

The shape and tempo of language evolution

Short Title: The shape and tempo of language evolution

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SUMMARY

There are approximately 7000 languages spoken in the world today. This diversity reflects the legacy of thousands of years of cultural evolution. How far back we can trace this history depends largely on the rate at which the different components of language evolve. Rates of lexical evolution are widely thought to impose an upper limit of 6-10 thousand years on reliably identifying language relationships. In contrast, it has been argued that certain structural elements of language are much more stable. Just as biologists use highly conserved genes to uncover the deepest branches in the tree of life, highly stable linguistic features hold the promise of identifying deep relationships between the world's languages. Here we present the first global network of languages based on this typological information. We evaluate the relative evolutionary rates of both typological and lexical features in the Austronesian and Indo-European language families. The first indications are that typological features evolve at similar rates to basic vocabulary but their evolution is substantially less treelike. Our results suggest that, whilst rates of vocabulary change are correlated between the two language families, the rates of evolution of typological features and structural sub-types show no consistent relationship across families.

KEYWORDS: language structure, language evolution, typology, evolution, phylogenetics, networks

Introduction:

How far back can we trace the history of languages? The traditional comparative method in historical linguistics uses systematic sound correspondences between homologous (“cognate”) words to infer relatedness between languages. Most linguists argue that this approach can only be used to make inferences about languages that diversified within the last 6,000 to 10,000 years (Nichols 1992; Ringe 1995; Kaufman & Golla 2000). Beyond this time, however, it becomes impossible to distinguish accurately whether any signal in the data represents descent from a common ancestor or false similarities due to chance and borrowing between languages.

Some authors have claimed that certain typological features that describe the structures present in a language, such as ergativity, head marking, and numeral classifiers, are more stable than the lexicon (Nichols 1992; Nichols 1994). If some typological features are consistently stable within language families, and resistant to borrowing, then they might hold the key to uncovering relationships at far deeper levels than previously possible. For example, Nichols (1994) uses typological features to argue for a spread of languages and cultures around the Pacific Rim, connecting Australia, Papua New Guinea, Asia, Russia, Siberia, Alaska and the western coasts of North and South America. If this is correct, then these typological features must be reflecting time depths of at least 16,000 years and possibly as deep as 50,000 years ago (Nichols 1994). A recent phylogenetic study of phonological and morphosyntactic features in non-Austronesian languages of Island Melanesia argued that typological traits reveal phylogenetic signal consistent with deep (~10,000 year) historical relationships (Dunn et al. 2005). One explanation for this stability is that the evolution of typological features is more constrained than that of the

lexicon because structural traits function as an interrelated system with strong dependencies between components (“un système où tout se tient”, variously attributed to Antoine Meillet, and Ferdinand de Saussure, Peeters 1990).

However, the lack of comprehensive world-wide typological data has meant that it has been difficult to assess the overall shape and tempo of changes in language structure. The recently published World Atlas of Language Structures remedies this problem (WALS, Haspelmath et al. 2005). WALS includes information about 141 typological features from 2,561 languages. Here we report the results of phylogenetic analyses of the typological data in the WALS. First, we explore the global pattern of typological data using a network method to assess evidence for deep signal in the data. Second, we quantify the fit of typological and lexical features onto known family trees for two of the world’s largest and best-studied language families – Indo-European and Austronesian. Third, we infer the rates of evolution of typological and lexical features within these families and compare rates between families.

Materials and Methods.

a) Typological Data

From the 141 characters in the World Atlas of Language Structures (WALS, Haspelmath et al. 2005), we discarded the three characters belonging to the “Sign Languages” and “Other” categories, leaving 138 characters for analysis (Table S1). We extracted three datasets from WALS. The first dataset was a "world-wide" dataset that included all languages in WALS with less than 25% missing data (Table S2). Unfortunately the WALS database has incomplete data for many languages and feature classes, so this left a

total of 99 languages in this world-wide dataset. The second and third datasets were comprised of 20 Austronesian and 20 Indo-European languages that we had sufficient lexical data for and that were well described in the WALS database. To maximise the phylogenetic signal in the typological data, we recoded 49 of the 138 characters by splitting up aggregate categories and combining feature states with few members (See supplementary material and Table S3).

b) Lexical Data

Lexical cognate data for the languages in WALS were taken from two sources (Tables S4 and S5). The Austronesian lexical data were extracted from the Austronesian Basic Vocabulary Database (Greenhill et al. 2008, <http://language.psy.auckland.ac.nz/austronesian>). This database project contains 210-item wordlists and cognate information from over 650 Austronesian languages. The Indo-European lexical data came from a published dataset of 200 item basic-vocabulary wordlists and cognate information from 95 Indo-European languages (Dyen et al. 1992). Both the Austronesian and Indo-European databases are comprised of items of basic vocabulary (terms for body parts, kinship terms, colors, simple verbs, numbers etc) that are thought to be highly stable over time and resistant to being borrowed between languages (Swadesh 1952).

c) NeighbourNet Analysis

The world-wide NeighbourNet was constructed using *SplitsTree* v4.8 using Uncorrected-P distances (Bryant and Moulton 2004; Bryant et al. 2005). To reduce the noise in the networks, splits were filtered according to a weight threshold of 0.002. NeighbourNets

were also constructed for each typological/lexical dataset using the same method, and splits were filtered to a threshold of 0.001 (Figure S5).

d) Character Fit Analysis

We constructed family trees for the Indo-European and Austronesian language families (Figure S4) from the standard Ethnologue classification (Gordon 2005) and previous research on Indo-European (Gray & Atkinson 2003; Atkinson & Gray 2005), and Austronesian (Blust 1999; Lynch et al. 2002; Gray et al. 2009). To measure the fit of each character onto these trees, we calculated the Retention Index (RI, Archie 1989, Farris 1989) for all characters in the four datasets (Austronesian Lexicon, Austronesian Typology, Indo-European Lexicon and Indo-European Typology) using PAUP* v.4b10 (Swofford 2002). We selected the RI for this comparison as it does not require us to estimate branch-lengths as likelihood-based character-fit analyses would. RIs are only available for characters that are parsimony informative (constant characters or characters with all unique states do not provide information on the fit of the data to a tree). RIs were calculated for 113/210 characters in the Austronesian lexicon, 109/138 characters in the Austronesian typology, 183/200 characters in the Indo-European lexicon, and 116/138 characters in the Indo-European typology.

e) Rates Analysis

To calculate the rate estimates, trees with branch-lengths proportional to the amount of change between each language are required. We used a Bayesian phylogenetic approach implemented in the program *BayesPhylogenies* (Pagel & Meade 2004) to produce a posterior distribution of phylogenetic trees from the binary-coded lexical cognate data. The analysis used a two-rate model of cognate evolution that allows cognates to be

gained and lost at different rates. The Markov chain ran for 10 million generations, and burn-in was set to 5 million generations after inspection of log likelihood plots of the parameters. The tree topologies were constrained to match the classification trees (Figure S4), so that each tree sample varied only in their estimate of the branch-lengths. Trees were sampled every 5,000 generations from the chain, leaving a total of 1,000 post-burn-in trees.

By constraining the tree topology to established language groupings we minimize any bias that might result from estimating the tree topology from the lexical cognate data. The use of the lexical data to estimate the branch-lengths is in line with arguments that lexical phylogenies based on basic vocabulary provide good estimators of the underlying cultural history (Mace & Pagel 1994). Moreover, the site-specific likelihoods (indicating the fit of the data under the model of evolution) calculated on the trees with branch-lengths derived from the typological data were essentially identical to those obtained with lexical branch-lengths (Spearman's $\rho=0.997$, $p<0.001$) – in other words, there is no reason to think that the use of lexical branch lengths biases our results.

Maximum likelihood rate estimates, μ , were calculated from these posterior tree distributions using *BayesTraits* (Pagel et al. 2004). *BayesTraits* implements a continuous time Markov model that allows characters to change between states over small time intervals. This can be used to reconstruct how traits with discrete, finite states evolve on the trees in the posterior distribution. Estimates of μ were obtained for all four datasets (Austronesian lexicon, Austronesian typology, Indo-European lexicon and Indo-European typology). Traits with greater than 50% missing data were excluded from the analyses.

For constant characters, the maximum likelihood rate estimate is zero. However, for any trait that can vary, the true rate is always non-zero. We can infer a rate for constant characters by plotting the observed number of states against the rate estimates for each feature within each of the four datasets. We fitted an exponential curve to the data and used this to provide a predicted rate for constant characters in each data set – the point on the curve where the observed number of states is one. The results we report include the estimated rates for non-constant characters and the inferred rate for constant characters. We also repeated all rate analyses setting the constant rate to the minimum estimated maximum likelihood rate among the variable characters. This had no appreciable effect on the results we report.

Results

The Global Pattern of Typological Diversity

To explore global patterns of typological signal, we used a phylogenetic network technique, NeighbourNet (Bryant & Moulton 2004; Bryant et al. 2005), to visualise the relationships implied by these data (Figure 1). In these networks the length of the branches are proportional to the amount of divergence between languages. Box-like structures represent the conflicting signals when typological features support incompatible language groupings. If typological features are deeply stable then we would expect the groupings in the network to reflect known linguistic history and contain few boxes of conflicting signal. In contrast, if the typological features tend to diffuse between adjacent languages in a linguistic area or evolve too rapidly to reveal deep signal, we would expect to see a star-like network with many boxes and clusters reflecting geographical proximity or chance resemblances.

The network in Figure 1 correctly groups some of the languages into known language families, with Indo-European, Altaic, Nakh-Daghestanian being the most distinct. The network also groups a number of sub-families together – such as the Pama-Nyungan languages (Kayardild, Martuthunira, and Ngiyambaa), the Bantu languages (Luvale, Swahili, and Zulu), the Oceanic languages (Maori, Fijian, and Rapanui), the Semitic languages (Hebrew and Arabic), and the Cushitic languages (Irakw and Oromo Harar). However, other well-known families are not recovered, including Sino-Tibetan, Uralic, and Trans-New Guinea. The Austronesian language family also does not form a monophyletic group. Additionally, the network shows evidence of substantial conflicting signal between structural elements (box-like structures) and does not accurately recover many attested phylogenetic relationships within the major language families. For example, in Indo-European the network links German to French, when German is more closely related to English (Beekes 1995).

The network does, however, show evidence for some higher-level clusters in the data. The first of these (cluster 1, labeled on Figure 1) includes the languages from continental Eurasia, which could be interpreted as indicating ancient common ancestry. This cluster groups the Indo-European languages with the Uralic languages (Finnish and Hungarian), consistent with the proposed macro-family Indo-Uralic. These two families are joined in this cluster by the Altaic language family (Turkish, Evenki, and Khalkha), the Dravidian language Kannada, and a number of languages from the Caucasus region: the Nakh-Daghestanian family (Ingush, Lezgian, and Hunzib), Abkhaz (Northwest Caucasian), and Georgian (Kartvelian). If typological features do indeed evolve slowly enough to reveal

deep history, then this cluster may represent the controversial Nostratic macrofamily (Renfrew & Nettle 1999). However, the inclusion of languages such as Alambalak (from Papua New Guinea), Awa Pit (from Colombia), Quechua (from Ecuador) and the isolate Basque are incompatible with this proposal. A second large cluster (cluster 2, labeled on Figure 1) includes the Australian languages, the Austronesian languages, and some languages from the African families of Afro-Asiatic and Niger-Congo. This second cluster does not correspond to any known macro-family proposals or geographical regions, however, Austronesian languages are placed next to some other non-Austronesian languages from South-East Asia (Thai, Vietnamese, and Mandarin).

The left side of the network (Fig. 1) contains a subset of the languages of Australia, and distinguishes between the Pama-Nyungan languages (Kayardild, Martuthunira, Ngiyambaa), and others from different families (Gooniyandi, Mangarayi). However, two other languages from the northern tip of Australia (Tiwi and Maung) are not included but placed in the second cluster. Another interesting subset here may also hint at some deeper links – most of the languages of North America are linked together in this network (Lakhota, Slave, Maricopa, and Koasati). However, this grouping rather unusually includes a language from Paraguay – Guaraní – and does not include other North American languages of Yaqui and Kutenai.

Modeling structural and lexical evolution on trees

The existence of high-level clusters in the WALS data is consistent with the proposal that some typological features evolve slowly enough to identify deep historical relationships. However, phylogenetic networks cannot distinguish between similarity due to common

ancestry and similarity due to areal diffusion or chance resemblances arising through independent innovation. To evaluate the claim that some typological features of language are highly stable, we compared the shape and tempo of typological and lexical evolution by modeling their replacement through time on two language family trees that have well-established internal subgroupings: Indo-European (Beekes 1995; Gray & Atkinson 2003, Atkinson & Gray 2006), and Austronesian (Blust 1999; Lynch et al. 2002; Gray et al. 2009). If some typological features are highly stable and good indicators of common ancestry, then we would expect them a) to fit well with established language groupings and b) to show slower rates of change than lexical features as a whole. We extracted typological data from the WALS for the twenty most well-attested languages in each of the two families, removing the languages with the least data. We assembled lexical datasets for the same twenty languages from published databases of Indo-European (Dyen et al. 1992) and Austronesian (Greenhill et al. 2008) vocabulary.

We assessed the shape of language evolution in these data by estimating the fit of the typological and lexical data onto the established family trees using the Retention Index (RI, Archie 1989; Farris 1989). A stable, well fitting character will have an RI approaching one, whilst an unstable or rapidly evolving character will have an RI approaching zero. Histograms of the RIs for the lexical and typological features in the Indo-European and Austronesian datasets are shown in Figure 2a. In the lexical data, the mean RI for each character was 0.84 (s.d.=0.31) for the Austronesian and 0.89 (s.d.=0.21) for the Indo-European. The mean RI per character of the typological data was much lower at 0.36 (s.d.=0.33) for the Austronesian and 0.32 (s.d.=0.33) for the Indo-European. In both families the lexical data were a significantly better fit to the expected family trees

than the typological data (Mann-Whitney: Austronesian $U=8331$, $p<0.001$, Indo-European $U=13086.5$, $p<0.001$). These differences in fit are also evident in networks of the typological and lexical data (Figure 2a inset) where the lexical networks clearly show a much more tree-like signal than the typological networks. Unfortunately the RI is unable to estimate the fit of constant characters on the trees. The characters that are constant in both language families ($n=6$) are potential candidates for deep relationship indicators. However, closer inspection of these characters shows that four of them are only constant due to large numbers of missing data (with only $\sim 12.5\%$ of the states assigned across the 40 languages). The two characters that are constant in both families and have appreciable amounts of data are “N-M Pronouns” (with 15 of the 40 languages showing “No N-M pronouns”, and the remainder missing data), and “Order of Adverbial Subordinator and Clause” (with 38 of 40 languages belonging to the state “adverbial subordinators which are separate words and which appear at the beginning of the subordinate clause”).

It could be argued that the analysis of character fit is biased in favor of the lexical cognate data since historical linguistics often uses lexical information to infer linguistic relationships. Indeed, some subgroups are defined by major lexical innovation, such as Eastern Malayo-Polynesian (Blust 2009). In other cases, however, subgroups are defined by phonological and morphological innovations (Durie and Ross 1996, Blust 2009). For example, the Proto-Nuclear Polynesian subgroup is demarcated by many morphological innovations, the Oceanic subfamily is defined by the phonological merger of $*p$ and $*b$, and Central-Eastern Malayo-Polynesian is identified by the lowering of high vowels and four shared grammatical morphemes (Blust 2009). The subgroups we use here represent

the best available estimate of the true underlying language tree, drawing on a consilience of evidence from both lexical and structural data (Durie and Ross 1996, Blust 2009). Any bias in favour of the cognate data is therefore expected to be minimal.

To estimate rates of change we calculated the maximum likelihood estimate for the rate of evolution across the posterior distribution of trees in each family. Figure 3 shows a comparison of the distributions of rates for Indo-European and Austronesian lexical and typological characters. In both families, the distributions of lexical and typological rates are comparable. The similar ranges evident in these plots indicate that there is in fact no substantial difference between the slowest rates of lexical and typological change in either family. Austronesian rates for lexical features were on average slightly higher than rates for typological features (Mann-Whitney: Austronesian $U = 5961$, $p < 0.001$) whilst in the Indo-European data lexical and typological rates were not significantly different (Mann-Whitney: Indo-European $U=6718$, $p>0.05$). The bimodal distribution for Austronesian lexicon indicates that its higher average rate is due to a relatively high number of rapidly evolving words.

Whilst we find no clear difference between overall rates of lexical and typological change, some subsets of typological features may nonetheless change slowly enough to infer deep relationships. For example, Nichols (1992) claims that ergativity, head marking, and numeral classifiers are among the most stable structural features of language. The WALS project groups the typological data into nine feature classes describing different aspects of language structure. Figure 3 shows the inferred rate distributions grouped according to the nine typological feature-classes defined in the

WALS database, together with the lexical rate estimates. This plot highlights considerable variation in the rates of evolution between feature-classes and between families. For example, characters in the Nominal Syntax feature class have some of the highest rates in Austronesian but lowest in Indo-European, whilst the reverse is true for Complex Sentence structures. A univariate ANOVA shows that, when controlling for language family, there is no effect of typological feature class on rates of feature evolution ($F=1.27$, $p=0.26$).

Finally, we examined the relationship between rates of change for individual lexical and typological features across language families. Identifying specific features that are consistently stable across families has the potential to greatly improve our ability to detect and evaluate deep inter-family relationships. In addition, the kinds of regularities identified may point to constraints on the process of language evolution itself. In agreement with previous research (Pagel 2000; Pagel & Meade 2006), we find that rates of lexical change are correlated across language families (Spearman's $\rho = 0.37$, $p < 0.001$). By contrast, there is no significant correlation in rates of typological feature change between Indo-European and Austronesian (Spearman's $\rho = 0.17$, $p=0.10$). Although non-significant, this relationship is positive, suggesting a small number of structural features may still be consistently stable. We can identify nine features that have rates in the slowest 0.20 quantile in both language families: The Velar Nasal, Case Syncretism, Numeral Bases, Pronominal and Adnominal Demonstratives, The Optative, Coding of Nominal Plurality, Glottalized Consonants, Syllable Structure and Suppletion According to Tense and Aspect. These traits could be seen as candidates for investigating deep time-scales, however caution is needed in interpreting these results. First, a chi-

square test reveals that finding nine traits in the slowest 0.20 percentile in both families does not differ significantly from chance ($\chi^2 = 3.487$, $p=0.062$), and the same applies using the 0.05 percentile ($\chi^2 = 2.34$, $p=0.13$). Second, many of these characters reflect shared absence in the majority of the languages in our sample. For example, for the character The Optative, WALS only has data for 30/40 of the languages in our sample, and 28 of these are marked as “Inflectional optative absent”. Likewise, in the character The Velar Nasal, the Austronesian languages show their well-known bias for nasal substitution (Blust 2004), with 11 of the 20 languages having initial velar nasals, 8 languages missing data, and only Kilivila showing an absence. However, in the 12 Indo-European languages with data, the most prominent state (10/12) is “no velar nasal”. Together with the absence of any correlation in the typological rates of evolution between the families, these patterns do not support the existence of a set of universally stable typological features.

Discussion

There is considerable interest in the possibility that analyses of typological features may enable us to “push back the time barrier” beyond the apparent 6,000-10,000 year upper limit of the comparative method (Gray 2005). It has been suggested that typology can reveal historical signal dating back at least this far (Dunn et al. 2005; Dunn et al. 2008), or even tens of thousands of years earlier (Nichols 1994). The network analysis of WALS structural features reported in figure 1 points to some intriguing possible deep relationships, perhaps most notably the cluster linking together many of the major language families of Eurasia. However, our analysis of rates of evolution failed to identify any typological features that evolve at consistently slower rates than the basic

lexicon. If the signal in the lexicon does stretch back as far as 10,000 years (Nichols 1992; Ringe 1995; Kaufman & Golla 2000), then our results suggest that typological data is constrained by a similar time horizon (e.g. Dunn et al. 2005; Dunn et al. 2007; Dunn et al. 2008).

Beyond the difficulty of identifying consistently stable typological features, our findings suggest two further challenges to inferring deep ancestral relationships from structural language data. First, the typological features show relatively high rates of homoplasy. The classification of lexical data into cognate sets relies on isomorphism between sound and meaning within a vast possible state space of the items under comparison. The coupling of these two aspects reduces the possibility of chance similarity (Meillet 1948). In contrast, there is a “poverty of choice” of possible typological states (Harrison 2003). For example, there are only six permutations for the ordering of the subject, object, and verb that a language can use. Accordingly, there is a 1/6 chance that any two languages share the same ordering – in fact, since some configurations are much more likely than others even this probability is an underestimate. This means that, even for a given rate of change, shared typological features are a less reliable indication of common ancestry than shared basic vocabulary, and are more likely to produce spurious relationships.

A second issue with identifying slowly-evolving typological features is diffusion between geographically-proximate languages (Matras et al. 2006). This can occur through processes like language shift (Thomason & Kaufman 1988) – where speakers of one language change to another due to societal influences, yet retain morphology or phonology from their original language, or metatypy (Ross 1996) – where a language

rearranges some aspect of typology (e.g. morphosyntax) due to contact between languages without explicit borrowing between the languages, usually as an outcome of intimate cultural contact. Our results show substantial non-tree-like signal in the typological data and a poor fit with known language relationships within the Austronesian and Indo-European language families. On a global scale, Figure 1 shows some putative geographic clusters like the “Nostratic” grouping in Eurasia. In this “Nostratic” cluster Hindi does not group correctly with Indo-European but is located with its geographic neighbour, the Dravidian language, Kannada, suggesting that the similarities seen here may indeed be due to diffusion. Likewise, a grouping of Indonesian, Thai, Vietnamese, and Mandarin may be the result of areal diffusion in the South-East Asian region (Bisang 2006; Matras et al. 2006). The areal diffusion of typological features – like lexical borrowing – does make it harder to identify common ancestry.

Diffusion and chance resemblances are serious challenges for historical inference based on typological data. The problem of diffusion can be lessened if known instances of diffusion are identified and removed (Ross 1996, Dunn et al 2008), and the data are analysed with methods that are robust to the effect of diffusion (Greenhill et al. 2009). For example, the WALS contains information about word order (subject, object and verb), but additional distinctions can be made between word order for different kinds of clauses (e.g. main vs subordinate clauses) or between clausal and nominal objects. By identifying these and other more specific character states it may be possible to increase the historical signal in typological data (Reesink et al, 2009), although rates of evolution will then necessarily increase. In addition, the WALS data is unfortunately sparse,

containing only 138 characters (compared to the ~200 well-attested items of lexicon), and with many languages missing information – perhaps more signal will be evident in a more complete dataset. The identification of more specific character states may maximise the historical signal in typological data.

Whilst we were unable to identify a set of consistently stable typological features, rates of lexical evolution in one family were a good predictor of rates in the other. This fits with previous work showing that rates of change in lexical items are highly correlated across the Indo-European, Austronesian and Bantu language families (Pagel 2000; Pagel & Meade 2006). Recent work has also shown that rates of lexical change are predictable based on frequency of use and part of speech (Pagel et al. 2007) and that some meanings have a lexical ‘half life’ – the time after which there is a 50% chance that the word is replaced – in excess of 20,000 years. These extremely slow and predictable rates of lexical change mean that basic vocabulary may be a more practical choice for investigating questions of deeper language origins.

Finally, our findings highlight how little we know about the shape and tempo of language change. Contrary to what might be intuitively expected, our results indicate that dependencies between structural elements of language appear to do little to slow down rates of structural change, or to limit the diffusion of features between languages. In addition, we find that rates of structural evolution are specific to each language family, whilst lexical rates are correlated across families. One explanation for this observation may be that the frequency of use of different structural elements is an important determinant of rates of structural change, just as is the case for lexical change (Pagel et al.

2007). Whilst frequency of word-use is relatively constant across languages, the way structures are used depends on what other structural constraints operate in a language (Meillet 1948). This may explain the variation we see in rates of structural evolution between language families. In future, model-based approaches like those outlined here could be used to test hypotheses about macro-scale language change, and so shed light on the basic mechanisms driving the shape and tempo of language evolution.

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References

- Archie, J. W. 1989 Homoplasy Excess Ratios: New Indices for Measuring Levels of Homoplasy in Phylogenetic Systematics and a Critique of the Consistency Index. *Systematic Zoology* **38**, 253-269.
- Atkinson, Q. D. & Gray, R. D. 2006 How old is the Indo-European language family? Progress or more moths to the flame? In *Phylogenetic Methods and the Prehistory of Languages* (ed. P. Forster & C. Renfrew), pp. 91-109. Cambridge: McDonald Institute for Archaeological Research.
- Beekes, R. S. P. 1995 *Comparative Indo-European Linguistics*. Amsterdam: John Benjamins.
- Bisang, W. 2006 Contact-induced convergence: Typology and areality. In *Encyclopedia of Language and Linguistics*, vol. 3 (ed. K. Brown), pp. 88-101. Oxford: Elsevier.
- Blust, R. A. 1999 Subgrouping, circularity and extinction: Some issues in Austronesian comparative linguistics. In *Selected Papers from the Eighth International Conference on Austronesian Linguistics* vol. 1 (ed. E. Zeitoun & P. Li), pp. 31-94. Taipei, Taiwan: Academia Sinica.
- Blust, R. 2004. Austronesian nasal substitution: A survey. *Oceanic Linguistics* 43, 73-148.
- Blust, R. 2009. The Austronesian Languages. Canberra: Pacific Linguistics.
- Bryant, D. & Moulton, V. 2004 Neighbour-Net: An agglomerative method for the construction of phylogenetic networks. *Molecular Biology and Evolution* 21, 255-265.

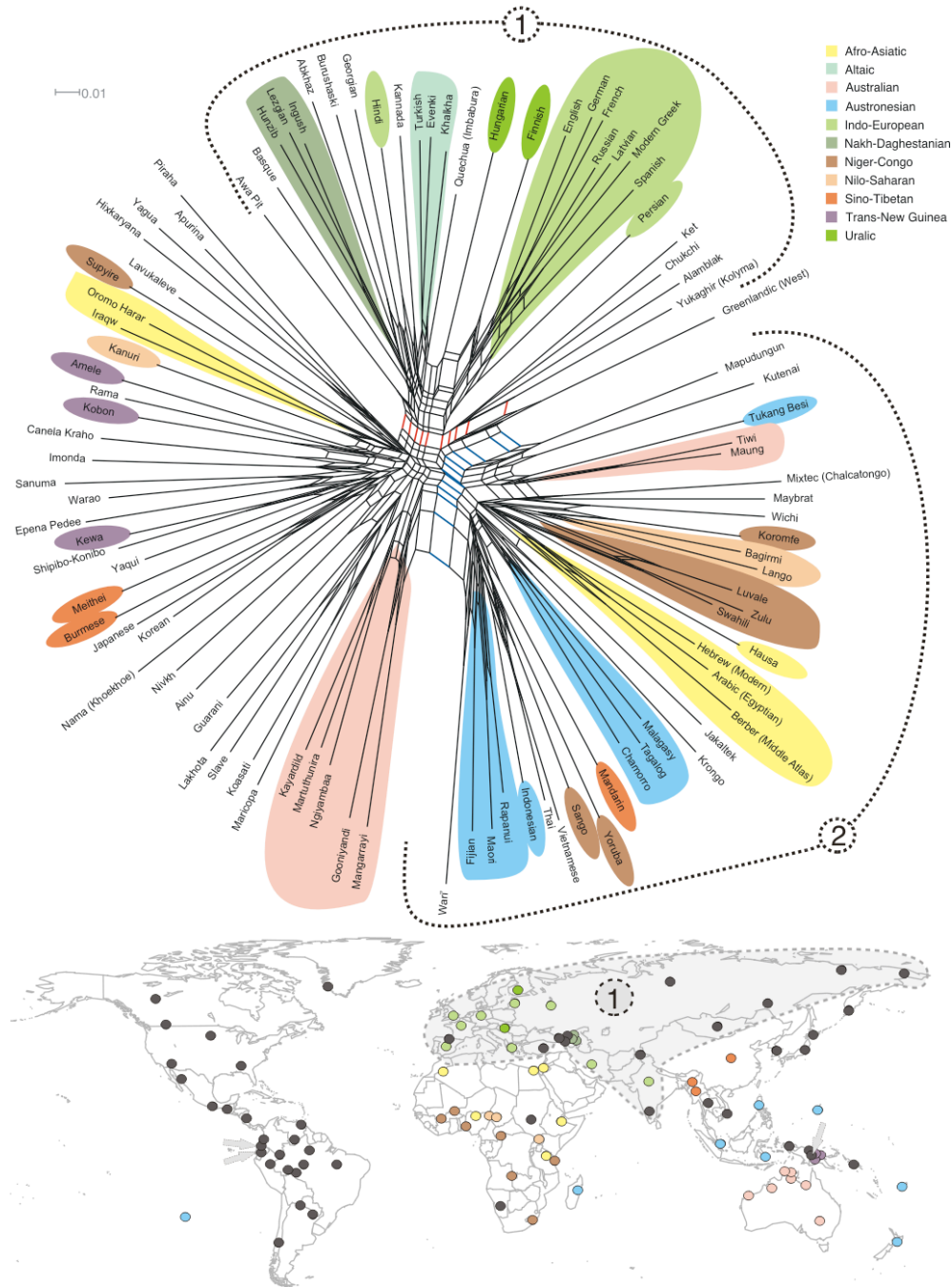
- Bryant, D., Filimon, F. & Gray, R. D. 2005 Untangling our past: Languages, trees, splits and networks. In *The Evolution of Cultural Diversity: Phylogenetic Approaches* (ed. R. Mace, C. J. Holden & S. Shennan), pp. 67-84. London: UCL Press.
- Dunn, M., Foley, R., Levinson, S., Reesink, G. & Terrill, A. 2007 Statistical reasoning in the evaluation of typological diversity in Island Melanesia. *Oceanic Linguistics* **46**, 388-403.
- Dunn, M., Levinson, S. C., Lindström, E., Reesink, G. & Terrill, A. 2008 Structural Phylogeny in Historical Linguistics: Methodological Explorations Applied in Island Melanesia. *Language* **84**, 710-759.
- Dunn, M., Terrill, A., Reesink, G., Foley, R. A. & Levinson, S. C. 2005 Structural phylogenetics and the reconstruction of ancient language history. *Science* **309**, 2072-5.
- Durie, M., & Ross, M. 1996. The comparative method reviewed : regularity and irregularity in language change. New York: Oxford University Press.
- Dyen, I., Kruskal, J. B. & Black, P. 1992 An Indoeuropean Classification: A Lexicostatistical Experiment. *Transactions of the American Philosophical Society* **82**, iii-132.
- Farris, J. 1989 The retention index and the rescaled consistency index. *Cladistics* **5**, 417-419.
- Gordon, R. G. J. (ed.) 2005 *Ethnologue: Languages of the World*. Dallas, Texas: SIL International.
- Gray, R. 2005 Pushing the time barrier in the quest for language roots. *Science* **309**, 2007-2008.

- Gray, R. D. & Atkinson, Q. D. 2003 Language-tree divergence times support the Anatolian theory of Indo-European origin. *Nature* **426**, 435-439.
- Gray, R. D., Drummond, A. J. & Greenhill, S. J. 2009 Language Phylogenies Reveal Expansion Pulses and Pauses in Pacific Settlement. *Science* **323**, 479-483.
- Greenhill, S. J., Blust, R. & Gray, R. D. 2008 The Austronesian Basic Vocabulary Database: from bioinformatics to lexomics. *Evolutionary Bioinformatics* **4**, 271-283.
- Greenhill, S. J., Currie, T. E. & Gray, R. D. 2009 Does horizontal transmission invalidate cultural phylogenies? In *Proceedings of the Royal Society B*.
- Harrison, S. P. 2003 On the limits of the comparative method. In *The handbook of historical linguistics* (ed. B. D. Joseph & R. D. Janda), pp. 213-243. Malden: Blackwell.
- Haspelmath, M., Dryer, M., Gil, D. & Comrie, B. (ed.) 2005 *The World Atlas of Language Structures*. Oxford: Oxford University Press.
- Kaufman, T. & Golla, V. 2000 Language Groupings in the New World: their reliability and usability in cross-disciplinary studies. In *America Past, America Present: Genes and Languages in the Americas and Beyond* (ed. C. Renfrew), pp. 47-57. Cambridge: The McDonald Institute for Archaeological Research.
- Lynch, J., Ross, M. & Crowley, T. 2002 *The Oceanic languages*. Curzon language family series. Richmond [England]: Curzon.
- Mace, R. & Pagel, M. 1994 The Comparative Method in Anthropology. *Current Anthropology* **35**, 549-564.
- Matras, Y., McMahon, A. & Vincent, N. 2006 *Linguistic areas: Convergence in historical and typological perspective*. New York: Palgrave.

- Meillet, A. 1948 *Linguistique historique et linguistique générale*. Paris: Champion.
- Nichols, J. 1992 *Linguistic diversity in space and time*. Chicago: University of Chicago Press.
- Nichols, J. 1994 The spread of language around the Pacific Rim. *Evolutionary Anthropology* **3**, 206-215.
- Pagel, M. 2000 Maximum-likelihood models for glottochronology and for constructing linguistic phylogenies. In *Time Depth in Historical Linguistics*, vol. 1 (ed. C. Renfrew, A. McMahon & L. Trask), pp. 189-207. Cambridge: McDonald Institute for Archaeological Research.
- Pagel, M., Atkinson, Q. D. & Meade, A. 2007 Frequency of word-use predicts rates of lexical evolution throughout Indo-European history. *Nature* **449**, 717-720.
- Pagel, M. & Meade, A. 2004 A phylogenetic mixture model for detecting pattern-heterogeneity in gene sequence or character-state data. *Systematic biology* **53**, 571-81.
- Pagel, M. & Meade, A. 2006 Estimating Rates of Lexical Replacement on Phylogenetic Trees of Languages. In *Phylogenetic methods and the prehistory of languages* (ed. C. Renfrew & P. Forster), pp. 173-182. Cambridge: McDonald institute Monographs.
- Pagel, M., Meade, A. & Barker, D. 2004 Bayesian estimation of ancestral character states on phylogenies. *Systematic Biology* **53**, 673-84.
- Peeters, B. 1990 Encore une fois 'où tout se tient'. *Historiographia Linguistica* **17**, 427-463.

- Reesink, G., Singer R., & Dunn, M. (2009) Explaining the Linguistic Diversity of Sahul Using Population Models. *PLoS Biology*, **7**, e1000241. doi:10.1371/journal.pbio.1000241
- Renfrew, C. & Nettle, D. 1999 *Nostratic: Examining a linguistic macrofamily*. Cambridge: McDonald Institute for Archaeological Research.
- Ringe, D. 1995 'Nostratic' and the factor of chance. *Diachronica* **12**, 55-74.
- Ross, M. D. 1996 Contact-induced change and the comparative method: cases from Papua New Guinea. In *The comparative method reviewed: regularity and irregularity in language change* (ed. M. Durie & M. D. Ross), pp. 180-217. New York: Oxford University Press.
- Swadesh, M. 1952 Lexico-Statistic Dating of Prehistoric Ethnic Contacts. *Proceedings of the American Philosophical Society* **96**, 452-463.
- Swofford, D. L. 2002 PAUP*: Phylogenetic Analysis Using Parsimony (* and Other Methods): Sinauer Associates.
- Thomason, S. G. & Kaufman, T. 1988 *Language contact, creolization, and genetic linguistics*. Berkeley: University of California Press.

Figures.



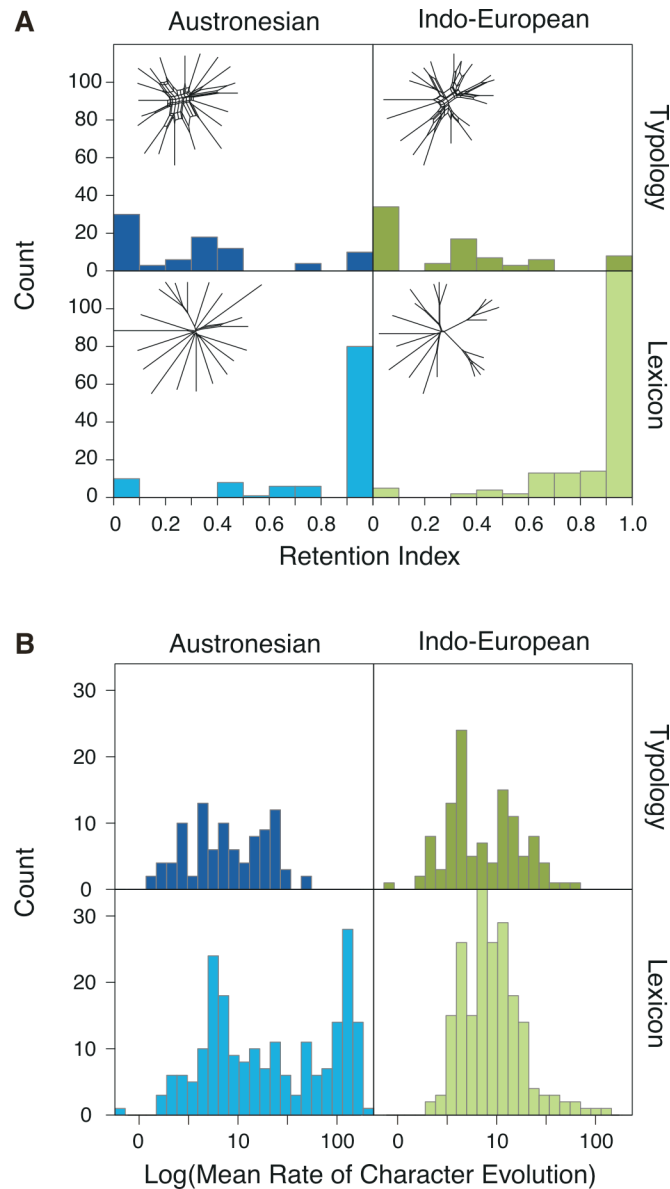


Figure 2: Histograms comparing the Austronesian and Indo-European lexical and typological data. 2A shows the retention index (RI) for each character state in the typological and lexical characters on the established classification trees. NeighbourNets of each dataset are inset in 2A, showing the greater amounts of conflicting signal found in both typological datasets. 2B shows the frequency distribution of observed rates of evolution across all the characters in both families and data types.

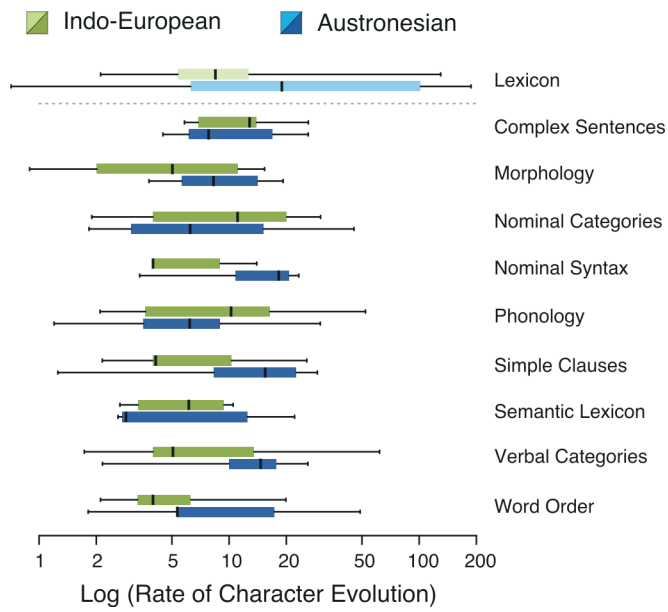


Figure 3: Boxplots showing the observed rate of change by feature class, across the Austronesian/Indo-European language families. Values closer to zero are evolving more slowly.

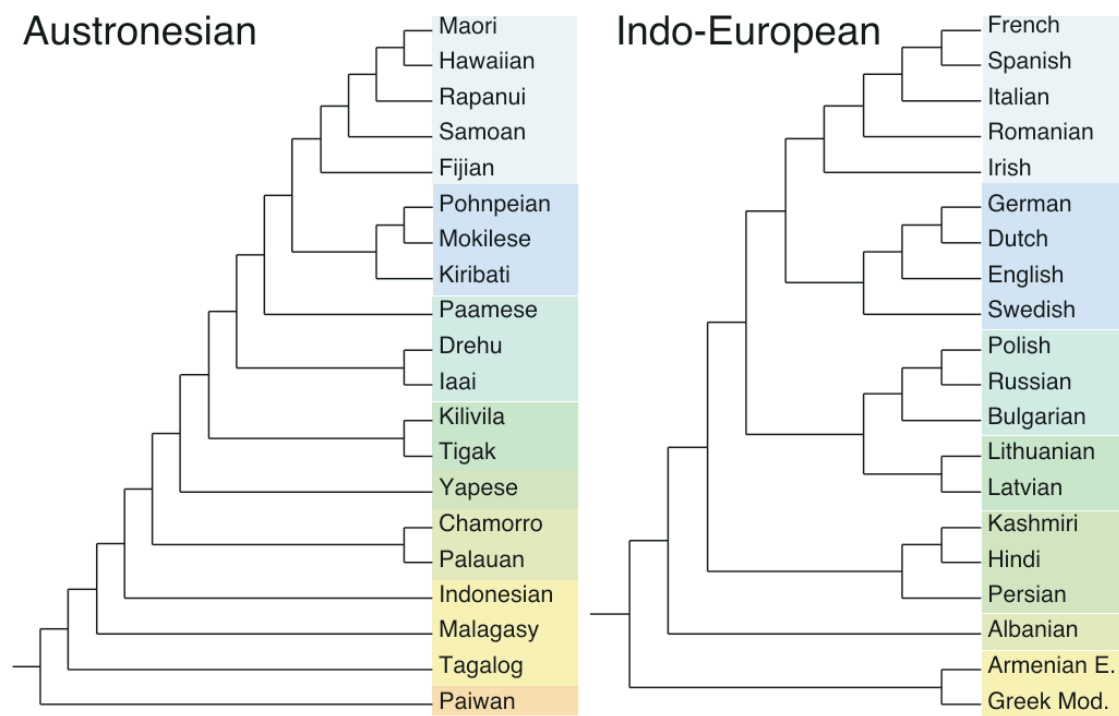


Figure S4: Trees for the Austronesian and Indo-European language families derived from linguistic classification. Languages are color-coded according to accepted subgroups.

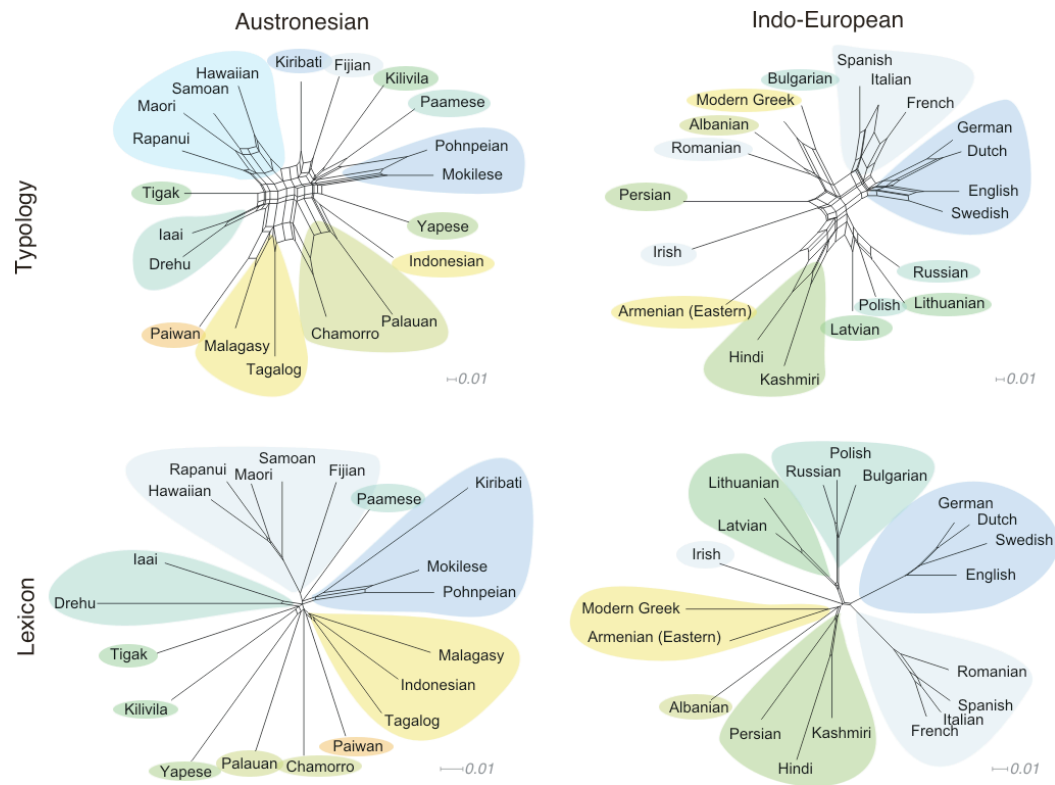


Figure S5: NeighbourNets for each lexical and typological dataset. Colors represent accepted subgroups.