \times 2$ | 4 | 4 | All levels present. |
| Total |  | 243 | 242 |  |

Exhibit C.1.2 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (res.qdu.nr) Model Group 1: Northeast

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 60 | 60 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| State (Count) | 9 | 8 | 8 | All levels present. |
| State (Binary) | 9 | 8 | 8 | All levels present. |
| Quarter (Count) | 4 | 3 | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 | 3 | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 | 4 | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 133 | 127 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ |  | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ | 1 | 1 | All levels present. |
| State $\times$ Age | $9 \times 5$ | 32 | 32 | All levels present. |
| State $\times$ Race | $9 \times 5$ | 32 | 26 | Coll. $(6,3) \&(6,4),(1,3) \&(1,4)$, <br> $(5,3) \&(5,4) \&(5,5),(8,3) \&$ <br> $(8,4) \&(8,5) ;$ conv. |
| State $\times$ Gender | $9 \times 2$ | 8 | 8 | All levels present. |
| State $\times$ Hispanicity | $9 \times 2$ | 8 | 8 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Hispanicity $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Hispanicity $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 50 | 49 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $3 \times 5 \times 3$ | 16 | 16 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $3 \times 2 \times 2$ | 2 | 2 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $3 \times 3 \times 2$ | 4 | 3 | Coll. (1,2,1) \& (1,3,1); conv. |
| State/Region $\times$ Race (3 Levels) $\times$ Gender | $3 \times 3 \times 2$ | 4 | 4 | All levels present. |
| Total |  | 243 | 236 |  |

Exhibit C.1.3 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (res.qdu.ps) Model Group 1: Northeast

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 60 | 60 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| State (Count) | 9 | 8 | 8 | All levels present. |
| State (Binary) | 9 | 8 | 8 | All levels present. |
| Quarter (Count) | 4 | 3 | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 | 3 | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 | 4 | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 133 | 129 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ | 4 | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ | 1 | 1 | All levels present. |
| State $\times$ Age | $9 \times 5$ | 32 | 32 | All levels present. |
| State $\times$ Race | $9 \times 5$ | 32 | 28 | Coll. $(1,3) \&(1,4),(4,3) \&(4,4)$, $(5,3) \&(5,4),(9,3) \&(9,4) ;$ conv. |
| State $\times$ Gender | $9 \times 2$ | 8 | 8 | All levels present. |
| State $\times$ Hispanicity | $9 \times 2$ | 8 | 8 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Hispanicity $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Hispanicity $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 50 | 50 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $3 \times 5 \times 3$ | 16 | 16 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $3 \times 2 \times 2$ | 2 | 2 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $3 \times 3 \times 2$ | 4 | 3 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Gender | $3 \times 3 \times 2$ | 4 | 4 | All levels present. |
| Total |  | 243 | 239 |  |

## Appendix C.2: Model Group 2: Midwest

(Illinois, Indiana, Iowa, Kansas, Michigan, Minnesota, Missouri, Nebraska, North Dakota, Ohio, South Dakota, Wisconsin)

Table C.2a 2012 QDU Weight GEM Modeling Summary (Model Group 2: Midwest)

| Modeling Step ${ }^{1}$ | Extreme Weight Proportions |  |  | UWE ${ }^{\mathbf{2}}$ | \# Covariates ${ }^{3}$ | Bounds ${ }^{4}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | \% Unweighted | \% Weighted | \% Outwinsor |  |  | Nominal | Realized |
| sel.qdu.ps | 1.78 | 1.86 | 0.42 | 2.6793 | 300 | (0.42, 1.10) | (0.42, 1.10) |
|  | 1.10 | 0.98 | 0.16 | 2.7109 | 300 | (0.20, 3.27) | (0.20, 3.27) |
|  |  |  |  |  |  | (0.90, 2.07) | (0.90, 2.07) |
| res.qdu.nr | 1.25 | 1.26 | 0.20 | 2.7756 | 300 | (1.00, 1.30) | (1.00, 1.30) |
|  | 0.76 | 0.90 | 0.12 | 2.9433 | 299 | (1.00, 3.43) | (1.00, 3.33) |
|  |  |  |  |  |  | (1.10, 1.56) | (1.10, 1.55) |
| res.qdu.ps | 0.76 | 0.90 | 0.12 | 2.9433 | 300 | (0.91, 1.20) | (0.92, 1.20) |
|  | 0.81 | 0.97 | 0.05 | 2.9433 | 300 | (0.77, 1.21) | (0.78, 1.21) |
|  |  |  |  |  |  | (0.90, 1.03) | (0.90, 1.03) |

$\mathrm{GEM}=$ generalized exponential model; $\mathrm{QDU}=$ questionnaire dwelling unit.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.
${ }^{2}$ Unequal weighting effect (UWE) defined as $1+[(n-1) / n]^{*} C V^{2}$, where $C V=$ coefficient of variation of weights.
${ }^{3}$ Number of proposed covariates on top line and number finalized after modeling.
4 There are six sets of bounds for each modeling step. Nominal bounds are used in defining maximum/minimum values for the GEM adjustment factors. The realized bound is the actual adjustment produced by the modeling. The set of three bounds listed for each step correspond to the high extreme values, the nonextreme values, and the low extreme values.

Table C.2b 2012 Distribution of Weight Adjustment Factors and Weight Products (Model Group 2: Midwest)

|  | SDU Weight | QDU Design Weight |  | sel.qdu.ps ${ }^{1}$ |  | res.qdu.nr ${ }^{1}$ |  | res.qdu.ps ${ }^{1}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1-10 | duwght11 | 1-11 | duwght12 | 1-12 | duwght13 | 1-13 | duwght14 | 1-14 |
| Minimum | 33 | 1.00 | 40 | 0.19 | 23 | 0.43 | 30 | 0.56 | 27 |
| 1\% | 97 | 1.00 | 107 | 0.61 | 109 | 1.00 | 128 | 0.96 | 128 |
| 5\% | 147 | 1.00 | 217 | 0.78 | 210 | 1.02 | 242 | 0.98 | 241 |
| 10\% | 268 | 1.00 | 373 | 0.85 | 350 | 1.05 | 385 | 0.99 | 382 |
| 25\% | 452 | 1.00 | 513 | 0.94 | 507 | 1.13 | 600 | 1.00 | 601 |
| Median | 541 | 1.13 | 737 | 1.00 | 762 | 1.21 | 911 | 1.00 | 913 |
| 75\% | 707 | 3.64 | 1,773 | 1.07 | 1,751 | 1.32 | 2,041 | 1.00 | 2,047 |
| 90\% | 1,150 | 6.65 | 3,754 | 1.16 | 3,805 | 1.44 | 4,917 | 1.01 | 4,913 |
| 95\% | 1,385 | 7.56 | 5,434 | 1.25 | 5,460 | 1.53 | 6,977 | 1.02 | 6,977 |
| 99\% | 1,832 | 11.84 | 9,852 | 1.59 | 10,371 | 1.81 | 14,335 | 1.05 | 14,315 |
| Maximum | 6,654 | 13.41 | 43,516 | 3.56 | 25,191 | 3.33 | 35,530 | 1.29 | 35,530 |
| $n$ | 16,984 | - | 16,984 | - | 16,984 | - | 13,687 | - | 13,687 |
| Mean | 633 | 2.50 | 1,544 | 1.01 | 1,558 | 1.24 | 1,933 | 1.00 | 1,933 |
| Max/Mean | 11 | - | 28 | - | 16 | - | 18 | - | 18 |

QDU = questionnaire dwelling unit; SDU = screener dwelling unit.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.

## Model Group 2 Overview

## Selected Questionnaire Dwelling Unit-Level Poststratification

All 300 proposed effects were kept in the model.

## Respondent Questionnaire Dwelling Unit-Level Nonresponse

All main and three-factor effects were maintained at proposed levels. Two-factor effects were modified for State by Race, combining American Indian or Alaska Native and Asian for North Dakota.

## Respondent Questionnaire Dwelling Unit-Level Poststratification

All 300 proposed effects were kept in the model.

Exhibit C.2.1 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (sel.qdu.ps) Model Group 2: Midwest

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 66 | 66 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| State (Count) | 9 | 11 | 11 | All levels present. |
| State (Binary) | 9 | 11 | 11 | All levels present. |
| Quarter (Count) | 4 |  | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 | 3 | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 | 4 | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 163 | 163 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ | 4 | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ | 1 | 1 | All levels present. |
| State $\times$ Age | $12 \times 5$ | 44 | 44 | All levels present. |
| State $\times$ Race | $12 \times 5$ | 44 | 44 | All levels present. |
| State $\times$ Gender | $12 \times 2$ | 11 | 11 | All levels present. |
| State $\times$ Hispanicity | $12 \times 2$ | 11 | 11 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Hispanicity or Latino $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Hispanicity or Latino $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 71 | 71 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $4 \times 5 \times 2$ | 12 | 12 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $4 \times 5 \times 2$ | 12 | 12 | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $4 \times 5 \times 3$ | 24 | 24 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $4 \times 2 \times 2$ | 3 | 3 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $4 \times 3 \times 2$ | 6 | 6 | All levels present. |
| State $/$ Region $\times$ Race (3 Levels) $\times$ Gender | $4 \times 3 \times 2$ | 6 | 6 | All levels present. |
| Total |  | 300 | 300 |  |

Exhibit C.2.2 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (res.qdu.nr) Model Group 2: Midwest

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 66 | 66 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| State (Count) | 9 | 11 | 11 | All levels present. |
| State (Binary) | 9 | 11 | 11 | All levels present. |
| Quarter (Count) | 4 | 3 | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 | 3 | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 | 4 | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 163 | 162 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ | 4 | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ | 1 | 1 | All levels present. |
| State $\times$ Age | $12 \times 5$ | 44 | 44 | All levels present. |
| State $\times$ Race | $12 \times 5$ | 44 | 43 | Coll. $(9,3) \&(9,4) ;$ conv. |
| State $\times$ Gender | $12 \times 2$ | 11 | 11 | All levels present. |
| State $\times$ Hispanicity | $12 \times 2$ | 11 | 11 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Hispanicity $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 3 | All levels present. |
| \% Hispanicity $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 71 | 71 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $4 \times 5 \times 2$ | 12 | 12 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $4 \times 5 \times 2$ | 12 | 12 | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $4 \times 5 \times 3$ | 24 | 24 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $4 \times 2 \times 2$ | 3 | 3 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $4 \times 3 \times 2$ | 6 | 6 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Gender | $4 \times 3 \times 2$ | 6 | 6 | All levels present. |
| Total |  | 300 | 299 |  |

Exhibit C.2.3 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (res.qdu.ps) Model Group 2: Midwest

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 66 | 66 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 |  | 2 | All levels present. |
| State (Count) | 9 | 11 | 11 | All levels present. |
| State (Binary) | 9 | 11 | 11 | All levels present. |
| Quarter (Count) | 4 | 3 | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 | 3 | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 | 4 | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 163 | 163 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ | 4 | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ | 1 | 1 | All levels present. |
| State $\times$ Age | $12 \times 5$ | 44 | 44 | All levels present. |
| State $\times$ Race | $12 \times 5$ | 44 | 44 | All levels present. |
| State $\times$ Gender | $12 \times 2$ | 11 | 11 | All levels present. |
| State $\times$ Hispanicity | $12 \times 2$ | 11 | 11 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Hispanicity $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Hispanicity $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 71 | 71 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $4 \times 5 \times 2$ | 12 | 12 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $4 \times 5 \times 2$ | 12 | 12 | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $4 \times 5 \times 3$ | 24 | 24 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $4 \times 2 \times 2$ | 3 | 3 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $4 \times 3 \times 2$ | 6 | 6 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Gender | $4 \times 3 \times 2$ | 6 | 6 | All levels present. |
| Total |  | 300 | 300 |  |

## Appendix C.3: Model Group 3: South

(Alabama, Arkansas, Delaware, District of Columbia, Florida, Georgia, Kentucky, Louisiana, Maryland, Mississippi, North Carolina, Oklahoma, South Carolina, Tennessee, Texas, Virginia, West Virginia)

Table C.3a 2012 QDU Weight GEM Modeling Summary (Model Group 3: South)

| Modeling Step ${ }^{1}$ | Extreme Weight Proportions |  |  | UWE ${ }^{\mathbf{2}}$ | \# Covariates ${ }^{3}$ | $\text { Bounds }{ }^{4}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | \% Unweighted | \% Weighted | \% Outwinsor |  |  | Nominal | Realized |
| sel.qdu.ps | 1.50 | 1.91 | 0.38 | 2.5352 | 339 | (0.42, 1.10) | (0.42, 1.10) |
|  | 1.17 | 1.51 | 0.17 | 2.5423 | 338 | $(0.46,4.94)$ | (0.49, 4.93) |
|  |  |  |  |  |  | $(0.90,4.99)$ | $(0.90,4.99)$ |
| res.qdu.nr | 1.24 | 1.57 | 0.20 | 2.5925 | 339 | (1.00, 1.70) | (1.00, 1.70) |
|  | 1.10 | 1.49 | 0.24 | 2.8943 | 339 | (1.00, 5.00) | (1.00, 5.00) |
|  |  |  |  |  |  | (1.30, 4.97) | (1.30, 4.79) |
| res.qdu.ps | 1.10 | 1.49 | 0.24 | 2.8943 | 339 | (0.20, 1.50) | (0.94, 1.50) |
|  | 1.09 | 1.51 | 0.12 | 2.8950 | 338 | $(0.20,5.00)$ | (0.79, 2.79) |
|  |  |  |  |  |  | (0..90, 5.00) | (0.95, 1.07) |

GEM = generalized exponential model; QDU = questionnaire dwelling unit.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.
${ }^{2}$ Unequal weighting effect (UWE) defined as $1+[(n-1) / n]^{*} C V^{2}$, where $C V=$ coefficient of variation of weights.
${ }^{3}$ Number of proposed covariates on top line and number finalized after modeling.
${ }^{4}$ There are six sets of bounds for each modeling step. Nominal bounds are used in defining maximum/minimum values for the GEM adjustment factors. The realized bound is the actual adjustment produced by the modeling. The set of three bounds listed for each step correspond to the high extreme values, the nonextreme values, and the low extreme values.

Table C.3b 2012 Distribution of Weight Adjustment Factors and Weight Products (Model Group 3: South)

|  | SDU Weight | QDU Design Weight |  | sel.qdu.ps ${ }^{1}$ |  | res.qdu.nr ${ }^{1}$ |  | res.qdu.ps ${ }^{1}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1-10 | duwght11 | 1-11 | duwght12 | 1-12 | duwght13 | 1-13 | duwght14 | 1-14 |
| Minimum | 9 | 1.00 | 12 | 0.25 | 10 | 0.52 | 20 | 0.49 | 25 |
| 1\% | 59 | 1.00 | 75 | 0.65 | 71 | 1.00 | 75 | 0.94 | 75 |
| 5\% | 112 | 1.00 | 216 | 0.80 | 213 | 1.02 | 219 | 0.98 | 220 |
| 10\% | 253 | 1.00 | 425 | 0.85 | 412 | 1.05 | 449 | 0.99 | 444 |
| 25\% | 641 | 1.00 | 817 | 0.92 | 813 | 1.10 | 911 | 1.00 | 912 |
| Median | 923 | 1.15 | 1,284 | 1.00 | 1,300 | 1.18 | 1,479 | 1.00 | 1,482 |
| 75\% | 1,313 | 3.50 | 2,761 | 1.09 | 2,787 | 1.29 | 2,980 | 1.00 | 2,975 |
| 90\% | 1,716 | 6.54 | 6,014 | 1.19 | 6,214 | 1.43 | 7,732 | 1.01 | 7,712 |
| 95\% | 2,144 | 7.92 | 8,523 | 1.27 | 8,924 | 1.53 | 11,901 | 1.02 | 11,911 |
| 99\% | 3,118 | 12.02 | 14,217 | 1.59 | 14,457 | 1.82 | 19,919 | 1.06 | 19,913 |
| Maximum | 8,236 | 21.12 | 45,296 | 5.31 | 44,158 | 5.00 | 64,497 | 2.79 | 63,870 |
| $n$ | 18,345 | - | 18,345 | - | 18,345 | - | 15,019 | - | 15,019 |
| Mean | 1,013 | 2.55 | 2,405 | 1.02 | 2,444 | 1.22 | 2,985 | 1.00 | 2,985 |
| Max/Mean | 8 | - | 19 | - | 18 | - | 22 | - | 21 |

QDU = questionnaire dwelling unit; SDU = screener dwelling unit.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.

## Model Group 3 Overview

## Selected Questionnaire Dwelling Unit-Level Poststratification

Out of 339 proposed effects, 338 were kept in the model. Two-factor effects were modified for State by Race, combining American Indian or Alaska Native and Asian for Delaware.

## Respondent Questionnaire Dwelling Unit-Level Nonresponse

All 339 proposed effects were kept in the model.

## Respondent Questionnaire Dwelling Unit-Level Poststratification

Out of 339 proposed effects, 338 were kept in the model. Two-factor effects were modified for State by Race, combining American Indian or Alaska Native and Asian for Delaware.

Exhibit C.3.1 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (sel.qdu.ps) Model Group 3: South

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 76 | 76 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| State (Count) | 17 | 16 | 16 | All levels present. |
| State (Binary) | 17 | 16 | 16 | All levels present. |
| Quarter (Count) | 4 | 3 | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 | 3 | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 | 4 | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 213 | 212 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ | 4 | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ | 1 | 1 | All levels present. |
| State $\times$ Age | $17 \times 5$ | 64 | 64 | All levels present. |
| State $\times$ Race | $17 \times 5$ | 64 | 63 | Coll. $(3,4) \&(3,5) ;$ conv. |
| State $\times$ Gender | $17 \times 2$ | 16 | 16 | All levels present. |
| State $\times$ Hispanicity | $17 \times 2$ | 16 | 16 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Hispanicity $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Hispanicity $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 50 | 50 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $3 \times 5 \times 3$ | 16 | 16 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $3 \times 2 \times 2$ | 2 | 2 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $3 \times 3 \times 2$ | 4 | 4 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Gender | $3 \times 3 \times 2$ | 4 | 4 | All levels present. |
| Total |  | 339 | 338 |  |

Exhibit C.3.2 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (res.qdu.nr) Model Group 3: South

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 76 | 76 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Population Density | 4 |  |  | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| State (Count) | 17 | 16 | 16 | All levels present. |
| State (Binary) | 17 | 16 | 16 | All levels present. |
| Quarter (Count) | 4 | 3 | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 |  | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 | 4 | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 213 | 213 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ | 4 | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ | 1 | 1 | All levels present. |
| State $\times$ Age | $17 \times 5$ | 64 | 64 | All levels present. |
| State $\times$ Race | $17 \times 5$ | 64 | 64 | All levels present. |
| State $\times$ Gender | $17 \times 2$ | 16 | 16 | All levels present. |
| State $\times$ Hispanicity | $17 \times 2$ | 16 | 16 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Hispanicity $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Hispanicity $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 50 | 50 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $3 \times 5 \times 3$ | 16 | 16 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $3 \times 2 \times 2$ | 2 | 2 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $3 \times 3 \times 2$ | 4 | 4 | All levels present. |
| State/Region $\times$ Race ( 3 levels) $\times$ Gender | $3 \times 3 \times 2$ | 4 | 4 | All levels present. |
| Total |  | 339 | 339 |  |

Exhibit C.3.3 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (res.qdu.ps) Model Group 3: South

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 76 | 76 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| State (Count) | 17 | 16 | 16 | All levels present. |
| State (Binary) | 17 | 16 | 16 | All levels present. |
| Quarter (Count) | 4 | 3 | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 | 3 | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 |  | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 213 | 212 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ | 4 | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ | 1 | 1 | All levels present. |
| State $\times$ Age | $17 \times 5$ | 64 | 64 | All levels present. |
| State $\times$ Race | $17 \times 5$ | 64 | 63 | Coll. $(3,3) \&(3,4) ;$ conv. |
| State $\times$ Gender | $17 \times 2$ | 16 | 16 | All levels present. |
| State $\times$ Hispanicity | $17 \times 2$ | 16 | 16 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Hispanicity $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Hispanicity $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 50 | 50 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $3 \times 5 \times 3$ | 16 | 16 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $3 \times 2 \times 2$ | 2 | 2 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $3 \times 3 \times 2$ | 4 | 4 | All levels present. |
| State/Region $\times$ Race (3 levels) $\times$ Gender | $3 \times 3 \times 2$ | 4 | 4 | All levels present. |
| Total |  | 339 | 338 |  |

# Appendix C.4: Model Group 4: West 

(Alaska, Arizona, California, Colorado, Idaho, Hawaii, Montana, Nevada, New Mexico, Oregon, Utah, Washington, Wyoming)

Table C.4a 2012 QDU Weight GEM Modeling Summary (Model Group 4: West)

| Modeling Step ${ }^{1}$ | Extreme Weight Proportions |  |  | UWE ${ }^{2}$ | \# Covariates ${ }^{3}$ | Bounds ${ }^{4}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | \% Unweighted | \% Weighted | \% Outwinsor |  |  | Nominal | Realized |
| sel.qdu.ps | 2.09 | 3.52 | 0.99 | 3.1857 | 270 | (0.49, 2.50) | (0.49, 2.50) |
|  | 1.44 | 2.60 | 0.62 | 3.1282 | 268 | $(0.34,3.27)$ | (0.34, 3.27) |
|  |  |  |  |  |  | (0.90, 1.66) | (0.90, 1.66) |
| res.qdu.nr | 1.51 | 3.11 | 0.72 | 3.1508 | 270 | (1.00, 2.90) | (1.00, 2.90) |
|  | 1.10 | 3.36 | 0.66 | 3.5391 | 267 | (1.00, 5.00) | (1.00, 5.00) |
|  |  |  |  |  |  | (1.30, 1.34) | (1.30, 1.34) |
| res.qdu.ps | 1.10 | 3.36 | 0.66 | 3.5391 | 270 | (0.82, 2.70) | (0.84, 2.70) |
|  | 0.91 | 2.98 | 0.57 | 3.5469 | 268 | (0.71, 2.70) | $(0.75,1.71)$ |
|  |  |  |  |  |  | (0.90, 1.02) | (1.02, 1.02) |

GEM = generalized exponential model; QDU = questionnaire dwelling unit.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.
${ }^{2}$ Unequal weighting effect (UWE) defined as $1+[(n-1) / n]^{*} C V^{2}$, where $C V=$ coefficient of variation of weights.
${ }^{3}$ Number of proposed covariates on top line and number finalized after modeling.
${ }^{4}$ There are six sets of bounds for each modeling step. Nominal bounds are used in defining maximum/minimum values for the GEM adjustment factors. The realized bound is the actual adjustment produced by the modeling. The set of three bounds listed for each step correspond to the high extreme values, the nonextreme values, and the low extreme values.

Table C.4b 2012 Distribution of Weight Adjustment Factors and Weight Products (Model Group 4: West)

|  | SDU Weight | QDU Design Weight |  | sel.qdu.ps ${ }^{1}$ |  | res.qdu.nr ${ }^{1}$ |  | res.qdu.ps ${ }^{1}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1-10 | duwght11 | 1-11 | duwght12 | 1-12 | duwght13 | 1-13 | duwght14 | 1-14 |
| Minimum | 20 | 1.00 | 20 | 0.22 | 13 | 0.45 | 41 | 0.38 | 44 |
| 1\% | 87 | 1.00 | 96 | 0.61 | 92 | 1.00 | 98 | 0.87 | 101 |
| 5\% | 112 | 1.00 | 141 | 0.74 | 132 | 1.02 | 155 | 0.96 | 155 |
| 10\% | 142 | 1.00 | 182 | 0.79 | 178 | 1.05 | 208 | 0.98 | 209 |
| 25\% | 251 | 1.00 | 426 | 0.88 | 416 | 1.11 | 468 | 0.99 | 469 |
| Median | 707 | 1.16 | 1,103 | 0.98 | 1,083 | 1.20 | 1,228 | 1.00 | 1,222 |
| 75\% | 1,480 | 3.32 | 2,176 | 1.09 | 2,200 | 1.32 | 2,645 | 1.01 | 2,627 |
| 90\% | 1,952 | 6.05 | 5,347 | 1.21 | 5,195 | 1.47 | 6,245 | 1.02 | 6,246 |
| 95\% | 2,211 | 7.90 | 8,100 | 1.31 | 8,010 | 1.57 | 10,398 | 1.04 | 10,482 |
| 99\% | 2,994 | 12.50 | 15,867 | 1.62 | 15,947 | 1.92 | 21,382 | 1.18 | 21,060 |
| Maximum | 8,444 | 14.78 | 48,558 | 3.77 | 35,038 | 5.00 | 53,491 | 1.71 | 59,335 |
| $n$ | 12,676 | - | 12,676 | - | 12,676 | - | 10,227 | - | 10,227 |
| Mean | 918 | 2.43 | 2,109 | 1.00 | 2,080 | 1.24 | 2,578 | 1.00 | 2,578 |
| Max/Mean | 9 | - | 23 | - | 17 | - | 21 | - | 23 |

QDU = questionnaire dwelling unit; SDU = screener dwelling unit.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.

## Model Group 4 Overview

## Selected Questionnaire Dwelling Unit-Level Poststratification

Out of 270 proposed effects, 268 were kept in the model. All main effects were maintained in full. Two-factor effects were modified for Rent/Housing by Percent Black or African American, combining 50-100\% and $10-<50 \%$ for the first and the fourth quintiles.

## Respondent Questionnaire Dwelling Unit-Level Nonresponse

Out of 270 proposed effects, 267 were kept in the model. All main effects were maintained in full. Two-factor effects were modified for Rent/Housing by Percent Black or African American, combining $50-100 \%$ and $10-<50 \%$ for the first and the fourth quintiles. Threefactor effects were modified for State by Race by Hispanicity, combining Black or African American with Other for California.

## Respondent Questionnaire Dwelling Unit-Level Poststratification

This step used the same set of effects as the selected questionnaire dwelling unit-level poststratification.

Exhibit C.4.1 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (sel.qdu.ps) Model Group 4: West

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 68 | 68 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| State (Count) | 13 | 12 | 12 | All levels present. |
| State (Binary) | 13 | 12 | 12 | All levels present. |
| Quarter (Count) | 4 | 3 | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 | 3 | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 | 4 | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 173 | 171 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ | 4 | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ | 1 | 1 | All levels present. |
| State $\times$ Age | $13 \times 5$ | 48 | 48 | All levels present. |
| State $\times$ Race | $13 \times 5$ | 48 | 48 | All levels present. |
| State $\times$ Gender | $13 \times 2$ | 12 | 12 | All levels present. |
| State $\times$ Hispanicity | $13 \times 2$ | 12 | 12 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 6 | Coll. $(1,1) \&(2,1)$; zero. Coll. $(1,4) \&(2,4) ;$ sing. |
| \% Hispanicity $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Hispanicity $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 29 | 29 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $2 \times 5 \times 2$ | 4 | 4 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $2 \times 5 \times 2$ | 4 | 4 | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $2 \times 5 \times 3$ | 8 | 8 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $2 \times 2 \times 2$ | 1 | 1 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $2 \times 3 \times 2$ | 2 | 2 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Gender | $2 \times 3 \times 2$ | 2 | 2 | All levels present. |
| Total |  | 270 | 268 |  |

Exhibit C.4.2 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (res.qdu.nr) Model Group 4: West

| Variables | Levels | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 68 | 68 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Household Type | 7 | 6 | 6 | All levels present. |
| Household Size | 1 | 1 | 1 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| \% Black or African American | 3 | 2 | 2 | All levels present. |
| \% Hispanic or Latino | 35 |  | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| State (Count) | 13 | 12 | 12 | All levels present. |
| State (Binary) | 13 | 12 | 12 | All levels present. |
| Quarter (Count) | 4 | 3 | 3 | All levels present. |
| Quarter (Binary) | 4 | 3 | 3 | All levels present. |
| Age Group | 5 | 4 | 4 | All levels present. |
| Race | 5 | 4 | 4 | All levels present. |
| Hispanicity | 2 | 1 | 1 | All levels present. |
| Gender | 2 | 1 | 1 | All levels present. |
| Two-Factor Effects |  | 173 | 171 |  |
| Age $\times$ Race (3 Levels) | $5 \times 3$ | 8 | 8 | All levels present. |
| Age $\times$ Hispanicity | $5 \times 2$ | 4 | 4 | All levels present. |
| Age $\times$ Gender | $5 \times 2$ | 4 | 4 | All levels present. |
| Race (3 Levels) $\times$ Hispanicity | $3 \times 2$ | 2 | 2 | All levels present. |
| Race (3 Levels) $\times$ Gender | $3 \times 2$ | 2 | 2 | All levels present. |
| Hispanicity $\times$ Gender | $2 \times 2$ |  | 1 | All levels present. |
| State $\times$ Age | $13 \times 5$ | 48 | 48 | All levels present. |
| State $\times$ Race | $13 \times 5$ | 48 | 48 | All levels present. |
| State $\times$ Gender | $13 \times 2$ | 12 | 12 | All levels present. |
| State $\times$ Hispanicity | $13 \times 2$ | 12 | 12 | All levels present. |
| \% Black or African American $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Black or African American $\times$ Rent/Housing | $3 \times 5$ | 8 | 6 | Coll. $(1,1) \&(2,1)$; zero. Coll. $(1,4) \&(2,4) ;$ sing. |
| \% Hispanicity $\times$ \% Owner-Occupied | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Hispanicity $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ Rent/Housing | $3 \times 5$ | 8 | 8 | All levels present. |
| Three-Factor Effects |  | 29 | 28 |  |
| Race (3 Levels) $\times$ Age $\times$ Gender | $3 \times 5 \times 2$ | 8 | 8 | All levels present. |
| State/Region $\times$ Age $\times$ Gender | $2 \times 5 \times 2$ | 4 | 4 | All levels present. |
| State/Region $\times$ Age $\times$ Hispanicity | $2 \times 5 \times 2$ | 4 | 4 | All levels present. |
| State/Region $\times$ Age $\times$ Race (3 Levels) | $2 \times 5 \times 3$ | 8 | 8 | All levels present. |
| State/Region $\times$ Hispanicity $\times$ Gender | $2 \times 2 \times 2$ | 1 | 1 | All levels present. |
| State/Region $\times$ Race (3 Levels) $\times$ Hispanicity | $2 \times 3 \times 2$ | 2 | 1 | Coll. (3,2,1) \& (3,3,1); conv. |
| State $/$ Region $\times$ Race (3 Levels) $\times$ Gender | $2 \times 3 \times 2$ | 2 | 2 | All levels present. |
| Total |  | 270 | 267 |  |

## Exhibit C.4.3 Covariates for 2012 NSDUH Questionnaire Dwelling Unit Weights (res.qdu.ps) Model Group 4: West

This step used the same set of covariates as the selected questionnaire dwelling unit poststratification.

# Appendix D: Evaluation of Calibration Weights: Questionnaire Dwelling Unit-Level Response Rates 

Table D. 12012 NSDUH QDU-Level Response Rates

| Domain | Selected QDU | Respondent QDU | \% Interview Response Rate ${ }^{1}$ |
| :---: | :---: | :---: | :---: |
| Total | 60,621 | 48,850 | 75.87 |
| Census Region |  |  |  |
| Northeast | 12,616 | 9,917 | 73.10 |
| South | 18,345 | 15,019 | 76.39 |
| Midwest | 16,984 | 13,687 | 77.25 |
| West | 12,676 | 10,227 | 75.87 |
| Quarter |  |  |  |
| Quarter 1 | 14,387 | 11,628 | 75.80 |
| Quarter 2 | 16,164 | 12,975 | 76.14 |
| Quarter 3 | 15,762 | 12,743 | 76.00 |
| Quarter 4 | 14,308 | 11,504 | 75.55 |
| Household Type |  |  |  |
| 12-17, 18-25, $26+$ | 5,811 | 5,004 | 85.32 |
| 12-17, 18-25 | 88 | 72 | 78.98 |
| 12-17, 26+ | 17,350 | 14,625 | 84.38 |
| 18-25, 26+ | 12,807 | 10,290 | 79.79 |
| 12-17 | 29 | 24 | 81.94 |
| 18-25 | 6,367 | 5,447 | 85.41 |
| 26+ | 18,169 | 13,388 | 72.47 |
| Race/Ethnicity of Householder |  |  |  |
| Hispanic or Latino White | 7,920 | 6,574 | 78.41 |
| Hispanic or Latino Black or African American | 187 | 165 | 78.73 |
| Hispanic or Latino Other | 415 | 355 | 83.80 |
| Non-Hispanic or Latino White | 40,460 | 32,098 | 74.78 |
| Non-Hispanic or Latino Black or African American | 7,135 | 6,092 | 81.01 |
| Non-Hispanic or Latino Other | 4,504 | 3,566 | 72.15 |
| \% Hispanic or Latino in Segment |  |  |  |
| 50-100\% | 4,137 | 3,397 | 78.21 |
| 10-<50\% | 11,129 | 9,003 | 75.80 |
| <10\% | 45,355 | 36,450 | 75.66 |
| \% Black or African American in Segment |  |  |  |
| 50-100\% | 4,381 | 3,749 | 80.55 |
| 10-<50\% | 9,185 | 7,524 | 76.96 |
| <10\% | 47,055 | 37,577 | 75.17 |
| \% Owner-Occupied DUs in Segment |  |  |  |
| 50-100\% | 45,901 | 36,704 | 75.09 |
| 10-<50\% | 11,429 | 9,421 | 77.84 |
| <10\% | 3,291 | 2,725 | 79.98 |
| Combined Median Rent/Housing Value |  |  |  |
| $1^{\text {st }}$ Quintile | 9,920 | 8,288 | 78.90 |
| $2^{\text {nd }}$ Quintile | 12,960 | 10,661 | 77.97 |
| $3{ }^{\text {rd }}$ Quintile | 13,713 | 11,009 | 75.45 |
| $4^{\text {th }}$ Quintile | 13,171 | 10,428 | 74.99 |
| $5^{\text {th }}$ Quintile | 10,857 | 8,464 | 73.22 |
| Population Density |  |  |  |
| Large MSA | 25,508 | 20,223 | 74.24 |
| Medium to Small MSA | 30,307 | 24,705 | 77.53 |
| Non-MSA, Urban | 1,249 | 1,028 | 80.00 |
| Non-MSA, Rural | 3,557 | 2,894 | 77.06 |
| Group Quarters |  |  |  |
| Group | 774 | 729 | 93.20 |
| Non-Group | 59,847 | 48,121 | 75.76 |
| Household Size |  |  |  |
| One | 7,063 | 5,567 | 75.11 |
| Two | 21,700 | 16,974 | 73.66 |
| Three | 17,330 | 14,071 | 78.45 |
| Four or More | 14,528 | 12,238 | 82.73 |

[^16]
# Appendix E: Evaluation of Calibration Weights: Questionnaire Dwelling Unit-Level Proportions of Extreme Values and Outwinsors 

Table E. 12012 NSDUH Selected QDU-Level Proportions of Extreme Values and Outwinsors

| Domain | $\boldsymbol{n}$ | SDU-Level Weights ${ }^{1}$ (SDUWT: WT1*...*WT9) |  |  | Before sel.qdu.ps ${ }^{1}$(SDUWT*DUWT10) |  |  | After sel.qdu.ps ${ }^{1}$(SDUWT*DUWT10*DUWT11) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | \% Unweighted | $\begin{gathered} \% \\ \text { Weighted }^{2} \\ \hline \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | $\begin{gathered} \% \\ \text { Unweighted } \end{gathered}$ | $\begin{gathered} \% \\ \text { Weighted }^{2} \\ \hline \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | $\begin{gathered} \% \\ \text { Unweighted } \end{gathered}$ | $\begin{gathered} \% \\ \text { Weighted }^{2} \\ \hline \end{gathered}$ | $\stackrel{\%}{\text { Outwinsor }}$ |
| Total | 60,621 | 2.62 | 5.38 | 1.44 | 1.88 | 2.60 | 0.66 | 1.41 | 2.17 | 0.43 |
| Census Region |  |  |  |  |  |  |  |  |  |  |
| Northeast | 12,616 | 2.85 | 6.54 | 2.01 | 2.33 | 3.79 | 1.09 | 2.15 | 4.48 | 1.06 |
| South | 18,345 | 2.15 | 4.78 | 1.13 | 1.50 | 1.91 | 0.38 | 1.17 | 1.51 | 0.17 |
| Midwest | 16,984 | 2.47 | 4.31 | 0.91 | 1.78 | 1.86 | 0.42 | 1.10 | 0.98 | 0.16 |
| West | 12,676 | 3.26 | 6.46 | 1.98 | 2.09 | 3.52 | 0.99 | 1.44 | 2.60 | 0.62 |
| Quarter |  |  |  |  |  |  |  |  |  |  |
| Quarter 1 | 14,387 | 2.91 | 5.41 | 1.48 | 1.97 | 2.64 | 0.70 | 1.46 | 2.27 | 0.50 |
| Quarter 2 | 16,164 | 2.36 | 5.15 | 1.41 | 1.68 | 2.28 | 0.59 | 1.39 | 2.35 | 0.49 |
| Quarter 3 | 15,762 | 2.33 | 4.78 | 1.12 | 1.76 | 2.49 | 0.54 | 1.21 | 1.95 | 0.33 |
| Quarter 4 | 14,308 | 2.93 | 6.19 | 1.75 | 2.12 | 3.00 | 0.81 | 1.60 | 2.10 | 0.39 |
| Household Type |  |  |  |  |  |  |  |  |  |  |
| 12-17, 18-25, 26+ | 5,811 | 2.63 | 5.43 | 1.52 | 2.63 | 5.43 | 1.52 | 2.27 | 5.20 | 1.09 |
| 12-17, 18-25 | 88 | 4.55 | 7.61 | 1.34 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 12-17, 26+ | 17,350 | 2.66 | 5.20 | 1.42 | 2.64 | 5.18 | 1.42 | 1.98 | 4.21 | 0.89 |
| 18-25, 26+ | 12,807 | 2.96 | 6.18 | 1.63 | 2.46 | 5.57 | 1.47 | 1.56 | 4.39 | 1.12 |
| 12-17 | 29 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 18-25 | 6,367 | 2.42 | 4.70 | 0.96 | 2.18 | 4.44 | 0.91 | 1.66 | 3.42 | 0.61 |
| 26+ | 18,169 | 2.39 | 5.12 | 1.43 | 0.40 | 1.35 | 0.32 | 0.40 | 1.16 | 0.18 |
| Race/Ethnicity of Householder |  |  |  |  |  |  |  |  |  |  |
| Hispanic or Latino White | 7,920 | 2.45 | 5.03 | 1.41 | 1.93 | 2.74 | 0.80 | 1.40 | 2.28 | 0.41 |
| Hispanic or Latino Black or African American | 187 | 57.75 | 82.43 | 35.31 | 48.13 | 60.13 | 21.66 | 40.64 | 61.93 | 18.69 |
| Hispanic or Latino Other | 415 | 34.70 | 62.52 | 23.51 | 27.95 | 44.43 | 11.95 | 17.83 | 31.05 | 6.77 |
| Non-Hispanic or Latino White | 40,460 | 1.26 | 2.31 | 0.43 | 0.80 | 1.01 | 0.18 | 0.49 | 0.68 | 0.08 |
| Non-Hispanic or Latino Black or African American | 7,135 | 3.90 | 7.48 | 1.82 | 3.10 | 4.13 | 1.04 | 3.27 | 4.20 | 0.80 |
| Non-Hispanic or Latino Other | 4,504 | 7.88 | 11.84 | 2.62 | 5.20 | 5.19 | 1.23 | 3.60 | 4.59 | 0.96 |

Table E. 12012 NSDUH Selected QDU-Level Proportions of Extreme Values and Outwinsors (continued)

| Domain | $n$ | SDU-Level Weights ${ }^{1}$ (SDUWT: WT1*...*WT9) |  |  | $\begin{gathered} \text { Before sel.qdu.ps }^{1} \\ \text { (SDUWT*DUWT10) } \end{gathered}$ |  |  | After sel.qdu.ps ${ }^{1}$(SDUWT*DUWT10*DUWT11) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\%$ Unweighted | $\begin{gathered} \% \\ \text { Weighted }^{2} \\ \hline \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | $\%$ Unweighted | $\begin{gathered} \% \\ \text { Weighted }^{2} \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | \% <br> Unweighted | $\begin{gathered} \% \\ \text { Weighted }^{2} \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 4,137 | 2.13 | 5.12 | 1.88 | 1.84 | 3.56 | 1.13 | 1.55 | 3.46 | 1.00 |
| 10-<50\% | 11,129 | 3.80 | 8.36 | 2.68 | 2.95 | 4.25 | 1.26 | 2.40 | 3.64 | 0.81 |
| <10\% | 45,355 | 2.37 | 4.37 | 0.94 | 1.62 | 2.01 | 0.43 | 1.15 | 1.59 | 0.25 |
| \% Black or African American in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 4,381 | 3.54 | 8.02 | 2.26 | 3.15 | 4.93 | 1.21 | 3.99 | 6.13 | 1.37 |
| 10- $50 \%$ | 9,185 | 2.77 | 6.22 | 1.97 | 2.19 | 3.75 | 1.15 | 2.01 | 3.58 | 0.79 |
| <10\% | 47,055 | 2.50 | 4.89 | 1.22 | 1.70 | 2.13 | 0.50 | 1.05 | 1.47 | 0.26 |
| \% Owner-Occupied DUs in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 45,901 | 2.27 | 4.70 | 1.25 | 1.59 | 2.10 | 0.52 | 1.06 | 1.55 | 0.27 |
| 10-<50\% | 11,429 | 3.73 | 7.01 | 1.84 | 2.80 | 3.95 | 1.00 | 1.95 | 3.05 | 0.74 |
| <10\% | 3,291 | 3.65 | 8.34 | 2.33 | 2.61 | 4.85 | 1.40 | 4.35 | 7.78 | 1.53 |
| Combined Median <br> Rent/Housing Value |  |  |  |  |  |  |  |  |  |  |
| $1^{\text {st }}$ Quintile | 9,920 | 2.46 | 4.70 | 1.26 | 1.88 | 2.29 | 0.60 | 1.76 | 2.24 | 0.42 |
| $2^{\text {nd }}$ Quintile | 12,960 | 2.35 | 5.03 | 1.55 | 1.87 | 2.70 | 0.66 | 1.40 | 1.94 | 0.47 |
| $3{ }^{\text {rd }}$ Quintile | 13,713 | 2.68 | 5.38 | 1.47 | 1.90 | 2.65 | 0.65 | 1.24 | 2.07 | 0.39 |
| $4^{\text {th }}$ Quintile | 13,171 | 2.73 | 5.46 | 1.50 | 1.90 | 2.55 | 0.70 | 1.31 | 2.30 | 0.50 |
| $5^{\text {th }}$ Quintile | 10,857 | 2.88 | 6.09 | 1.33 | 1.83 | 2.72 | 0.66 | 1.43 | 2.28 | 0.35 |
| Population Density |  |  |  |  |  |  |  |  |  |  |
| Large MSA ${ }^{1}$ | 25,508 | 2.93 | 6.67 | 1.86 | 2.16 | 3.43 | 0.90 | 1.81 | 2.91 | 0.62 |
| Medium to Small MSA ${ }^{1}$ | 30,307 | 2.47 | 4.22 | 1.06 | 1.75 | 1.84 | 0.46 | 1.17 | 1.53 | 0.24 |
| Non-MSA, ${ }^{1}$ Urban | 1,249 | 1.28 | 1.23 | 0.16 | 0.96 | 0.41 | 0.06 | 0.88 | 0.54 | 0.09 |
| Non-MSA, ${ }^{1}$ Rural | 3,557 | 2.08 | 2.19 | 0.38 | 1.27 | 1.37 | 0.12 | 0.79 | 0.54 | 0.12 |
| Group Quarters |  |  |  |  |  |  |  |  |  |  |
| Group | 774 | 4.01 | 7.46 | 0.51 | 2.71 | 3.60 | 0.36 | 2.97 | 6.11 | 1.34 |
| Non-Group | 59,847 | 2.60 | 5.35 | 1.45 | 1.87 | 2.60 | 0.66 | 1.39 | 2.14 | 0.42 |
| Household Size |  |  |  |  |  |  |  |  |  |  |
| One | 7,063 | 2.10 | 5.10 | 1.33 | 0.93 | 1.17 | 0.28 | 0.72 | 1.02 | 0.14 |
| Two | 21,700 | 2.35 | 5.09 | 1.38 | 1.35 | 2.17 | 0.51 | 1.01 | 1.70 | 0.32 |
| Three | 17,330 | 2.80 | 5.40 | 1.50 | 2.44 | 4.60 | 1.28 | 1.51 | 3.49 | 0.80 |
| Four or More | 14,528 | 3.06 | 5.85 | 1.47 | 2.46 | 4.72 | 1.21 | 2.21 | 4.72 | 1.00 |

[^17]Table E. 22012 NSDUH Respondent QDU-Level Proportions of Extreme Values and Outwinsors

| Domain | $n$ | Before res.qdu.nr ${ }^{1}$(SDUWT*DUWT10*DUWT11) |  |  | After res.qdu.nr(SDUWT*DUWT10*...*DUWT12) |  |  | Final Weight: After res.qdu.ps ${ }^{1}$ (SDUWT*DUWT10*...*DUWT13) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | \% Unweighted | $\begin{gathered} \hline \% \\ \text { Weighted }^{2} \\ \hline \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | \% Unweighted | $\begin{gathered} \hline \% \\ \text { Weighted }^{2} \\ \hline \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | $\%$ <br> Unweighted | $\begin{gathered} \hline \% \\ \text { Weighted }^{2} \\ \hline \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ |
| Total | 48,850 | 1.52 | 2.42 | 0.50 | 1.17 | 2.32 | 0.42 | 1.17 | 2.20 | 0.31 |
| Census Region |  |  |  |  |  |  |  |  |  |  |
| Northeast | 9,917 | 2.33 | 4.92 | 1.27 | 1.92 | 4.56 | 0.86 | 2.06 | 4.21 | 0.68 |
| South | 15,019 | 1.24 | 1.57 | 0.20 | 1.10 | 1.49 | 0.24 | 1.09 | 1.51 | 0.12 |
| Midwest | 13,687 | 1.25 | 1.26 | 0.20 | 0.76 | 0.90 | 0.12 | 0.81 | 0.97 | 0.05 |
| West | 10,227 | 1.51 | 3.11 | 0.72 | 1.10 | 3.36 | 0.66 | 0.91 | 2.98 | 0.57 |
| Quarter |  |  |  |  |  |  |  |  |  |  |
| Quarter 1 | 11,628 | 1.49 | 2.55 | 0.56 | 1.18 | 2.58 | 0.43 | 1.22 | 2.49 | 0.32 |
| Quarter 2 | 12,975 | 1.53 | 2.81 | 0.58 | 1.21 | 2.43 | 0.51 | 1.16 | 2.24 | 0.40 |
| Quarter 3 | 12,743 | 1.37 | 2.08 | 0.38 | 0.94 | 2.01 | 0.33 | 0.97 | 1.91 | 0.22 |
| Quarter 4 | 11,504 | 1.70 | 2.24 | 0.47 | 1.36 | 2.28 | 0.40 | 1.36 | 2.16 | 0.29 |
| Household Type |  |  |  |  |  |  |  |  |  |  |
| 12-17, 18-25, 26+ | 5,004 | 2.16 | 5.05 | 1.03 | 1.60 | 4.54 | 1.05 | 1.58 | 4.10 | 0.75 |
| 12-17, 18-25 | 72 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 12-17, 26+ | 14,625 | 2.18 | 4.63 | 0.99 | 1.42 | 3.71 | 0.86 | 1.41 | 3.42 | 0.67 |
| 18-25, $26+$ | 10,290 | 1.73 | 4.53 | 1.18 | 1.62 | 4.91 | 1.17 | 1.66 | 4.67 | 0.91 |
| 12-17 | 24 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 18-25 | 5,447 | 1.43 | 3.37 | 0.69 | 1.19 | 2.40 | 0.34 | 1.10 | 2.14 | 0.30 |
| 26+ | 13,388 | 0.44 | 1.31 | 0.22 | 0.39 | 1.50 | 0.18 | 0.41 | 1.46 | 0.11 |
| Race/Ethnicity of Householder |  |  |  |  |  |  |  |  |  |  |
| Hispanic or Latino White | 6,574 | 1.55 | 2.76 | 0.40 | 1.11 | 2.30 | 0.43 | 1.22 | 2.36 | 0.34 |
| Hispanic or Latino Black or African American | 165 | 40.61 | 65.50 | 19.97 | 38.18 | 57.52 | 16.88 | 40.00 | 55.69 | 15.87 |
| Hispanic or Latino Other | 355 | 16.90 | 30.32 | 7.01 | 10.14 | 22.63 | 4.43 | 9.30 | 16.59 | 2.82 |
| Non-Hispanic or Latino White | 32,098 | 0.53 | 0.66 | 0.09 | 0.50 | 0.81 | 0.07 | 0.40 | 0.68 | 0.02 |
| Non-Hispanic or Latino Black or African American | 6,092 | 3.33 | 4.69 | 0.95 | 1.63 | 3.26 | 0.63 | 1.89 | 3.32 | 0.48 |
| Non-Hispanic or Latino Other | 3,566 | 3.95 | 5.81 | 1.33 | 3.93 | 8.90 | 1.64 | 4.21 | 9.25 | 1.14 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 3,397 | 1.65 | 3.36 | 0.97 | 1.32 | 2.70 | 0.77 | 1.47 | 2.70 | 0.65 |
| 10-<50\% | 9,003 | 2.63 | 4.27 | 0.95 | 1.79 | 3.98 | 0.82 | 1.88 | 3.77 | 0.61 |
| <10\% | 36,450 | 1.23 | 1.76 | 0.31 | 1.00 | 1.78 | 0.26 | 0.97 | 1.68 | 0.18 |

Table E. 2012 NSDUH Respondent QDU-Level Proportions of Extreme Values and Outwinsors (continued)

| Domain | $n$ | Before res.qdu.nr ${ }^{1}$(SDUWT*DUWT10*DUWT11) |  |  | After res.qdu.nr(SDUWT*DUWT10*...*DUWT12) |  |  | Final Weight: After res.qdu.ps ${ }^{1}$ (SDUWT*DUWT10*...*DUWT13) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | \% <br> Unweighted | $\begin{gathered} \% \\ \text { Weighted }^{2} \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | $\begin{gathered} \% \\ \text { Unweighted } \\ \hline \end{gathered}$ | $\begin{gathered} \% \\ \text { Weighted }^{2} \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | \% Unweighted | $\begin{gathered} \hline \% \\ \text { Weighted }^{2} \\ \hline \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ |
| \% Black or African American in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 3,749 | 4.27 | 7.07 | 1.64 | 2.37 | 6.14 | 1.03 | 2.69 | 5.77 | 0.77 |
| 10-<50\% | 7,524 | 2.21 | 3.91 | 0.87 | 1.66 | 3.50 | 0.80 | 1.70 | 3.25 | 0.65 |
| <10\% | 37,577 | 1.11 | 1.60 | 0.30 | 0.95 | 1.69 | 0.27 | 0.91 | 1.62 | 0.18 |
| \% Owner-Occupied DUs in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 36,704 | 1.18 | 1.83 | 0.35 | 1.01 | 2.00 | 0.34 | 0.98 | 1.91 | 0.24 |
| 10-<50\% | 9,421 | 1.98 | 3.19 | 0.77 | 1.33 | 2.82 | 0.56 | 1.54 | 2.79 | 0.46 |
| <10\% | 2,725 | 4.48 | 7.59 | 1.61 | 2.79 | 5.11 | 0.98 | 2.46 | 4.33 | 0.74 |
| Combined Median Rent/Housing Value |  |  |  |  |  |  |  |  |  |  |
| $1^{\text {st }}$ Quintile | 8,288 | 1.91 | 2.67 | 0.44 | 0.98 | 1.47 | 0.32 | 0.92 | 1.37 | 0.25 |
| $2^{\text {nd }}$ Quintile | 10,661 | 1.52 | 2.19 | 0.52 | 0.92 | 1.77 | 0.43 | 1.01 | 1.76 | 0.42 |
| $3^{\text {rd }}$ Quintile | 11,009 | 1.38 | 2.24 | 0.48 | 1.15 | 2.40 | 0.38 | 1.13 | 2.09 | 0.20 |
| $4^{\text {th }}$ Quintile | 10,428 | 1.44 | 2.74 | 0.61 | 1.18 | 2.74 | 0.47 | 1.23 | 2.70 | 0.38 |
| $5{ }^{\text {th }}$ Quintile | 8,464 | 1.42 | 2.30 | 0.41 | 1.68 | 2.90 | 0.45 | 1.59 | 2.76 | 0.25 |
| Population Density |  |  |  |  |  |  |  |  |  |  |
| Large MSA ${ }^{1}$ | 20,223 | 1.96 | 3.19 | 0.72 | 1.71 | 3.52 | 0.63 | 1.67 | 3.26 | 0.46 |
| Medium to Small MSA ${ }^{1}$ | 24,705 | 1.27 | 1.83 | 0.31 | 0.81 | 1.03 | 0.20 | 0.86 | 1.08 | 0.15 |
| Non-MSA, ${ }^{1}$ Urban | 1,028 | 0.78 | 0.53 | 0.10 | 0.58 | 0.33 | 0.05 | 0.49 | 0.32 | 0.05 |
| Non-MSA, ${ }^{1}$ Rural | 2,894 | 0.83 | 0.60 | 0.16 | 0.66 | 1.80 | 0.20 | 0.59 | 1.64 | 0.07 |
| Group Quarters |  |  |  |  |  |  |  |  |  |  |
| Group | 729 | 2.88 | 6.48 | 1.42 | 1.51 | 2.75 | 0.15 | 1.37 | 3.82 | 1.11 |
| Non-Group | 48,121 | 1.50 | 2.39 | 0.49 | 1.16 | 2.32 | 0.42 | 1.17 | 2.19 | 0.30 |
| Household Size |  |  |  |  |  |  |  |  |  |  |
| One | 5,567 | 0.70 | 1.13 | 0.17 | 0.79 | 1.43 | 0.15 | 0.65 | 1.30 | 0.11 |
| Two | 16,974 | 1.10 | 2.01 | 0.42 | 0.95 | 2.00 | 0.31 | 0.98 | 1.89 | 0.21 |
| Three | 14,071 | 1.69 | 3.67 | 0.78 | 1.17 | 2.95 | 0.72 | 1.28 | 3.06 | 0.65 |
| Four or More | 12,238 | 2.27 | 4.85 | 1.07 | 1.64 | 4.72 | 1.04 | 1.54 | 4.26 | 0.65 |

${ }^{1} \mathrm{DU}=$ dwelling unit, MSA $=$ metropolitan statistical area, $\mathrm{NR}=$ nonresponse adjustment, $\mathrm{PS}=$ poststratification adjustment, QDU = questionnaire dwelling unit, Res $=$ Respondent, SDU $=$ screener dwelling unit.
${ }^{2}$ Weighted extreme value proportion: $100^{*} \sum_{k} w_{e k} / \sum_{k} w_{k}$, where $w_{e k}$ denotes the weight for extreme values, and $w_{k}$ denotes the weight for both extreme values and nonextreme values.
${ }^{3}$ Outwinsor weight proportion: $100^{*} \sum_{k}\left(w_{e k}-b_{k}\right) / \sum_{k} w_{k}$, where $b_{k}$ denotes the winsorized weight.

# Appendix F: Evaluation of Calibration Weights: Questionnaire Dwelling Unit-Level Slippage Rates 

Table F. 12012 NSDUH QDU-Level Slippage Rates

| Domain | $n$ | Initial Total (I) ${ }^{1}$ | Final Total (F) ${ }^{\mathbf{2}}$ | Control from SDU Weights (C) | $(I-C) / C \%$ | $(F-C) / C \%$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Total | 48,850 | 119,147,983 | 119,147,983 | 119,147,983 | 0.00 | -0.00 |
| Census Region |  |  |  |  |  |  |
| Northeast | 9,917 | 21,504,846 | 21,504,846 | 21,504,846 | 0.00 | -0.00 |
| South | 15,019 | 44,826,838 | 44,826,838 | 44,826,838 | 0.00 | -0.00 |
| Midwest | 13,687 | 26,453,818 | 26,453,818 | 26,453,818 | 0.00 | -0.00 |
| West | 10,227 | 26,362,481 | 26,362,481 | 26,362,481 | 0.00 | -0.00 |
| Quarter |  |  |  |  |  |  |
| Quarter 1 | 11,628 | 29,829,432 | 29,829,432 | 29,829,432 | 0.00 | -0.00 |
| Quarter 2 | 12,975 | 29,634,723 | 29,634,723 | 29,634,723 | 0.00 | -0.00 |
| Quarter 3 | 12,743 | 29,744,017 | 29,744,017 | 29,744,017 | 0.00 | -0.00 |
| Quarter 4 | 11,504 | 29,939,811 | 29,939,811 | 29,939,811 | 0.00 | -0.00 |
| Household Type |  |  |  |  |  |  |
| 12-17, 18-25, 26+ | 5,004 | 5,062,357 | 5,062,357 | 5,062,357 | 0.00 | 0.00 |
| 12-17, 18-25 | 72 | 66,238 | 66,238 | 66,238 | 0.00 | 0.00 |
| 12-17, 26+ | 14,625 | 13,691,239 | 13,691,239 | 13,691,239 | 0.00 | -0.00 |
| 18-25, 26+ | 10,290 | 13,956,293 | 13,956,293 | 13,956,293 | 0.00 | -0.00 |
| 12-17 | 24 | 22,258 | 22,258 | 22,258 | 0.00 | 0.00 |
| 18-25 | 5,447 | 5,779,323 | 5,779,323 | 5,779,323 | 0.00 | -0.00 |
| 26+ | 13,388 | 80,570,274 | 80,570,274 | 80,570,274 | 0.00 | -0.00 |
| Race/Ethnicity of Householder |  |  |  |  |  |  |
| Hispanic or Latino White | 6,574 | 13,508,511 | 13,508,511 | 13,508,511 | -0.00 | -0.00 |
| Hispanic or Latino Black or African American | 165 | 725,656 | 725,656 | 725,656 | -0.00 | -0.00 |
| Hispanic or Latino Other | 355 | 1,005,979 | 1,005,979 | 1,005,979 | 0.00 | -0.00 |
| Non-Hispanic or Latino White | 32,098 | 81,939,462 | 81,939,462 | 81,939,462 | 0.00 | -0.00 |
| Non-Hispanic or Latino Black or African American | 6,092 | 14,332,726 | 14,332,726 | 14,332,726 | 0.00 | -0.00 |
| Non-Hispanic or Latino Other | 3,566 | 7,635,648 | 7,635,648 | 7,635,648 | 0.00 | -0.00 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |
| 50-100\% | 3,397 | 8,634,013 | 8,634,013 | 8,634,013 | 0.00 | -0.00 |
| 10-<50\% | 9,003 | 25,715,926 | 25,715,926 | 25,715,926 | 0.00 | -0.00 |
| <10\% | 36,450 | 84,798,044 | 84,798,044 | 84,798,044 | 0.00 | -0.00 |
| \% Black or African American in Segment |  |  |  |  |  |  |
| 50-100\% | 3,749 | 8,810,957 | 8,810,957 | 8,810,957 | 0.00 | -0.00 |
| 10-<50\% | 7,524 | 20,110,346 | 20,110,346 | 20,110,346 | -0.00 | -0.00 |
| <10\% | 37,577 | 90,226,680 | 90,226,679 | 90,226,679 | 0.00 | -0.00 |
| \% Owner-Occupied DUs <br> in Segment |  |  |  |  |  |  |
| 50-100\% | 36,704 | 90,409,242 | 90,409,242 | 90,409,242 | 0.00 | -0.00 |
| $10-<50 \%$ | 9,421 | 22,264,244 | 22,264,244 | 22,264,244 | 0.00 | -0.00 |
| $<10 \%$ | 2,725 | 6,474,497 | 6,474,497 | 6,474,497 | 0.00 | -0.00 |

Table F. 12012 NSDUH QDU-Level Slippage Rates (continued)

| Domain | $n$ | Initial Total ( $)^{1}$ | Final Total ( $F)^{\mathbf{2}}$ | Control from SDU Weights (C) | $(I-C) / C \%$ | ( $F-C$ )/C\% |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Combined Median Rent/Housing Value |  |  |  |  |  |  |
|  |  |  |  |  |  |  |
| $1^{\text {st }}$ Quintile | 8,288 | 17,328,372 | 17,328,372 | 17,328,372 | 0.00 | -0.00 |
| $2^{\text {nd }}$ Quintile | 10,661 | 23,902,320 | 23,902,320 | 23,902,320 | 0.00 | -0.00 |
| $3^{\text {rd }}$ Quintile | 11,009 | 24,983,991 | 24,983,991 | 24,983,991 | 0.00 | -0.00 |
| $4^{\text {th }}$ Quintile | 10,428 | 27,292,895 | 27,292,895 | 27,292,895 | 0.00 | -0.00 |
| $5^{\text {th }}$ Quintile | 8,464 | 25,640,405 | 25,640,405 | 25,640,405 | 0.00 | -0.00 |
| Population Density |  |  |  |  |  |  |
| Large MSA | 20,223 | 60,608,068 | 60,608,068 | 60,608,068 | 0.00 | -0.00 |
| Medium to Small MSA | 24,705 | 50,803,819 | 50,803,818 | 50,803,818 | 0.00 | -0.00 |
| Non-MSA, Urban | 1,028 | 2,013,232 | 2,013,232 | 2,013,232 | 0.00 | -0.00 |
| Non-MSA, Rural | 2,894 | 5,722,865 | 5,722,865 | 5,722,865 | 0.00 | -0.00 |
| Group Quarters |  |  |  |  |  |  |
| Group | 729 | 766,775 | 766,775 | 766,775 | 0.00 | -0.00 |
| Non-Group | 48,121 | 118,381,208 | 118,381,208 | 118,381,208 | 0.00 | -0.00 |
| Household Size |  |  |  |  |  |  |
| One | 5,567 | 31,354,646 | 31,358,657 | 31,327,228 | 0.09 | 0.10 |
| Two | 16,974 | 54,585,381 | 54,556,667 | 54,622,573 | -0.07 | -0.12 |
| Three | 14,071 | 19,065,921 | 19,111,984 | 19,137,226 | -0.37 | -0.13 |
| Four or More | 12,238 | 14,142,035 | 14,120,675 | 14,060,955 | 0.58 | 0.42 |

$\mathrm{DU}=$ dwelling unit, MSA = metropolitan statistical area, $\mathrm{QDU}=$ questionnaire dwelling unit, $\mathrm{SDU}=$ screener dwelling unit.
${ }^{1}$ WT1*...*WT9*DUWT10*...*DUWT12 (before QDU poststratification).
${ }^{2}$ WT1*...*WT9*DUWT10*...*DUWT13 (after QDU poststratification).

# Appendix G: Evaluation of Calibration Weights: Questionnaire Dwelling Unit-Level Weight Summary Statistics 

Table G. 12012 NSDUH Selected QDU-Level Weight Summary Statistics

| Domain | $n$ | SDU-Level Weights ${ }^{1}$ (SDUWT: WT1*...*WT9) |  |  |  |  |  | Before sel.qdu.ps ${ }^{1}$(SDUWT*DUWT10) |  |  |  |  |  | After sel.qdu.ps ${ }^{1}$(SDUWT*DUWT10*DUWT11) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{\mathbf{2}}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ |
| Total | 60,621 | 9 | 436 | 664 | 1,098 | 8,444 | 1.56 | 12 | 552 | 1,013 | 2,115 | 48,558 | 2.86 | 8 | 542 | 1,015 | 2,137 | 44,158 | 2.85 |
| Census Region |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Northeast | 12,616 | 18 | 305 | 603 | 834 | 7,128 | 1.57 | 18 | 503 | 856 | 1,954 | 30,862 | 2.91 | 8 | 483 | 850 | 1,931 | 29,264 | 2.91 |
| South | 18,345 | 9 | 641 | 923 | 1,313 | 8,236 | 1.39 | 12 | 817 | 1,284 | 2,761 | 45,296 | 2.54 | 10 | 813 | 1,300 | 2,787 | 44,158 | 2.54 |
| Midwest | 16,984 | 33 | 452 | 541 | 707 | 6,654 | 1.35 | 40 | 513 | 737 | 1,773 | 43,516 | 2.68 | 23 | 507 | 762 | 1,751 | 25,191 | 2.71 |
| West | 12,676 | 20 | 251 | 707 | 1,480 | 8,444 | 1.75 | 20 | 426 | 1,103 | 2,176 | 48,558 | 3.19 | 13 | 416 | 1,083 | 2,200 | 35,038 | 3.13 |
| Quarter |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Quarter 1 | 14,387 | 21 | 495 | 717 | 1,140 | 8,082 | 1.54 | 21 | 610 | 1,079 | 2,216 | 45,473 | 2.88 | 10 | 597 | 1,065 | 2,231 | 35,038 | 2.85 |
| Quarter 2 | 16,164 | 18 | 402 | 601 | 1,045 | 8,236 | 1.56 | 18 | 505 | 951 | 1,961 | 30,862 | 2.80 | 8 | 498 | 955 | 1,980 | 27,818 | 2.84 |
| Quarter 3 | 15,762 | 9 | 431 | 617 | 1,087 | 7,138 | 1.57 | 12 | 520 | 981 | 2,068 | 45,296 | 2.85 | 10 | 510 | 975 | 2,077 | 44,158 | 2.85 |
| Quarter 4 | 14,308 | 16 | 471 | 699 | 1,111 | 8,444 | 1.57 | 16 | 590 | 1,035 | 2,274 | 48,558 | 2.88 | 13 | 585 | 1,062 | 2,280 | 30,041 | 2.84 |
| Household Type |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 12-17, 18-25, $26+$ | 5,811 | 18 | 463 | 679 | 1,156 | 8,444 | 1.59 | 18 | 463 | 679 | 1,156 | 8,446 | 1.59 | 13 | 454 | 680 | 1,164 | 9,750 | 1.59 |
| 12-17, 18-25 | 88 | 42 | 357 | 592 | 958 | 2,459 | 1.53 | 42 | 357 | 592 | 958 | 2,460 | 1.53 | 27 | 350 | 559 | 1,099 | 2,529 | 1.61 |
| 12-17, 26+ | 17,350 | 16 | 412 | 626 | 1,059 | 8,351 | 1.55 | 16 | 414 | 628 | 1,065 | 8,379 | 1.56 | 10 | 400 | 631 | 1,080 | 8,263 | 1.56 |
| 18-25, $26+$ | 12,807 | 9 | 486 | 709 | 1,188 | 8,236 | 1.57 | 12 | 584 | 891 | 1,410 | 9,795 | 1.53 | 8 | 569 | 905 | 1,441 | 11,391 | 1.54 |
| 12-17 | 29 | 83 | 319 | 641 | 951 | 2,460 | 1.52 | 84 | 322 | 648 | 957 | 2,485 | 1.52 | 116 | 322 | 579 | 1,205 | 2,291 | 1.56 |
| 18-25 | 6,367 | 21 | 364 | 652 | 1,069 | 6,940 | 1.58 | 21 | 424 | 784 | 1,237 | 7,503 | 1.52 | 30 | 408 | 768 | 1,226 | 6,586 | 1.56 |
| $26+$ | 18,169 | 20 | 435 | 656 | 1,080 | 8,082 | 1.53 | 74 | 1,893 | 3,509 | 5,773 | 48,558 | 1.72 | 53 | 1,865 | 3,491 | 5,815 | 44,158 | 1.71 |
| Race/Ethnicity of Householder |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Hispanic or Latino White | 7,920 | 19 | 505 | 809 | 1,317 | 8,444 | 1.44 | 20 | 596 | 1,102 | 1,839 | 27,837 | 2.38 | 8 | 589 | 1,110 | 1,858 | 26,638 | 2.47 |
| Hispanic or Latino Black or African American | 187 | 18 | 807 | 1,559 | 2,869 | 8,351 | 1.79 | 18 | 947 | 2,014 | 4,659 | 37,254 | 2.79 | 20 | 896 | 1,913 | 4,342 | 29,264 | 2.76 |
| Hispanic or Latino Other | 415 | 9 | 189 | 615 | 1,448 | 7,653 | 2.68 | 12 | 269 | 935 | 2,525 | 45,473 | 4.61 | 10 | 298 | 964 | 2,404 | 33,561 | 4.15 |
| Non-Hispanic or Latino White | 40,460 | 18 | 416 | 625 | 1,016 | 8,082 | 1.50 | 18 | 537 | 992 | 2,214 | 48,558 | 2.86 | 10 | 530 | 985 | 2,217 | 44,158 | 2.87 |
| Non-Hispanic or Latino Black or African American | 7,135 | 29 | 549 | 757 | 1,191 | 7,579 | 1.50 | 29 | 667 | 1,046 | 2,111 | 43,516 | 2.80 | 13 | 656 | 1,078 | 2,186 | 31,554 | 2.69 |
| Non-Hispanic or Latino Other | 4,504 | 20 | 244 | 595 | 1,222 | 8,359 | 1.79 | 20 | 383 | 924 | 2,043 | 39,281 | 3.04 | 21 | 362 | 896 | 1,942 | 35,038 | 3.03 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 4,137 | 26 | 590 | 1,011 | 1,441 | 6,622 | 1.34 | 26 | 769 | 1,331 | 2,091 | 22,955 | 2.35 | 33 | 763 | 1,348 | 2,139 | 26,638 | 2.43 |
| 10-<50\% | 11,129 | 9 | 565 | 876 | 1,420 | 8,444 | 1.49 | 12 | 712 | 1,317 | 2,491 | 43,516 | 2.57 | 10 | 712 | 1,330 | 2,557 | 30,041 | 2.56 |
| <10\% | 45,355 | 16 | 384 | 605 | 963 | 8,082 | 1.57 | 16 | 512 | 916 | 2,005 | 48,558 | 3.00 | 8 | 503 | 913 | 1,999 | 44,158 | 2.98 |

Table G. 12012 NSDUH Selected QDU-Level Weight Summary Statistics (continued)

| Domain | $n$ | SDU-Level Weights ${ }^{1}$ (SDUWT: WT1*...*WT9) |  |  |  |  |  | Before sel.qdu.ps ${ }^{1}$ (SDUWT*DUWT10) |  |  |  |  |  | After sel.qdu.ps ${ }^{1}$(SDUWT*DUWT10*DUWT11) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{\mathbf{2}}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{\mathbf{2}}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{\mathbf{2}}$ | Max | UWE $^{3}$ |
| \% Black or African American in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 4,381 | 9 | 530 | 735 | 1,115 | 7,874 | 1.57 | 12 | 648 | 1,002 | 2,038 | 27,138 | 2.99 | 10 | 650 | 1,050 | 2,184 | 31,554 | 2.87 |
| 10-<50\% | 9,185 | 22 | 533 | 802 | 1,286 | 8,236 | 1.48 | 22 | 688 | 1,192 | 2,412 | 43,516 | 2.65 | 13 | 680 | 1,197 | 2,432 | 30,807 | 2.59 |
| <10\% | 47,055 | 18 | 400 | 627 | 1,055 | 8,444 | 1.57 | 18 | 520 | 975 | 2,070 | 48,558 | 2.89 | 8 | 511 | 970 | 2,071 | 44,158 | 2.91 |
| \% Owner-Occupied DUs ${ }^{1}$ in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 45,901 | 18 | 434 | 648 | 1,074 | 8,444 | 1.55 | 18 | 546 | 1,005 | 2,131 | 48,558 | 2.85 | 8 | 540 | 1,011 | 2,143 | 44,158 | 2.84 |
| 10-<50\% | 11,429 | 16 | 433 | 698 | 1,115 | 8,236 | 1.57 | 16 | 567 | 1,006 | 2,055 | 43,516 | 2.98 | 13 | 541 | 1,000 | 2,040 | 35,038 | 2.97 |
| <10\% | 3,291 | 9 | 473 | 810 | 1,317 | 8,351 | 1.59 | 12 | 599 | 1,135 | 2,145 | 37,254 | 2.64 | 10 | 585 | 1,150 | 2,315 | 27,818 | 2.64 |
| Combined Median Rent/Housing Value |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $1^{\text {st }}$ Quintile | 9,920 | 19 | 385 | 594 | 903 | 8,351 | 1.58 | 19 | 507 | 876 | 1,879 | 26,477 | 2.82 | 8 | 502 | 900 | 1,909 | 31,554 | 2.84 |
| $2^{\text {nd }}$ Quintile | 12,960 | 18 | 402 | 622 | 1,024 | 8,444 | 1.63 | 18 | 512 | 944 | 1,995 | 30,624 | 2.90 | 10 | 514 | 948 | 2,013 | 29,325 | 2.85 |
| $3^{\text {rd }}$ Quintile | 13,713 | 9 | 370 | 612 | 1,008 | 8,082 | 1.61 | 12 | 502 | 920 | 1,928 | 43,136 | 3.04 | 10 | 494 | 912 | 1,899 | 30,807 | 3.05 |
| $4^{\text {th }}$ Quintile | 13,171 | 16 | 474 | 732 | 1,182 | 7,853 | 1.53 | 16 | 609 | 1,110 | 2,291 | 45,296 | 2.81 | 21 | 585 | 1,083 | 2,279 | 44,158 | 2.81 |
| $5^{\text {th }}$ Quintile | 10,857 | 25 | 534 | 816 | 1,257 | 8,359 | 1.42 | 25 | 666 | 1,227 | 2,577 | 48,558 | 2.64 | 19 | 676 | 1,255 | 2,624 | 33,561 | 2.63 |
| Population Density |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Large MSA ${ }^{1}$ | 25,508 | 9 | 578 | 849 | 1,343 | 8,359 | 1.41 | 12 | 731 | 1,298 | 2,664 | 48,558 | 2.55 | 10 | 728 | 1,311 | 2,650 | 35,038 | 2.54 |
| Medium to Small MSA ${ }^{1}$ | 30,307 | 18 | 315 | 558 | 886 | 8,444 | 1.62 | 18 | 458 | 838 | 1,765 | 43,516 | 3.06 | 8 | 450 | 843 | 1,770 | 30,041 | 3.06 |
| Non-MSA, ${ }^{1}$ Urban | 1,249 | 20 | 207 | 512 | 791 | 3,628 | 1.63 | 20 | 382 | 770 | 1,684 | 16,479 | 2.97 | 45 | 359 | 792 | 1,711 | 17,924 | 3.05 |
| Non-MSA, ${ }^{1}$ Rural | 3,557 | 18 | 199 | 469 | 802 | 4,220 | 1.71 | 19 | 335 | 772 | 1,692 | 45,296 | 3.44 | 23 | 330 | 759 | 1,648 | 44,158 | 3.51 |
| Group Quarters |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Group | 774 | 25 | 219 | 459 | 986 | 4,098 | 1.90 | 25 | 275 | 592 | 1,269 | 30,862 | 4.92 | 27 | 264 | 571 | 1,225 | 20,208 | 3.38 |
| Non-Group | 59,847 | 9 | 439 | 665 | 1,099 | 8,444 | 1.56 | 12 | 556 | 1,019 | 2,136 | 48,558 | 2.85 | 8 | 548 | 1,021 | 2,153 | 44,158 | 2.84 |
| Household Size |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| One | 7,063 | 23 | 404 | 638 | 1,029 | 6,552 | 1.52 | 69 | 998 | 2,660 | 6,271 | 48,558 | 2.16 | 55 | 976 | 2,629 | 6,225 | 44,158 | 2.16 |
| Two | 21,700 | 19 | 428 | 649 | 1,055 | 8,236 | 1.54 | 20 | 706 | 1,490 | 3,530 | 45,473 | 2.21 | 14 | 704 | 1,481 | 3,501 | 33,561 | 2.20 |
| Three | 17,330 | 18 | 446 | 672 | 1,105 | 8,444 | 1.55 | 18 | 475 | 785 | 1,353 | 24,578 | 2.07 | 21 | 468 | 789 | 1,369 | 15,882 | 2.05 |
| Four or More | 14,528 | 9 | 442 | 694 | 1,185 | 8,359 | 1.60 | 12 | 454 | 729 | 1,268 | 14,597 | 1.78 | 8 | 439 | 728 | 1,262 | 13,876 | 1.79 |

[^18]Table G. 22012 NSDUH Respondent QDU-Level Weight Summary Statistics

| Domain | $n$ | Before res.qdu.nr ${ }^{1}$(SDUWT*DUWT10*DUWT11) |  |  |  |  |  | After res.qdu.nr ${ }^{1}$(SDUWT*DUWT10*... ${ }^{*}$ DUWT12) |  |  |  |  |  | Final Weight: After res.qdu.ps ${ }^{1}$ (SDUWT*DUWT10*...*DUWT13) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{\mathbf{2}}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ |
| Total | 48,850 | 8 | 527 | 971 | 1,975 | 44,158 | 2.90 | 8 | 627 | 1,175 | 2,485 | 64,497 | 3.21 | 6 | 626 | 1,178 | 2,487 | 63,870 | 3.22 |
| Census Region |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Northeast | 9,917 | 8 | 450 | 803 | 1,716 | 29,264 | 3.02 | 8 | 521 | 988 | 2,287 | 47,848 | 3.34 | 6 | 513 | 985 | 2,308 | 45,340 | 3.34 |
| South | 15,019 | 10 | 788 | 1,254 | 2,452 | 44,158 | 2.59 | 20 | 911 | 1,479 | 2,980 | 64,497 | 2.89 | 25 | 912 | 1,482 | 2,975 | 63,870 | 2.89 |
| Midwest | 13,687 | 23 | 502 | 739 | 1,634 | 25,191 | 2.78 | 30 | 600 | 911 | 2,041 | 35,530 | 2.94 | 27 | 601 | 913 | 2,047 | 35,530 | 2.94 |
| West | 10,227 | 41 | 397 | 1,024 | 2,093 | 35,038 | 3.15 | 41 | 468 | 1,228 | 2,645 | 53,491 | 3.54 | 44 | 469 | 1,222 | 2,627 | 59,335 | 3.55 |
| Quarter |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Quarter 1 | 11,628 | 10 | 581 | 1,024 | 2,051 | 35,038 | 2.93 | 10 | 693 | 1,249 | 2,586 | 53,491 | 3.22 | 8 | 693 | 1,251 | 2,566 | 59,335 | 3.22 |
| Quarter 2 | 12,975 | 8 | 484 | 914 | 1,850 | 27,818 | 2.91 | 8 | 577 | 1,102 | 2,336 | 37,706 | 3.18 | 6 | 577 | 1,106 | 2,336 | 37,778 | 3.19 |
| Quarter 3 | 12,743 | 10 | 498 | 930 | 1,935 | 44,158 | 2.86 | 15 | 587 | 1,106 | 2,410 | 64,497 | 3.20 | 13 | 588 | 1,108 | 2,415 | 63,870 | 3.20 |
| Quarter 4 | 11,504 | 13 | 571 | 1,020 | 2,096 | 30,014 | 2.88 | 14 | 681 | 1,230 | 2,660 | 47,848 | 3.21 | 10 | 678 | 1,231 | 2,658 | 45,340 | 3.21 |
| Household Type |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 12-17, 18-25, $26+$ | 5,004 | 13 | 451 | 680 | 1,156 | 9,750 | 1.57 | 14 | 508 | 781 | 1,340 | 9,750 | 1.61 | 10 | 506 | 781 | 1,342 | 9,750 | 1.60 |
| 12-17, 18-25 | 72 | 27 | 301 | 536 | 1,099 | 2,529 | 1.69 | 27 | 329 | 701 | 1,375 | 3,411 | 1.68 | 22 | 321 | 697 | 1,349 | 3,856 | 1.74 |
| 12-17, $26+$ | 14,625 | 10 | 398 | 632 | 1,081 | 8,263 | 1.57 | 10 | 458 | 747 | 1,266 | 10,398 | 1.60 | 8 | 457 | 748 | 1,267 | 12,788 | 1.61 |
| 18-25, $26+$ | 10,290 | 8 | 558 | 903 | 1,430 | 9,930 | 1.54 | 8 | 673 | 1,112 | 1,778 | 15,051 | 1.60 | 6 | 673 | 1,115 | 1,775 | 20,359 | 1.62 |
| 12-17 | 24 | 116 | 279 | 611 | 1,208 | 1,725 | 1.48 | 124 | 295 | 836 | 1,446 | 2,127 | 1.46 | 121 | 301 | 847 | 1,435 | 2,135 | 1.47 |
| 18-25 | 5,447 | 30 | 403 | 761 | 1,221 | 6,586 | 1.57 | 33 | 455 | 892 | 1,434 | 6,277 | 1.57 | 34 | 454 | 889 | 1,432 | 12,046 | 1.59 |
| $26+$ | 13,388 | 53 | 1,827 | 3,412 | 5,696 | 44,158 | 1.73 | 57 | 2,383 | 4,556 | 7,834 | 64,497 | 1.78 | 44 | 2,384 | 4,544 | 7,837 | 63,870 | 1.78 |
| Race/Ethnicity of Householder |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Hispanic or Latino White | 6,574 | 8 | 576 | 1,077 | 1,743 | 26,638 | 2.49 | 8 | 648 | 1,251 | 2,105 | 35,972 | 2.86 | 6 | 648 | 1,254 | 2,098 | 36,038 | 2.88 |
| Hispanic or Latino Black or African American | 165 | 20 | 796 | 1,789 | 3,943 | 29,264 | 2.99 | 20 | 885 | 2,126 | 4,681 | 47,848 | 3.39 | 25 | 1,009 | 2,201 | 4,493 | 45,340 | 3.38 |
| Hispanic or Latino Other | 355 | 10 | 296 | 982 | 2,419 | 33,561 | 4.10 | 28 | 313 | 1,051 | 2,782 | 44,608 | 4.82 | 30 | 330 | 1,045 | 2,888 | 38,613 | 4.48 |
| Non-Hispanic or Latino White | 32,098 | 10 | 513 | 942 | 2,043 | 44,158 | 2.94 | 10 | 621 | 1,165 | 2,641 | 64,497 | 3.20 | 8 | 620 | 1,168 | 2,641 | 63,870 | 3.20 |
| Non-Hispanic or Latino Black or African American | 6,092 | 34 | 647 | 1,048 | 2,037 | 28,858 | 2.69 | 36 | 718 | 1,182 | 2,379 | 34,671 | 3.01 | 36 | 720 | 1,190 | 2,382 | 34,627 | 3.02 |
| Non-Hispanic or Latino Other | 3,566 | 23 | 344 | 833 | 1,782 | 35,038 | 3.05 | 27 | 408 | 1,040 | 2,373 | 53,491 | 3.63 | 22 | 397 | 1,028 | 2,369 | 59,335 | 3.69 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 3,397 | 33 | 731 | 1,300 | 2,006 | 26,638 | 2.48 | 33 | 845 | 1,510 | 2,440 | 35,972 | 2.78 | 35 | 850 | 1,502 | 2,433 | 36,038 | 2.79 |
| 10-<50\% | 9,003 | 10 | 689 | 1,277 | 2,377 | 29,264 | 2.58 | 20 | 801 | 1,528 | 3,041 | 47,848 | 2.89 | 23 | 801 | 1,536 | 3,041 | 45,340 | 2.89 |
| <10\% | 36,450 | 8 | 488 | 880 | 1,851 | 44,158 | 3.04 | 8 | 580 | 1,069 | 2,333 | 64,497 | 3.35 | 6 | 580 | 1,069 | 2,333 | 63,870 | 3.35 |

Table G. 22012 NSDUH Respondent QDU-Level Weight Summary Statistics (continued)

| Domain | $n$ | Before res.qdu.nr ${ }^{1}$ <br> (SDUWT*DUWT10*DUWT11) |  |  |  |  |  | $\begin{gathered} \text { After res.qu.qur } \\ \text { (SDUWT*DUWT10*.... }{ }^{\text {DU }} \end{gathered}$ |  |  |  |  |  | Final Weight: After res.qdu.ps ${ }^{1}$ (SDUWT*DUWT10*...*DUWT13) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ |
| \% Black or African American in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 3,749 | 10 | 639 | 1,022 | 1,988 | 28,858 | 2.89 | 27 | 707 | 1,156 | 2,281 | 40,699 | 3.28 | 30 | 702 | 1,161 | 2,277 | 38,613 | 3.27 |
| 10-<50\% | 7,524 | 20 | 665 | 1,139 | 2,230 | 29,264 | 2.64 | 20 | 775 | 1,363 | 2,762 | 47,848 | 2.97 | 23 | 775 | 1,365 | 2,758 | 45,340 | 2.97 |
| <10\% | 37,577 | 8 | 494 | 929 | 1,922 | 44,158 | 2.96 | 8 | 590 | 1,139 | 2,442 | 64,497 | 3.26 | 6 | 590 | 1,140 | 2,444 | 63,870 | 3.26 |
| \% Owner-Occupied DUs ${ }^{1}$ in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 36,704 | 8 | 526 | 967 | 1,972 | 44,158 | 2.89 | 8 | 632 | 1,180 | 2,510 | 64,497 | 3.19 | 6 | 631 | 1,181 | 2,510 | 63,870 | 3.19 |
| 10-50\% | 9,421 | 23 | 517 | 954 | 1,924 | 35,038 | 3.04 | 26 | 594 | 1,130 | 2,309 | 53,491 | 3.38 | 17 | 594 | 1,131 | 2,312 | 59,335 | 3.40 |
| <10\% | 2,725 | 10 | 574 | 1,109 | 2,223 | 27,818 | 2.67 | 20 | 646 | 1,292 | 2,720 | 39,554 | 2.93 | 25 | 634 | 1,293 | 2,718 | 37,754 | 2.93 |
| Combined Median Rent/Housing Value |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $1^{\text {st }}$ Quintile | 8,288 | 8 | 487 | 871 | 1,784 | 28,858 | 2.86 | 8 | 563 | 1,014 | 2,149 | 35,972 | 3.16 | 6 | 564 | 1,010 | 2,139 | 36,038 | 3.17 |
| $2^{\text {nd }}$ Quintile | 10,661 | 10 | 499 | 915 | 1,897 | 26,430 | 2.92 | 10 | 582 | 1,085 | 2,290 | 40,699 | 3.24 | 8 | 582 | 1,086 | 2,287 | 38,613 | 3.24 |
| $3^{\text {rd }}$ Quintile | 11,009 | 10 | 481 | 883 | 1,765 | 29,155 | 3.06 | 20 | 569 | 1,069 | 2,237 | 39,581 | 3.41 | 20 | 568 | 1,072 | 2,242 | 39,310 | 3.41 |
| $4^{\text {th }}$ Quintile | 10,428 | 23 | 567 | 1,044 | 2,104 | 44,158 | 2.89 | 30 | 678 | 1,271 | 2,683 | 64,497 | 3.15 | 26 | 673 | 1,271 | 2,689 | 63,870 | 3.16 |
| $5^{\text {th }}$ Quintile | 8,464 | 19 | 646 | 1,194 | 2,378 | 33,561 | 2.67 | 20 | 792 | 1,491 | 3,123 | 44,608 | 2.93 | 18 | 789 | 1,491 | 3,109 | 40,087 | 2.93 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Large MSA ${ }^{1}$ | 20,223 | 10 | 704 | 1,243 | 2,431 | 35,038 | 2.60 | 20 | 843 | 1,499 | 3,145 | 53,491 | 2.88 | 17 | 843 | 1,502 | 3,128 | 59,335 | 2.88 |
| Medium to Small MSA ${ }^{1}$ | 24,705 | 8 | 438 | 816 | 1,667 | 29,155 | 3.10 | 8 | 514 | 975 | 2,054 | 39,581 | 3.39 | 6 | 513 | 977 | 2,056 | 39,310 | 3.39 |
| Non-MSA, ${ }^{1}$ Urban | 1,028 | 45 | 341 | 777 | 1,603 | 17,924 | 3.10 | 50 | 382 | 893 | 1,944 | 22,067 | 3.35 | 50 | 382 | 898 | 1,954 | 22,001 | 3.35 |
| Non-MSA, ${ }^{1}$ Rural | 2,894 | 23 | 322 | 711 | 1,560 | 44,158 | 3.51 | 23 | 370 | 865 | 1,942 | 64,497 | 3.87 | 20 | 372 | 865 | 1,944 | 63,870 | 3.86 |
| Group Quarters |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Group | 729 | 27 | 267 | 561 | 1,225 | 20,208 | 3.45 | 27 | 283 | 590 | 1,343 | 20,208 | 3.57 | 22 | 279 | 572 | 1,290 | 19,652 | 3.69 |
| Non-Group | 48,121 | 8 | 531 | 979 | 1,991 | 44,158 | 2.89 | 8 | 633 | 1,184 | 2,516 | 64,497 | 3.20 | 6 | 632 | 1,187 | 2,515 | 63,870 | 3.20 |
| Household Size |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| One | 5,567 | 63 | 917 | 2,379 | 5,997 | 44,158 | 2.22 | 72 | 1,114 | 2,949 | 8,054 | 64,497 | 2.34 | 83 | 1,114 | 2,965 | 8,041 | 63,870 | 2.35 |
| Two | 16,974 | 14 | 670 | 1,361 | 3,229 | 33,561 | 2.25 | 15 | 805 | 1,679 | 4,247 | 44,608 | 2.44 | 13 | 803 | 1,678 | 4,234 | 38,613 | 2.43 |
| Three | 14,071 | 23 | 459 | 775 | 1,341 | 15,882 | 2.01 | 27 | 543 | 938 | 1,617 | 23,531 | 2.32 | 22 | 544 | 940 | 1,621 | 24,616 | 2.34 |
| Four or More | 12,238 | 8 | 431 | 721 | 1,245 | 13,876 | 1.77 | 8 | 497 | 845 | 1,459 | 24,188 | 2.02 | 6 | 493 | 844 | 1,461 | 20,300 | 2.03 |

[^19]
# Appendix H: GEM Modeling Summary for the Pair Weights 

## Appendix H: GEM Modeling Summary for the Pair Weights

This appendix summarizes each model group throughout all stages of weight calibration modeling. Unlike much of the other information presented in this report, this section provides a model-specific overview of weight calibration, as opposed to a domain-specific one.

For 2012, modeling involved taking two model groups through four adjustment steps: (1) selected pair poststratification, (2) pair nonresponse adjustment, (3) responding pair poststratification, and (4) responding pair extreme value adjustment.

Model-specific summary statistics are shown in Tables H.1a through H.2b. Included in these tables, for each stage of modeling, are the number of factor effects included in the final model; the high, low, and nonextreme weight bounds set to provide the upper and lower limits for the generalized exponential model (GEM) macro; the weighted, unweighted, and winsorized weight proportions; the unequal weighting effect (UWE); and weight distributions. The UWE provides an approximate partial measure of variance and provides a summary of how much impact a particular stage of modeling has on the distribution of the new product of weights. At each stage in the modeling, these summary statistics were calculated and utilized to help evaluate the quality of the weight component under the model chosen.

Occurrences of small sample sizes and exact linear combinations in the realized data led to situations whereby modeling inclusion of all originally proposed levels of covariates in the model was not possible. The text and exhibits in Sections H. 1 and H. 2 summarize the decisions made with regard to final covariates included in each model. For the list of proposed initial covariates considered at each stage of modeling, see Exhibit H.2. For the list of realized final model covariates, see Exhibits H.1.1 to H.2.4. For guidelines on interpreting these exhibits, see Appendix C.

## Final Model Explanatory Variables

For brevity, numeric abbreviations for factor levels are established in Exhibit 4.2 (included here as Exhibit H. 1 for easy reference). A complete list of all variables and associated levels used at any stage of modeling is provided. Note that not all factors or levels are present in all stages of modeling, and the initial set of variables is the same across model groups but may change for an adjustment step of modeling. The initial candidates are found in any of the proposed variables columns for a particular stage of weight adjustment.

## Exhibit H. 1 Definitions of Levels for Pair-Level Calibration Modeling Variables

## Group Quarter Indicator

1: College Dorm, 2: Other Group Quarter, 3: Non-Group Quarter ${ }^{1}$

## Household Size

2: DU with 2 Persons, ${ }^{1}$ 3: DU with 3 Persons, 4 : DU with $\geq 4$ Persons
Pair Age (15 Levels)
1: 12-17 and 12-17, ${ }^{1} 2: 12-17$ and 18-25, 3: 12-17 and 26-34, 4: 12-17 and 35-49, 5: 12-17 and 50+, 6: 18-25 and 18-25, 7: 18-25 and 26-34, 8: 18-25 and 35-49, 9: 18-25 and 50+, 10: 26-34 and 26-34, 11:26-34 and 3549, 12: 26-34 and 50+, 13: 35-49 and 35-49, 14: 35-49 and 50+, 15: 50+ and 50+
Pair Age (6 Levels)
1: 12-17 and 12-17, ${ }^{1} 2: 12-17$ and $18-25,3: 12-17$ and $26+, 4: 18-25$ and $18-25,5: 18-25$ and $26+, 6: 26+$ and 26+
Pair Age (3 Levels)
1: 12-17 and 12-17, ${ }^{1} 2: 12-17$ and 18+, 3: 18+ and 18+
Pair Gender
1: Male and Female, ${ }^{1}$ 2: Female and Female, 3: Male and Male
Pair Race/Ethnicity ( 10 Levels)
1: White and White, ${ }^{1} 2$ : White and Black or African American, 3: White and Hispanic or Latino, 4: White and Other, 5: Black or African American and Black or African American, 6: Black or African American and Hispanic or Latino, 7: Black or African American and Other, 8: Hispanic or Latino and Hispanic or Latino, 9: Hispanic or Latino and Other, 10: Other and Other

## Pair Race/Ethnicity (5 Levels)

1: Two or More Races Pair, 2: Hispanic or Latino Pair, 3: Black or African American Pair, 4: White Pair, ${ }^{1}$ 5: Other Pair

Pair Race/Ethnicity (4 Levels)
1: Two or More Races Pair or Other and Other, 2: Hispanic or Latino Pair, 3: Black or African American Pair,
4: White Pair ${ }^{1}$
Percentage of Owner-Occupied Dwelling Units in Segment (\% Owner-Occupied)
1: $50-100 \%^{1}, 2: 10-<50 \%, 3: 0-<10 \%$
Percentage of Segments That Are Black or African American 1: $50-100 \%, 2: 10-<50 \%, 3: 0-<10 \%{ }^{1}$
Percentage of Segments That Are Hispanic or Latino 1: $50-100 \%$, $2: 10-<50 \%, 3: 0-<10 \%{ }^{1}$
Segment-Combined Median Rent and Housing Value (Rent/Housing) ${ }^{2}$
1: First Quintile, 2: Second Quintile, 3: Third Quintile, 4: Fourth Quintile, 5: Fifth Quintile ${ }^{1}$

## Population Density

 1: MSA $1,000,000$ or More, 2 : MSA Less than $1,000,000,3$ : Non-MSA Urban, 4: Non-MSA Rural ${ }^{1}$Quarter
1: Quarter 1, 2: Quarter 2, 3: Quarter 3, 4: Quarter 4 ${ }^{1}$
Race/Ethnicity of Householder
1: Hispanic or Latino White, ${ }^{1}$ 2: Hispanic or Latino Black or African American, 3: Hispanic or Latino Other, 4: Non-Hispanic or Latino White, 5: Non-Hispanic or Latino Black or African American, 6: Non-Hispanic or Latino Other

## State/Region

Model Group 1: 1: Connecticut, Maine, Massachusetts, New Hampshire, New Jersey, Rhode Island, Vermont; 2: Alabama, Arkansas, Delaware, District of Columbia, Georgia, Kentucky, Louisiana, Maryland, Mississippi, North Carolina, Oklahoma, South Carolina, Tennessee, Virginia, West Virginia; ${ }^{1}$ 3: New York; 4: Pennsylvania; 5: Florida; 6: Texas
Model Group 2: 1: Indiana, Iowa, Kansas, Minnesota, Missouri, Nebraska, North Dakota, South Dakota, Wisconsin; ${ }^{1}$ 2: Alaska, Arizona, Colorado, Idaho, Hawaii, Montana, Nevada, New Mexico, Oregon, Utah, Washington, Wyoming; 3: Michigan; 4: Illinois; 5: Ohio; 6: California

Exhibit H. 1 Definitions of Levels for Pair-Level Calibration Modeling Variables (continued)
States ${ }^{3}$
Model Group 1: 1: Alabama, 2: Arkansas, 3: Connecticut, 4: Delaware, 5: District of Columbia, 6: Florida, 7: Georgia, 8: Kentucky, 9: Louisiana, 10: Maine, 11: Maryland, ${ }^{1}$ 12: Massachusetts, 13: Mississippi, 14: New Hampshire, 15: New Jersey, 16: New York, 17: North Carolina, 18: Oklahoma, 19: Pennsylvania, 20: Rhode Island, 21: South Carolina, 22: Tennessee, 23: Texas, 24: Vermont, 25: Virginia, 26: West Virginia
Model Group 2: 1: Alaska, 2: Arizona, ${ }^{1}$ 3: California, 4: Colorado, 5: Idaho, 6: Illinois, 7: Indiana, 8: Iowa, 9: Hawaii, 10: Kansas, 11: Michigan, 12: Minnesota, 13: Missouri, 14: Montana, 15: Nebraska, 16: Nevada, 17: New Mexico, 18: North Dakota, 19: Ohio, 20: Oregon, 21: South Dakota, 22: Utah, 23: Washington, 24: Wisconsin, 25: Wyoming

Pair Relationship Associated with Multiplicity
1: Parent-Child (12-14)*
2: Parent-Child (12-17)*
3: Parent-Child (12-10)*
4: Parent*-Child (12-14)
5: Parent*-Child (12-17)
6: Parent*-Child (12-20)
7: Sibling (12-14)-Sibling (15-17)*
8: Sibling (12-17)-Sibling (18-25)*
9: Spouse-Spouse/Partner-Partner
10: Spouse-Spouse/Partner-Partner with Children (Younger than 18)
DU = dwelling unit, MSA = metropolitan statistical area.
${ }^{1}$ The reference level for this variable. This is the level against which effects of other factor levels are measured.
${ }^{2}$ Segment-Combined Median Rent and Housing Value is a composite measure based on rent, housing value, and percentage owner-occupied.
${ }^{3}$ The States or district assigned to a particular model is based on combined census regions.

* The pair member focused on.

Exhibit H. 2 Covariates for 2012 NSDUH Pair Weights

| Variables | Level | Proposed |
| :---: | :---: | :---: |
| One-Factor Effects |  |  |
| Intercept | 1 | 1 |
| State | Model-specific |  |
| Quarter | 4 | 3 |
| Population Density | 3 | 2 |
| Group Quarter | 3 | 2 |
| Household Size | 3 | 2 |
| Pair Age | 15 | 14 |
| Pair Gender | 4 | 2 |
| Pair Race/Ethnicity | 10 | 9 |
| Race/Ethnicity of Householder | 6 | 5 |
| Rent/Housing | 5 | 4 |
| Segment \% Black or African American | 3 | 2 |
| Segment \% Hispanic or Latino | 3 | 2 |
| \% Owner-Occupied | 3 | 2 |
| Pair Relationship | Model-specific |  |
| Two-Factor Effects |  |  |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Age (6 Levels) | $5 \times 6$ | 20 |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Gender | $5 \times 3$ | 8 |
| Pair Gender $\times$ Pair Age (6 Levels) | $3 \times 6$ | 10 |
| State/Region $\times$ Pair Race/Ethnicity (5 Levels) | Model-specific |  |
| State/Region $\times$ Pair Age (6 Levels) | Model-specific |  |
| State/Region $\times$ Pair Gender | Model-specific |  |
| Rent/Housing $\times$ \% Black or African American | $5 \times 3$ | 8 |
| Rent/Housing $\times$ \% Hispanic or Latino | $5 \times 3$ | 8 |
| Rent/Housing $\times \%$ Owner-Occupied | $5 \times 3$ | 8 |
| \% Owner-Occupied $\times$ \% Black or African American | $3 \times 3$ | 4 |
| $\%$ Owner-Occupied $\times \%$ Hispanic or Latino | $3 \times 3$ | 4 |
| Three-Factor Effects |  |  |
| Pair Race/Ethnicity (4 Levels) $\times$ Pair Gender $\times$ Pair Age (3 Levels) | $4 \times 3 \times 3$ | 12 |

# Appendix H.1: Model Group 1: Northeast and South 

(Alabama, Arkansas, Connecticut, Delaware, District of Columbia, Florida, Georgia, Kentucky, Louisiana, Maine, Massachusetts, Maryland, Mississippi, New
Hampshire, New Jersey, New York, North Carolina, Oklahoma, Pennsylvania, Rhode Island, South Carolina, Tennessee, Texas, Vermont, Virginia, West Virginia)

Table H.1a 2012 Pair Weight GEM Modeling Summary (Model Group 1: Northeast and South)

| Modeling Step ${ }^{1}$ | Extreme Weight Proportions |  |  | UWE ${ }^{2}$ | \# Covariates ${ }^{3}$ | Bounds ${ }^{4}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | \% Unweighted | \% Weighted | \% Winsorized |  |  | Nominal | Realized |
| sel.pr.ps | 5.38 | 23.39 | 12.45 | 67.5167 | 213 | (0.26, 1.10) | $(0.26,1.10)$ |
|  | 1.86 | 7.18 | 1.41 | 8.5932 | 201 | (0.21, 3.98) | (0.21, 3.96) |
|  |  |  |  |  |  | (0.90, 1.60) | $(0.90,1.60)$ |
| res.pr.nr | 1.74 | 6.26 | 1.31 | 8.9725 | 213 | (1.00, 1.60) | $(1.00,1.60)$ |
|  | 2.09 | 9.12 | 1.85 | 9.3991 | 213 | (1.00, 5.00) | (1.00, 5.00) |
|  |  |  |  |  |  | N/A | N/A |
| res.pr.ps | 2.09 | 9.39 | 1.81 | 9.3991 | 223 | (0.39, 1.20) | $(0.39,1.20)$ |
|  | 1.47 | 5.44 | 0.55 | 9.6725 | 211 | (0.25, 1.89) | $(0.25,1.89)$ |
|  |  |  |  |  |  | N/A | N/A |
| res.pr.ev | 1.47 | 5.44 | 0.55 | 9.6725 | 223 | (0.94, 1.30) | $(0.96,1.30)$ |
|  | 0.81 | 5.70 | 0.19 | 9.6484 | 211 | (0.80, 1.44) | $(0.85,1.28)$ |
|  |  |  |  |  |  | N/A | N/A |

GEM = generalized exponential model.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.
${ }^{2}$ Unequal weighting effect (UWE) defined as $1+[(n-1) / n]^{*} C V^{2}$, where $C V=$ coefficient of variation of weights.
${ }^{3}$ Number of proposed covariates on top line and number finalized after modeling.
${ }^{4}$ Nominal bounds are used in defining maximum/minimum values for the GEM adjustment factors. The realized bound is the actual adjustment produced by the modeling. The first set of bounds listed is for high extreme values, the second is for nonextreme values, and the third is for low extreme values.

Table H.1b 2012 Distribution of Weight Adjustment Factors and Weight Products (Model Group 1: Northeast and South)

|  | SDU <br> Weight | Pair Selection Prob |  | sel.pr.ps ${ }^{1}$ |  | res.pr.nr ${ }^{1}$ |  | res.pr.ps ${ }^{1}$ |  | res.pr.ev ${ }^{1}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1-10 | pairwt11 | 1-11 | pairwt12 | 1-12 | pairwt13 | 1-13 | pairwt14 | 1-14 | pairwt15 | 1-15 |
| Minimum | 9 | 1.02 | 20 | 0.01 | 8 | 0.41 | 10 | 0.17 | 4 | 0.61 | 3 |
| 1\% | 63 | 1.11 | 131 | 0.22 | 62 | 0.85 | 65 | 0.30 | 53 | 0.88 | 51 |
| 5\% | 127 | 1.20 | 355 | 0.31 | 216 | 1.00 | 229 | 0.49 | 189 | 0.95 | 184 |
| 10\% | 214 | 1.31 | 645 | 0.41 | 410 | 1.01 | 433 | 0.62 | 382 | 0.96 | 385 |
| 25\% | 498 | 1.51 | 1,307 | 0.63 | 1,075 | 1.04 | 1,145 | 0.84 | 1,091 | 0.98 | 1,085 |
| Median | 781 | 5.52 | 3,306 | 0.94 | 3,088 | 1.16 | 3,462 | 1.01 | 3,367 | 1.00 | 3,360 |
| 75\% | 1,185 | 11.85 | 8,863 | 1.30 | 8,660 | 1.42 | 10,165 | 1.15 | 9,887 | 1.01 | 9,890 |
| 90\% | 1,629 | 22.86 | 18,678 | 1.75 | 21,257 | 1.88 | 29,270 | 1.37 | 28,709 | 1.03 | 28,478 |
| 95\% | 2,096 | 28.63 | 28,096 | 2.13 | 33,961 | 2.33 | 51,092 | 1.52 | 51,237 | 1.04 | 51,790 |
| 99\% | 3,206 | 53.82 | 60,653 | 2.89 | 75,394 | 3.96 | 135,350 | 1.69 | 138,677 | 1.11 | 139,613 |
| Maximum | 7,138 | 4,403.70 | 6,068,764 | 3.96 | 1,053,229 | 5.00 | 1,323,156 | 1.89 | 1,234,015 | 1.28 | 1,237,337 |
| $n$ | 13,619 | - | 13,619 | - | 13,619 | - | 9,723 | - | 9,723 | - | 9,723 |
| Mean | 905 | 10.01 | 8,700 | 1.03 | 8,841 | 1.34 | 12,384 | 1.00 | 12,384 | 1.00 | 12,384 |
| Max/Mean | 8 | - | 698 | - | 119 | - | 107 | - | 100 | - | 100 |

SDU = screener dwelling unit.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.

## Model Group 1 Overview

## Selected Pair-Level Poststratification

In the selected pair-level poststratification step, 201 of 213 proposed factors were retained in the final model. All main and two-factor effects were retained. All three-factor effects were dropped due to convergence problems.

## Respondent Pair-Level Nonresponse

In the respondent pair-level nonresponse step, all proposed factors were retained in the final model.

## Respondent Pair-Level Poststratification

In the respondent pair-level poststratification step, 211 of 223 proposed factors were retained in the final model. All main and two-factor effects were retained. All three-factor effects were dropped due to convergence problems.

## Respondent Pair-Level Extreme Value Adjustment

This step used exactly the same variables as in the respondent pair-level nonresponse and poststratification steps.

Exhibit H.1.1 Covariates for 2012 NSDUH Pair Weights (sel.pr.ps) Model Group 1: Northeast and South

| Variables | Level | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 76 | 76 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| State | 26 | 25 | 25 | All levels present. |
| Quarter | 4 | 3 | 3 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Household Size | 3 | 2 | 2 | All levels present. |
| Pair Age | 15 | 14 | 14 | All levels present. |
| Pair Gender | 3 | 2 | 2 | All levels present. |
| Pair Race/Ethnicity | 10 | 9 | 9 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Segment \% Black or African American | 3 | 2 | 2 | All levels present. |
| Segment \% Hispanic or Latino | 3 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| Two-Factor Effects |  | 125 | 125 |  |
| Pair Race/Ethnicity ( 5 Levels) $\times$ Pair Age (6 Levels) | $5 \times 6$ | 20 | 20 | All levels present. |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Gender | $5 \times 3$ | 8 | 8 | All levels present. |
| Pair Gender $\times$ Pair Age (6 Levels) | $3 \times 6$ | 10 | 10 | All levels present. |
| State/Region $\times$ Pair Race/Ethnicity (5 Levels) | $6 \times 5$ | 20 | 20 | All levels present. |
| State/Region $\times$ Pair Age (6 Levels) | $6 \times 6$ | 25 | 25 | All levels present. |
| State/Region $\times$ Pair Gender | $6 \times 3$ | 10 | 10 | All levels present. |
| Rent/Housing $\times$ \% Black or African American | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Hispanic or Latino | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Owner-Occupied | $5 \times 3$ | 8 | 8 | All levels present. |
| \% Owner-Occupied $\times$ \% Black or African American | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Owner-Occupied $\times$ \% Hispanic or Latino | $3 \times 3$ | 4 | 4 | All levels present. |
| Three-Factor Effects |  | 12 | 0 |  |
| Pair Race/Ethnicity (4 Levels) $\times$ Pair Gender $\times$ Pair Age (3 Levels) | $4 \times 3 \times 3$ | 12 | 0 | Drop all; conv. |
| Total |  | 213 | 201 |  |

Exhibit H.1.2 Covariates for 2012 NSDUH Pair Weights (res.pr.nr) Model Group 1: Northeast and South

| Variables | Level | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 76 | 76 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| State | 26 | 25 | 25 | All levels present. |
| Quarter | 4 | 3 | 3 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Household Size | 3 | 2 | 2 | All levels present. |
| Pair Age | 15 | 14 | 14 | All levels present. |
| Pair Gender | 3 | 2 | 2 | All levels present. |
| Pair Race/Ethnicity | 10 | 9 | 9 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Segment \% Black or African American | 3 | 2 | 2 | All levels present. |
| Segment \% Hispanic or Latino | 3 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| Two-Factor Effects |  | 125 | 125 |  |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Age (6 Levels) | $5 \times 6$ | 20 | 20 | All levels present. |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Gender | $5 \times 3$ | 8 | 8 | All levels present. |
| Pair Gender $\times$ Pair Age (6 Levels) | $3 \times 6$ | 10 | 10 | All levels present. |
| State/Region $\times$ Pair Race/Ethnicity (5 Levels) | $6 \times 5$ | 20 | 20 | All levels present. |
| State/Region $\times$ Pair Age (6 Levels) | $6 \times 6$ | 25 | 25 | All levels present. |
| State/Region $\times$ Pair Gender | $6 \times 3$ | 10 | 10 | All levels present. |
| Rent/Housing $\times$ \% Black or African American | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Hispanic or Latino | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Owner-Occupied | $5 \times 3$ | 8 | 8 | All levels present. |
| $\%$ Owner-Occupied $\times \%$ Black or African American | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Owner-Occupied $\times$ \% Hispanic or Latino | $3 \times 3$ | 4 | 4 | All levels present. |
| Three-Factor Effects |  | 12 | 12 |  |
| Pair Race/Ethnicity (4 Levels) $\times$ Pair Gender $\times$ Pair Age (3 Levels) | $4 \times 3 \times 3$ | 12 | 12 | All levels present. |
| Total |  | 213 | 213 |  |

Exhibit H.1.3 Covariates for 2012 NSDUH Pair Weights (res.pr.ps) Model Group 1: Northeast and South

| Variables | Level | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 86 | 86 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| State | 26 | 25 | 25 | All levels present. |
| Quarter | 4 | 3 | 3 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Household Size | 3 | 2 | 2 | All levels present. |
| Pair Age | 15 | 14 | 14 | All levels present. |
| Pair Gender | 3 | 2 | 2 | All levels present. |
| Pair Race/Ethnicity | 10 | 9 |  | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Segment \% Black or African American | 3 | 2 | 2 | All levels present. |
| Segment \% Hispanic or Latino | 3 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| Pair Relationship | 10 | 10 | 10 | All levels present. |
| Two-Factor Effects |  | 125 | 125 |  |
| Pair Race/Ethnicity ( 5 Levels) $\times$ Pair Age ( 6 Levels) | $5 \times 6$ | 20 | 20 | All levels present. |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Gender | $5 \times 3$ | 8 | 8 | All levels present. |
| Pair Gender $\times$ Pair Age (6 Levels) | $3 \times 6$ | 10 | 10 | All levels present. |
| State/Region $\times$ Pair Race/Ethnicity (5 Levels) | $6 \times 5$ | 20 | 20 | All levels present. |
| State/Region $\times$ Pair Age (6 Levels) | $6 \times 6$ | 25 | 25 | All levels present. |
| State/Region $\times$ Pair Gender | $6 \times 3$ | 10 | 10 | All levels present. |
| Rent/Housing $\times$ \% Black or African American | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Hispanic or Latino | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Owner-Occupied | $5 \times 3$ | 8 | 8 | All levels present. |
| $\%$ Owner-Occupied $\times \%$ Black or African American | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Owner-Occupied $\times$ \% Hispanic or Latino | $3 \times 3$ | 4 | 4 | All levels present. |
| Three-Factor Effects |  | 12 | 0 |  |
| Pair Race/Ethnicity (4 Levels) $\times$ Pair Gender $\times$ Pair Age (3 Levels) | $4 \times 3 \times 3$ | 12 | 0 | Drop all; conv. |
| Total |  | 223 | 211 |  |

Exhibit H.1.4 Covariates for 2012 NSDUH Pair Weights (res.pr.ev) Model Group 1: Northeast and South

This step used the same variables as the respondent pair-level poststratification step in Exhibit H.1.3.

# Appendix H.2: Model Group 2: Midwest and West 

(Alaska, Arizona, California, Colorado, Hawaii, Idaho, Illinois, Indiana, Iowa, Kansas, Michigan, Minnesota, Missouri, Montana, Nebraska, Nevada, New Mexico, North Dakota, Ohio, Oregon, South Dakota, Utah, Washington, Wisconsin, Wyoming)

Table H.2a 2012 Pair Weight GEM Modeling Summary (Model Group 2: Midwest and West)

| Modeling Step ${ }^{1}$ | Extreme Weight Proportions |  |  | UWE ${ }^{2}$ | \# Covariates ${ }^{3}$ | Bounds ${ }^{4}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | \% Unweighted | \% Weighted | \% Winsorized |  |  | Nominal | Realized |
| sel.pr.ps | 5.02 | 17.79 | 6.19 | 16.0574 | 212 | (0.41, 2.00) | (0.42, 2.00) |
|  | 2.08 | 6.47 | 1.21 | 10.0819 | 200 | (0.24, 3.46) | (0.25, 3.45) |
|  |  |  |  |  |  | (0.90, 1.73) | (.0.90, 1.73) |
| res.pr.nr | 2.31 | 9.13 | 1.90 | 10.7706 | 212 | (1.01, 2.70) | (1.01, 2.70) |
|  | 2.87 | 11.04 | 2.21 | 11.1015 | 212 | (1.00, 5.00) | (1.00, 5.00) |
|  |  |  |  |  |  | N/A | N/A |
| res.pr.ps | 2.83 | 12.68 | 3.78 | 11.1015 | 222 | $(0.56,1.90)$ | $(0.58,1.90)$ |
|  | 2.25 | 12.52 | 2.14 | 9.2899 | 210 | (0.43, 2.50) | (0.45, 2.49) |
|  |  |  |  |  |  | N/A | N/A |
| res.pr.ev | 2.25 | 12.52 | 2.14 | 9.2899 | 222 | (0.80, 1.80) | (0.94, 1.77) |
|  | 1.04 | 7.73 | 0.48 | 8.7981 | 210 | (0.70, 1.80) | (0.80, 1.30) |
|  |  |  |  |  |  | N/A | N/A |

GEM = generalized exponential model.
${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.
${ }^{2}$ Unequal weighting effect (UWE) defined as $1+[(n-1) / n]^{*} C V^{2}$, where $C V=$ coefficient of variation of weights.
${ }^{3}$ Number of proposed covariates on top line and number finalized after modeling.
${ }^{4}$ Nominal bounds are used in defining maximum/minimum values for the GEM adjustment factors. The realized bound is the actual adjustment produced by the modeling. The first set of bounds listed is for high extreme values, the second is for nonextreme values, and the third is for low extreme values.

Table H.2b 2012 Distribution of Weight Adjustment Factors and Weight Products (Model Group 2: Midwest and West)

|  | $\underset{\text { Weight }}{\text { SDU }}$ | Pair Selection |  | sel.pr.ps ${ }^{1}$ |  | res.pr.nr ${ }^{1}$ |  | res.pr.ps ${ }^{1}$ |  | res.pr.ev ${ }^{1}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1-10 | pairwt11 | 1-11 | pairwt12 | 1-12 | pairwt13 | 1-13 | pairwt14 | 1-14 | pairwt15 | 1-15 |
| Minimum | 20 | 1.02 | 64 | 0.06 | 29 | 0.50 | 30 | 0.11 | 18 | 0.52 | 16 |
| 1\% | 92 | 1.10 | 161 | 0.31 | 132 | 0.99 | 153 | 0.50 | 124 | 0.83 | 118 |
| 5\% | 132 | 1.21 | 336 | 0.47 | 293 | 1.00 | 334 | 0.59 | 292 | 0.89 | 279 |
| 10\% | 175 | 1.32 | 550 | 0.58 | 470 | 1.01 | 518 | 0.66 | 474 | 0.92 | 462 |
| 25\% | 405 | 1.49 | 964 | 0.78 | 984 | 1.04 | 1,115 | 0.82 | 1,073 | 0.96 | 1,056 |
| Median | 577 | 5.26 | 2,692 | 1.01 | 2,564 | 1.15 | 2,913 | 0.97 | 2,888 | 0.99 | 2,861 |
| 75\% | 1,045 | 11.43 | 7,082 | 1.28 | 7,159 | 1.43 | 8,363 | 1.15 | 8,574 | 1.02 | 8,610 |
| 90\% | 1,691 | 21.68 | 15,760 | 1.59 | 17,689 | 1.93 | 22,988 | 1.45 | 23,090 | 1.06 | 23,367 |
| 95\% | 1,995 | 28.31 | 24,872 | 1.79 | 27,387 | 2.45 | 41,048 | 1.66 | 40,453 | 1.09 | 40,772 |
| 99\% | 2,664 | 54.34 | 57,015 | 2.42 | 67,482 | 4.09 | 118,207 | 2.03 | 122,494 | 1.14 | 121,848 |
| Maximum | 8,444 | 1,603.54 | 1,594,823 | 3.45 | 1,028,039 | 5.00 | 1,550,385 | 2.49 | 804,009 | 1.30 | 639,872 |
| $n$ | 13,416 | - | 13,416 | - | 13,416 | - | 9,736 | - | 9,736 | - | 9,736 |
| Mean | 787 | 9.09 | 7,314 | 1.06 | 7,569 | 1.36 | 10,430 | 1.02 | 10,430 | 0.99 | 10,430 |
| Max/Mean | 11 | - | 218 | - | 136 | - | 149 | - | 77 | - | 61 |

[^20]
## Model Group 2 Overview

## Selected Pair-Level Poststratification

In the selected pair-level poststratification step, 200 of 212 proposed factors were retained in the final model. All main and two-factor effects were retained at proposed levels. None of the 12 three-factor effects were kept in the model due to convergence problems.

## Respondent Pair-Level Nonresponse

In the respondent pair-level nonresponse step, all 212 proposed factors were retained in the final model.

## Respondent Pair-Level Poststratification

In the respondent pair-level poststratification step, 210 of 222 proposed factors were retained in the final model, as in the selected pair-level poststratification step.

## Respondent Pair-Level Extreme Value Adjustment

In the respondent pair-level extreme value adjustment step, 210 of 222 proposed factors were retained in the final model, as in the respondent pair-level poststratification step.

Exhibit H.2.1 Covariates for 2012 NSDUH Pair Weights (sel.pr.ps) Model Group 2: Midwest and West

| Variables | Level | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 75 | 75 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| State | 25 | 24 | 24 | All levels present. |
| Quarter | 4 | 3 | 3 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Household Size | 3 | 2 | 2 | All levels present. |
| Pair Age | 15 | 14 | 14 | All levels present. |
| Pair Gender | 3 | 2 | 2 | All levels present. |
| Pair Race/Ethnicity | 10 | 9 | 9 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Segment \% Black or African American | 3 | 2 | 2 | All levels present. |
| Segment \% Hispanic or Latino | 3 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| Two-Factor Effects |  | 125 | 125 |  |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Age (6 Levels) | $5 \times 6$ | 20 | 20 | All levels present. |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Gender | $5 \times 3$ | 8 | 8 | All levels present. |
| Pair Gender $\times$ Pair Age (6 Levels) | $3 \times 6$ | 10 | 10 | All levels present. |
| State/Region $\times$ Pair Race/Ethnicity (5 Levels) | $6 \times 5$ | 20 | 20 | All levels present. |
| State/Region $\times$ Pair Age (6 Levels) | $6 \times 6$ | 25 | 25 | All levels present. |
| State/Region $\times$ Pair Gender | $6 \times 3$ | 10 | 10 | All levels present. |
| Rent/Housing $\times$ \% Black or African American | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Hispanic or Latino | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Owner-Occupied | $5 \times 3$ | 8 | 8 | All levels present. |
| $\%$ Owner-Occupied $\times \%$ Black or African American | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Owner-Occupied $\times$ \% Hispanic or Latino | $3 \times 3$ | 4 | 4 | All levels present. |
| Three-Factor Effects |  | 12 | 0 |  |
| Pair Race/Ethnicity (4 Levels) $\times$ Pair Gender $\times$ Pair Age (3 Levels) | $4 \times 3 \times 3$ | 12 | 0 | Drop all; conv. |
| Total |  | 212 | 200 |  |

Exhibit H.2.2 Covariates for 2012 NSDUH Pair Weights (res.pr.nr) Model Group 2: Midwest and West

| Variables | Level | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 75 | 75 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| State | 25 | 24 | 24 | All levels present. |
| Quarter | 4 | 3 | 3 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Household Size | 3 | 2 | 2 | All levels present. |
| Pair Age | 15 | 14 | 14 | All levels present. |
| Pair Gender | 3 | 2 | 2 | All levels present. |
| Pair Race/Ethnicity | 10 | 9 | 9 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Segment \% Black or African American | 3 | 2 | 2 | All levels present. |
| Segment \% Hispanic or Latino | 3 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| Two-Factor Effects |  | 125 | 125 |  |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Age (6 Levels) | $5 \times 6$ | 20 | 20 | All levels present. |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Gender | $5 \times 3$ | 8 | 8 | All levels present. |
| Pair Gender $\times$ Pair Age (6 Levels) | $3 \times 6$ | 10 | 10 | All levels present. |
| State/Region $\times$ Pair Race/Ethnicity (5 Levels) | $6 \times 5$ | 20 | 20 | All levels present. |
| State/Region $\times$ Pair Age (6 Levels) | $6 \times 6$ | 25 | 25 | All levels present. |
| State/Region $\times$ Pair Gender | $6 \times 3$ | 10 | 10 | All levels present. |
| Rent/Housing $\times$ \% Black or African American | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Hispanic or Latino | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Owner-Occupied | $5 \times 3$ | 8 | 8 | All levels present. |
| $\%$ Owner-Occupied $\times \%$ Black or African American | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Owner-Occupied $\times$ \% Hispanic or Latino | $3 \times 3$ | 4 | 4 | All levels present. |
| Three-Factor Effects |  | 12 | 12 |  |
| Pair Race/Ethnicity (4 Levels) $\times$ Pair Gender $\times$ Pair Age (3 Levels) | $4 \times 3 \times 3$ | 12 | 12 | All levels present. |
| Total |  | 212 | 212 |  |

Exhibit H.2.3 Covariates for 2012 NSDUH Pair Weights (res.pr.ps) Model Group 2: Midwest and West

| Variables | Level | Proposed | Final | Comments |
| :---: | :---: | :---: | :---: | :---: |
| One-Factor Effects |  | 85 | 85 |  |
| Intercept | 1 | 1 | 1 | All levels present. |
| State | 26 | 24 | 24 | All levels present. |
| Quarter | 4 | 3 | 3 | All levels present. |
| Population Density | 4 | 3 | 3 | All levels present. |
| Group Quarter | 3 | 2 | 2 | All levels present. |
| Household Size | 3 | 2 | 2 | All levels present. |
| Pair Age | 15 | 14 | 14 | All levels present. |
| Pair Gender | 3 | 2 | 2 | All levels present. |
| Pair Race/Ethnicity | 10 | 9 | 9 | All levels present. |
| Race/Ethnicity of Householder | 6 | 5 | 5 | All levels present. |
| Rent/Housing | 5 | 4 | 4 | All levels present. |
| Segment \% Black or African American | 3 | 2 | 2 | All levels present. |
| Segment \% Hispanic or Latino | 3 | 2 | 2 | All levels present. |
| \% Owner-Occupied | 3 | 2 | 2 | All levels present. |
| Pair Relationship | 10 | 10 | 10 | All levels present. |
| Two-Factor Effects |  | 125 | 125 |  |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Age ( 6 Levels) | $5 \times 6$ | 20 | 20 | All levels present. |
| Pair Race/Ethnicity (5 Levels) $\times$ Pair Gender | $5 \times 3$ | 8 | 8 | All levels present. |
| Pair Gender $\times$ Pair Age (6 Levels) | $3 \times 6$ | 10 | 10 | All levels present. |
| State/Region $\times$ Pair Race/Ethnicity (5 Levels) | $6 \times 5$ | 20 | 20 | All levels present. |
| State/Region $\times$ Pair Age (6 Levels) | $6 \times 6$ | 25 | 25 | All levels present. |
| State/Region $\times$ Pair Gender | $6 \times 3$ | 10 | 10 | All levels present. |
| Rent/Housing $\times$ \% Black or African American | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times$ \% Hispanic or Latino | $5 \times 3$ | 8 | 8 | All levels present. |
| Rent/Housing $\times \%$ Owner-Occupied | $5 \times 3$ | 8 | 8 | All levels present. |
| $\%$ Owner-Occupied $\times \%$ Black or African American | $3 \times 3$ | 4 | 4 | All levels present. |
| \% Owner-Occupied $\times$ \% Hispanic or Latino | $3 \times 3$ | 4 | 4 | All levels present. |
| Three-Factor Effects |  | 12 | 0 |  |
| Pair Race/Ethnicity (4 Levels) $\times$ Pair Gender $\times$ Pair Age (3 Levels) | $4 \times 3 \times 3$ | 12 | 0 | Drop all; conv. |
| Total |  | 222 | 210 |  |

Exhibit H.2.4 Covariates for 2012 NSDUH Pair Weights (res.pr.ev) Model Group 2: Midwest and West

This step used the same variables as the respondent pair-level poststratification step in Exhibit H.2.3.

# Appendix I: Evaluation of Calibration Weights: Pair-Level Response Rates 

Table I. 12012 NSDUH Person Pair-Level Response Rates

| Domain | Selected Pairs | Respondent Pairs | \% Interview Response Rate ${ }^{1}$ |
| :---: | :---: | :---: | :---: |
| Total | 27,035 | 19,459 | 63.76 |
| Pair Age Group |  |  |  |
| 12-17, 12-17 | 4,507 | 3,666 | 81.51 |
| 12-17, 18-25 | 3,627 | 2,778 | 77.17 |
| 12-17, 26-34 | 825 | 653 | 77.42 |
| 12-17, 35-49 | 3,813 | 2,816 | 73.86 |
| 12-17, 50+ | 851 | 627 | 74.42 |
| 18-25, 18-25 | 5,476 | 3,976 | 73.18 |
| 18-25, 26-34 | 1,079 | 742 | 67.14 |
| 18-25, 35-49 | 1,582 | 1,058 | 65.84 |
| 18-25, 50+ | 1,074 | 650 | 58.92 |
| 26-34, 26-34 | 880 | 597 | 71.06 |
| 26-34, 35-49 | 469 | 305 | 52.70 |
| 26-34, 50+ | 315 | 178 | 53.89 |
| 35-49, 35-49 | 833 | 489 | 61.19 |
| 35-49, 50+ | 466 | 266 | 50.91 |
| 50+, 50+ | 1,238 | 658 | 52.22 |
| Pair Race/Ethnicity |  |  |  |
| Hispanic or Latino | 4,176 | 3,084 | 69.40 |
| Black or African American | 2,760 | 2,210 | 71.36 |
| White | 16,039 | 11,284 | 61.92 |
| Other | 1,791 | 1,175 | 50.45 |
| White \& Black or African American | 217 | 166 | 68.55 |
| White \& Hispanic or Latino | 921 | 692 | 71.01 |
| White \& Other | 737 | 552 | 63.97 |
| Black or African American \& Hispanic or Latino | 123 | 91 | 69.54 |
| Black or African American \& Other | 126 | 99 | 37.40 |
| Hispanic or Latino \& Other | 145 | 106 | 58.42 |
| Pair Gender |  |  |  |
| Male, Male | 5,928 | 4,173 | 60.83 |
| Female, Female | 5,813 | 4,425 | 67.33 |
| Male, Female | 15,294 | 10,861 | 63.58 |
| Household Size |  |  |  |
| Two | 6,855 | 4,660 | 59.76 |
| Three | 7,398 | 5,248 | 60.52 |
| Four or More | 12,782 | 9,551 | 67.43 |

Table I. 12012 NSDUH Person Pair-Level Response Rates (continued)

| Domain | Selected Pairs | Respondent Pairs | \% Interview Response Rate ${ }^{1}$ |
| :---: | :---: | :---: | :---: |
| Census Region |  |  |  |
| Northeast | 5,685 | 3,856 | 58.61 |
| South | 7,934 | 5,867 | 65.95 |
| Midwest | 7,515 | 5,455 | 66.14 |
| West | 5,901 | 4,281 | 62.52 |
| Quarter |  |  |  |
| Quarter 1 | 6,334 | 4,613 | 64.66 |
| Quarter 2 | 7,183 | 5,115 | 63.01 |
| Quarter 3 | 7,012 | 5,093 | 64.85 |
| Quarter 4 | 6,506 | 4,638 | 62.52 |
| \% Hispanic or Latino in Segment |  |  |  |
| 50-100\% | 2,105 | 1,537 | 65.52 |
| 10-<50\% | 5,112 | 3,726 | 64.88 |
| <10\% | 19,818 | 14,196 | 63.07 |
| \% Black or African American in Segment |  |  |  |
| 50-100\% | 1,905 | 1,482 | 64.05 |
| 10-<50\% | 4,030 | 2,974 | 66.88 |
| <10\% | 21,100 | 15,003 | 63.01 |
| \% Owner-Occupied DUs in Segment |  |  |  |
| 50-100\% | 20,778 | 14,866 | 63.43 |
| 10-<50\% | 4,889 | 3,600 | 64.61 |
| <10\% | 1,368 | 993 | 71.03 |
| Combined Median Rent/Housing Value |  |  |  |
| $1^{\text {st }}$ Quintile | 4,264 | 3,227 | 68.88 |
| $2^{\text {nd }}$ Quintile | 5,758 | 4,250 | 67.80 |
| $3^{\text {rd }}$ Quintile | 6,136 | 4,364 | 65.47 |
| $4^{\text {th }}$ Quintile | 6,013 | 4,276 | 58.78 |
| $5{ }^{\text {th }}$ Quintile | 4,864 | 3,342 | 60.58 |
| Population Density |  |  |  |
| Large MSA | 11,649 | 8,192 | 61.00 |
| Medium to Small MSA | 13,392 | 9,833 | 67.16 |
| Non-MSA, Urban | 511 | 380 | 71.70 |
| Non-MSA, Rural | 1,483 | 1,054 | 65.73 |
| Group Quarters |  |  |  |
| Group | 386 | 312 | 79.80 |
| Non-Group | 26,649 | 19,147 | 63.70 |

[^21]
# Appendix J: Evaluation of Calibration Weights: Pair-Level Proportions of Extreme Values and Outwinsors 

Table J. 12012 NSDUH Selected Pair-Level Proportions of Extreme Values and Outwinsors

| Domain | $n$ | SDU-Level Weights ${ }^{1}$ (SDUWT: WT1*...*WT9) |  |  | Before sel.pr.ps ${ }^{1}$ (SDUWT*PRWT10) |  |  | $\begin{gathered} \text { After sel.pr.ps }{ }^{1} \\ \text { (SDUWT*PRWT10*PRWT11) } \end{gathered}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{gathered} \% \\ \text { Unweighted } \end{gathered}$ | $\%$ <br> Weighted ${ }^{2}$ | $\begin{gathered} \% \\ \text { Outwinsor }{ }^{3} \end{gathered}$ | \% <br> Unweighted | $\begin{gathered} \% \\ \text { Weighted }{ }^{2} \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | \% <br> Unweighted | $\%$ <br> Weighted ${ }^{2}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ |
| Total | 27,035 | 2.14 | 5.22 | 1.44 | 5.34 | 26.13 | 15.62 | 2.15 | 15.29 | 7.51 |
| Pair Age Group |  |  |  |  |  |  |  |  |  |  |
| 12-17, 12-17 | 4,507 | 1.84 | 4.16 | 1.29 | 3.64 | 17.26 | 6.55 | 0.55 | 3.06 | 0.64 |
| 12-17, 18-25 | 3,627 | 2.21 | 5.76 | 1.79 | 8.08 | 28.38 | 11.52 | 1.96 | 7.45 | 1.53 |
| 12-17, 26-34 | 825 | 1.82 | 4.03 | 1.52 | 1.82 | 5.22 | 1.36 | 0.85 | 2.56 | 0.36 |
| 12-17, 35-49 | 3,813 | 1.60 | 3.70 | 0.89 | 2.12 | 9.32 | 2.46 | 0.87 | 3.65 | 0.67 |
| 12-17, 50+ | 851 | 2.12 | 6.13 | 1.48 | 2.23 | 10.42 | 3.28 | 0.59 | 0.89 | 0.17 |
| 18-25, 18-25 | 5,476 | 2.12 | 4.99 | 1.22 | 8.16 | 27.90 | 10.87 | 3.67 | 12.83 | 2.39 |
| 18-25, 26-34 | 1,079 | 3.61 | 7.88 | 1.97 | 3.15 | 13.23 | 4.97 | 2.32 | 8.97 | 1.53 |
| 18-25, 35-49 | 1,582 | 3.41 | 9.38 | 3.19 | 6.76 | 26.19 | 11.00 | 2.40 | 7.55 | 1.43 |
| 18-25, 50+ | 1,074 | 2.33 | 5.87 | 1.61 | 3.82 | 16.98 | 6.23 | 1.30 | 3.90 | 0.59 |
| 26-34, 26-34 | 880 | 2.73 | 5.15 | 1.11 | 3.64 | 21.67 | 12.88 | 2.84 | 21.62 | 10.40 |
| 26-34, 35-49 | 469 | 2.77 | 5.47 | 1.26 | 7.46 | 43.03 | 29.62 | 5.97 | 36.72 | 19.45 |
| 26-34, 50+ | 315 | 2.22 | 5.94 | 2.07 | 4.44 | 22.48 | 12.39 | 1.59 | 16.78 | 12.07 |
| 35-49, 35-49 | 833 | 2.04 | 6.54 | 2.01 | 6.24 | 44.99 | 30.94 | 4.80 | 45.93 | 30.71 |
| 35-49, 50+ | 466 | 1.50 | 4.26 | 0.59 | 1.93 | 20.54 | 15.31 | 2.15 | 21.08 | 13.47 |
| 50+, 50+ | 1,238 | 1.62 | 2.89 | 0.40 | 8.24 | 43.73 | 33.58 | 4.36 | 21.24 | 8.98 |
| Pair Race/Ethnicity |  |  |  |  |  |  |  |  |  |  |
| Hispanic or Latino | 4,176 | 3.86 | 10.45 | 3.98 | 5.17 | 35.09 | 24.22 | 2.95 | 22.46 | 11.57 |
| Black or African American | 2,760 | 3.44 | 5.97 | 1.01 | 6.63 | 33.29 | 22.33 | 3.22 | 20.24 | 10.19 |
| White | 16,039 | 0.68 | 1.49 | 0.19 | 4.65 | 21.19 | 12.31 | 1.35 | 11.75 | 6.17 |
| Other | 1,791 | 6.03 | 12.30 | 2.98 | 7.76 | 28.26 | 10.20 | 3.35 | 15.98 | 5.08 |
| White \& Black or African American | 217 | 6.45 | 9.80 | 1.92 | 9.68 | 16.84 | 6.21 | 13.82 | 19.84 | 3.28 |
| White \& Hispanic or Latino | 921 | 2.71 | 5.44 | 1.05 | 6.84 | 21.82 | 7.40 | 3.04 | 11.85 | 3.66 |
| White \& Other | 737 | 3.53 | 6.72 | 1.57 | 4.75 | 18.77 | 8.09 | 2.04 | 14.84 | 7.83 |
| Black or African American \& Hispanic or Latino | 123 | 17.89 | 43.36 | 18.62 | 17.89 | 66.55 | 44.19 | 6.50 | 39.53 | 17.08 |
| Black or African American \& Other | 126 | 5.56 | 12.08 | 2.35 | 5.56 | 40.75 | 27.57 | 1.59 | 34.09 | 21.52 |
| Hispanic or Latino \& Other | 145 | 8.28 | 16.29 | 5.15 | 8.97 | 21.25 | 9.63 | 6.90 | 7.90 | 1.71 |
| Pair Gender |  |  |  |  |  |  |  |  |  |  |
| Male, Male | 5,928 | 2.14 | 4.77 | 1.29 | 7.25 | 31.61 | 20.02 | 2.65 | 11.86 | 5.08 |
| Female, Female | 5,813 | 2.34 | 5.51 | 1.66 | 5.33 | 18.19 | 7.47 | 2.31 | 11.75 | 3.90 |
| Male, Female | 15,294 | 2.07 | 5.28 | 1.42 | 4.61 | 26.61 | 16.48 | 1.90 | 17.24 | 9.21 |
| Household Size |  |  |  |  |  |  |  |  |  |  |
| Two | 6,855 | 1.74 | 4.34 | 1.11 | 0.83 | 2.47 | 0.67 | 0.44 | 2.29 | 0.31 |
| Three | 7,398 | 2.10 | 5.32 | 1.69 | 2.61 | 27.75 | 19.75 | 1.82 | 16.75 | 7.66 |
| Four or More | 12,782 | 2.39 | 5.59 | 1.47 | 9.35 | 37.59 | 21.18 | 3.25 | 20.99 | 11.01 |

Table J. 12012 NSDUH Selected Pair-Level Proportions of Extreme Values and Outwinsors (continued)

| Domain | $n$ | SDU-Level Weights ${ }^{1}$ (SDUWT: WT1*...*WT9) |  |  | Before sel.pr.ps ${ }^{1}$ (SDUWT*PRWT10) |  |  | After sel.pr.ps ${ }^{1}$ <br> (SDUWT*PRWT10*PRWT11) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | \% <br> Unweighted | $\%$ <br> Weighted ${ }^{2}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | \% <br> Unweighted | $\%$ <br> Weighted ${ }^{2}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | \% <br> Unweighted | $\%$ <br> Weighted ${ }^{2}$ | $\begin{gathered} \text { \% } \\ \text { Outwinsor }^{3} \end{gathered}$ |
| Census Region |  |  |  |  |  |  |  |  |  |  |
| Northeast | 5,685 | 2.62 | 6.98 | 2.28 | 5.82 | 21.94 | 9.44 | 2.36 | 15.05 | 6.33 |
| South | 7,934 | 1.78 | 4.49 | 1.05 | 5.29 | 30.24 | 20.63 | 1.79 | 15.23 | 7.69 |
| Midwest | 7,515 | 1.77 | 3.61 | 0.66 | 5.59 | 26.11 | 14.87 | 2.13 | 14.43 | 7.58 |
| West | 5,901 | 2.64 | 6.43 | 2.11 | 4.64 | 22.62 | 12.66 | 2.46 | 16.23 | 8.07 |
| Quarter |  |  |  |  |  |  |  |  |  |  |
| Quarter 1 | 6,334 | 2.07 | 4.60 | 1.28 | 5.32 | 29.27 | 20.03 | 2.68 | 13.34 | 5.68 |
| Quarter 2 | 7,183 | 2.03 | 5.25 | 1.58 | 4.82 | 18.50 | 7.62 | 2.05 | 11.60 | 3.85 |
| Quarter 3 | 7,012 | 2.05 | 4.87 | 1.15 | 5.03 | 25.81 | 14.78 | 1.64 | 19.10 | 10.98 |
| Quarter 4 | 6,506 | 2.43 | 6.14 | 1.77 | 6.29 | 30.19 | 19.30 | 2.29 | 17.08 | 9.51 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 2,105 | 2.28 | 5.26 | 1.89 | 4.42 | 36.21 | 27.31 | 2.23 | 17.79 | 8.80 |
| 10-<50\% | 5,112 | 3.66 | 9.15 | 3.09 | 5.97 | 33.28 | 22.09 | 3.01 | 18.47 | 9.83 |
| $<10 \%$ | 19,818 | 1.74 | 3.74 | 0.76 | 5.28 | 21.35 | 10.80 | 1.92 | 13.74 | 6.47 |
| \% Black or African American in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 1,905 | 3.67 | 8.74 | 2.71 | 6.72 | 27.02 | 13.67 | 3.78 | 19.57 | 8.12 |
| 10-<50\% | 4,030 | 2.68 | 6.54 | 2.12 | 6.23 | 36.11 | 25.20 | 2.70 | 16.10 | 7.76 |
| <10\% | 21,100 | 1.90 | 4.52 | 1.14 | 5.05 | 23.47 | 13.34 | 1.90 | 14.68 | 7.40 |
| \% Owner-Occupied DUs ${ }^{1}$ in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 20,778 | 1.90 | 4.79 | 1.30 | 5.09 | 27.19 | 16.70 | 2.19 | 15.79 | 7.99 |
| 10-<50\% | 4,889 | 2.88 | 6.37 | 1.88 | 6.14 | 22.44 | 12.07 | 2.39 | 14.14 | 5.88 |
| <10\% | 1,368 | 3.14 | 7.07 | 1.97 | 6.36 | 18.81 | 7.21 | 0.66 | 2.43 | 0.34 |
| Combined Median |  |  |  |  |  |  |  |  |  |  |
| Rent/Housing Value |  |  |  |  |  |  |  |  |  |  |
| $1^{\text {st }}$ Quintile | 4,264 | 1.74 | 4.47 | 1.35 | 5.42 | 39.11 | 29.10 | 2.11 | 20.67 | 12.05 |
| $2^{\text {nd }}$ Quintile | 5,758 | 2.15 | 5.91 | 1.99 | 5.45 | 22.10 | 9.42 | 2.57 | 16.93 | 7.54 |
| $3^{\text {rd }}$ Quintile | 6,136 | 2.22 | 4.62 | 1.18 | 5.30 | 23.81 | 13.29 | 1.65 | 14.94 | 8.36 |
| $4^{\text {th }}$ Quintile | 6,013 | 2.33 | 5.21 | 1.39 | 5.32 | 26.61 | 17.44 | 2.06 | 14.47 | 7.09 |
| $5{ }^{\text {th }}$ Quintile | 4,864 | 2.16 | 5.64 | 1.28 | 5.24 | 20.76 | 10.26 | 2.43 | 11.49 | 4.17 |
| Population Density |  |  |  |  |  |  |  |  |  |  |
| Large MSA ${ }^{1}$ | 11,649 | 2.83 | 6.69 | 1.90 | 5.55 | 27.46 | 16.79 | 2.59 | 16.62 | 8.29 |
| Medium to Small MSA ${ }^{1}$ | 13,392 | 1.68 | 3.67 | 0.97 | 5.39 | 19.31 | 7.73 | 1.92 | 12.88 | 5.52 |
| Non-MSA, ${ }^{1}$ Urban | 511 | 0.78 | 0.68 | 0.13 | 3.33 | 72.80 | 70.94 | 1.37 | 23.83 | 19.61 |
| Non-MSA, ${ }^{1}$ Rural | 1,483 | 1.35 | 2.73 | 0.51 | 3.98 | 26.19 | 19.15 | 1.01 | 17.82 | 12.43 |
| Group Quarters |  |  |  |  |  |  |  |  |  |  |
| Group | 386 | 2.59 | 3.81 | 0.78 | 8.03 | 23.60 | 6.46 | 6.99 | 19.26 | 3.27 |
| Non-Group | 26,649 | 2.14 | 5.23 | 1.45 | 5.31 | 26.14 | 15.65 | 2.08 | 15.28 | 7.53 |

[^22]Table J. 2012 NSDUH Respondent Pair-Level Proportions of Extreme Values and Outwinsors

| Domain | $n$ | Before res.pr.nr ${ }^{1}$(SDUWT*PRWT10*PRWT11) |  |  | After res.pr.nr ${ }^{1}$(SDUWT*PRWT10*... ${ }^{*}$ PRWT12) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | \% Unweighted | \% Weighted ${ }^{2}$ | \% Outwinsor ${ }^{3}$ | \% Unweighted | \% Weighted ${ }^{2}$ | \% Outwinsor ${ }^{3}$ |
| Total | 19,459 | 2.16 | 15.29 | 7.43 | 2.61 | 15.86 | 6.50 |
| Pair Age Group |  |  |  |  |  |  |  |
| 12-17, 12-17 | 3,666 | 0.46 | 2.36 | 0.35 | 0.38 | 3.16 | 0.97 |
| 12-17, 18-25 | 2,778 | 2.12 | 8.17 | 1.63 | 1.73 | 7.94 | 2.34 |
| 12-17, 26-34 | 653 | 1.07 | 4.90 | 0.74 | 1.38 | 7.70 | 1.95 |
| 12-17, 35-49 | 2,816 | 0.96 | 4.78 | 0.78 | 1.07 | 2.89 | 0.55 |
| 12-17, 50+ | 627 | 0.32 | 0.72 | 0.22 | 0.64 | 3.03 | 0.66 |
| 18-25, 18-25 | 3,976 | 3.65 | 13.28 | 2.63 | 4.45 | 17.62 | 3.91 |
| 18-25, 26-34 | 742 | 2.96 | 11.93 | 1.64 | 4.04 | 17.04 | 3.72 |
| 18-25, 35-49 | 1,058 | 2.46 | 7.42 | 1.52 | 4.73 | 15.74 | 3.50 |
| 18-25, 50+ | 650 | 1.23 | 2.95 | 0.54 | 2.00 | 5.21 | 0.72 |
| 26-34, 26-34 | 597 | 3.85 | 25.89 | 13.06 | 2.68 | 27.33 | 11.77 |
| 26-34, 35-49 | 305 | 6.89 | 29.07 | 11.40 | 10.49 | 34.39 | 14.18 |
| 26-34, 50+ | 178 | 2.25 | 20.31 | 12.72 | 1.69 | 10.63 | 7.06 |
| 35-49, 35-49 | 489 | 5.32 | 49.78 | 35.62 | 7.77 | 48.74 | 30.19 |
| 35-49, 50+ | 266 | 3.38 | 26.20 | 12.42 | 4.14 | 16.72 | 6.22 |
| 50+, 50+ | 658 | 3.80 | 21.67 | 11.23 | 4.86 | 17.34 | 6.29 |
| Pair Race/Ethnicity |  |  |  |  |  |  |  |
| Hispanic or Latino | 3,084 | 2.95 | 26.88 | 14.07 | 3.44 | 23.61 | 11.28 |
| Black or African American | 2,210 | 3.21 | 18.30 | 8.75 | 2.08 | 15.61 | 7.34 |
| White | 11,284 | 1.27 | 10.54 | 5.41 | 2.03 | 12.64 | 5.04 |
| Other | 1,175 | 3.32 | 18.31 | 5.76 | 6.47 | 26.06 | 7.61 |
| White \& Black or African American | 166 | 14.46 | 20.29 | 4.46 | 10.84 | 15.25 | 3.46 |
| White \& Hispanic or Latino | 692 | 3.76 | 13.37 | 4.64 | 0.87 | 7.53 | 4.25 |
| White \& Other | 552 | 1.99 | 6.27 | 1.52 | 1.09 | 13.41 | 2.08 |
| Black or African American \& Hispanic or Latino | 91 | 5.49 | 29.46 | 19.46 | 7.69 | 27.76 | 12.80 |
| Black or African American \& Other | 99 | 2.02 | 8.26 | 0.30 | 10.10 | 20.77 | 7.12 |
| Hispanic or Latino \& Other | 106 | 8.49 | 12.57 | 2.94 | 2.83 | 10.03 | 1.10 |
| Pair Gender |  |  |  |  |  |  |  |
| Male, Male | 4,173 | 2.64 | 10.72 | 3.25 | 3.50 | 15.22 | 4.05 |
| Female, Female | 4,425 | 2.31 | 8.63 | 2.15 | 2.10 | 9.70 | 2.24 |
| Male, Female | 10,861 | 1.92 | 18.50 | 10.11 | 2.47 | 17.77 | 8.38 |
| Household Size |  |  |  |  |  |  |  |
| Two | 4,660 | 0.45 | 2.58 | 0.43 | 0.67 | 4.13 | 1.44 |
| Three | 5,248 | 1.91 | 17.38 | 8.33 | 2.76 | 19.63 | 7.37 |
| Four or More | 9,551 | 3.14 | 19.92 | 10.09 | 3.47 | 19.72 | 8.56 |

Table J. 22012 NSDUH Respondent Pair-Level Proportions of Extreme Values and Outwinsors (continued)

| Domain | $n$ | Before res.pr.nr ${ }^{1}$(SDUWT*PRWT10*PRWT11) |  |  | After res.pr.nr ${ }^{1}$(SDUWT*PRWT10*... ${ }^{*}$ PRWT12) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | \% Unweighted | \% Weighted ${ }^{2}$ | \% Outwinsor ${ }^{3}$ | \% Unweighted | \% Weighted ${ }^{2}$ | \% Outwinsor ${ }^{3}$ |
| Census Region |  |  |  |  |  |  |  |
| Northeast | 3,856 | 2.26 | 14.23 | 6.16 | 2.52 | 17.44 | 6.83 |
| South | 5,867 | 1.62 | 13.45 | 7.09 | 2.01 | 13.98 | 5.98 |
| Midwest | 5,455 | 2.11 | 15.39 | 8.50 | 2.88 | 19.13 | 8.55 |
| West | 4,281 | 2.90 | 18.66 | 7.89 | 3.15 | 14.76 | 5.35 |
| Quarter |  |  |  |  |  |  |  |
| Quarter 1 | 4,613 | 2.64 | 13.92 | 5.48 | 3.14 | 12.82 | 4.46 |
| Quarter 2 | 5,115 | 2.09 | 10.20 | 3.27 | 2.64 | 13.67 | 3.28 |
| Quarter 3 | 5,093 | 1.45 | 18.44 | 11.32 | 1.65 | 18.64 | 9.28 |
| Quarter 4 | 4,638 | 2.54 | 18.57 | 9.55 | 3.08 | 18.26 | 8.94 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |  |
| 50-100\% | 1,537 | 2.93 | 21.09 | 10.38 | 2.80 | 17.33 | 7.53 |
| 10-<50\% | 3,726 | 3.46 | 21.35 | 10.55 | 3.65 | 19.43 | 8.57 |
| <10\% | 14,196 | 1.74 | 12.08 | 5.77 | 2.31 | 14.34 | 5.58 |
| \% Black or African American in Segment |  |  |  |  |  |  |  |
| 50-100\% | 1,482 | 3.71 | 15.16 | 4.95 | 3.78 | 15.80 | 4.57 |
| 10-<50\% | 2,974 | 3.19 | 19.03 | 9.57 | 3.30 | 17.37 | 7.61 |
| <10\% | 15,003 | 1.81 | 14.40 | 7.16 | 2.35 | 15.52 | 6.44 |
| \% Owner-Occupied DUs ${ }^{1}$ in Segment |  |  |  |  |  |  |  |
| 50-100\% | 14,866 | 2.07 | 15.63 | 7.94 | 2.52 | 15.96 | 6.65 |
| 10-<50\% | 3,600 | 2.86 | 14.94 | 5.66 | 3.36 | 15.87 | 6.06 |
| <10\% | 993 | 1.01 | 4.00 | 0.63 | 1.21 | 11.38 | 3.21 |
| Combined Median Rent/Housing Value |  |  |  |  |  |  |  |
| $1^{\text {st }}$ Quintile | 3,227 | 2.17 | 22.21 | 14.19 | 2.14 | 23.88 | 13.52 |
| $2^{\text {nd }}$ Quintile | 4,250 | 2.49 | 18.39 | 8.55 | 2.49 | 18.83 | 7.49 |
| $3^{\text {rd }}$ Quintile | 4,364 | 1.81 | 16.57 | 8.97 | 2.47 | 17.25 | 7.64 |
| $4^{\text {th }}$ Quintile | 4,276 | 2.03 | 9.98 | 2.44 | 3.04 | 10.26 | 2.39 |
| $5{ }^{\text {th }}$ Quintile | 3,342 | 2.36 | 11.42 | 5.10 | 2.81 | 12.86 | 4.56 |
| Population Density |  |  |  |  |  |  |  |
| Large MSA ${ }^{1}$ | 8,192 | 2.70 | 16.16 | 7.80 | 3.33 | 15.99 | 6.41 |
| Medium to Small MSA ${ }^{1}$ | 9,833 | 1.88 | 13.18 | 5.37 | 2.16 | 14.08 | 5.06 |
| Non-MSA, ${ }^{1}$ Urban | 380 | 1.58 | 33.21 | 27.29 | 1.84 | 33.42 | 22.73 |
| Non-MSA, ${ }^{1}$ Rural | 1,054 | 0.85 | 18.99 | 16.13 | 1.42 | 25.72 | 16.42 |
| Group Quarters |  |  |  |  |  |  |  |
| Group | 312 | 8.33 | 25.50 | 3.58 | 5.77 | 23.85 | 6.74 |
| Non-Group | 19,147 | 2.06 | 15.25 | 7.44 | 2.55 | 15.83 | 6.50 |

${ }^{1}$ This step used demographic variables from screener data for all responding person pairs; $\mathrm{DU}=$ dwelling unit, $\mathrm{MSA}=$ metropolitan statistical area, $\mathrm{NR}=$ nonresponse adjustment, $\mathrm{PR}=$ pair, $\mathrm{Res}=$
respondent, SDU $=$ screener dwelling unit.
${ }^{2}$ Weighted extreme value proportion: $100^{*} \sum_{k} w_{e k} / \sum_{k} w_{k}$, where $w_{e k}$ denotes the weight for extreme values, and $w_{k}$ denotes the weight for both extreme values and nonextreme values.
${ }^{3}$ Outwinsor weight proportion: $100^{*} \sum_{k}\left(w_{e k}-b_{k}\right) / \sum_{k} w_{k}$, where $b_{k}$ denotes the winsorized weight.

Table J. 32012 NSDUH Respondent Pair-Level Proportions of Extreme Values and Outwinsors

| Domain | $n$ | Before res.pr.ps ${ }^{1}$(SDUWT*PRWT10*... ${ }^{*}$ PRWT12) |  |  | After res.pr.ps ${ }^{1}$ (SDUWT*PRWT10*...*PRWT13) |  |  | Final Weight: After res.pr.ev ${ }^{1}$ (SDUWT*PRWT10*...*PRWT14) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | \% <br> Unweighted | $\%$ <br> Weighted ${ }^{2}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | \% <br> Unweighted | $\%$ <br> Weighted ${ }^{2}$ | $\begin{gathered} \% \\ \text { Outwinsor } \end{gathered}$ | \% Unweighted | \% <br> Weighted ${ }^{2}$ | $\begin{gathered} \% \\ \text { Outwinsor } \end{gathered}$ |
| Total | 19,459 | 2.46 | 10.90 | 2.71 | 1.86 | 8.68 | 1.28 | 0.93 | 5.01 | 0.33 |
| Pair Age Group |  |  |  |  |  |  |  |  |  |  |
| 12-17, 12-17 | 3,668 | 0.41 | 3.31 | 1.07 | 0.41 | 1.82 | 0.25 | 0.27 | 1.52 | 0.16 |
| 12-17, 18-25 | 2,759 | 1.70 | 7.95 | 2.36 | 0.91 | 2.82 | 0.58 | 0.69 | 2.47 | 0.36 |
| 12-17, 26-34 | 658 | 1.37 | 7.58 | 1.92 | 1.22 | 3.40 | 0.26 | 0.61 | 1.83 | 0.12 |
| 12-17, 35-49 | 2,812 | 1.07 | 2.84 | 0.53 | 1.21 | 3.50 | 0.58 | 0.71 | 2.08 | 0.20 |
| 12-17, 50+ | 631 | 0.63 | 3.04 | 0.65 | 0.95 | 5.54 | 1.15 | 0.48 | 2.34 | 0.23 |
| 18-25, 18-25 | 3,901 | 4.43 | 17.14 | 3.78 | 3.05 | 10.07 | 1.19 | 0.95 | 3.40 | 0.24 |
| 18-25, 26-34 | 794 | 3.78 | 18.47 | 7.16 | 3.78 | 13.35 | 1.57 | 1.76 | 5.64 | 0.41 |
| 18-25, 35-49 | 1,053 | 5.03 | 16.30 | 3.45 | 5.13 | 14.84 | 2.08 | 2.28 | 6.87 | 0.62 |
| 18-25, 50+ | 660 | 1.97 | 5.13 | 0.75 | 2.58 | 10.37 | 1.51 | 1.52 | 3.63 | 0.23 |
| 26-34, 26-34 | 604 | 1.66 | 10.16 | 3.30 | 0.50 | 7.55 | 1.26 | 0.50 | 7.77 | 0.36 |
| 26-34, 35-49 | 320 | 8.13 | 15.90 | 5.30 | 1.25 | 3.16 | 0.33 | 0.94 | 2.71 | 0.12 |
| 26-34, 50+ | 177 | 2.26 | 13.69 | 2.88 | 0.56 | 5.68 | 0.28 | 0.00 | 0.00 | 0.00 |
| 35-49, 35-49 | 487 | 5.95 | 24.71 | 8.32 | 1.23 | 7.32 | 1.46 | 0.82 | 5.68 | 0.19 |
| 35-49, 50+ | 272 | 2.57 | 12.66 | 2.81 | 1.84 | 9.61 | 1.24 | 0.74 | 3.39 | 0.18 |
| 50+, 50+ | 663 | 4.37 | 9.62 | 1.32 | 5.28 | 14.60 | 2.27 | 4.07 | 11.55 | 0.69 |
| Pair Race/Ethnicity |  |  |  |  |  |  |  |  |  |  |
| Hispanic or Latino | 3,134 | 3.06 | 16.96 | 3.48 | 1.79 | 9.82 | 1.35 | 1.40 | 5.41 | 0.52 |
| Black or African American | 2,139 | 2.10 | 7.08 | 1.49 | 1.92 | 6.00 | 0.63 | 0.70 | 1.36 | 0.14 |
| White | 11,006 | 1.92 | 8.98 | 2.34 | 1.33 | 6.65 | 1.04 | 0.31 | 3.11 | 0.15 |
| Other | 1,116 | 6.36 | 22.29 | 6.39 | 6.72 | 25.53 | 3.65 | 5.73 | 22.66 | 1.50 |
| White \& Black or African American | 167 | 8.38 | 14.47 | 3.01 | 10.78 | 19.96 | 3.60 | 5.99 | 11.36 | 0.65 |
| White \& Hispanic or Latino | 721 | 1.25 | 2.96 | 0.51 | 1.94 | 11.77 | 1.94 | 0.55 | 8.94 | 0.15 |
| White \& Other | 715 | 0.98 | 2.51 | 0.70 | 0.28 | 0.80 | 0.02 | 0.00 | 0.00 | 0.00 |
| Black or African American \& Hispanic or Latino | 109 | 9.17 | 15.51 | 7.09 | 5.50 | 6.74 | 1.02 | 7.34 | 7.51 | 0.66 |
| Black or African American \& Other | 198 | 6.06 | 23.35 | 8.75 | 1.52 | 8.09 | 2.89 | 0.51 | 5.37 | 1.63 |
| Hispanic or Latino \& Other | 154 | 2.60 | 1.72 | 0.79 | 0.65 | 1.00 | 0.16 | 0.00 | 0.00 | 0.00 |
| Pair Gender |  |  |  |  |  |  |  |  |  |  |
| Male, Male | 4,162 | 3.41 | 13.03 | 3.17 | 2.33 | 7.66 | 1.10 | 0.96 | 3.83 | 0.25 |
| Female, Female | 4,430 | 2.19 | 9.64 | 2.19 | 1.81 | 7.70 | 1.10 | 0.93 | 4.16 | 0.44 |
| Male, Female | 10,867 | 2.21 | 10.65 | 2.73 | 1.70 | 9.23 | 1.38 | 0.91 | 5.57 | 0.31 |
| Household Size |  |  |  |  |  |  |  |  |  |  |
| Two | 4,660 | 0.64 | 3.96 | 1.39 | 0.43 | 1.72 | 0.21 | 0.34 | 1.53 | 0.14 |
| Three | 5,248 | 2.69 | 15.06 | 4.04 | 2.02 | 13.08 | 2.13 | 1.28 | 9.73 | 0.61 |
| Four or More | 9,551 | 3.22 | 12.17 | 2.68 | 2.47 | 9.83 | 1.36 | 1.02 | 4.27 | 0.27 |

Table J. 32012 NSDUH Respondent Pair-Level Proportions of Extreme Values and Outwinsors (continued)

| Domain | $n$ | $\begin{gathered} \text { Before res.pr.ps }{ }^{1} \\ \text { (SDUWT*PRWT10*...*PRWT12) } \end{gathered}$ |  |  | After res.pr.ps ${ }^{1}$ (SDUWT*PRWT10*...*PRWT13) |  |  | Final Weight: After res.pr.ev ${ }^{1}$ (SDUWT*PRWT10*...*PRWT14) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | \% <br> Unweighted | \% Weighted ${ }^{2}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | $\begin{gathered} \% \\ \text { Unweighted } \end{gathered}$ | $\%$ <br> Weighted ${ }^{2}$ | $\stackrel{\%}{\text { Outwinsor }^{3}}$ | $\%$ <br> Unweighted | \% Weighted ${ }^{2}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ |
| Census Region |  |  |  |  |  |  |  |  |  |  |
| Northeast | 3,856 | 2.54 | 12.15 | 2.44 | 1.92 | 7.94 | 0.95 | 1.32 | 5.11 | 0.39 |
| South | 5,867 | 1.79 | 7.96 | 1.49 | 1.18 | 4.14 | 0.35 | 0.48 | 1.46 | 0.09 |
| Midwest | 5,455 | 2.64 | 14.44 | 4.77 | 2.42 | 11.90 | 2.00 | 1.06 | 6.70 | 0.47 |
| West | 4,281 | 3.08 | 11.28 | 2.99 | 2.03 | 13.00 | 2.25 | 1.00 | 8.56 | 0.49 |
| Quarter |  |  |  |  |  |  |  |  |  |  |
| Quarter 1 | 4,613 | 2.99 | 8.43 | 2.00 | 1.99 | 6.66 | 1.02 | 0.89 | 4.69 | 0.30 |
| Quarter 2 | 5,115 | 2.35 | 9.58 | 1.92 | 1.62 | 7.03 | 1.10 | 0.92 | 3.44 | 0.28 |
| Quarter 3 | 5,093 | 1.49 | 9.89 | 3.06 | 1.57 | 12.63 | 1.87 | 0.98 | 7.23 | 0.46 |
| Quarter 4 | 4,638 | 3.13 | 15.67 | 3.86 | 2.31 | 8.33 | 1.11 | 0.91 | 4.65 | 0.26 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 1,537 | 2.47 | 8.16 | 1.91 | 1.63 | 7.42 | 1.05 | 1.11 | 4.54 | 0.30 |
| 10-<50\% | 3,726 | 3.38 | 14.96 | 3.61 | 2.74 | 13.61 | 1.92 | 1.83 | 8.96 | 0.67 |
| <10\% | 14,196 | 2.22 | 9.86 | 2.52 | 1.66 | 7.09 | 1.08 | 0.67 | 3.65 | 0.21 |
| \% Black or African American in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 1,482 | 3.58 | 11.09 | 2.83 | 2.50 | 8.02 | 0.97 | 1.62 | 3.27 | 0.34 |
| 10- $50 \%$ | 2,974 | 3.03 | 10.73 | 2.90 | 2.35 | 9.89 | 1.39 | 1.68 | 5.52 | 0.35 |
| <10\% | 15,003 | 2.24 | 10.91 | 2.66 | 1.70 | 8.46 | 1.28 | 0.71 | 5.06 | 0.32 |
| \% Owner-Occupied DUs ${ }^{1}$ in Segment |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 14,866 | 2.36 | 10.81 | 2.54 | 1.59 | 8.49 | 1.26 | 0.73 | 4.76 | 0.30 |
| 10-<50\% | 3,600 | 3.25 | 11.32 | 3.53 | 2.86 | 9.19 | 1.26 | 1.53 | 5.55 | 0.39 |
| <10\% | 993 | 1.11 | 11.07 | 3.09 | 2.32 | 12.84 | 2.43 | 1.61 | 11.60 | 1.12 |
| Combined Median Rent/Housing Value |  |  |  |  |  |  |  |  |  |  |
| $1^{\text {st }}$ Quintile | 3,227 | 2.05 | 16.76 | 5.92 | 1.77 | 12.66 | 2.44 | 0.90 | 6.50 | 0.54 |
| $2^{\text {nd }}$ Quintile | 4,250 | 2.24 | 7.56 | 1.39 | 1.53 | 6.50 | 0.80 | 0.73 | 4.36 | 0.27 |
| $3^{\text {rd }}$ Quintile | 4,364 | 2.29 | 11.09 | 2.09 | 1.60 | 6.66 | 0.83 | 0.69 | 2.28 | 0.15 |
| $4^{\text {th }}$ Quintile | 4,276 | 2.92 | 9.78 | 2.64 | 2.27 | 7.94 | 1.11 | 1.05 | 4.95 | 0.28 |
| $5{ }^{\text {th }}$ Quintile | 3,342 | 2.78 | 11.32 | 2.57 | 2.18 | 10.99 | 1.62 | 1.35 | 7.41 | 0.46 |
| Population Density |  |  |  |  |  |  |  |  |  |  |
| Large MSA ${ }^{1}$ | 8,192 | 3.15 | 12.57 | 2.97 | 2.45 | 9.78 | 1.33 | 1.39 | 5.05 | 0.36 |
| Medium to Small MSA ${ }^{1}$ | 9,833 | 2.02 | 7.23 | 1.47 | 1.38 | 6.92 | 1.07 | 0.59 | 4.69 | 0.28 |
| Non-MSA, ${ }^{1}$ Urban | 380 | 1.84 | 4.24 | 1.68 | 1.05 | 0.63 | 0.21 | 0.26 | 0.09 | 0.00 |
| Non-MSA, ${ }^{1}$ Rural | 1,054 | 1.42 | 25.72 | 11.58 | 1.99 | 13.15 | 2.95 | 0.66 | 9.05 | 0.39 |
| Group Quarters |  |  |  |  |  |  |  |  |  |  |
| Group | 312 | 6.73 | 23.62 | 6.88 | 8.97 | 20.47 | 3.20 | 2.88 | 10.63 | 0.39 |
| Non-Group | 19,147 | 2.39 | 10.85 | 2.70 | 1.74 | 8.63 | 1.27 | 0.89 | 4.99 | 0.33 |

Table J. 32012 NSDUH Respondent Pair-Level Proportions of Extreme Values and Outwinsors (continued)

| Domain | $n$ | $\begin{gathered} \text { Before res.pr.ps }{ }^{1} \\ \text { (SDUWT*PRWT10*...*PRWT12) } \end{gathered}$ |  |  | After res.pr.ps ${ }^{1}$ (SDUWT*PRWT10*...*PRWT13) |  |  | Final Weight: After res.pr.ev ${ }^{1}$ (SDUWT*PRWT10*...*PRWT14) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\%$ <br> Unweighted | $\%$ <br> Weighted ${ }^{2}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | $\begin{gathered} \% \\ \text { Unweighted } \end{gathered}$ | $\begin{gathered} \% \\ \text { Weighted }{ }^{2} \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ | $\%$ <br> Unweighted | $\begin{gathered} \% \\ \text { Weighted }^{2} \end{gathered}$ | $\begin{gathered} \% \\ \text { Outwinsor }^{3} \end{gathered}$ |
| Pair Relationship Domain ${ }^{4}$ |  |  |  |  |  |  |  |  |  |  |
| Parent-Child (12-14) | 1,981 | 0.91 | 3.04 | 0.71 | 1.31 | 3.74 | 0.58 | 0.56 | 1.53 | 0.11 |
| Parent-Child (12-17) | 3,748 | 0.96 | 3.49 | 0.74 | 1.17 | 3.83 | 0.57 | 0.59 | 1.89 | 0.18 |
| Parent-Child (12-20) | 4,443 | 1.67 | 6.67 | 1.28 | 2.21 | 8.89 | 1.33 | 1.04 | 3.48 | 0.34 |
| Sibling (12-14)-Sibling (15-17) | 2,191 | 0.46 | 4.07 | 1.43 | 0.46 | 2.39 | 0.36 | 0.32 | 2.01 | 0.24 |
| Sibling (12-17)-Sibling (18-25) | 2,477 | 1.53 | 7.04 | 1.82 | 0.89 | 2.81 | 0.54 | 0.61 | 2.11 | 0.31 |
| Spouse-Spouse/Partner-Partner | 3,664 | 2.07 | 9.93 | 2.85 | 1.83 | 8.06 | 1.30 | 0.98 | 6.75 | 0.31 |
| Spouse-Spouse/Partner-Partner with Children (Younger Than 18) | 1,640 | 1.77 | 14.25 | 3.18 | 2.87 | 8.20 | 1.34 | 0.91 | 5.10 | 0.15 |

${ }^{1}$ This step used demographic variables from questionnaire data for all responding person pairs; $\mathrm{DU}=$ dwelling unit, $\mathrm{EV}=$ extreme value adjustment, $\mathrm{MSA}=\mathrm{metropolitan}$ statistical area, $\mathrm{PR}=\mathrm{pair}, \mathrm{PS}=$ poststratification adjustment, Res $=$ respondent, $\mathrm{SDU}=$ screener dwelling unit.
${ }^{2}$ Weighted extreme value proportion: $100 * \sum_{k} w_{e k} / \sum_{k} w_{k}$, where $w_{e k}$ denotes the weight for extreme values, and $w_{k}$ denotes the weight for both extreme values and nonextreme values.
${ }^{3}$ Outwinsor weight proportion: $100^{*} \sum_{k}\left(w_{e k}-b_{k}\right) / \sum_{k} w_{k}$, where $b_{k}$ denotes the winsorized weight.
${ }^{4}$ Parent-child (15-17) was not included here since extreme values were not controlled with this domain

# Appendix K: Evaluation of Calibration Weights: Pair-Level Slippage Rates 

Table K. $1 \quad 2012$ NSDUH Respondent Pair-Level Slippage Rates

| Domain | $n$ | $\begin{gathered} \text { Initial } \\ \text { Total }(I)^{1} \\ \hline \end{gathered}$ | $\begin{gathered} \text { Final } \\ \text { Total }(F)^{2} \\ \hline \end{gathered}$ | Control Total from SDU (C) | $(I-C) / C \%$ | $(F-C) / C \%$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Total | 19,459 | 221,952,167 | 221,952,167 | 221,952,167 | -0.00 | -0.00 |
| Pair Age Group |  |  |  |  |  |  |
| 12-17, 12-17 | 3,668 | 7,255,585 | 7,213,926 | 7,213,926 | 0.58 | -0.00 |
| 12-17, 18-25 | 2,759 | 8,114,464 | 8,172,406 | 8,172,406 | -0.71 | -0.00 |
| 12-17, 26-34 | 658 | 5,360,122 | 5,281,913 | 5,281,913 | 1.48 | -0.00 |
| 12-17, 35-49 | 2,812 | 29,558,177 | 29,592,182 | 29,592,182 | -0.11 | -0.00 |
| 12-17, 50+ | 631 | 11,948,454 | 11,999,450 | 11,999,450 | -0.42 | -0.00 |
| 18-25, 18-25 | 3,901 | 12,207,700 | 12,545,793 | 12,545,793 | -2.69 | -0.00 |
| 18-25, 26-34 | 794 | 8,056,376 | 7,391,400 | 7,391,400 | 9.00 | -0.00 |
| 18-25, 35-49 | 1,053 | 17,787,086 | 17,720,889 | 17,720,889 | 0.37 | -0.00 |
| 18-25, 50+ | 660 | 18,580,624 | 18,321,739 | 18,321,739 | 1.41 | -0.00 |
| 26-34, 26-34 | 604 | 9,996,036 | 10,662,196 | 10,662,196 | -6.25 | 0.00 |
| 26-34, 35-49 | 320 | 8,782,458 | 8,567,386 | 8,567,386 | 2.51 | 0.00 |
| 26-34, 50+ | 177 | 10,685,258 | 11,157,395 | 11,157,395 | -4.23 | 0.00 |
| 35-49, 35-49 | 487 | 18,497,856 | 18,578,012 | 18,578,012 | -0.43 | 0.00 |
| 35-49, 50+ | 272 | 16,664,510 | 16,640,517 | 16,640,517 | 0.14 | -0.00 |
| 50+, 50+ | 663 | 38,457,461 | 38,106,962 | 38,106,962 | 0.92 | 0.00 |
| Pair Race/Ethnicity |  |  |  |  |  |  |
| Hispanic or Latino | 3,134 | 38,960,071 | 38,601,525 | 38,601,525 | 0.93 | -0.00 |
| Black or African American | 2,139 | 24,049,201 | 24,291,619 | 24,291,619 | -1.00 | -0.00 |
| White | 11,006 | 121,400,147 | 123,391,386 | 123,391,386 | -1.61 | 0.00 |
| Other | 1,116 | 15,936,917 | 16,500,923 | 16,500,923 | -3.42 | -0.00 |
| White \& Black or African American | 167 | 1,811,631 | 2,028,598 | 2,028,598 | -10.70 | 0.00 |
| White \& Hispanic or Latino | 721 | 7,863,668 | 8,002,876 | 8,002,876 | -1.74 | 0.00 |
| White \& Other | 715 | 6,519,769 | 5,364,414 | 5,364,414 | 21.54 | 0.00 |
| Black or African American \& Hispanic or Latino | 109 | 1,375,212 | 1,561,670 | 1,561,670 | -11.94 | -0.00 |
| Black or African American \& Other | 198 | 2,236,103 | 1,038,323 | 1,038,323 | 115.36 | -0.00 |
| Hispanic or Latino \& Other | 154 | 1,799,447 | 1,170,834 | 1,170,834 | 53.69 | 0.00 |
| Pair Gender |  |  |  |  |  |  |
| Male, Male | 4,162 | 39,706,191 | 39,891,754 | 39,891,754 | -0.47 | -0.00 |
| Female, Female | 4,430 | 39,862,298 | 39,874,794 | 39,874,794 | -0.03 | -0.00 |
| Male, Female | 10,867 | 142,383,679 | 142,185,618 | 142,185,618 | 0.14 | -0.00 |
| Pair Relationship Domain ${ }^{\text {3,4,5 }}$ |  |  |  |  |  |  |
| Parent-Child (12-14)* | 1,981 | 11,386,620 | 12,505,760 | 12,505,760 | -8.95 | -0.00 |
| Parent-Child (12-17)* | 3,748 | 23,617,276 | 25,045,406 | 25,045,406 | -5.70 | -0.00 |
| Parent-Child (15-17)* | 1,767 | 12,230,656 | 12,539,646 | 12,539,646 | -2.46 | -0.00 |
| Parent-Child (12-20)* | 4,443 | 31,454,617 | 34,378,755 | 34,378,755 | -8.51 | -0.00 |
| Parent*-Child (12-14) | 1,981 | 18,350,450 | 19,342,709 | 19,342,709 | -5.13 | -0.00 |
| Parent*-Child (12-17) | 3,748 | 30,953,375 | 32,235,193 | 32,235,193 | -3.98 | -0.00 |
| Parent*-Child (15-17) | 1,767 | 19,026,718 | 19,334,414 | 19,163,807 | -0.72 | 0.89 |
| Parent*-Child (12-20) | 4,443 | 38,356,007 | 40,347,812 | 40,347,812 | -4.94 | -0.00 |
| Sibling (12-14)-Sibling (15-17)* | 2,191 | 3,929,975 | 4,012,028 | 4,012,028 | -2.05 | -0.00 |
| Sibling (12-17)-Sibling (18-25)* | 2,477 | 5,978,196 | 6,172,680 | 6,172,680 | -3.15 | -0.00 |
| Spouse-Spouse/PartnerPartner | 3,664 | 70,208,199 | 71,421,020 | 71,421,020 | -1.70 | -0.00 |
| Spouse-Spouse/Partner- <br> Partner with Children <br> (Younger Than 18) | 1,640 | 22,125,791 | 29,600,927 | 29,600,927 | -25.25 | -0.00 |

Table K. 12012 NSDUH Respondent Pair-Level Slippage Rates (continued)

| Domain | $n$ | $\begin{gathered} \text { Initial } \\ \text { Total }(I)^{1} \end{gathered}$ | $\begin{gathered} \text { Final } \\ \text { Total }(F)^{2} \end{gathered}$ | Control Total from SDU (C) | $(I-C) / C \%$ | $(F-C) / C \%$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Household Size |  |  |  |  |  |  |
| Two | 4,660 | 54,622,573 | 54,622,573 | 54,622,573 | -0.00 | -0.00 |
| Three | 5,248 | 57,411,679 | 57,411,679 | 57,411,679 | -0.00 | -0.00 |
| Four or More | 9,551 | 109,917,914 | 109,917,914 | 109,917,914 | -0.00 | -0.00 |
| Census Region |  |  |  |  |  |  |
| Northeast | 3,856 | 41,025,207 | 41,025,207 | 41,025,207 | -0.00 | -0.00 |
| South | 5,867 | 79,383,669 | 79,383,669 | 79,383,669 | 0.00 | -0.00 |
| Midwest | 5,455 | 44,933,034 | 44,933,034 | 44,933,034 | -0.00 | -0.00 |
| West | 4,281 | 56,610,256 | 56,610,256 | 56,610,256 | -0.00 | -0.00 |
| Quarter |  |  |  |  |  |  |
| Quarter 1 | 4,613 | 54,726,611 | 54,726,611 | 54,726,611 | -0.00 | -0.00 |
| Quarter 2 | 5,115 | 55,791,727 | 55,791,727 | 55,791,727 | -0.00 | -0.00 |
| Quarter 3 | 5,093 | 55,933,469 | 55,933,469 | 55,933,469 | -0.00 | -0.00 |
| Quarter 4 | 4,638 | 55,500,360 | 55,500,360 | 55,500,360 | -0.00 | -0.00 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |
| 50-100\% | 1,537 | 23,192,230 | 23,192,230 | 23,192,230 | -0.00 | -0.00 |
| 10- $5 \mathbf{5 0 \%}$ | 3,726 | 52,879,054 | 52,879,055 | 52,879,055 | -0.00 | -0.00 |
| <10\% | 14,196 | 145,880,883 | 145,880,883 | 145,880,883 | -0.00 | -0.00 |
| \% Black or African American in Segment |  |  |  |  |  |  |
| 50-100\% | 1,482 | 16,732,625 | 16,732,625 | 16,732,625 | -0.00 | -0.00 |
| 10- $5 \mathbf{5 0 \%}$ | 2,974 | 38,150,858 | 38,150,858 | 38,150,858 | -0.00 | -0.00 |
| <10\% | 15,003 | 167,068,684 | 167,068,684 | 167,068,684 | -0.00 | -0.00 |
| \% Owner-Occupied DUs in Segment |  |  |  |  |  |  |
| 50-100\% | 14,866 | 182,125,104 | 182,125,104 | 182,125,104 | -0.00 | -0.00 |
| 10-<50\% | 3,600 | 35,952,245 | 35,952,245 | 35,952,245 | -0.00 | -0.00 |
| <10\% | 993 | 3,874,817 | 3,874,817 | 3,874,817 | -0.00 | -0.00 |
| Combined Median |  |  |  |  |  |  |
| $1^{\text {st }} \text { Quintile }$ | 3,227 | 30,573,140 | 30,573,140 | 30,573,140 | -0.00 | -0.00 |
| $2^{\text {nd }}$ Quintile | 4,250 | 44,525,897 | 44,525,897 | 44,525,897 | -0.00 | -0.00 |
| $3^{\text {rd }}$ Quintile | 4,364 | 46,327,833 | 46,327,833 | 46,327,833 | -0.00 | -0.00 |
| $4^{\text {th }}$ Quintile | 4,276 | 53,795,269 | 53,795,269 | 53,795,269 | -0.00 | -0.00 |
| $5{ }^{\text {th }}$ Quintile | 3,342 | 46,730,029 | 46,730,029 | 46,730,029 | -0.00 | -0.00 |
| Population Density |  |  |  |  |  |  |
| Large MSA | 8,192 | 122,562,456 | 122,562,456 | 122,562,456 | -0.00 | -0.00 |
| Medium to Small MSA | 9,833 | 87,447,088 | 87,447,088 | 87,447,088 | -0.00 | -0.00 |
| Non-MSA, Urban | 380 | 2,868,028 | 2,868,028 | 2,868,028 | 0.00 | 0.00 |
| Non-MSA, Rural | 1,054 | 9,074,595 | 9,074,595 | 9,074,595 | -0.00 | -0.00 |
| Group Quarters |  |  |  |  |  |  |
| Group | 312 | 798,082 | 798,082 | 798,082 | 0.00 | 0.00 |
| Non-Group | 19,147 | 221,154,085 | 221,154,085 | 221,154,085 | -0.00 | -0.00 |

DU $=$ dwelling unit, MSA $=$ metropolitan statistical area, $\mathrm{SDU}=$ screener dwelling unit.
${ }^{1}$ WT1*...*WT10*PRWT11*...*PRWT13 (before person pair poststratification).
${ }^{2} \mathrm{WT} 1 *$...*WT10*PRWT11*... PRWT 14 (after person pair poststratification).
${ }^{3}$ The member of the pair that is the focus is designated with an asterisk (*).
${ }^{4}$ The parent-child (15-17) pair domains were not controlled for within the modeling and thus have higher slippage rates than the other domains listed. However, since these domains are a subset of other controlled domains, the rates are not large.
${ }^{5}$ Slippage rates were not calculated for the sibling-sibling domains with the younger child as the focus since no household counts for this domain were calculated and are required to construct the appropriate controls totals.

# Appendix L: Evaluation of Calibration Weights: Pair-Level Weight Summary Statistics 

Table L. 12012 NSDUH Selected Pair-Level Weight Summary Statistics

| Domain | $n$ | SDU-Level Weights ${ }^{1}$ (SDUWT: WT1*...*WT9) |  |  |  |  |  | Before sel.pr.ps ${ }^{1}$(SDUWT*PRWT10) |  |  |  |  |  | After sel.pr.ps ${ }^{1}$(SDUWT*PRWT10*PRWT11) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{\mathbf{2}}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ |
| Total | 27,035 | 9 | 441 | 678 | 1,134 | 8,444 | 1.58 | 20 | 1,136 | 2,941 | 7,922 | 6,068,764 | 46.74 | 8 | 1,023 | 2,791 | 7,871 | 1,053,229 | 9.27 |
| Pair Age Group |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 12-17, 12-17 | 4,507 | 16 | 399 | 609 | 1,042 | 7,853 | 1.62 | 20 | 582 | 1,017 | 1,841 | 56,843 | 3.05 | 8 | 516 | 998 | 1,969 | 25,287 | 2.40 |
| 12-17, 18-25 | 3,627 | 18 | 476 | 694 | 1,182 | 8,444 | 1.60 | 26 | 744 | 1,363 | 2,544 | 52,622 | 2.84 | 15 | 709 | 1,451 | 2,854 | 31,529 | 2.22 |
| 12-17, 26-34 | 825 | 32 | 397 | 675 | 1,131 | 5,004 | 1.53 | 248 | 2,487 | 4,469 | 7,333 | 70,973 | 2.29 | 143 | 1,949 | 3,759 | 7,455 | 69,227 | 2.63 |
| 12-17, 35-49 | 3,813 | 20 | 435 | 626 | 1,064 | 8,351 | 1.52 | 141 | 2,952 | 5,274 | 8,974 | 168,153 | 2.41 | 114 | 2,321 | 4,657 | 9,259 | 117,119 | 2.55 |
| 12-17, 50+ | 851 | 56 | 472 | 727 | 1,164 | 7,097 | 1.56 | 878 | 5,875 | 10,166 | 17,187 | 219,397 | 2.32 | 229 | 4,637 | 9,115 | 17,385 | 156,917 | 2.33 |
| 18-25, 18-25 | 5,476 | 9 | 432 | 704 | 1,207 | 7,653 | 1.62 | 32 | 798 | 1,392 | 2,461 | 41,249 | 2.70 | 15 | 589 | 1,253 | 2,936 | 39,303 | 2.52 |
| 18-25, 26-34 | 1,079 | 19 | 477 | 738 | 1,208 | 5,421 | 1.57 | 164 | 2,798 | 4,588 | 7,820 | 179,640 | 3.04 | 105 | 1,819 | 3,421 | 8,319 | 139,322 | 3.17 |
| 18-25, 35-49 | 1,582 | 34 | 491 | 711 | 1,196 | 7,847 | 1.61 | 489 | 3,467 | 6,420 | 12,481 | 245,571 | 3.24 | 180 | 2,804 | 5,903 | 13,746 | 105,508 | 2.60 |
| 18-25, 50+ | 1,074 | 28 | 508 | 704 | 1,240 | 6,541 | 1.54 | 377 | 6,270 | 9,858 | 17,777 | 412,510 | 3.25 | 282 | 5,644 | 10,249 | 20,203 | 187,095 | 2.45 |
| 26-34, 26-34 | 880 | 20 | 442 | 711 | 1,134 | 5,065 | 1.49 | 334 | 5,246 | 8,233 | 13,604 | 789,953 | 7.71 | 167 | 3,867 | 6,865 | 12,454 | 313,383 | 5.05 |
| 26-34, 35-49 | 469 | 43 | 478 | 731 | 1,206 | 3,895 | 1.46 | 571 | 6,016 | 9,618 | 16,559 | 1,594,823 | 18.62 | 412 | 5,391 | 9,152 | 16,689 | 440,141 | 5.90 |
| 26-34, 50+ | 315 | 43 | 466 | 757 | 1,126 | 7,250 | 1.57 | 1,140 | 11,808 | 19,298 | 29,840 | 554,381 | 3.33 | 488 | 10,916 | 21,855 | 40,294 | 701,763 | 3.70 |
| 35-49, 35-49 | 833 | 36 | 445 | 649 | 1,142 | 6,883 | 1.59 | 553 | 5,602 | 9,452 | 15,385 | 1,317,081 | 13.17 | 439 | 5,576 | 10,612 | 18,055 | 1,053,229 | 11.42 |
| 35-49, 50+ | 466 | 59 | 446 | 666 | 1,098 | 3,202 | 1.48 | 1,471 | 9,255 | 15,553 | 27,455 | 858,997 | 5.46 | 1,037 | 9,888 | 20,007 | 39,458 | 868,188 | 4.56 |
| 50+, 50+ | 1,238 | 32 | 456 | 664 | 1,083 | 3,612 | 1.43 | 1,016 | 12,714 | 19,083 | 33,172 | 6,068,764 | 42.37 | 684 | 12,424 | 20,906 | 35,262 | 856,498 | 3.21 |
| Pair Race/Ethnicity |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Hispanic or Latino | 4,176 | 9 | 493 | 815 | 1,390 | 8,444 | 1.61 | 23 | 1,375 | 3,155 | 8,373 | 6,068,764 | 102.68 | 16 | 1,168 | 3,013 | 8,416 | 1,053,229 | 11.77 |
| Black or African American | 2,760 | 42 | 579 | 796 | 1,251 | 5,724 | 1.45 | 59 | 1,341 | 3,285 | 8,689 | 1,594,823 | 23.31 | 17 | 1,240 | 3,266 | 8,011 | 586,052 | 9.23 |
| White | 16,039 | 38 | 428 | 636 | 1,031 | 4,310 | 1.50 | 54 | 1,066 | 2,834 | 7,640 | 5,298,149 | 39.50 | 28 | 988 | 2,654 | 7,599 | 1,028,039 | 9.12 |
| Other | 1,791 | 19 | 256 | 585 | 1,242 | 5,171 | 1.82 | 26 | 976 | 2,726 | 7,732 | 272,377 | 5.93 | 14 | 774 | 2,543 | 7,854 | 293,645 | 6.53 |
| White \& Black or African American | 217 | 53 | 523 | 816 | 1,211 | 3,937 | 1.47 | 193 | 1,697 | 3,581 | 8,188 | 57,878 | 2.48 | 68 | 2,166 | 4,858 | 11,831 | 97,680 | 2.63 |
| White \& Hispanic or Latino | 921 | 18 | 430 | 701 | 1,231 | 6,940 | 1.63 | 20 | 1,148 | 3,423 | 9,307 | 324,611 | 5.04 | 8 | 1,009 | 3,223 | 8,847 | 277,585 | 4.94 |
| White \& Other | 737 | 20 | 255 | 535 | 981 | 5,026 | 1.72 | 91 | 1,014 | 2,737 | 8,205 | 306,037 | 6.09 | 29 | 777 | 2,428 | 7,257 | 453,420 | 8.99 |
| Black or African American \& Hispanic or Latino | 123 | 38 | 696 | 1,166 | 1,904 | 7,579 | 1.81 | 39 | 1,931 | 5,048 | 10,354 | 789,953 | 18.15 | 27 | 1,805 | 4,597 | 9,866 | 285,003 | 7.52 |
| Black or African American \& Other | 126 | 36 | 380 | 563 | 994 | 3,202 | 1.67 | 65 | 1,112 | 2,716 | 6,804 | 371,337 | 17.37 | 15 | 726 | 2,375 | 5,646 | 334,168 | 16.81 |
| Hispanic or Latino \& Other | 145 | 20 | 263 | 559 | 1,137 | 5,421 | 1.91 | 85 | 942 | 2,697 | 7,865 | 72,609 | 3.69 | 120 | 1,026 | 3,407 | 8,538 | 88,347 | 3.58 |

Table L. 12012 NSDUH Selected Pair-Level Weight Summary Statistics (continued)

| Domain | $n$ | SDU-Level Weights ${ }^{1}$ (SDUWT: WT1*...*WT9) |  |  |  |  |  | Before sel.pr.ps ${ }^{1}$ (SDUWT*PRWT10) |  |  |  |  |  | After sel.pr.ps ${ }^{1}$(SDUWT*PRWT10*PRWT11) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ |
| Pair Gender |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Male, Male | 5,928 | 18 | 446 | 705 | 1,137 | 7,154 | 1.56 | 20 | 1,020 | 2,442 | 6,430 | 5,298,149 | 99.54 | 8 | 995 | 2,517 | 6,644 | 856,498 | 8.36 |
| Female, Female | 5,813 | 9 | 426 | 651 | 1,126 | 8,359 | 1.61 | 23 | 1,032 | 2,588 | 6,877 | 348,424 | 5.05 | 15 | 924 | 2,455 | 6,648 | 527,688 | 6.63 |
| Male, Female | 15,294 | 18 | 447 | 676 | 1,136 | 8,444 | 1.58 | 31 | 1,242 | 3,295 | 9,094 | 6,068,764 | 40.82 | 14 | 1,078 | 3,126 | 9,128 | 1,053,229 | 9.71 |
| Household Size |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Two | 6,855 | 20 | 428 | 663 | 1,087 | 6,940 | 1.54 | 38 | 1,320 | 4,305 | 11,582 | 105,073 | 2.48 | 15 | 862 | 2,867 | 10,535 | 192,611 | 3.36 |
| Three | 7,398 | 18 | 447 | 677 | 1,117 | 8,444 | 1.58 | 20 | 1,270 | 3,134 | 6,962 | 5,298,149 | 79.82 | 8 | 1,109 | 3,072 | 7,076 | 1,028,039 | 10.54 |
| Four or More | 12,782 | 9 | 442 | 689 | 1,176 | 8,359 | 1.60 | 23 | 1,014 | 2,371 | 7,069 | 6,068,764 | 52.53 | 14 | 1,069 | 2,633 | 7,458 | 1,053,229 | 11.36 |
| Census Region |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Northeast | 5,685 | 18 | 335 | 605 | 843 | 7,128 | 1.61 | 20 | 998 | 2,548 | 7,060 | 451,843 | 6.11 | 8 | 784 | 2,357 | 6,775 | 586,033 | 8.38 |
| South | 7,934 | 9 | 659 | 958 | 1,348 | 7,138 | 1.40 | 23 | 1,545 | 4,149 | 10,607 | 6,068,764 | 79.32 | 15 | 1,332 | 3,723 | 10,044 | 1,053,229 | 8.39 |
| Midwest | 7,515 | 33 | 451 | 548 | 728 | 4,474 | 1.36 | 64 | 896 | 2,504 | 6,270 | 1,594,823 | 23.29 | 29 | 960 | 2,354 | 5,998 | 1,028,039 | 11.16 |
| West | 5,901 | 20 | 281 | 749 | 1,548 | 8,444 | 1.74 | 68 | 1,124 | 2,913 | 9,026 | 1,246,856 | 11.04 | 35 | 1,015 | 2,955 | 9,102 | 868,188 | 8.75 |
| Quarter |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Quarter1 | 6,334 | 21 | 496 | 733 | 1,173 | 7,847 | 1.54 | 26 | 1,205 | 2,948 | 7,968 | 6,068,764 | 90.25 | 14 | 1,182 | 3,093 | 8,706 | 674,933 | 6.79 |
| Quarter2 | 7,183 | 18 | 409 | 619 | 1,092 | 7,853 | 1.58 | 31 | 1,069 | 2,743 | 7,529 | 789,953 | 6.39 | 16 | 966 | 2,669 | 7,602 | 456,537 | 5.56 |
| Quarter3 | 7,012 | 9 | 434 | 631 | 1,141 | 7,138 | 1.59 | 20 | 1,077 | 2,911 | 7,852 | 1,246,856 | 12.56 | 8 | 895 | 2,545 | 7,289 | 868,188 | 12.34 |
| Quarter 4 | 6,506 | 16 | 476 | 711 | 1,145 | 8,444 | 1.60 | 23 | 1,179 | 3,230 | 8,450 | 5,298,149 | 63.92 | 15 | 1,071 | 2,967 | 8,090 | 1,053,229 | 12.19 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 2,105 | 55 | 583 | 1,046 | 1,488 | 6,365 | 1.36 | 85 | 1,810 | 4,352 | 10,864 | 6,068,764 | 114.14 | 42 | 1,465 | 3,748 | 10,433 | 674,933 | 7.42 |
| 10-<50\% | 5,112 | 9 | 570 | 895 | 1,480 | 8,444 | 1.52 | 23 | 1,571 | 3,777 | 10,099 | 5,298,149 | 57.95 | 16 | 1,455 | 3,901 | 10,200 | 1,053,229 | 9.93 |
| <10\% | 19,818 | 16 | 389 | 613 | 987 | 7,853 | 1.58 | 20 | 1,008 | 2,647 | 7,213 | 1,594,823 | 11.23 | 8 | 914 | 2,467 | 7,090 | 1,028,039 | 8.91 |
| \% Black or African American in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 1,905 | 9 | 556 | 763 | 1,172 | 7,128 | 1.59 | 23 | 1,312 | 3,254 | 8,310 | 371,337 | 6.18 | 15 | 1,180 | 3,228 | 8,308 | 527,688 | 7.76 |
| 10- $50 \%$ | 4,030 | 22 | 541 | 832 | 1,337 | 7,847 | 1.50 | 31 | 1,436 | 3,529 | 9,224 | 6,068,764 | 95.73 | 27 | 1,237 | 3,357 | 9,426 | 800,116 | 8.03 |
| $<10 \%$ | 21,100 | 18 | 408 | 639 | 1,089 | 8,444 | 1.58 | 20 | 1,071 | 2,821 | 7,657 | 5,298,149 | 33.09 | 8 | 979 | 2,674 | 7,618 | 1,053,229 | 9.72 |
| \% Owner-Occupied DUs ${ }^{1}$ in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 20,778 | 18 | 441 | 663 | 1,116 | 8,444 | 1.57 | 20 | 1,155 | 3,052 | 8,163 | 6,068,764 | 52.77 | 8 | 1,117 | 3,059 | 8,514 | 1,053,229 | 9.35 |
| 10- $50 \%$ | 4,889 | 16 | 429 | 715 | 1,145 | 7,154 | 1.57 | 23 | 1,076 | 2,680 | 7,260 | 1,594,823 | 15.44 | 15 | 914 | 2,535 | 7,070 | 586,033 | 7.27 |
| <10\% | 1,368 | 9 | 475 | 826 | 1,415 | 8,351 | 1.63 | 29 | 1,079 | 2,343 | 6,980 | 179,640 | 3.88 | 15 | 409 | 1,123 | 2,923 | 66,093 | 4.36 |
| Combined Median Rent/Housing Value |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $1^{\text {st }}$ Quintile | 4,264 | 19 | 392 | 609 | 966 | 8,351 | 1.62 | 26 | 991 | 2,568 | 6,844 | 6,068,764 | 131.76 | 15 | 864 | 2,304 | 6,422 | 1,028,039 | 15.85 |
| $2{ }^{\text {nd }}$ Quintile | 5,758 | 18 | 411 | 641 | 1,108 | 8,444 | 1.68 | 20 | 1,058 | 2,749 | 7,573 | 573,547 | 6.64 | 8 | 959 | 2,665 | 7,281 | 586,052 | 8.64 |
| $3{ }^{\text {rd }}$ Quintile | 6,136 | 9 | 389 | 617 | 1,036 | 6,940 | 1.60 | 29 | 1,050 | 2,804 | 7,334 | 1,246,856 | 13.44 | 15 | 868 | 2,487 | 6,999 | 1,053,229 | 10.97 |
| $4^{\text {th }}$ Quintile | 6,013 | 16 | 478 | 750 | 1,210 | 7,853 | 1.54 | 23 | 1,266 | 3,248 | 8,839 | 5,298,149 | 65.38 | 14 | 1,161 | 3,121 | 8,853 | 868,188 | 8.61 |
| $5{ }^{\text {th }}$ Quintile | 4,864 | 25 | 530 | 809 | 1,264 | 8,359 | 1.44 | 31 | 1,361 | 3,320 | 9,477 | 1,594,823 | 11.59 | 15 | 1,314 | 3,596 | 10,130 | 701,763 | 5.48 |

Table L. 12012 NSDUH Selected Pair-Level Weight Summary Statistics (continued)

| Domain | $n$ | SDU-Level Weights ${ }^{1}$ (SDUWT: WT1*...*WT9) |  |  |  |  |  | Before sel.pr.ps ${ }^{1}$ (SDUWT*PRWT10) |  |  |  |  |  | After sel.pr.ps ${ }^{1}$(SDUWT*PRWT10*PRWT11) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | $\mathrm{UWE}^{3}$ |
| Population Density |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Large MSA ${ }^{1}$ | 11,649 | 9 | 586 | 869 | 1,395 | 8,359 | 1.42 | 23 | 1,593 | 3,974 | 10,354 | 5,298,149 | 33.22 | 15 | 1,538 | 3,988 | 10,441 | 1,053,229 | 8.13 |
| Medium to Small MSA ${ }^{1}$ | 13,392 | 18 | 319 | 561 | 914 | 8,444 | 1.66 | 20 | 920 | 2,352 | 6,521 | 573,547 | 6.09 | 8 | 783 | 2,176 | 6,263 | 736,748 | 8.04 |
| Non-MSA, ${ }^{1}$ Urban | 511 | 20 | 223 | 492 | 798 | 3,628 | 1.71 | 79 | 787 | 2,016 | 5,989 | 6,068,764 | 256.70 | 31 | 607 | 1,591 | 4,576 | 674,933 | 30.68 |
| Non-MSA, ${ }^{1}$ Rural | 1,483 | 20 | 190 | 472 | 829 | 4,220 | 1.77 | 64 | 720 | 1,878 | 5,562 | 1,317,081 | 35.03 | 35 | 597 | 1,693 | 5,171 | 1,028,039 | 23.96 |
| Group Quarters |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Group | 386 | 25 | 235 | 422 | 934 | 3,853 | 1.99 | 38 | 455 | 983 | 1,865 | 21,866 | 3.52 | 15 | 349 | 977 | 2,295 | 23,423 | 3.36 |
| Non-Group | 26,649 | 9 | 446 | 680 | 1,136 | 8,444 | 1.58 | 20 | 1,155 | 2,984 | 8,033 | 6,068,764 | 46.38 | 8 | 1,042 | 2,843 | 7,985 | 1,053,229 | 9.20 |

${ }^{1}$ This step used demographic variables from screener data for all selected person pairs; DU = dwelling unit, MSA = metropolitan statistical area, PR = pair, PS = poststratification, SDU = screener dwelling unit, Sel = selected.
${ }^{2}$ Q1 and Q3 refer to the first and third quartile of the weight distribution.
${ }^{3}$ Unequal weighting effect (UWE) is defined as $1+[(n-1) / n] * C V^{2}$, where $C V=$ coefficient of variation of weights.

Table L. 2012 NSDUH Respondent Pair-Level Weight Summary Statistics (res.pr.nr)

| Domain | $n$ | Before res.pr.nr ${ }^{1}$(SDUWT*PRWT10*PRWT11) |  |  |  |  |  | After res.pr.nr ${ }^{1}$(SDUWT*PRWT10*...*PRWT12) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{\mathbf{2}}$ | Max | UWE ${ }^{3}$ |
| Total | 19,459 | 8 | 965 | 2,556 | 7,018 | 1,053,229 | 9.79 | 10 | 1,129 | 3,171 | 9,270 | 1,550,385 | 10.18 |
| Pair Age Group |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 12-17, 12-17 | 3,666 | 8 | 515 | 1,007 | 1,973 | 17,520 | 2.32 | 10 | 590 | 1,169 | 2,375 | 34,664 | 2.58 |
| 12-17, 18-25 | 2,778 | 15 | 716 | 1,477 | 2,953 | 21,346 | 2.17 | 15 | 854 | 1,813 | 3,689 | 37,794 | 2.39 |
| 12-17, 26-34 | 653 | 143 | 1,901 | 3,700 | 7,161 | 69,227 | 2.69 | 143 | 2,053 | 4,448 | 8,869 | 151,613 | 3.23 |
| 12-17, 35-49 | 2,816 | 174 | 2,326 | 4,648 | 9,277 | 117,119 | 2.59 | 176 | 2,867 | 5,831 | 12,576 | 148,919 | 2.68 |
| 12-17, 50+ | 627 | 229 | 4,611 | 8,883 | 17,324 | 156,917 | 2.42 | 258 | 5,329 | 10,817 | 22,909 | 230,684 | 2.70 |
| 18-25, 18-25 | 3,976 | 15 | 587 | 1,252 | 2,936 | 39,303 | 2.56 | 17 | 645 | 1,528 | 3,893 | 44,826 | 2.84 |
| 18-25, 26-34 | 742 | 105 | 1,745 | 3,332 | 8,192 | 91,433 | 3.02 | 113 | 1,973 | 3,892 | 10,805 | 186,312 | 3.68 |
| 18-25, 35-49 | 1,058 | 180 | 2,675 | 5,658 | 13,746 | 105,508 | 2.64 | 183 | 3,407 | 8,114 | 21,008 | 184,727 | 2.83 |
| 18-25, 50+ | 650 | 282 | 5,534 | 9,839 | 19,855 | 148,367 | 2.33 | 289 | 7,718 | 15,800 | 34,492 | 296,605 | 2.53 |
| 26-34, 26-34 | 597 | 396 | 4,040 | 6,960 | 12,877 | 313,383 | 5.22 | 402 | 4,375 | 7,653 | 15,191 | 481,267 | 6.43 |
| 26-34, 35-49 | 305 | 412 | 4,837 | 8,543 | 15,131 | 382,959 | 4.50 | 412 | 7,194 | 13,271 | 29,726 | 776,131 | 5.09 |
| 26-34, 50+ | 178 | 989 | 10,601 | 21,015 | 38,170 | 701,763 | 4.02 | 991 | 12,816 | 31,414 | 78,144 | 919,088 | 3.14 |
| 35-49, 35-49 | 489 | 439 | 5,391 | 10,342 | 18,217 | 1,053,229 | 13.26 | 463 | 7,345 | 14,279 | 29,236 | 1,550,385 | 10.91 |
| 35-49, 50+ | 266 | 1,037 | 9,310 | 17,241 | 34,884 | 800,116 | 4.44 | 1,038 | 12,826 | 28,815 | 75,129 | 1,024,681 | 3.17 |
| 50+, 50+ | 658 | 753 | 13,107 | 20,833 | 34,426 | 674,933 | 3.29 | 1,612 | 20,499 | 36,877 | 69,712 | 876,058 | 2.68 |
| Pair Race/Ethnicity |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Hispanic or Latino | 3,084 | 22 | 1,110 | 2,795 | 7,446 | 1,053,229 | 14.92 | 25 | 1,237 | 3,235 | 9,954 | 1,323,156 | 13.93 |
| Black or African American | 2,210 | 21 | 1,175 | 3,050 | 7,271 | 586,052 | 8.76 | 21 | 1,270 | 3,630 | 8,954 | 876,058 | 9.62 |
| White | 11,284 | 28 | 927 | 2,403 | 6,635 | 1,028,039 | 8.77 | 38 | 1,107 | 3,054 | 8,977 | 1,550,385 | 9.94 |
| Other | 1,175 | 14 | 641 | 2,074 | 6,224 | 249,217 | 6.58 | 14 | 883 | 2,804 | 9,328 | 337,536 | 6.94 |
| White \& Black or African American | 166 | 68 | 2,086 | 4,813 | 11,751 | 58,915 | 2.20 | 86 | 2,298 | 6,930 | 16,409 | 77,753 | 2.55 |
| White \& Hispanic or Latino | 692 | 8 | 1,010 | 3,072 | 8,420 | 277,585 | 5.25 | 10 | 1,047 | 3,268 | 10,111 | 481,267 | 7.61 |
| White \& Other | 552 | 29 | 760 | 2,335 | 6,975 | 118,112 | 4.53 | 29 | 816 | 2,564 | 8,087 | 355,652 | 8.32 |
| Black or African American \& Hispanic or Latino | 91 | 27 | 1,805 | 4,597 | 9,157 | 285,003 | 8.23 | 27 | 1,965 | 5,172 | 12,415 | 285,060 | 6.03 |
| Black or African American \& Other | 99 | 15 | 722 | 2,323 | 4,297 | 28,574 | 2.69 | 17 | 1,070 | 4,113 | 11,659 | 106,101 | 3.28 |
| Hispanic or Latino \& Other | 106 | 120 | 1,132 | 3,393 | 6,819 | 37,041 | 2.69 | 126 | 1,363 | 4,407 | 10,976 | 104,183 | 3.84 |
| Pair Gender |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Male, Male | 4,173 | 8 | 941 | 2,356 | 5,891 | 342,223 | 5.71 | 10 | 1,130 | 2,898 | 8,040 | 488,930 | 7.38 |
| Female, Female | 4,425 | 15 | 882 | 2,362 | 6,170 | 249,217 | 4.74 | 17 | 1,024 | 2,850 | 7,945 | 285,780 | 6.13 |
| Male, Female | 10,861 | 14 | 1,012 | 2,769 | 7,844 | 1,053,229 | 11.30 | 14 | 1,170 | 3,457 | 10,446 | 1,550,385 | 11.15 |
| Household Size |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Two | 4,660 | 15 | 789 | 2,337 | 8,759 | 192,611 | 3.54 | 17 | 856 | 2,654 | 11,109 | 292,988 | 5.19 |
| Three | 5,248 | 8 | 992 | 2,765 | 6,318 | 1,028,039 | 11.78 | 10 | 1,189 | 3,534 | 8,651 | 1,550,385 | 12.87 |
| Four or More | 9,551 | 14 | 1,046 | 2,514 | 6,792 | 1,053,229 | 11.40 | 14 | 1,264 | 3,187 | 8,995 | 1,323,156 | 11.36 |

Table L. 22012 NSDUH Respondent Pair-Level Weight Summary Statistics (res.pr.nr) (continued)

| Domain | $n$ | Before res.pr.nr $^{1}$(SDUWT*PRWT10*PRWT11) |  |  |  |  |  | After res.pr.nr ${ }^{1}$(SDUWT*PRWT10*... ${ }^{\text {PRWT12 }}$ ) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ |
| Census Region |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Northeast | 3,856 | 8 | 691 | 2,056 | 5,591 | 586,033 | 8.98 | 10 | 813 | 2,697 | 8,171 | 1,124,135 | 10.75 |
| South | 5,867 | 15 | 1,238 | 3,355 | 8,957 | 1,053,229 | 8.65 | 17 | 1,403 | 4,051 | 11,527 | 1,323,156 | 8.68 |
| Midwest | 5,455 | 29 | 922 | 2,184 | 5,400 | 1,028,039 | 13.98 | 30 | 1,083 | 2,645 | 6,808 | 1,550,385 | 14.22 |
| West | 4,281 | 35 | 986 | 2,738 | 8,023 | 800,116 | 8.29 | 36 | 1,161 | 3,392 | 10,802 | 1,024,681 | 8.68 |
| Quarter |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Quarter1 | 4,613 | 14 | 1,138 | 2,868 | 7,787 | 674,933 | 6.90 | 14 | 1,351 | 3,582 | 10,437 | 854,597 | 7.22 |
| Quarter2 | 5,115 | 22 | 912 | 2,501 | 6,737 | 285,003 | 4.94 | 25 | 1,071 | 3,131 | 8,922 | 417,070 | 6.38 |
| Quarter3 | 5,093 | 8 | 830 | 2,311 | 6,395 | 800,116 | 12.88 | 10 | 963 | 2,769 | 8,219 | 1,024,681 | 12.79 |
| Quarter 4 | 4,638 | 15 | 1,000 | 2,635 | 7,081 | 1,053,229 | 14.19 | 15 | 1,189 | 3,328 | 9,797 | 1,550,385 | 14.07 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 1,537 | 42 | 1,399 | 3,267 | 8,917 | 674,933 | 8.97 | 42 | 1,594 | 4,086 | 11,989 | 854,597 | 9.02 |
| 10-<50\% | 3,726 | 34 | 1,311 | 3,552 | 8,987 | 1,053,229 | 11.03 | 35 | 1,482 | 4,402 | 11,789 | 1,323,156 | 10.37 |
| <10\% | 14,196 | 8 | 860 | 2,272 | 6,236 | 1,028,039 | 8.79 | 10 | 1,024 | 2,844 | 8,288 | 1,550,385 | 9.90 |
| \% Black or African American in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 1,482 | 15 | 1,076 | 2,881 | 7,396 | 342,223 | 5.53 | 17 | 1,232 | 3,602 | 10,074 | 488,930 | 6.27 |
| 10-<50\% | 2,974 | 29 | 1,178 | 3,074 | 8,215 | 800,116 | 9.87 | 29 | 1,287 | 3,799 | 10,451 | 1,024,681 | 9.57 |
| $<10 \%$ | 15,003 | 8 | 923 | 2,428 | 6,752 | 1,053,229 | 10.14 | 10 | 1,089 | 3,042 | 8,915 | 1,550,385 | 10.70 |
| \% Owner-Occupied DUs ${ }^{1}$ in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 14,866 | 8 | 1,055 | 2,758 | 7,528 | 1,053,229 | 9.93 | 10 | 1,248 | 3,505 | 10,121 | 1,550,385 | 9.97 |
| 10-<50\% | 3,600 | 15 | 855 | 2,336 | 6,170 | 586,033 | 7.50 | 17 | 991 | 2,731 | 8,123 | 1,124,135 | 9.53 |
| <10\% | 993 | 15 | 419 | 1,065 | 2,810 | 66,093 | 4.60 | 17 | 458 | 1,220 | 3,258 | 133,448 | 6.99 |
| Combined Median <br> Rent/Housing Value |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\mathbf{1}^{\text {st }}$ Quintile | 3,227 | 15 | 853 | 2,199 | 5,701 | 1,028,039 | 20.71 | 15 | 955 | 2,526 | 6,979 | 1,550,385 | 21.52 |
| $2^{\text {nd }}$ Quintile | 4,250 | 8 | 911 | 2,453 | 6,481 | 586,052 | 9.89 | 10 | 1,048 | 2,919 | 8,231 | 1,124,135 | 12.13 |
| $3{ }^{\text {rd }}$ Quintile | 4,364 | 15 | 822 | 2,323 | 6,197 | 1,053,229 | 12.07 | 17 | 994 | 2,839 | 8,440 | 1,323,156 | 11.06 |
| $4^{\text {th }}$ Quintile | 4,276 | 14 | 1,077 | 2,772 | 7,763 | 313,383 | 4.68 | 14 | 1,285 | 3,558 | 10,977 | 355,652 | 6.06 |
| $5^{\text {th }}$ Quintile | 3,342 | 15 | 1,216 | 3,258 | 8,732 | 701,763 | 6.12 | 17 | 1,488 | 4,266 | 12,366 | 919,088 | 6.64 |
| Population Density |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Large MSA ${ }^{1}$ | 8,192 | 15 | 1,418 | 3,593 | 8,988 | 1,053,229 | 8.36 | 17 | 1,717 | 4,648 | 12,779 | 1,323,156 | 8.06 |
| Medium to Small MSA ${ }^{1}$ | 9,833 | 8 | 749 | 2,019 | 5,606 | 586,052 | 7.84 | 10 | 874 | 2,423 | 7,178 | 876,058 | 9.61 |
| Non-MSA, ${ }^{1}$ Urban | 380 | 48 | 623 | 1,562 | 4,150 | 674,933 | 42.44 | 52 | 742 | 1,980 | 5,241 | 854,597 | 35.92 |
| Non-MSA, ${ }^{1}$ Rural | 1,054 | 35 | 593 | 1,682 | 4,927 | 1,028,039 | 34.20 | 38 | 717 | 1,978 | 6,264 | 1,550,385 | 35.14 |
| Group Quarters |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Group | 312 | 15 | 349 | 872 | 1,924 | 23,423 | 3.63 | 17 | 429 | 1,061 | 2,898 | 34,734 | 4.01 |
| Non-Group | 19,147 | 8 | 983 | 2,606 | 7,093 | 1,053,229 | 9.72 | 10 | 1,150 | 3,234 | 9,414 | 1,550,385 | 10.09 |

[^23]Table L. 32012 NSDUH Respondent Pair-Level Weight Summary Statistics (res.pr.ps and res.pr.ev)

| Domain | $n$ | Before res.pr.ps ${ }^{1}$(SDUWT*PRWT10*...*PRWT12) |  |  |  |  |  | After res.pr.ps ${ }^{1}$(SDUWT*PRWT10*...*PRWT13) |  |  |  |  |  | Final Weight: After res.pr.ev ${ }^{1}$ (SDUWT*PRWT10*...*PRWT14) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ |
| Total | 19,459 | 10 | 1,129 | 3,171 | 9,270 | 1,550,385 | 10.18 | 4 | 1,082 | 3,123 | 9,214 | 1,234,015 | 9.58 | 3 | 1,068 | 3,112 | 9,208 | 1,237,337 | 9.36 |
| Pair Age Group |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 12-17, 12-17 | 3,668 | 10 | 593 | 1,177 | 2,391 | 34,664 | 2.58 | 4 | 542 | 1,168 | 2,393 | 30,611 | 2.50 | 3 | 534 | 1,162 | 2,395 | 32,280 | 2.52 |
| 12-17, 18-25 | 2,759 | 15 | 847 | 1,791 | 3,683 | 37,794 | 2.42 | 6 | 815 | 1,821 | 3,816 | 25,541 | 2.31 | 6 | 813 | 1,829 | 3,856 | 24,364 | 2.31 |
| 12-17, 26-34 | 658 | 143 | 2,030 | 4,412 | 8,943 | 151,613 | 3.25 | 89 | 1,938 | 4,295 | 9,294 | 102,559 | 3.01 | 87 | 1,884 | 4,201 | 9,418 | 112,015 | 3.09 |
| 12-17, 35-49 | 2,812 | 176 | 2,886 | 5,847 | 12,501 | 148,919 | 2.69 | 130 | 2,567 | 5,540 | 12,475 | 142,220 | 2.83 | 105 | 2,514 | 5,526 | 12,426 | 143,455 | 2.82 |
| 12-17, 50+ | 631 | 258 | 5,228 | 10,765 | 22,907 | 230,684 | 2.72 | 184 | 4,418 | 10,190 | 23,400 | 196,486 | 2.77 | 179 | 4,462 | 10,151 | 23,711 | 192,088 | 2.71 |
| 18-25, 18-25 | 3,901 | 17 | 644 | 1,524 | 3,865 | 44,826 | 2.83 | 15 | 580 | 1,562 | 4,185 | 44,635 | 2.77 | 14 | 568 | 1,556 | 4,193 | 32,458 | 2.71 |
| 18-25, 26-34 | 794 | 52 | 1,764 | 3,722 | 10,598 | 323,710 | 4.78 | 40 | 1,516 | 3,814 | 9,782 | 153,541 | 3.94 | 36 | 1,448 | 3,764 | 9,723 | 151,757 | 3.98 |
| 18-25, 35-49 | 1,053 | 183 | 3,389 | 8,145 | 21,593 | 184,727 | 2.82 | 172 | 3,071 | 7,900 | 21,308 | 177,102 | 2.80 | 158 | 3,082 | 7,802 | 21,369 | 154,716 | 2.75 |
| 18-25, 50+ | 660 | 289 | 7,649 | 15,463 | 33,878 | 296,605 | 2.59 | 432 | 6,714 | 15,392 | 33,732 | 299,472 | 2.64 | 402 | 6,780 | 15,470 | 34,493 | 279,736 | 2.57 |
| 26-34, 26-34 | 604 | 207 | 3,971 | 7,400 | 14,503 | 481,267 | 6.58 | 232 | 3,844 | 8,467 | 16,566 | 727,153 | 7.05 | 216 | 3,645 | 8,417 | 16,631 | 624,327 | 6.46 |
| 26-34, 35-49 | 320 | 412 | 6,938 | 12,531 | 27,491 | 776,131 | 5.14 | 137 | 5,215 | 12,267 | 24,629 | 1,004,156 | 7.79 | 130 | 5,035 | 11,784 | 24,653 | 1,011,952 | 7.98 |
| 26-34, 50+ | 177 | 991 | 12,555 | 30,990 | 77,130 | 919,088 | 3.28 | 812 | 12,545 | 30,285 | 91,722 | 633,208 | 2.64 | 756 | 12,324 | 30,567 | 91,602 | 549,404 | 2.55 |
| 35-49, 35-49 | 487 | 463 | 7,217 | 14,279 | 29,236 | 1,550,385 | 10.96 | 164 | 6,271 | 15,001 | 32,702 | 1,234,015 | 8.51 | 155 | 6,301 | 15,391 | 32,738 | 1,237,337 | 8.28 |
| 35-49, 50+ | 272 | 1,038 | 12,849 | 27,776 | 71,681 | 1,024,681 | 3.21 | 671 | 10,431 | 27,472 | 74,724 | 760,439 | 2.95 | 620 | 10,194 | 26,379 | 74,748 | 612,413 | 2.84 |
| 50+, 50+ | 663 | 1,612 | 20,423 | 36,884 | 70,196 | 876,058 | 2.67 | 1,647 | 18,150 | 35,352 | 69,493 | 818,783 | 2.72 | 1,632 | 18,609 | 35,696 | 71,197 | 815,689 | 2.68 |
| Pair Race/Ethnicity |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Hispanic or Latino | 3,134 | 25 | 1,251 | 3,295 | 10,003 | 1,323,156 | 13.84 | 15 | 1,218 | 3,280 | 9,992 | 1,234,015 | 13.21 | 15 | 1,201 | 3,234 | 9,988 | 1,237,337 | 12.94 |
| Black or African American | 2,139 | 21 | 1,286 | 3,663 | 9,282 | 876,058 | 9.71 | 15 | 1,253 | 3,637 | 9,153 | 818,783 | 9.44 | 15 | 1,247 | 3,647 | 9,107 | 815,689 | 9.47 |
| White | 11,006 | 29 | 1,102 | 3,054 | 8,995 | 1,550,385 | 10.05 | 23 | 1,093 | 3,116 | 9,136 | 1,010,083 | 8.72 | 21 | 1,078 | 3,108 | 9,117 | 1,011,952 | 8.50 |
| Other | 1,116 | 14 | 958 | 3,017 | 10,103 | 337,536 | 6.66 | 4 | 914 | 3,081 | 10,892 | 374,107 | 7.06 | 3 | 926 | 3,165 | 10,780 | 346,839 | 6.88 |
| White \& Black or African American | 167 | 86 | 1,809 | 6,079 | 14,148 | 77,753 | 2.61 | 124 | 1,826 | 7,641 | 16,843 | 100,462 | 2.57 | 123 | 1,810 | 7,636 | 14,991 | 99,568 | 2.59 |
| White \& Hispanic or Latino | 721 | 10 | 1,065 | 2,958 | 9,380 | 481,267 | 8.22 | 6 | 856 | 2,589 | 9,030 | 727,153 | 11.44 | 6 | 822 | 2,527 | 9,100 | 624,327 | 10.20 |
| White \& Other | 715 | 16 | 881 | 2,689 | 7,897 | 240,232 | 6.02 | 8 | 608 | 1,904 | 6,289 | 347,902 | 8.03 | 8 | 568 | 1,854 | 6,261 | 370,874 | 8.47 |
| Black or African American \& Hispanic or Latino | 109 | 98 | 1,474 | 4,559 | 9,426 | 140,505 | 4.00 | 166 | 1,473 | 4,575 | 10,998 | 165,865 | 4.66 | 145 | 1,364 | 4,448 | 10,885 | 180,929 | 4.82 |
| Black or African American \& Other | 198 | 17 | 1,033 | 3,397 | 7,930 | 181,458 | 6.32 | 5 | 392 | 1,408 | 3,562 | 91,049 | 5.96 | 4 | 349 | 1,388 | 3,582 | 90,913 | 5.98 |
| Hispanic or Latino \& Other | 154 | 55 | 1,030 | 3,551 | 9,841 | 206,329 | 6.36 | 32 | 539 | 2,016 | 6,085 | 109,860 | 6.18 | 30 | 521 | 1,973 | 6,135 | 109,119 | 6.22 |

Table L. 32012 NSDUH Respondent Pair-Level Weight Summary Statistics (res.pr.ps and res.pr.ev) (continued)

| Domain | $n$ | Before res.pr.ps ${ }^{1}$(SDUWT*PRWT10*... ${ }^{*}$ PRWT12) |  |  |  |  |  | After res.pr.ps ${ }^{1}$(SDUWT*PRWT10*... ${ }^{*}$ PRWT13) |  |  |  |  |  | Final Weight: After res.pr.ev ${ }^{1}$ (SDUWT*PRWT10*...*PRWT14) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE $^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | $\mathbf{U W E}^{\mathbf{3}}$ |
| Pair Gender |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Male, Male | 4,162 | 10 | 1,127 | 2,892 | 8,027 | 488,930 | 7.41 | 6 | 1,078 | 2,843 | 8,046 | 501,086 | 7.64 | 6 | 1,057 | 2,848 | 8,133 | 502,540 | 7.60 |
| Female, Female | 4,430 | 17 | 1,023 | 2,852 | 7,945 | 285,780 | 6.13 | 13 | 965 | 2,752 | 7,804 | 374,107 | 6.57 | 12 | 961 | 2,741 | 7,772 | 346,839 | 6.53 |
| Male, Female | 10,867 | 14 | 1,170 | 3,460 | 10,465 | 1,550,385 | 11.13 | 4 | 1,141 | 3,386 | 10,381 | 1,234,015 | 10.20 | 3 | 1,121 | 3,357 | 10,468 | 1,237,337 | 9.92 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Two | 4,660 | 17 | 856 | 2,654 | 11,109 | 292,988 | 5.19 | 16 | 774 | 2,609 | 11,083 | 280,634 | 5.16 | 14 | 743 | 2,532 | 11,015 | 288,931 | 5.26 |
| Three | 5,248 | 10 | 1,189 | 3,534 | 8,651 | 1,550,385 | 12.87 | 5 | 1,150 | 3,349 | 8,560 | 1,010,083 | 10.50 | 4 | 1,158 | 3,356 | 8,647 | 1,002,962 | 10.00 |
| Four or More | 9,551 | 14 | 1,264 | 3,187 | 8,995 | 1,323,156 | 11.36 | 4 | 1,226 | 3,180 | 8,946 | 1,234,015 | 11.35 | 3 | 1,224 | 3,186 | 8,955 | 1,237,337 | 11.11 |
| Census Region |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Northeast | 3,856 | 10 | 813 | 2,697 | 8,171 | 1,124,135 | 10.75 | 4 | 754 | 2,559 | 8,056 | 1,010,083 | 11.36 | 3 | 746 | 2,522 | 8,063 | 1,011,952 | 11.29 |
| South | 5,867 | 17 | 1,403 | 4,051 | 11,527 | 1,323,156 | 8.68 | 5 | 1,362 | 3,928 | 11,269 | 1,234,015 | 8.81 | 4 | 1,364 | 3,931 | 11,279 | 1,237,337 | 8.81 |
| Midwest | 5,455 | 30 | 1,083 | 2,645 | 6,808 | 1,550,385 | 14.22 | 18 | 1,043 | 2,601 | 6,923 | 804,009 | 9.01 | 16 | 1,023 | 2,563 | 6,817 | 639,872 | 8.34 |
| West | 4,281 | 36 | 1,161 | 3,392 | 10,802 | 1,024,681 | 8.68 | 22 | 1,132 | 3,385 | 10,721 | 760,439 | 8.69 | 20 | 1,105 | 3,351 | 10,734 | 624,327 | 8.33 |
| Quarter |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Quarter1 | 4,613 | 14 | 1,351 | 3,582 | 10,437 | 854,597 | 7.22 | 4 | 1,274 | 3,548 | 10,107 | 805,478 | 7.47 | 3 | 1,247 | 3,494 | 10,168 | 806,173 | 7.45 |
| Quarter2 | 5,115 | 25 | 1,071 | 3,131 | 8,922 | 417,070 | 6.38 | 9 | 1,021 | 3,015 | 8,778 | 506,343 | 6.85 | 8 | 1,001 | 3,000 | 8,776 | 515,372 | 6.86 |
| Quarter3 | 5,093 | 10 | 963 | 2,769 | 8,219 | 1,024,681 | 12.79 | 6 | 917 | 2,717 | 8,257 | 884,682 | 12.27 | 6 | 910 | 2,724 | 8,293 | 890,858 | 11.63 |
| Quarter4 | 4,638 | 15 | 1,189 | 3,328 | 9,797 | 1,550,385 | 14.07 | 6 | 1,170 | 3,336 | 9,836 | 1,234,015 | 11.60 | 6 | 1,151 | 3,312 | 9,789 | 1,237,337 | 11.37 |
| \% Hispanic or Latino in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 1,537 | 42 | 1,594 | 4,086 | 11,989 | 854,597 | 9.02 | 34 | 1,508 | 4,019 | 12,166 | 884,682 | 9.65 | 33 | 1,482 | 3,980 | 12,144 | 890,858 | 9.73 |
| 10-<50\% | 3,726 | 35 | 1,482 | 4,402 | 11,789 | 1,323,156 | 10.37 | 9 | 1,434 | 4,263 | 11,836 | 1,234,015 | 9.54 | 8 | 1,399 | 4,222 | 11,854 | 1,237,337 | 9.27 |
| <10\% | 14,196 | 10 | 1,024 | 2,844 | 8,288 | 1,550,385 | 9.90 | 4 | 968 | 2,790 | 8,241 | 1,010,083 | 9.16 | 3 | 960 | 2,779 | 8,206 | 1,011,952 | 8.90 |
| \% Black or African American in Segment |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 1,482 | 17 | 1,232 | 3,602 | 10,074 | 488,930 | 6.27 | 5 | 1,175 | 3,589 | 10,217 | 501,086 | 6.35 | 4 | 1,165 | 3,573 | 10,398 | 502,540 | 6.32 |
| 10-<50\% | 2,974 | 29 | 1,287 | 3,799 | 10,451 | 1,024,681 | 9.57 | 18 | 1,257 | 3,738 | 10,228 | 805,478 | 8.61 | 16 | 1,267 | 3,750 | 10,205 | 806,173 | 8.21 |
| $<10 \%$ | 15,003 | 10 | 1,089 | 3,042 | 8,915 | 1,550,385 | 10.70 | 4 | 1,034 | 2,978 | 8,935 | 1,234,015 | 10.13 | 3 | 1,018 | 2,949 | 8,924 | 1,237,337 | 9.94 |
| \% Owner-Occupied    <br> DUs ${ }^{1}$ in Segment    |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 50-100\% | 14,866 | 10 | 1,248 | 3,505 | 10,121 | 1,550,385 | 9.97 | 4 | 1,216 | 3,410 | 9,934 | 1,234,015 | 9.37 | 3 | 1,209 | 3,401 | 9,941 | 1,237,337 | 9.12 |
| 10-<50\% | 3,600 | 17 | 991 | 2,731 | 8,123 | 1,124,135 | 9.53 | 9 | 929 | 2,762 | 8,444 | 1,010,083 | 9.02 | 8 | 909 | 2,733 | 8,434 | 1,002,962 | 9.05 |
| $<10 \%$ | 993 | 17 | 458 | 1,220 | 3,258 | 133,448 | 6.99 | 5 | 364 | 1,097 | 3,271 | 123,637 | 7.55 | 4 | 357 | 1,098 | 3,192 | 117,144 | 7.34 |
| Combined Median Rent/Housing Value |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $1^{\text {st }}$ Quintile | 3,227 | 15 | 955 | 2,526 | 6,979 | 1,550,385 | 21.52 | 6 | 943 | 2,599 | 7,106 | 1,004,156 | 15.44 | 6 | 933 | 2,629 | 7,213 | 1,011,952 | 14.03 |
| $2^{\text {nd }}$ Quintile | 4,250 | 10 | 1,048 | 2,919 | 8,231 | 1,124,135 | 12.13 | 6 | 1,020 | 2,887 | 8,310 | 1,010,083 | 12.22 | 6 | 1,008 | 2,860 | 8,337 | 1,002,962 | 12.24 |
| $3^{\text {rd }}$ Quintile | 4,364 | 17 | 994 | 2,839 | 8,440 | 1,323,156 | 11.06 | 5 | 944 | 2,762 | 8,242 | 1,234,015 | 11.13 | 4 | 921 | 2,724 | 8,185 | 1,237,337 | 11.29 |
| $4^{\text {th }}$ Quintile | 4,276 | 14 | 1,285 | 3,558 | 10,977 | 355,652 | 6.06 | 4 | 1,220 | 3,461 | 10,809 | 399,094 | 6.23 | 3 | 1,209 | 3,424 | 10,736 | 413,094 | 6.20 |
| $5^{\text {th }}$ Quintile | 3,342 | 17 | 1,488 | 4,266 | 12,366 | 919,088 | 6.64 | 18 | 1,409 | 4,103 | 12,186 | 727,153 | 6.73 | 16 | 1,419 | 4,111 | 12,119 | 624,327 | 6.41 |

Table L. 32012 NSDUH Respondent Pair-Level Weight Summary Statistics (res.pr.ps and res.pr.ev) (continued)

| Domain | $n$ | Before res.pr.ps ${ }^{1}$(SDUWT*PRWT10*...*PRWT12) |  |  |  |  |  | After res.pr.ps ${ }^{1}$(SDUWT*PRWT10*... ${ }^{*}$ PRWT13) |  |  |  |  |  | Final Weight: After res.pr.ev ${ }^{1}$ (SDUWT*PRWT10*...*PRWT14) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ | Min | Q1 ${ }^{2}$ | Med | Q3 ${ }^{2}$ | Max | UWE ${ }^{3}$ |
| Population Density |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Large MSA ${ }^{1}$ | 8,192 | 17 | 1,717 | 4,648 | 12,779 | 1,323,156 | 8.06 | 5 | 1,708 | 4,620 | 12,943 | 1,234,015 | 7.56 | 4 | 1,692 | 4,589 | 12,922 | 1,237,337 | 7.40 |
| Medium to Small MSA ${ }^{1}$ | 9,833 | 10 | 874 | 2,423 | 7,178 | 876,058 | 9.61 | 4 | 807 | 2,307 | 6,911 | 1,004,156 | 11.13 | 3 | 788 | 2,290 | 6,933 | 1,011,952 | 11.08 |
| Non-MSA, ${ }^{1}$ Urban | 380 | 52 | 742 | 1,980 | 5,241 | 854,597 | 35.92 | 16 | 762 | 1,878 | 5,022 | 805,478 | 32.27 | 16 | 737 | 1,852 | 5,022 | 806,173 | 32.41 |
| Non-MSA, ${ }^{1}$ Rural | 1,054 | 38 | 717 | 1,978 | 6,264 | 1,550,385 | 35.14 | 30 | 852 | 2,366 | 6,703 | 804,009 | 13.38 | 29 | 859 | 2,382 | 6,979 | 639,872 | 10.49 |
| Group Quarters |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Group | 312 | 17 | 429 | 1,061 | 2,898 | 34,734 | 4.01 | 25 | 408 | 910 | 2,667 | 35,420 | 3.79 | 25 | 418 | 877 | 2,730 | 28,121 | 3.70 |
| Non-Group | 19,147 | 10 | 1,150 | 3,234 | 9,414 | 1,550,385 | 10.09 | 4 | 1,109 | 3,185 | 9,334 | 1,234,015 | 9.49 | 3 | 1,093 | 3,175 | 9,352 | 1,237,337 | 9.28 |
| Pair Relationship Domain ${ }^{4}$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\begin{aligned} & \text { Parent-Child } \\ & (12-14) \end{aligned}$ | 1,981 | 83 | 2,456 | 5,444 | 11,989 | 163,632 | 2.95 | 80 | 2,413 | 5,537 | 12,561 | 196,486 | 3.13 | 75 | 2,355 | 5,499 | 12,516 | 192,088 | 3.14 |
| $\begin{aligned} & \text { Parent-Child } \\ & (12-17) \end{aligned}$ | 3,748 | 15 | 2,755 | 5,933 | 12,968 | 230,684 | 2.95 | 6 | 2,630 | 5,869 | 13,546 | 196,486 | 3.04 | 6 | 2,547 | 5,798 | 13,561 | 192,088 | 3.04 |
| $\begin{aligned} & \text { Parent-Child } \\ & (12-20) \end{aligned}$ | 4,443 | 15 | 2,915 | 6,347 | 14,088 | 296,605 | 3.18 | 6 | 2,825 | 6,450 | 14,879 | 299,472 | 3.27 | 6 | 2,793 | 6,413 | 14,820 | 279,736 | 3.21 |
| Sibling (12-14)- <br> Sibling (15-17) | 2,191 | 14 | 588 | 1,175 | 2,404 | 34,664 | 2.58 | 4 | 562 | 1,193 | 2,423 | 30,611 | 2.44 | 3 | 555 | 1,184 | 2,449 | 32,280 | 2.47 |
| Sibling (12-17)- <br> Sibling (18-25) | 2,477 | 16 | 844 | 1,789 | 3,667 | 37,794 | 2.36 | 8 | 828 | 1,843 | 3,811 | 23,928 | 2.28 | 8 | 823 | 1,839 | 3,849 | 24,220 | 2.28 |
| Spouse-Spouse/ Partner-Partner | 3,664 | 30 | 1,097 | 4,481 | 16,693 | 1,550,385 | 10.77 | 16 | 1,108 | 4,420 | 17,392 | 1,234,015 | 9.77 | 14 | 1,065 | 4,375 | 17,434 | 1,237,337 | 9.62 |
| Spouse-Spouse/ Partner-Partner with Children (Younger Than 18) | 1,640 | 51 | 1,055 | 3,442 | 10,817 | 1,323,156 | 16.13 | 70 | 1,595 | 5,143 | 15,628 | 1,234,015 | 12.24 | 69 | 1,575 | 5,108 | 15,656 | 1,237,337 | 12.17 |

[^24]
## Appendix M: Pair Analysis Manual Excerpt

## Appendix M: Pair Analysis Manual Excerpt

This appendix provides background information on pair data analysis for the National Survey on Drug Use and Health (NSDUH). This excerpt is from Section 3.2 in How To Prepare and Analyze Pair Data in the National Survey on Drug Use and Health by Frechtel, Warren, and Porter (in press).

## Inferential Population and Multiplicities

There are different perspectives through which pair data can be analyzed: (1) with pairs as the focus, or (2) with one member of the pair as the focus. When the focus is on the pair, the PRANALWT variable can be used to weight the data directly with no adjustments. However, when the focus is on one member of the pair, an adjustment often needs to be made to the weight to account for this. For example, the analysis outlined in Table 1 is focused on the child's behavior when the father talked to him or her about substance use. However, if the analyst was interested in the child's behavior when a parent talked to him or her about substance use, regardless of the gender of the parent, the weight would need to be adjusted to account for the fact that the child may be a member of more than one parent-child pair. For the analysis in Table 1 , there is no multiple counting problem as long as an assumption is made that no child lives with more than one father.

Adjustments for the multiple counting problem are done using "multiplicities" (Chromy \& Singh, 2001). These multiplicities have been computed for the analyst and can be accounted for by simply adjusting the weight variable (described in Section 4.4). Analysts should exercise care or seek assistance in computing multiplicities for any pair types not listed. The process by which the multiplicities are created is described in detail in Chapter 10 of Frechtel et al. (2013). Table 5 lists the pair domains for which multiplicities can be used to perform person-level analyses.

Table 5. Pair Domains and Multiplicities

| Pair Domain |  |  | Multiplicity Variable |  |
| :---: | :---: | :---: | :---: | :---: |
| IRPRREL <br> Levels | Description | Focus ${ }^{1}$ | Name | Description |
| 1 | Parent-child, child aged 1214 | Parent | IRMPCP14 | Number of children aged 12-14 living with responding parent |
|  |  | Child | IRMPCC14 | Number of parents living with responding child aged 12-14 |
| 1,2 | Parent-child, child aged 1217 | Parent | IRMPCP17 | Number of children aged 12-17 living with responding parent |
|  |  | Child | IRMPCC17 | Number of parents living with responding child aged 12-17 |
| 1,2,3 | Parent-child, child aged 1220 | Parent | IRMPCP20 | Number of children aged 12-20 living with responding parent |
|  |  | Child | IRMPCC20 | Number of parents living with responding child aged 12-20 |
| 2 | Parent-child, child aged 1517 | Parent | IRMPCP57 | Number of children aged 15-17 living with responding parent |
|  |  | Child | IRMPCC57 | Number of parents living with responding child aged 15-17 |
| 5 | Sibling-sibling, older sibling aged 15-17, younger sibling aged 12-14 | Older sibling | IRMS1417 | Number of siblings aged 12-14 living with responding sibling aged 15-17 |
|  |  | Younger sibling | IRMS1714 | Number of siblings aged 15-17 living with responding sibling aged 12-14 |
| 6 | Sibling-sibling, older sibling aged 18-25, younger sibling aged 12-17 | Older sibling | IRMS1725 | Number of siblings aged 12-17 living with responding sibling aged $18-25$ |
|  |  | Younger sibling | IRMS2517 | Number of siblings aged 18-25 living with responding sibling aged 12-17 |
| 8,9 | Spouse-spouse and partnerpartner | No multiplicity necessary: assume only one spouse per person |  |  |
| 8 | Spouse-spouse and partnerpartner, with children aged 0-17 | No multiplicity necessary: assume only one spouse per person |  |  |

[^25]To help clarify this concept, some bulleted examples are included below. A full list of multiplicities and definitions is included in Table 5 above. The examples are based on the analysis in Table 6, which is the same as the analysis in Table 1 except that the parent can be a father or a mother.

Table 6. Example of Pair-Level Analysis Requiring Multiplicities

| Parent Reports Talking to | Child Used an Illicit Drug in the Past Year |  |
| :---: | :---: | :---: |
| Child Aged 12-17 about <br> Substance Use | Yes | No |
| Yes | $(1,1)$ | $(1,2)$ |
| No | $(2,1)$ | $(2,2)$ |

- To populate Table 6 with estimates related to children (i.e., row percentages), the proper weight is PRANALWT/IRMPCC17. For example, this weight could be used to estimate the probability that a child aged 12 to 17 used an illicit drug in the past year, given that a parent talked to him or her about substance use in the past year (cell $(1,1)$ in the table).
- If Table 6 instead showed estimates related to parents (i.e., column percentages), the proper weight would be PRANALWT/IRMPCP17. For example, this weight could be used to estimate the probability that a parent talked to his or her child aged 12 to 17 about substance use in the past year, given that his or her child used an illicit drug in the past year (again, cell $(1,1)$ in the table).
- As stated above, if estimates related to children are desired, but interest is restricted to either fathers or mothers, standard practice is to assume the presence of no more than one father and no more than one mother in the dwelling unit. In these cases, no multiplicity is necessary, and PRANALWT can be used as the analysis weight. For example, PRANALWT should be used to estimate the probability that a child aged 12 to 17 used an illicit drug in the past year, given that his or her mother talked to him or her about substance use in the past year.
- If Table 6 showed estimates related to parent-child pairs and did not focus on either member of the pair (i.e., cell percentages), then no multiplicity would be necessary. The variable PRANALWT is designed for this. An analysis like this is of questionable value because of the multiple counting problem, however. The inferential population includes all parent-child pairs where the child is aged 12 to 17 . Many persons are represented in more than one pair. A dwelling unit with two parents and three children aged 12 to 17 would represent six pairs in the population, with each parent being a member of three of the six and each child being a member of two of the six. A dwelling unit with one parent and one child aged 12 to 17 would represent only one pair in the population. Analyses that focus on only one member of the pair usually have a more natural interpretation.


[^0]:    ${ }^{1}$ The National Household Survey on Drug Abuse (NHSDA) was renamed the National Survey on Drug Use and Health (NSDUH) in the 2002 survey year.

[^1]:    ${ }^{2}$ This report presents information from the 2012 National Survey on Drug Use and Health (NSDUH). Prior to 2002, the survey was called the National Household Survey on Drug Abuse (NHSDA).

[^2]:    ${ }^{3}$ A circus owner had 50 elephants, and wanted to estimate the total weight to help him make arrangements for shipping. To save time, he only wanted to weigh Sambo (an average sized elephant), and use 50 times its weight as an estimate. However, the circus statistician, being highly conscious of the optimality and unbiasedness of the Horvitz-Thompson (HT) estimator, objected about the potential bias of his estimate because of the purposive selection. Instead, he suggested random selection of an elephant with a very high probability of 99/100 for Sambo, and the rest including Jumbo (the biggest in the herd) with probability $1 / 4900$ each. The circus owner was very unhappy with the statistician's response of 100/99 times the Sambo's weight as the estimate if Sambo got selected in this random draw, and was outraged with the response of 4900 times the Jumbo's weight if Jumbo happened to get selected. It was obvious to the owner that this new estimator was extremely poor, although he didn't know anything about its unbiasedness. The story had an unhappy ending with the circus statistician losing his job. To alleviate the instability of the HT-estimator, Hajek suggested to multiply it by 50 divided by inverse of the selection probability, which reduces simply to 50 times the weight of the selected elephant.
    ${ }^{4}$ Winsorization is a method of extreme value adjustment that replaces extreme values with the critical values used for defining low and high extreme values.
    ${ }^{5}$ Pair age in this case should not be confused with the modeling term, which has a finer level breakdown.

[^3]:    ${ }^{6}$ This report presents information from the 2012 National Survey on Drug Use and Health (NSDUH). Prior to 2002, the survey was called the National Household Survey on Drug Abuse (NHSDA).

[^4]:    ${ }^{7}$ Deep stratification refers to the stratification that was used in the sample design. In the case of the 2012 survey, deep stratification refers to the cross-classification of State sampling region by age group.

[^5]:    ${ }^{8}$ Pair age in this case should not be confused with the modeling term, which has a finer level breakdown.

[^6]:    QDU = questionnaire dwelling unit; $\mathrm{SE}=$ standard error.
    Note: Standard errors of prevalence estimates are provided in parentheses.
    ${ }^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last step of calibration, res.qdu.ps, and a full model for preceding steps.
    ${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

[^7]:    Note: Standard errors of prevalence estimates are provided in parentheses.
    ${ }_{2}^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last two steps of calibration, res.pr.ps and res.pr.ev, and a full model for preceding steps.
    ${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

[^8]:    Note: Standard errors of prevalence estimates are provided in parentheses.
    ${ }^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last two steps of calibration, res.pr.ps and res.pr.ev, and a full model for preceding steps.
    ${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

[^9]:    Note: Standard errors of prevalence estimates are provided in parentheses.
    ${ }^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last two steps of calibration, res.pr.ps and res.pr.ev, and a full model for preceding steps.
    ${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

[^10]:    Note: Standard errors of prevalence estimates are provided in parentheses.
    ${ }^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last two steps of calibration, res.pr.ps and res.pr.ev, and a full model for preceding steps.
    ${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

[^11]:    Note: Standard errors of prevalence estimates are provided in parentheses.
    ${ }^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last two steps of calibration, res.pr.ps and res.pr.ev, and a full model for preceding steps.
    ${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

[^12]:    Note: Standard errors of prevalence estimates are provided in parentheses.
    ${ }^{1}$ Baseline refers to the weight obtained from using a main effects only model for the last two steps of calibration, res.pr.ps and res.pr.ev, and a full model for preceding steps.
    ${ }^{2}$ Final refers to the weight obtained using a full model throughout all steps of calibration.

[^13]:    ${ }^{1}$ A half step refers to halving the increment in the Newton-Raphson iterative process for fitting GEM.
    ${ }^{2}$ The National Household Survey on Drug Abuse (NHSDA) was renamed the National Survey on Drug Use and Health (NSDUH) in the 2002 survey year.

[^14]:    ${ }^{1}$ The District of Columbia is included among States.

[^15]:    ${ }^{1}$ This is the reference level for this variable. This is the level against which effects of other factor levels are measured.

[^16]:    $\mathrm{DU}=$ dwelling unit; MSA = metropolitan statistical area; QDU = questionnaire dwelling unit; SDU = screener dwelling unit.
    ${ }^{1}$ The weight used for calculating the response rate includes SDU- and QDU-level design weights, SDU nonresponse and poststratification adjustments, and selected QDU poststratification adjustment. This weight is the product of WT1*...*WT9*DUWT10*DUWT11.

[^17]:    ${ }^{1} \mathrm{DU}=$ dwelling unit, MSA = metropolitan statistical area, $\mathrm{PS}=$ poststratification adjustment, QDU = questionnaire dwelling unit, SDU = screener dwelling unit, Sel = selected.
    ${ }^{2}$ Weighted extreme value proportion: $100^{*} \sum_{k} w_{e k} / \sum_{k} w_{k}$, where $w_{e k}$ denotes the weight for extreme values, and $w_{k}$ denotes the weight for both extreme values and nonextreme values.
    ${ }^{3}$ Outwinsor weight proportion: $100^{*} \sum_{k}\left(w_{e k}-b_{k}\right) / \sum_{k} w_{k}$, where $b_{k}$ denotes the winsorized weight.

[^18]:    ${ }^{1} \mathrm{DU}=$ dwelling unit, MSA = metropolitan statistical area, $\mathrm{PS}=$ poststratification adjustment, $\mathrm{QDU}=$ questionnaire dwelling unit, $\mathrm{SDU}=$ screener dwelling unit, Sel $=$ selected.
    ${ }^{2} \mathrm{Q} 1$ and Q3 refer to the first and third quartile of the weight distribution.
    ${ }^{3}$ Unequal weighting effect (UWE) is defined as $1+[(n-1) / n]^{*} C V^{2}$, where $C V=$ coefficient of variation of weights.

[^19]:    $\mathrm{DU}=$ dwelling unit, MSA = metropolitan statistical area, $\mathrm{NR}=$ nonresponse adjustment, $\mathrm{PS}=$ poststratification adjustment, QDU = questionnaire dwelling unit, Res = respondent, SDU = screener
    dwelling unit, $\mathrm{Sel}=$ selected.
    ${ }^{2}$ Q1 and Q3 refer to the first and third quartile of the weight distribution.
    ${ }^{3}$ Unequal weighting effect (UWE) is defined as $1+[(n-1) / n]^{*} C V^{2}$, where $C V=$ coefficient of variation of weights.

[^20]:    SDU = screener dwelling unit.
    ${ }^{1}$ For a key to modeling abbreviations, see Chapter 6, Exhibit 6.1.

[^21]:    DU $=$ dwelling unit, MSA $=$ metropolitan statistical area.
    ${ }^{1}$ The weight used for calculating the response rate includes screener dwelling unit (SDU)- and pair-level design weights, SDU nonresponse and poststratification adjustments, and selected pair poststratification adjustment. This weight is the product of WT1*...*WT9*PRWT10*PRWT11.

[^22]:    This step used demographic variables from screener data for all selected person pairs; $\mathrm{DU}=$ dwelling unit, MSA = metropolitan statistical area, $\mathrm{PR}=$ pair, $\mathrm{PS}=$ poststratification adjustment,
    SDU = screener dwelling unit, $\mathrm{Sel}=$ selected.
    ${ }^{2}$ Weighted extreme value proportion: $100^{*} \sum_{k} w_{e k} \sum_{k} w_{k}$, where $w_{e k}$ denotes the weight for extreme values, and $w_{k}$ denotes the weight for both extreme values and nonextreme values
    Outwinsor weight proportion: $100 * \sum_{k}\left(w_{e k}-b_{k}\right) / \sum_{k} w_{k}$, where $b_{k}$ denotes the winsorized weight.

[^23]:     screener dwelling unit.
    ${ }^{2}$ Q1 and Q3 refer to the first and third quartile of the weight distribution.
    ${ }^{3}$ Unequal weighting effect (UWE) is defined as $1+[(n-1) / n]^{*} C V^{2}$, where $C V=$ coefficient of variation of weights.

[^24]:     adjustment, Res $=$ respondent, SDU $=$ screener dwelling unit.
    ${ }^{2}$ Q1 and Q3 refer to the first and third quartile of the weight distribution.
    ${ }^{3}$ Unequal weighting effect (UWE) is defined as $1+[(n-1) / n] * C V^{2}$, where $C V=$ coefficient of variation of weights.
    ${ }^{4}$ Parent-child (15-17) was not included here since extreme values were not controlled with this domain.

[^25]:    ${ }^{1}$ No weight adjustment is needed when the inferential focus is on the pair.

